

Implementation Notes for entropy estimation based on NIST SP 800-90B non-IID track

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1 Implementation notes for numerical computation of entropy estimates

1.1 Notes for the Markov estimate

In this section, we follow the convention using a column vector of probabilities \mathbf{P} and a left stochastic matrix \mathbf{T} .

$$\mathbf{P} \equiv \begin{bmatrix} p_0 \\ p_1 \end{bmatrix} \quad (1)$$

$$\mathbf{T} \equiv \begin{bmatrix} t_{0,0} & t_{0,1} \\ t_{1,0} & t_{1,1} \end{bmatrix} \quad (2)$$

Note here that 6.3.3 of NIST SP 800-90B[1] uses the convention of a row vector of probabilities and a right stochastic matrix. So the transition matrix components $t_{i,j}$ in this section will be interpreted as $P_{j,i}$ in 6.3.3 of NIST SP 800-90B[1].

Considering a sequence of binary-valued samples of length λ , there are following four possible combinations of the first sample value and the last sample value:

- a) $0 \rightarrow 0$
- b) $0 \rightarrow 1$
- c) $1 \rightarrow 0$
- d) $1 \rightarrow 1$

For each combination, the probability of occurrence of the most likely sequence is expressed by the following equation, by using the parameters μ and ν :

a)

$$f_a(\mathbf{T}, \mu, \nu) \equiv p_0 t_{00}^\nu t_{11}^{\lambda-1-2\mu-\nu} t_{01}^\mu t_{10}^\mu \quad (3)$$

$$\begin{cases} 0 \leq 2\mu < \lambda - 1 \\ 0 \leq \nu \leq \lambda - 1 \\ 0 \leq \lambda - 1 - 2\mu - \nu \leq \lambda - 1 \end{cases} \quad (4)$$

b)

$$f_b(\mathbf{T}, \mu, \nu) \equiv p_0 t_{00}^\nu t_{11}^{\lambda-2-2\mu-\nu} t_{01}^\mu t_{10}^{\mu+1} \quad (5)$$

$$\begin{cases} 0 \leq 2\mu < \lambda - 1 \\ 0 \leq \nu < \lambda - 1 \\ 0 \leq \lambda - 2 - 2\mu - \nu \leq \lambda - 2 \end{cases} \quad (6)$$

c)

$$f_c(\mathbf{T}, \mu, \nu) \equiv p_1 t_{00}^\nu t_{11}^{\lambda-2-2\mu-\nu} t_{01}^{\mu+1} t_{10}^\mu \quad (7)$$

$$\begin{cases} 0 \leq 2\mu < \lambda - 1 \\ 0 \leq \nu < \lambda - 1 \\ 0 \leq \lambda - 2 - 2\mu - \nu \leq \lambda - 2 \end{cases} \quad (8)$$

d)

$$f_d(\mathbf{T}, \mu, \nu) \equiv p_1 t_{00}^\nu t_{11}^{\lambda-1-2\mu-\nu} t_{01}^\mu t_{10}^\mu \quad (9)$$

$$\begin{cases} 0 \leq 2\mu < \lambda - 1 \\ 0 \leq \nu \leq \lambda - 1 \\ 0 \leq \lambda - 1 - 2\mu - \nu \leq \lambda - 1 \end{cases} \quad (10)$$

If we denote κ as the exponent of t_{11} , then the possible values of μ, ν , and κ are expressed as integer coordinates in the approximately triangular region (to be precise, truncated triangle) shown in Figure 1.

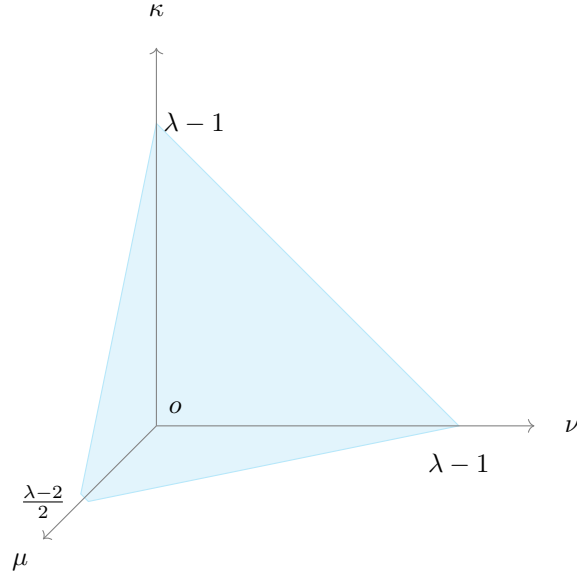


Figure 1 Parameter plane of the exponents κ, μ, ν

In order to calculate the extrema of f_a, f_b, f_c , and f_d , assuming that μ , and ν are continuous variables, we first consider the partial derivatives of $f_a(\mathbf{T}, \mu, \nu)$ with respect to μ , and ν .

$$\begin{aligned}\frac{\partial}{\partial \mu} f_a(\mathbf{T}, \mu, \nu) &= (-2 \ln t_{11} + \ln t_{01} + \ln t_{10}) f_a(\mathbf{T}, \mu, \nu) \\ &= \ln \frac{t_{01} t_{10}}{t_{11}^2} f_a(\mathbf{T}, \mu, \nu)\end{aligned}\tag{11}$$

$$\begin{aligned}\frac{\partial}{\partial \nu} f_a(\mathbf{T}, \mu, \nu) &= (\ln t_{00} - \ln t_{11}) f_a(\mathbf{T}, \mu, \nu) \\ &= \ln \frac{t_{00}}{t_{11}} f_a(\mathbf{T}, \mu, \nu)\end{aligned}\tag{12}$$

Taking into account eqs. (11) and (12), and the property that the value $f_a(\mathbf{T}, \mu, \nu)$ is non-negative, a rough phase diagram will be obtained as shown in Figure 2, by taking the horizontal axis as $\ln \frac{t_{00}}{t_{11}}$, and the vertical axis as $\frac{1}{2} \ln \frac{t_{01} t_{10}}{t_{11}^2}$.

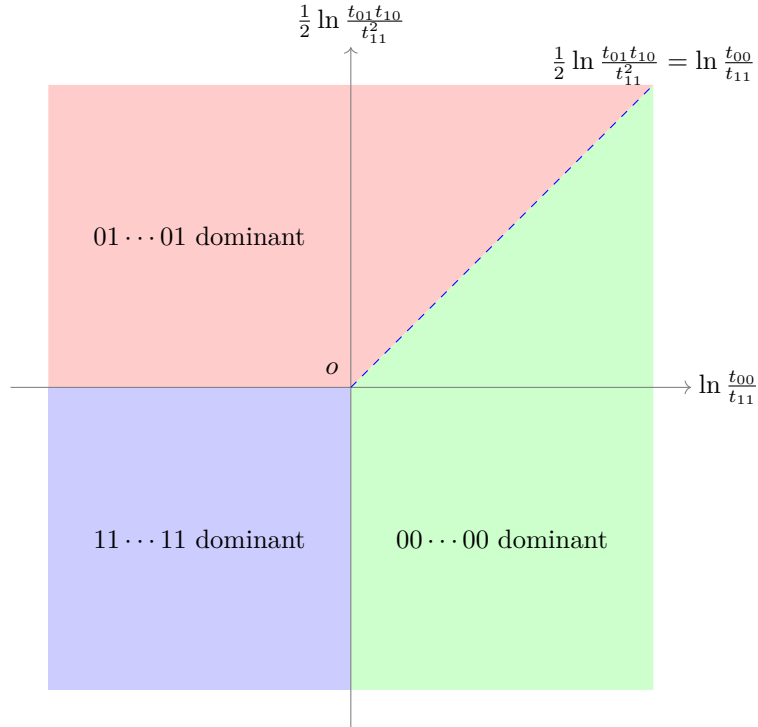


Figure 2 Phase diagram of most likely sequence based on $f_a(\mathbf{T}, \mu, \nu)$ with respect to components of stochastic matrix \mathbf{T}

First, if we consider the third quadrant, for example, the conditional probability t_{11} is greater than or equal to t_{00} and $\sqrt{t_{01} t_{10}}$, then the sequence $11 \cdots 11$ will be dominant.

Next, for the other quadrants, let us consider whether $00 \cdots 00$ or $01 \cdots 01$ is the dominant sequence. In particular, for the first quadrant, since right hand sides of eqs. (11) and (12) are positive, let us examine whether $00 \cdots 00$ or $0101 \cdots 0101$ is the dominant sequence. Since at least $11 \cdots 11$ is not a dominant sequence, we can limit the discussion with $\kappa = 0, 1$, the parameter region of possible values of μ and ν is restricted to a line segment in Figure 1. In this case, the total differentiation of f_a is then

expressed by the following equation:

$$\begin{aligned}\Delta f_a &= \left(\frac{\partial f_a}{\partial \mu} + \frac{\partial f_a}{\partial \nu} \frac{d\nu}{d\mu} \right) \Delta \mu \\ &= \left(\ln \frac{t_{01}t_{10}}{t_{00}^2} \right) \Delta \mu\end{aligned}\quad (13)$$

From this equation, if RHS of eq. (13) is positive, f_a increases with respect to μ and $01 \cdots 01$ becomes the dominant sequence, and conversely, if RHS of eq. (13) is negative, $00 \cdots 00$ becomes the dominant sequence. This will be determined by the following coefficient:

$$\ln \frac{t_{01}t_{10}}{t_{00}^2} = \ln \frac{t_{01}t_{10}}{t_{11}^2} - 2 \ln \frac{t_{00}}{t_{11}} \quad (14)$$

Figure 2 shows a dashed diagonal straight line in the first quadrant, on which the RHS of eq. (14) is zero. In other words, in the region above this line, $01 \cdots 01$ is the dominant sequence, and $00 \cdots 00$ is the dominant sequence in the region below the line.

From the above, we have to consider the following three dominant sequences:

- i) $00 \cdots 00$ dominant
- ii) $01 \cdots 01$ dominant
- iii) $11 \cdots 11$ dominant

For each combination of f_a, f_b, f_c, f_d , and above three dominant sequences i), ii), and iii), the following expressions are obtained giving the maximum probability with respect to μ and ν :

a)-i) $00 \cdots 00$ dominant

$$\max_{\mu, \nu} (f_a(\mathbf{T}, \mu, \nu)) = p_0 t_{00}^{\lambda-1} \quad (15)$$

a)-ii) $01 \cdots 01$ dominant

$$\max_{\mu, \nu} (f_a(\mathbf{T}, \mu, \nu)) = p_0 t_{00} t_{01}^{(\lambda-2)/2} t_{10}^{(\lambda-2)/2}, \text{ or} \quad (16)$$

$$\max_{\mu, \nu} (f_a(\mathbf{T}, \mu, \nu)) = p_0 t_{11} t_{01}^{(\lambda-2)/2} t_{10}^{(\lambda-2)/2} \quad (17)$$

a)-iii) $11 \cdots 11$ dominant

$$\max_{\mu, \nu} (f_a(\mathbf{T}, \mu, \nu)) = p_0 t_{11}^{\lambda-3} t_{01} t_{10} \quad (18)$$

b)-i) $00 \cdots 00$ dominant

$$\max_{\mu, \nu} (f_b(\mathbf{T}, \mu, \nu)) = p_0 t_{00}^{\lambda-2} t_{10} \quad (19)$$

b)-ii) $01 \cdots 01$ dominant

$$\max_{\mu, \nu} (f_b(\mathbf{T}, \mu, \nu)) = p_0 t_{01}^{(\lambda-2)/2} t_{10}^{\lambda/2} \quad (20)$$

b)-iii) $11 \cdots 11$ dominant

$$\max_{\mu, \nu} (f_b(\mathbf{T}, \mu, \nu)) = p_0 t_{11}^{\lambda-2} t_{10} \quad (21)$$

c)-i) $00 \cdots 00$ dominant

$$\max_{\mu, \nu} (f_c(\mathbf{T}, \mu, \nu)) = p_1 t_{00}^{\lambda-2} t_{01} \quad (22)$$

c)-ii) $01 \cdots 01$ dominant

$$\max_{\mu, \nu}(f_c(\mathbf{T}, \mu, \nu)) = p_1 t_{01}^{\lambda/2} t_{10}^{(\lambda-2)/2} \quad (23)$$

c)-iii) $11 \cdots 11$ dominant

$$\max_{\mu, \nu}(f_c(\mathbf{T}, \mu, \nu)) = p_1 t_{11}^{\lambda-2} t_{01} \quad (24)$$

d)-i) $00 \cdots 00$ dominant

$$\max_{\mu, \nu}(f_d(\mathbf{T}, \mu, \nu)) = p_1 t_{00}^{\lambda-3} t_{01} t_{10} \quad (25)$$

d)-ii) $01 \cdots 01$ dominant

$$\max_{\mu, \nu}(f_d(\mathbf{T}, \mu, \nu)) = p_1 t_{00} t_{01}^{(\lambda-2)/2} t_{10}^{(\lambda-2)/2}, \text{ or} \quad (26)$$

$$\max_{\mu, \nu}(f_d(\mathbf{T}, \mu, \nu)) = p_1 t_{11} t_{01}^{(\lambda-2)/2} t_{10}^{(\lambda-2)/2} \quad (27)$$

d)-iii) $11 \cdots 11$ dominant

$$\max_{\mu, \nu}(f_d(\mathbf{T}, \mu, \nu)) = p_1 t_{11}^{\lambda-1} \quad (28)$$

Note here that we considered the property that λ is even in the above expressions. Also note that all boundary conditions are considered in the above calculation, combination of first sample values and last sample values.

From the above, we have to compare 14 expressions listed in Table 1.

Table 1 Probabilities of the most or second most likely sequences of length λ

No.	Sequence	Probability	$-\log_2(\text{Probability})/\lambda$	Notes
1	0000...0000	$p_0 \times t_{00}^{\lambda-1}$	$-\left[\frac{\lambda-1}{\lambda} \log_2 t_{00} + \frac{1}{\lambda} \log_2 p_0\right]$	a, b
2	0101...0101001010...1010	$p_0 \times t_{00} \times t_{01}^{(\lambda-2)/2} \times t_{10}^{(\lambda-2)/2}$	$-\left[\frac{\lambda-2}{2\lambda} \log_2 (t_{01}t_{10}) + \frac{1}{\lambda} \log_2 (p_0 \times t_{00})\right]$	b
3	0101...0101101010...1010	$p_0 \times t_{11} \times t_{01}^{(\lambda-2)/2} \times t_{10}^{(\lambda-2)/2}$	$-\left[\frac{\lambda-2}{2\lambda} \log_2 (t_{01}t_{10}) + \frac{1}{\lambda} \log_2 (p_0 \times t_{11})\right]$	
4	0111...1110	$p_0 \times t_{11}^{\lambda-3} \times t_{01} \times t_{10}$	$-\left[\frac{\lambda-3}{\lambda} \log_2 t_{11} + \frac{1}{\lambda} \log_2 (p_0 \times t_{01} \times t_{10})\right]$	
5	0000...0001	$p_0 \times t_{00}^{\lambda-2} \times t_{10}$	$-\left[\frac{\lambda-2}{\lambda} \log_2 t_{00} + \frac{1}{\lambda} \log_2 (p_0 \times t_{10})\right]$	
6	0101...0101	$p_0 \times t_{01}^{(\lambda-2)/2} \times t_{10}^{\lambda/2}$	$-\left[\frac{\lambda-2}{2\lambda} \log_2 (t_{01}t_{10}) + \frac{1}{\lambda} \log_2 (p_0 \times t_{10})\right]$	a, b
7	0111...1111	$p_0 \times t_{11}^{\lambda-2} \times t_{10}$	$-\left[\frac{\lambda-2}{\lambda} \log_2 t_{11} + \frac{1}{\lambda} \log_2 (p_0 \times t_{10})\right]$	a, c
8	1000...0000	$p_1 \times t_{00}^{\lambda-2} \times t_{01}$	$-\left[\frac{\lambda-2}{\lambda} \log_2 t_{00} + \frac{1}{\lambda} \log_2 (p_1 \times t_{01})\right]$	a, b
9	1010...1010	$p_1 \times t_{01}^{\lambda/2} \times t_{10}^{(\lambda-2)/2}$	$-\left[\frac{\lambda-2}{2\lambda} \log_2 (t_{01}t_{10}) + \frac{1}{\lambda} \log_2 (p_1 \times t_{01})\right]$	a, b
10	1111...1110	$p_1 \times t_{11}^{\lambda-2} \times t_{01}$	$-\left[\frac{\lambda-2}{\lambda} \log_2 t_{11} + \frac{1}{\lambda} \log_2 (p_1 \times t_{01})\right]$	
11	1000...0001	$p_1 \times t_{00}^{\lambda-3} \times t_{01} \times t_{10}$	$-\left[\frac{\lambda-3}{\lambda} \log_2 t_{00} + \frac{1}{\lambda} \log_2 (p_1 \times t_{01} \times t_{10})\right]$	
12	1010...1010100101...0101	$p_1 \times t_{00} \times t_{01}^{(\lambda-2)/2} \times t_{10}^{(\lambda-2)/2}$	$-\left[\frac{\lambda-2}{2\lambda} \log_2 (t_{01}t_{10}) + \frac{1}{\lambda} \log_2 (p_1 \times t_{00})\right]$	
13	1010...1010110101...0101	$p_1 \times t_{11} \times t_{01}^{(\lambda-2)/2} \times t_{10}^{(\lambda-2)/2}$	$-\left[\frac{\lambda-2}{2\lambda} \log_2 (t_{01}t_{10}) + \frac{1}{\lambda} \log_2 (p_1 \times t_{11})\right]$	c
14	1111...1111	$p_1 \times t_{11}^{\lambda-1}$	$-\left[\frac{\lambda-1}{\lambda} \log_2 t_{11} + \frac{1}{\lambda} \log_2 p_1\right]$	a, b

^a Included in 6.3.3 of NIST SP 800-90B[1].

^b Explicitly included in paragraph 516 of AIS 20/31 Version 2.35 - DRAFT [2], by using the relation $m \equiv \lambda - 1$.

^c Implicitly included in paragraph 516 of AIS 20/31 Version 2.35 - DRAFT [2], by using the relation $m \equiv \lambda - 1$ and by relabeling the state space $\Omega = \{0, 1\}$.

1.2 Notes for the collision estimate

In this section, we try to rewrite the following equation using elementary functions:

$$F(1/z) = \Gamma(3, z)z^{-3} \exp(z) \quad (29)$$

In NIST SP 800-90B[1], it is documented to evaluate incomplete gamma function using continued fraction. However, we try to obtain simpler expression using known properties of incomplete gamma function without using continued fraction.

First we try to use the following equations (see 8.4.8 and 8.4.11 in [3]).

$$\Gamma(n+1, z) = n! \exp(-z) e_n(z), \quad (30)$$

$$e_n(z) = \sum_{k=0}^n \frac{z^k}{k!} \quad (31)$$

With these two equations, $F(1/z)$ can be rewritten to as follows:

$$\begin{aligned} F(1/z) &= \Gamma(3, z)z^{-3} \exp(z) \\ &= 2! \exp(-z) e_2(z) z^{-3} \exp(z) \\ &= 2z^{-3} e_2(z) \\ &= 2z^{-3} \sum_{k=0}^2 \frac{z^k}{k!} \\ &= 2z^{-3} \left(1 + z + \frac{z^2}{2} \right) \\ &= z^{-1} (2z^{-2} + 2z^{-1} + 1) \end{aligned} \quad (32)$$

By replacing the argument $1/z$ with q , the following expression can be obtained:

$$F(q) = q (2q^2 + 2q + 1) \quad (33)$$

From this expression, the range of $F(q)$ is $[0, \frac{5}{4}]$, with respect to the domain $q \in [0, 0.5]$.

If we denote $g(p)$ as the right hand side (RHS) of equation to be solved, in step 7 of 6.3.2 of NIST SP

800-90B[1], $g(p)$ can be rewritten by using Eq.(33).

$$\begin{aligned}
g(p) &\equiv pq^{-2} \left[1 + \frac{1}{2}(p^{-1} - q^{-1}) \right] F(q) - pq^{-1} \frac{1}{2}(p^{-1} - q^{-1}) \\
&= pq^{-2} \left[1 + \frac{1}{2}(p^{-1} - q^{-1}) \right] q (2q^2 + 2q + 1) - pq^{-1} \frac{1}{2}(p^{-1} - q^{-1}) \\
&= pq^{-1} \left[1 + \frac{1}{2}(p^{-1} - q^{-1}) \right] (2q^2 + 2q + 1) - pq^{-1} \frac{1}{2}(p^{-1} - q^{-1}) \\
&= pq^{-1} \left[1 + \frac{1}{2}(p^{-1} - q^{-1}) \right] - pq^{-1} \frac{1}{2}(p^{-1} - q^{-1}) \\
&\quad + pq^{-1} \left[1 + \frac{1}{2}(p^{-1} - q^{-1}) \right] (2q^2 + 2q) \\
&= pq^{-1} + 2p \left[1 + \frac{1}{2}(p^{-1} - q^{-1}) \right] (q + 1) \\
&= pq^{-1} + 2p \left[(q + 1) + \frac{1}{2}p^{-1}(q + 1) - \frac{1}{2}(1 + q^{-1}) \right] \\
&= pq^{-1} + p [2(q + 1) + p^{-1}(q + 1) - (1 + q^{-1})] \\
&= p [2(q + 1) + p^{-1}(q + 1) - 1] \\
&= p [2q + 1 + p^{-1}(q + 1)] \\
&= (2pq + p + q + 1) \\
&= (2pq + 2) \\
&= 2(pq + 1) \\
&= 2[p(1 - p) + 1] \\
&= 2 \left[-\left(p - \frac{1}{2}\right)^2 + \frac{5}{4} \right]
\end{aligned} \tag{34}$$

Figure 3 shows Eq.(34) graphically.

If \bar{X}' is in range $[2, \frac{5}{2}]$, the solution p of step 7 of 6.3.2 of NIST SP 800-90B[1] can be expressed by the following equation:

$$p = \frac{1}{2} + \sqrt{\frac{5}{4} - \frac{\bar{X}'}{2}}. \tag{35}$$

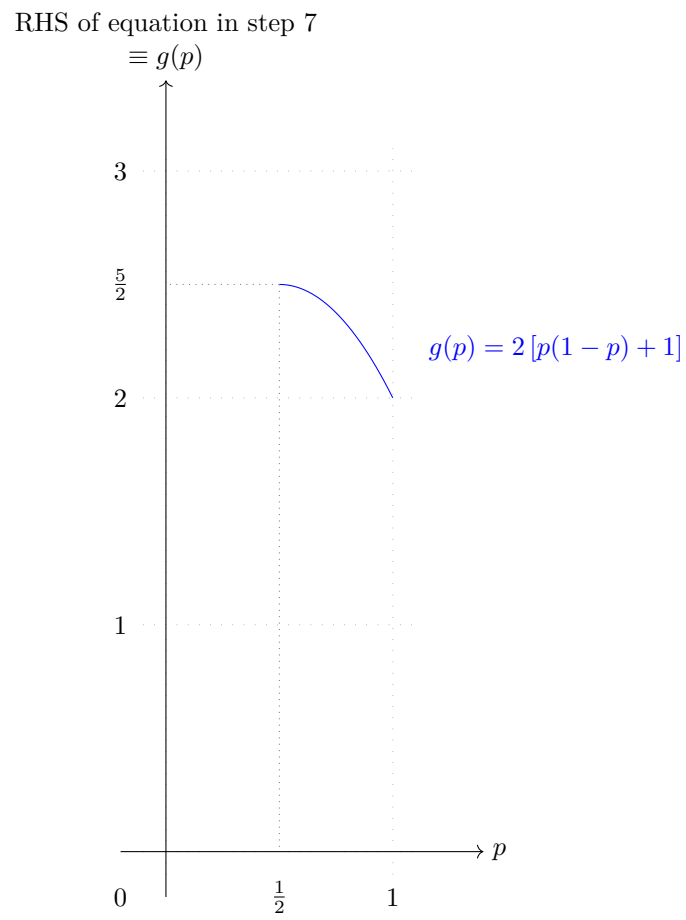


Figure 3 The right hand side of the equation in step 7 of 6.3.2 of NIST SP 800-90B

1.3 Notes for the compression estimate

As documented in [4], Eq.(36) should be replaced by Eq.37.

$$G(z) = \frac{1}{\nu} \sum_{t=d+1}^L \sum_{u=1}^t \log_2(u) F(z, t, u) \quad (36)$$

$$G(z) = \frac{1}{\nu} \sum_{t=d+1}^{\lfloor L/b \rfloor} \sum_{u=2}^t \log_2(u) F(z, t, u) \quad (37)$$

$F(z, t, u)$ in Eq.(37) can be expressed by the following equation:

$$F(z, t, u) = \begin{cases} z^2(1-z)^{u-1} & \text{if } u < t \\ z(1-z)^{t-1} & \text{if } u = t \end{cases} \quad (38)$$

We have to evaluate Eq.(37), but this expression requires relatively large computational complexity due to its nested summation. First Figure 4 shows parameters u, t , where nested summation must be taken to compute $G(z)$.

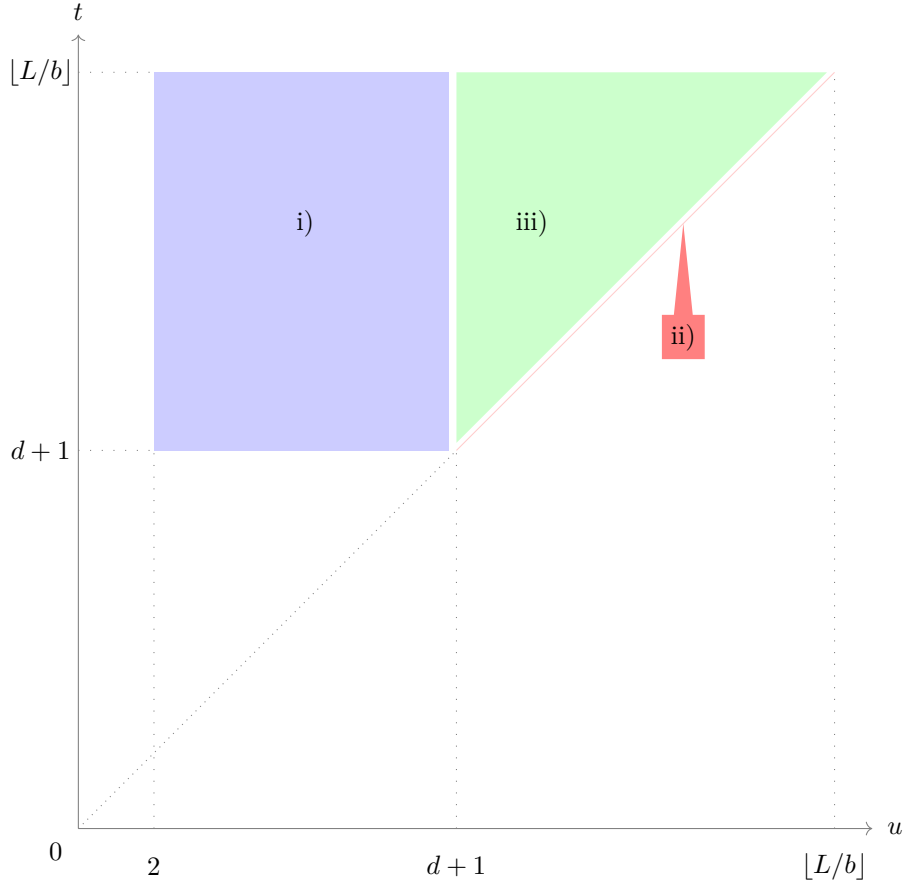


Figure 4 Parameters u, t , where nested summation must be taken

t -dependency of $F(z, t, u)$ must be considered when $u = t$. From the above, the parameters domain where summation must be taken can be divided into the following three groups:

i)

$$\begin{cases} 2 \leq u < d+1 \\ d+1 \leq t \leq \lfloor L/b \rfloor \end{cases} \quad (39)$$

ii)

$$\begin{cases} (u, t) = (\tau, \tau) \\ d+1 \leq \tau \leq \lfloor L/b \rfloor \end{cases} \quad (40)$$

iii)

$$\begin{cases} d+1 \leq u < t \\ d+1 < t \leq \lfloor L/b \rfloor \end{cases} \quad (41)$$

For groups i) and iii), the summation with respect to t can be taken first. By using the relation $\nu = \lfloor L/b \rfloor - d$, Eq.(37) can be rewritten to as follows:

$$\begin{aligned} G(z) &= \sum_{u=2}^d \log_2(u) z^2 (1-z)^{u-1} \\ &\quad + \frac{1}{\nu} \sum_{u=d+1}^{\lfloor L/b \rfloor} \log_2(u) z (1-z)^{u-1} \\ &\quad + \frac{1}{\nu} \sum_{u=d+1}^{\lfloor L/b \rfloor - 1} (\lfloor L/b \rfloor - u) \log_2(u) z^2 (1-z)^{u-1} \end{aligned} \quad (42)$$

With this expression, the computational complexity can be decreased from $\mathcal{O}(\nu^2)$ to $\mathcal{O}(\nu)$.

1.4 Notes for t-Tuple estimate, LRS estimate, MultiMMC Prediction and LZ78Y prediction estimate

In the following groups of entropy estimate, we have to handle t -tuples (pairs, triples, etc.)

- a) t-Tuple estimate (NIST SP 800-90B 6.3.5)
- b) Longest Repeated Substring (LRS) estimate (NIST SP 800-90B 6.3.6)
- c) Multi Most Common in Window Prediction estimate (NIST SP 800-90B 6.3.7)
- d) The MultiMMC Prediction estimate (NIST SP 800-90B 6.3.9)
- e) The LZ78Y Prediction estimate (NIST SP 800-90B 6.3.10)

In order to express t -tuples, bitset or multiprecision integer is used without using array.

1.5 Notes for Multi Most Common in Window prediction estimate

In step 3-a-i of 6.3.7 of NIST SP 800-90B, it is required to compute the mode in the previous window of w_j before s_i . As the order of w_j is 1000, the computational complexity is estimated to be about $\mathcal{O}(1000n)$, and expected to be time-consuming. We attempt to reduce time-complexity by preparing histograms of certain lengths in advance, and in step 3-a-i, we use the histograms that fit within the target window, and compute the unavailable parts on-demand basis.

- [1] Meltem Sönmez Turan, Elaine Barker, John Kelsey, Kerry A. McKay, Mary L. Baish, Mike Boyle *Recommendation for the Entropy Sources Used for Random Bit Generation*, NIST Special Publication 800-90B, Jan. 2018
- [2] M. Peter and W. Schindler, *A Proposal for Functionality Classes for Random Number Generators* Version 2.35 - DRAFT, September 2, 2022. https://www.bsi.bund.de/SharedDocs/Downloads/EN/BSI/Certification/Interpretations/AIS_31_Functionality_classes_for_random_number_generators_e.pdf?__blob=publicationFile&v=7
- [3] Franck W. J. Oliver, Daniel W. Lozier, Ronald F. Boisvert, Charles W. Clark, *NIST Handbook of Mathematical Functions*, National Institute of Standards and Technology, 2010
- [4] G. Sakurai, *Proposed list of corrections for NIST SP 800-90B 6.3 Estimators*, Dec. 2022 https://github.com/g-g-sakura/AnotherEntropyEstimationTool/blob/main/documentation/ProposedListOfCorrections_SP800-90B.pdf