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## Chapter 17

### The Theory of Probability

'Statistics is a branch of theology' -  
A Cambridge research fellow

Probability theory has its origins in questions about gambling. In a game of cards, or dice, when do I have the best chance of winning? What are the odds?<sup>1</sup>

Because games are usually finite, the methods needed to handle such questions are *combinatorial*, that is, based on counting arguments. For example, to find the chance of throwing three consecutive heads with a coin one lists the possibilities

HHH HHT HTH HTT  
TTH THT TTH TTT

which are 8 in number. Exactly 1 is favourable, so the probability is 1/8.

This of course makes the assumption that throws of *H* or *T* are equally likely. Now we can't define 'equally likely' by saying 'probability 1/2' until we have defined what we mean by 'probability 1/2': and we can't do that without defining 'equally likely'. Or at least, so it seems.

If we try to get round it by doing experiments, we run into another difficulty. If *H* and *T* are equally likely, then in a long series of throws we would expect to have approximately equal numbers of *H* and *T*. Not exactly equal, of course: they couldn't possibly be equal in an odd number of throws anyway, and in an even number of throws there would probably be a small discrepancy. Toss a coin 20 times and see if you get exactly 10 heads. (If you do, try several more times and see how often it happens!)

What we would hope is that 'in the limit' the ratio of the number of *H*s to the number of *T*s should 'tend to' 1/2. The trouble is that this 'limit' is not a limit in the usual sense of analysis. It is conceivable that we might throw a sequence consisting entirely of *H*s with a *fair* coin. It is, of course, unlikely. But to set up an idea of 'limit' which takes account of this possibility involves making precise what we mean by 'unlikely', which seems to require a definition of 'probability' again!

It wasn't until the 1930s that these difficulties were circumvented. This was achieved by developing an *axiomatic* probability theory. By divorcing the mathematics from its applications one can develop the mathematics without any logical qualms: then it can be tested experimentally to see if it fits the facts. Axiomatic probability theory succeeds for the same reason that axiomatic geometry succeeds.

### Combinatorial Probability

For the moment assume that we know what 'equally likely' means. Then a rough working definition of the probability  $p(E)$  of an event *E* is

$$p(E) = \frac{\text{number of ways in which } E \text{ can occur}}{\text{total number of possible occurrences}}$$

(provided all occurrences are equally likely).

Thus there are 36 ways of throwing 2 dice; and 5 of these give a total of 6 (namely 1+5, 2+4, 3+3, 4+2, 5+1).

$$\text{Therefore the probability that the total is 6 is } \frac{\text{number of ways of throwing 6}}{36}$$

which is 5/36.

Since the numbers involved are positive, and since the number of ways *E* can occur is at most equal to the total number of occurrences, we see that

$$0 \leq p(E) \leq 1.$$

If  $p(E) = 0$  then *E* is impossible; if  $p(E) = 1$  then *E* is certain.

The techniques of combinatorial probability centre around ways of combining events. Suppose we have two distinct events *E* and *F*. What is the probability that *either E or F* occurs?

Take the case of a die. *E* is the event '6 is thrown' and *F* the event '5 is thrown'. *E* or *F* is '5 or 6 is thrown' which obviously occurs 2 times out of 6. So

$$p(E \text{ or } F) = 1/3.$$

In general, let  $N(E)$  and  $N(F)$  be the number of ways in which *E* and *F* can occur, and  $T$  the total number of occurrences. Then

$$p(E \text{ or } F) = N(E \text{ or } F)/T.$$

What is  $N(E \text{ or } F)$ ? Suppose the events  $E$  and  $F$  do not 'overlap'. (I'll return to this point.) Then

$$N(E \text{ or } F) = N(E) + N(F)$$

so that

$$\begin{aligned} p(E \text{ or } F) &= (N(E) + N(F))/T \\ &= (N(E)/T + (N(F)/T)) \\ &= p(E) + p(F). \end{aligned} \quad (1)$$

If, however,  $E$  and  $F$  do overlap then  $N(E) + N(F)$  counts everything in the overlap *twice* whereas  $N(E \text{ or } F)$  only counts it *once*.

Suppose, for instance, that

$E$  = 'A prime number is thrown'

$F$  = 'An odd number is thrown'.

Then  $E$  occurs in three ways: 2, 3, 5. (Note: 1 is not prime.) And  $F$  occurs in three ways: 1, 3, 5. But  $E$  or  $F$  occurs in *four* ways: 1, 2, 3, 5. So

$$p(E) = 1/2 \quad p(F) = 1/2 \quad p(E \text{ or } F) = 2/3.$$

What happens in general is that

$N(E \text{ or } F) = N(E) + N(F) - N(E \text{ and } F)$  (2)  
because subtracting  $N(E \text{ and } F)$  puts right the double count in the overlap. In the above example,  $E$  and  $F$  occurs in two ways: 3, 5. So the equation gives

$$4 = 3 + 3 - 2$$

which is correct.

Dividing (2) by  $T$  we get

$$p(E \text{ or } F) = p(E) + p(F) - p(E \text{ and } F). \quad (3)$$

## Enter Set Theory

We can express these ideas much better in terms of sets. The possible *outcomes* when throwing a die form a set

$$X = \{1, 2, 3, 4, 5, 6\}.$$

The events  $E$  and  $F$  are represented by *subsets* of  $X$

$$E = \{2, 3, 5\}$$

$$F = \{1, 3, 5\}$$

as in Figure 172.

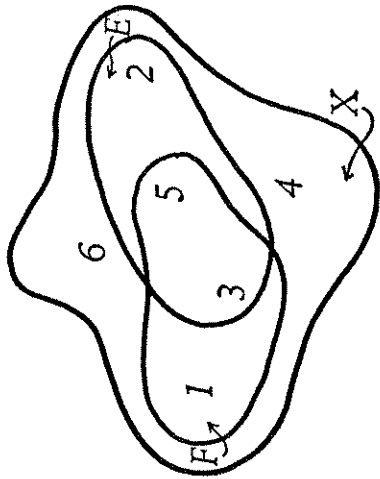


Figure 172

The event ' $E$  or  $F$ ' is the set  $\{1, 2, 3, 5\}$  which is the *union*  $E \cup F$ . The event ' $E$  and  $F$ ' is the set  $\{3, 5\}$  which is the *intersection*  $E \cap F$ . The probability  $p$  is a function defined on the set  $\mathcal{E}$  of all subsets of  $X$  with target  $R$ . In general we can say a little more about  $p$ : it has target  $[0, 1]$ , where this denotes the set of real numbers between 0 and 1.

Abstracting from this we obtain the idea of a *finite probability space*. This comprises

- (i) a finite set  $X$ ,
- (ii) the set  $\mathcal{E}$  of all subsets of  $X$
- (iii) a function  $p: \mathcal{E} \rightarrow [0, 1]$  with the property that

$$p(E \cup F) = p(E) + p(F) - p(E \cap F)$$

for all  $E, F \in \mathcal{E}$ .

Axiomatic probability theory works entirely in terms of probability spaces. However, if one wishes to consider infinite probability spaces, the definition has to be made more subtly. In many applications it is necessary to have infinite sets  $X$ : for example the height of a man can be any real number (within certain limits) so there are infinitely many possibilities.

## Independence

Another basic operation in probability theory deals with two trials in succession: what is the probability of event  $E$  occurring

on the first trial, event  $F$  on the second? For example, we throw a die twice: what are the chances that we throw first a 5, then a 2? Of the 36 possible combinations, only 1 is favourable: 5 followed by 2. So the probability is  $1/36$ .

If  $E$  and  $F$  were the events considered in the previous section, then  $E$  can occur first in 3 ways,  $F$  second in 3 ways. We can pair any occurrence of  $E$  with any of  $F$ , giving  $3 \times 3 = 9$  favourable outcomes. So the probability of  $E$  followed by  $F$  is  $9/36 = 1/4$ .

In general we must suppose that there are  $T_1$  possible outcomes of the first trial, of which  $N(E)$  are occurrences of  $E$ ; and  $T_2$  in the second,  $N(F)$  being occurrences of  $F$ . Then in the two trials together the total number of outcomes is  $T_1 \times T_2$ , because any of the  $T_1$  possibilities for the first can be followed by any of the  $T_2$  possibilities for the second. In the same way the number of ways in which  $E$  can occur first, followed by  $F$ , is  $N(E) \times N(F)$ . So

$$\begin{aligned} p(E \text{ followed by } F) &= \frac{N(E) \times N(F)}{T_1 T_2} \\ &= \frac{N(E)}{T_1} \times \frac{N(F)}{T_2} \\ &= p(E) \times p(F). \end{aligned} \quad (4)$$

In this calculation we must assume that  $E$  and  $F$  are independent: that the outcome of the first trial does not alter the probabilities in the second one.

This would not be the case if, say, the second event  $F$  was 'the total thrown is 4'. For if the first throw is 4 or more, the chance of success on the second is 0; if the first throw is 1, 2, or 3 the chance of success on the second is  $1/6$ .

The notion of independence can be formulated in terms of probability spaces. In applications, one takes as hypothesis the independence of the real-world events to be considered, applies the theory, and tests the result by experiment.

### Paradoxical Dice

Often our intuition about probabilities is wrong. Consider four dice  $A$ ,  $B$ ,  $C$ ,  $D$  marked

$A$ :	0	0	4	4	4	4
$B$ :	3	3	3	3	3	3
$C$ :	2	2	2	2	7	7
$D$ :	1	1	1	5	5	5

(The precise arrangement of faces does not matter.)

What is the probability that in a single throw die  $A$  will have a higher number showing than die  $B$ ?

$B$  always throws a 3. If  $A$  throws 4, which happens 4 times out of 6, he wins. If he throws 0, which happens 2 times out of 6, he loses. Therefore

$A$  beats  $B$  with probability  $2/3$ .

If  $B$  is thrown in competition with  $C$  it will win when  $C$  shows 2, lose when  $C$  shows 7. So

$B$  beats  $C$  with probability  $2/3$ .

If  $C$  plays against  $D$  matters are more complicated. With probability  $1/2$ ,  $D$  shows 1, and then  $C$  always wins; with probability  $1/2$   $D$  shows 5 and  $C$  wins by showing 7 with probability  $1/3$ . The probability that  $C$  will win is therefore

$$\frac{1}{2} \cdot 1 + \frac{1}{2} \cdot \frac{1}{3} = \frac{1}{2} + \frac{1}{6} = \frac{2}{3}.$$

Thus

$C$  beats  $D$  with probability  $2/3$ .

Finally, look at  $D$  versus  $A$ . If  $D$  shows 5, with probability  $1/2$ , then  $D$  always wins. If  $D$  shows 1, with probability  $1/2$ , then  $D$  wins if  $A$  shows 0, which has probability  $1/3$ . The probability that  $D$  will win is

$$\frac{1}{2} \cdot 1 + \frac{1}{2} \cdot \frac{1}{3} = \frac{2}{3}.$$

Thus

$D$  beats  $A$  with probability  $2/3$ .

Now a die which wins more often than not is clearly 'better' than one which loses more often than it wins. In these terms,

$A$  is better than  $B$ .

$B$  is better than  $C$ .

$C$  is better than  $D$ .

and  $D$  is better than  $A$ .

There is nothing wrong with these calculations. If you play the game in practice, and let your opponent choose his die, then you can always choose another that gives you odds of 2:1 for a win.

We expect that  $A$  better than  $B$  better than  $C$  better than  $D$  should mean  $A$  better than  $D$ . We're wrong. In the present context the meaning of 'better than' depends on the choice of dice: we are really playing four *different* games. It is as if we had four people playing games: Alfred beats Bertram at tennis, Bertram beats Charlotte at chess, Charlotte beats Dierdre at badminton - and Dierdre beats Alfred at shove-halfpenny.

Those economists who believe that commodities can be ordered by majority preference might take note of this phenomenon.

### Binomial Bias

Imagine a *biased* coin. Instead of coming down heads and tails with equal frequency, it has a preference for one particular side.

Such a coin provides a model for many probabilistic processes. If we are throwing a die and are interested only in whether a 6 turns up, we are effectively dealing with a biased coin such that  $p(\text{head}) = 1/6$ ,  $p(\text{tail}) = 5/6$ . If we are looking at the sex of newborn babies, we have  $p(\text{boy}) = 0.52$ ,  $p(\text{girl}) = 0.48$ .

In general we let

$$p = p(\text{head})$$

$$q = p(\text{tail})$$

and of course  $p + q = 1$ , because from (1) above

$$p(\text{head}) + p(\text{tail}) = p(\text{head or tail}) = 1.$$

Using the theory of independent events we easily find the following list of probabilities for sequences of heads and tails:

$H$	$p$	$HH$	$p^2$	$HHH$	$p^3$
$T$	$q$	$HT$	$pq$	$HHT$	$p^2q$
		$TH$	$pq$	$HTH$	$p^2q$
		$TT$	$q^2$	$HTT$	$pq^2$
				$THH$	$p^2q$
				$THT$	$pq^2$
				$TTH$	$pq^2$
				$TTT$	$q^3$

What is the probability that we throw a given number (0, 1, 2, or 3) of heads? We have to group together sequences with the same number of heads. Thus for 2 heads in 3 throws we get

$HHT$ ,  $HTH$ ,  $THH$ , each with probability  $p^2q$ , which gives a total probability of  $3p^2q$ . Similar calculations give another table:

		number of heads			
		0	1	2	3
number of throws	1	$q$	$p$		
	2	$q^2$	$2pq$	$p^2$	
	3	$q^3$	$3pq^2$	$3p^2q$	$p^3$

The rows of this table should look familiar: compare the expansions

$$(q+p)^1 = q+p$$

$$(q+p)^2 = q^2 + 2pq + p^2$$

$$(q+p)^3 = q^3 + 3pq^2 + 3p^2q + p^3$$

The terms on the right are exactly the entries in the table. The next row ought to come from

$$(q+p)^4 = q^4 + 4pq^3 + 6p^2q^2 + 4p^3q + p^4$$

and it is good practice to check that it does. In general, the entries in the  $n$ th row will be the terms of the expansion of  $(q+p)^n$ .

This is not a coincidence, and it is not difficult to explain it. To expand, say,  $(q+p)^5$  we must work out

$$(q+p)(q+p)(q+p)(q+p)(q+p)$$

The terms with exactly 3  $q$ s come from products like this:

$q$	$q$	$q$	$p$	$p$
$q$	$q$	$p$	$q$	$p$
$q$	$q$	$p$	$p$	$q$
$q$	$p$	$q$	$q$	$p$
$q$	$p$	$q$	$p$	$q$
$p$	$q$	$q$	$q$	$p$
$p$	$q$	$q$	$p$	$q$
$p$	$q$	$p$	$q$	$q$
$p$	$p$	$q$	$q$	$q$

These correspond exactly to the 10 possible sequences of 3 tails and 2 heads:

T T T H H H  
 T T T H T H  
 T T T H H T  
 T H T T H  
 T H T H T  
 T H H T T  
 H T T T H  
 H T T H T  
 H T H T T  
 H H T T T.

Obviously the same holds in general. If we write  $\binom{n}{r}$  for the number of sequences of  $n$   $H$ s and  $T$ s containing exactly  $r$   $H$ s and  $(n-r)$   $T$ s, then the probability of getting exactly  $r$  heads in  $n$  throws is

$$\binom{n}{r} p^r q^{n-r}.$$

It isn't too hard to work out what  $\binom{n}{r}$  is. If we choose the  $r$  positions for  $H$ s then everything is determined; so  $\binom{n}{r}$  is just the

number of ways of choosing  $r$  things from  $n$ . This, it can be shown, is given by

$$\binom{n}{r} = \frac{n(n-1)(n-2) \dots (n-r+1)}{r(r-1)(r-2) \dots 1}.$$

Thus for sequences of 2 heads and 3 tails we want

$$\binom{5}{2} = \frac{5 \cdot 4}{2 \cdot 1} = 10$$

which is correct.

The general expansion is

$$(q+p)^n = q^n + npq^{n-1} + \dots + \binom{n}{r} p^r q^{n-r} + \dots + p^n.$$

This is the *Binomial Theorem*, usually credited to Isaac Newton. It may or may not be coincidence that at one period Newton was Master of the Royal Mint.

The average number of heads obtainable in  $n$  throws can be calculated from this formula, and it turns out to be  $np$ . Thus the frequency with which heads occur is  $np/n$ , which is  $p$ . So we have come full circle to the idea of a probability as an 'average frequency of occurrence'. This theorem, which in a stronger form is called the *Law of Large Numbers*, shows how our mathematical model connects up with observations in the real world.

### Random Walks

In the final section of this chapter I want to discuss another type of problem arising in probability theory. It has applications to questions about electrons bouncing around inside crystals, and particles floating in a liquid.

Imagine a particle starting at position  $x = 0$  on the  $x$ -axis, at time  $t = 0$ . In time  $t = 1$  it moves to the point  $x = -1$  with probability  $1/2$ , or to the point  $x = +1$  with probability  $1/2$ . If it is in position  $x$  at time  $t$ , then at time  $t+1$  it moves either to  $x-1$  or  $x+1$ , each with probability  $1/2$ . What can we say about the particle's subsequent motion?

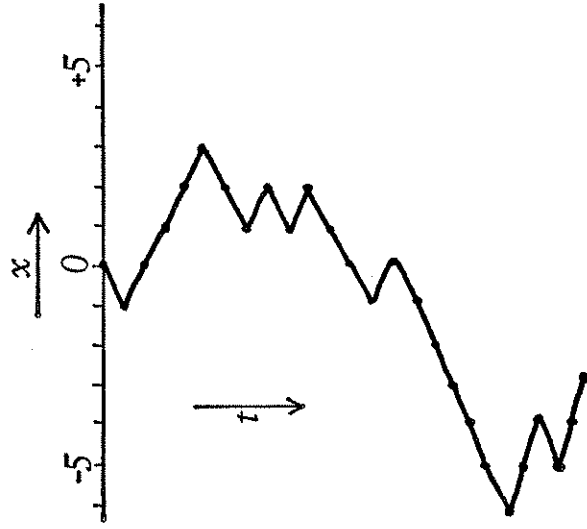


Figure 173

