

Performance Evaluation of Networks

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Part II – Queueing Theory

■ Part I – Markov chains

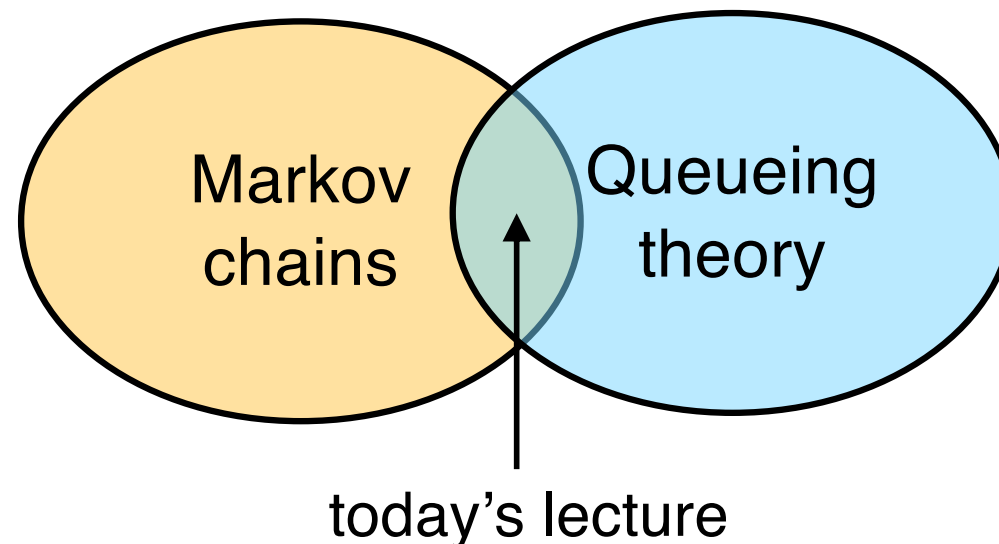
▶ Irreducible

- ◆ Discrete-time Markov chains (Chapter 1)
- ◆ Continuous-time Markov chains (Chapter 2)

▶ Absorbing

- ◆ Discrete-time and continuous-time (Chapter 3)

■ Part II – Queueing Theory



Part II – Queueing Theory

- What is a queue?
 - ▶ Supermarket, bank, postoffice, administrations, etc.
 - ▶ CPU, servers, clusters, etc.
 - ▶ Manufacturing, product lines, etc.
 - ▶ ...
- System with
 - ◆ at least one **service facility**
 - ◆ potentially a **waiting room** (finite or infinite)
 - ◆ customers
- Representation



Kendall's Notation

- Describe a queueing system

$$A / B / c / K$$

$A \rightarrow$ distribution of interarrivals

$B \rightarrow$ distribution of service times

$c \rightarrow$ number of servers

$K \rightarrow$ number of customers in system (omitted if infinite)

- Distributions often used

- ▶ Exponential $\rightarrow M$

- ▶ Deterministic $\rightarrow D$

- ▶ General $\rightarrow G$

- ▶ Erlang $\rightarrow E$

- ▶ Phase-type $\rightarrow PH$

What is Not in Kendall's Notation

- Service discipline (scheduling)
 - ▶ First-in-first-out (First-come-first-served)
 - ▶ Last-in-first-out (last-come-first-served)
 - ▶ Processor sharing
 - ▶ Random
 - ▶ Shortest-job-first
 - ▶ Shortest-processing-time-first
- Multiple waiting rooms / queues: Join discipline
 - ▶ Random
 - ▶ Join-shortest-queue
 - ▶ Best-out-of- d

$M / M / 1$ Queue

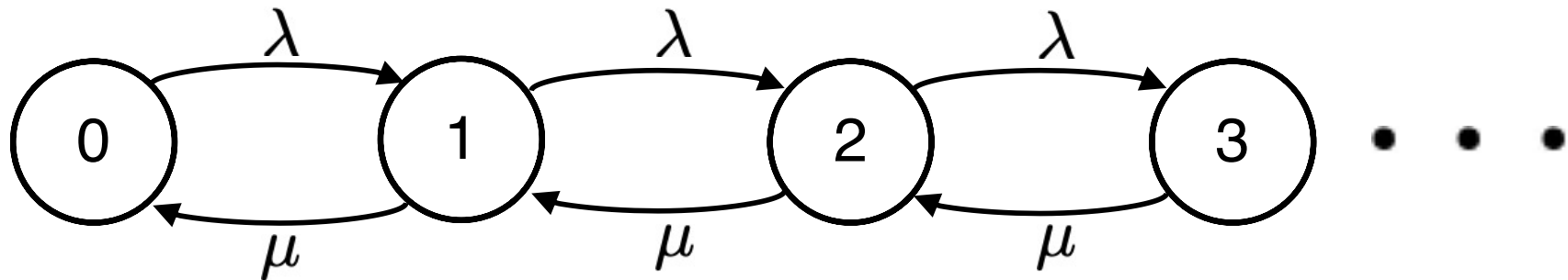
- Arrivals: Poisson rate $\lambda \rightarrow$ interarrival time $\text{Exp}(\lambda)$
- Service time $\text{Exp}(\mu)$
- Independence between arrivals and service times
- Service discipline \rightarrow not relevant (memoryless property)
- $X(t)$ number of customers in system at time t (queue size)
- **Proposition 12:** $\{X(t), t \geq 0\}$ is a birth-and-death process, birth rate λ , death rate μ
- Proof: construction rule #2

$$i \rightarrow i + 1 \quad \text{Exp}(\lambda)$$

$$i \rightarrow i - 1 \quad \text{Exp}(\mu) \quad i > 0$$

$M / M / 1$ Queue

- Transition diagram, $\mathcal{E} = \mathbb{N}$



- Global balance equations

$$\lambda \pi_{i-1} = \mu \pi_i, \quad i \geq 1$$

- System utilization $\rho = \frac{\lambda}{\mu}$

- Proposition 13: If $\rho < 1$ limiting/stationary distribution is

$$\pi_i = (1 - \rho) \rho^i, \quad i \geq 0$$

$M / M / 1$ Queue

■ Proof

$$\begin{aligned}\pi_i &= \rho \pi_{i-1} \\ &= \rho^i \pi_0\end{aligned}$$

■ Normalization

$$\sum_{i \geq 0} \pi_i = 1$$

$$\Leftrightarrow \pi_0 \sum_{i \geq 0} \rho^i = 1 \quad \text{sum of terms in geometric progression}$$

■ If $\rho < 1 \Rightarrow$

$$\pi_0 \frac{1}{1 - \rho} = 1$$

stability condition

$$\Rightarrow \pi_0 = 1 - \rho \quad (\rho \text{ system utilization})$$

$$\Rightarrow \pi_i = (1 - \rho) \rho^i, \quad i \geq 0$$

$M / M / 1$ Queue

- Let X stationary version of queue size $\rightarrow X \sim \text{Geom}(1 - \rho)$
 $1 - \rho$ probability to find system empty
 X number of « failed trials » before finding system empty

- Expected queue size

$$\begin{aligned} E[X] &= \sum_{i \geq 0} i \pi_i = (1 - \rho) \sum_{i \geq 0} i \rho^i \\ &= \rho(1 - \rho) \sum_{i \geq 1} i \rho^{i-1} = \rho(1 - \rho) \left(\sum_{i \geq 1} \rho^i \right)' \\ \Rightarrow \boxed{E[X] = \frac{\rho}{1 - \rho}} \end{aligned}$$

- Throughput (= rate of everything that goes **through** the system)

$$\text{Thpt} = \mu (1 - \pi_0) = \mu \rho = \lambda$$

Burke's Theorem

- Suppose $M / M / 1$ starts in steady-state
 - ▶ **Departure process** is Poisson with rate λ
 - ▶ Queue size at t independent of **departures before** time t
- Proof: use time-reversibility of $M / M / 1$
 - Forward chain** identical to **Backwards chain**
 - ✓ **Arrival process** is Poisson with rate λ
 - ✓ Queue size at t independent of **arrivals after** time t

$M / M / 1 / K$ Queue

- $M / M / 1$ finite waiting room \rightarrow at most K customers in system
- Queue size is a **finite** birth-and-death process

$$\pi_i = \rho^i \pi_0, \quad 0 \leq i \leq K$$

- Normalization
$$\sum_{i=0}^K \pi_i = 1 \Leftrightarrow \pi_0 \sum_{i=0}^K \rho^i = 1$$

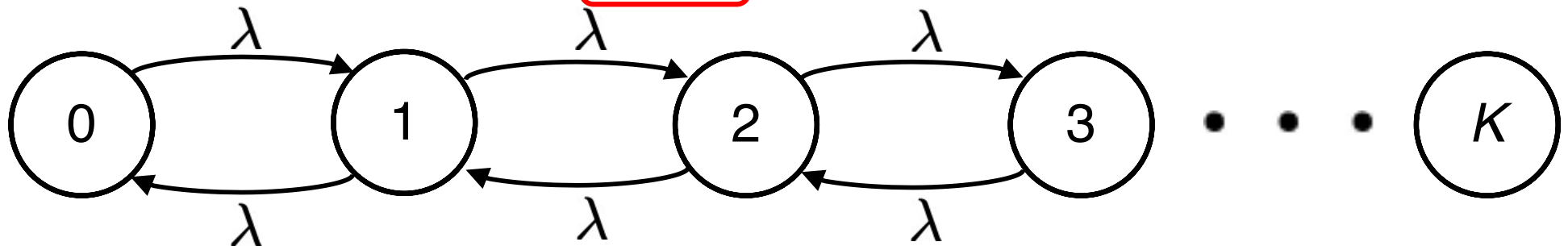
- If $\rho \neq 1 \Rightarrow \pi_0 \frac{1 - \rho^{K+1}}{1 - \rho} = 1 \Leftrightarrow \pi_0 = \frac{1 - \rho}{1 - \rho^{K+1}}$

$$\Rightarrow \pi_i = \frac{(1 - \rho) \rho^i}{1 - \rho^{K+1}}, \quad i = 0, 1, \dots, K$$

- If $\rho = 1 \Rightarrow \pi_0 = \frac{1}{K + 1} = \pi_i, \quad i = 0, 1, \dots, K$

$M / M / 1 / K$ Queue

- Transition diagram, $\rho = 1$ $\mathcal{E} = \{0, 1, \dots, K\}$



- We are equally likely to go left or right
- Stationary process X is **Uniform** between 0 and K

$$\pi_i = \frac{1}{K+1}, \quad i = 0, 1, \dots, K$$

- Expected queue size

$$E[X] = \sum_{i=0}^K i \pi_i = \frac{1}{K+1} \sum_{i=0}^K i = \frac{1}{K+1} \frac{K(K+1)}{2} = \frac{K}{2}$$

$M / M / 1 / K$ Queue

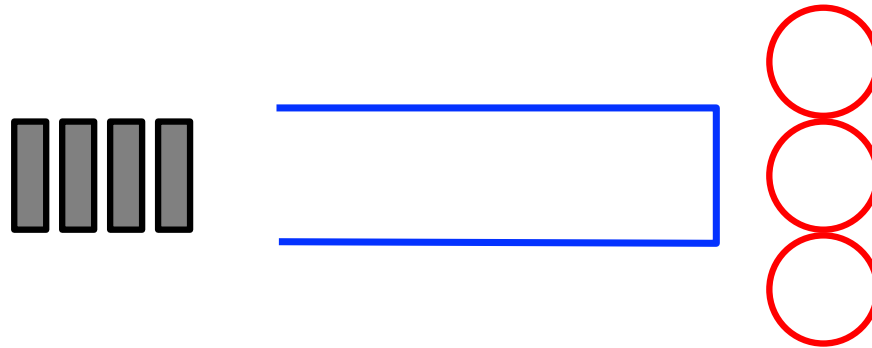
- Stationary distribution always exists!
no stability condition (finite system)
- Finite system \rightarrow customers may find system full!
loss probability ?
- **PASTA** : Poisson Arrivals See Time Averages
- Birth-and-death process is ergodic
 - ▶ Time averages = stationary distribution
- Loss probability = prob customer arrives and sees full queue
$$P_{\text{loss}} = \pi_K = \frac{(1 - \rho) \rho^K}{1 - \rho^{K+1}}$$
- Throughput $\text{Thpt} = \mu (1 - \pi_0) = \lambda (1 - \pi_K)$

$M / M / c$ Queue

- Poisson arrivals rate $\lambda \rightarrow$ interarrival time $\text{Exp}(\lambda)$
- Service time $\text{Exp}(\mu)$
- Multiple servers
- Infinite waiting room
- Queue size is a birth-and-death process
 - ▶ $i \rightarrow i + 1$ birth rate λ
 - ▶ $i \rightarrow i - 1$ death rate $\mu_i = i\mu, \quad i = 1, 2, \dots, c - 1$
 $= c\mu, \quad i \geq c$
- System utilization $\rho = \frac{\lambda}{c\mu}$
- Infinite system \rightarrow stability condition might be needed!

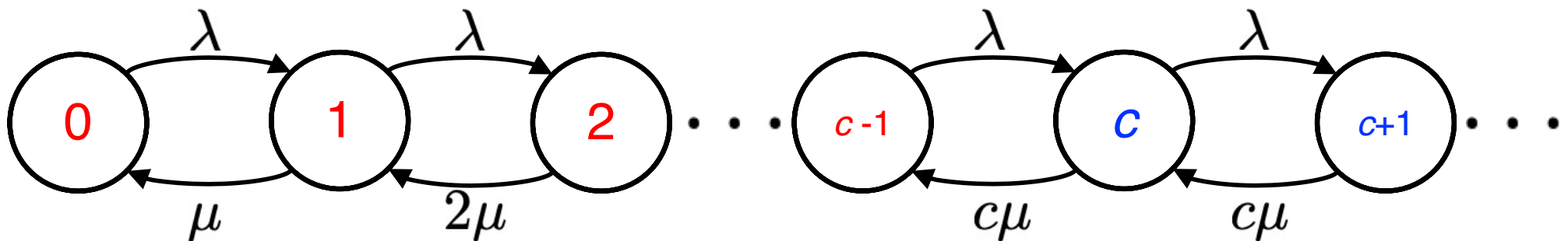
$M / M / c$ Queue

■ Representation



Example:
call center

■ Transition diagram, $\mathcal{E} = \mathbb{N}$



Customers served upon arrival

new customer needs to wait

$M / M / c$ Queue

- Stationary distribution: for $i = 1, 2, \dots$

$$\pi_i = \frac{\lambda_0 \lambda_1 \cdots \lambda_{i-1}}{\mu_1 \mu_2 \cdots \mu_i} \pi_0 = \begin{cases} \left(\frac{\lambda}{\mu}\right)^i \frac{1}{i!} \pi_0 & i = 0, 1, \dots, c \\ \left(\frac{\lambda}{\mu}\right)^i \frac{1}{c!} \frac{1}{c^{i-c}} \pi_0 & i \geq c \end{cases}$$

- Normalization: if $\rho < 1$ (stability condition)

$$\pi_0 = \left[\sum_{i=0}^{c-1} \left(\frac{\lambda}{\mu}\right)^i \frac{1}{i!} + \left(\frac{\lambda}{\mu}\right)^c \frac{1}{c!} \left(\frac{1}{1-\rho}\right) \right]^{-1}$$

- Probability of waiting (PASTA)

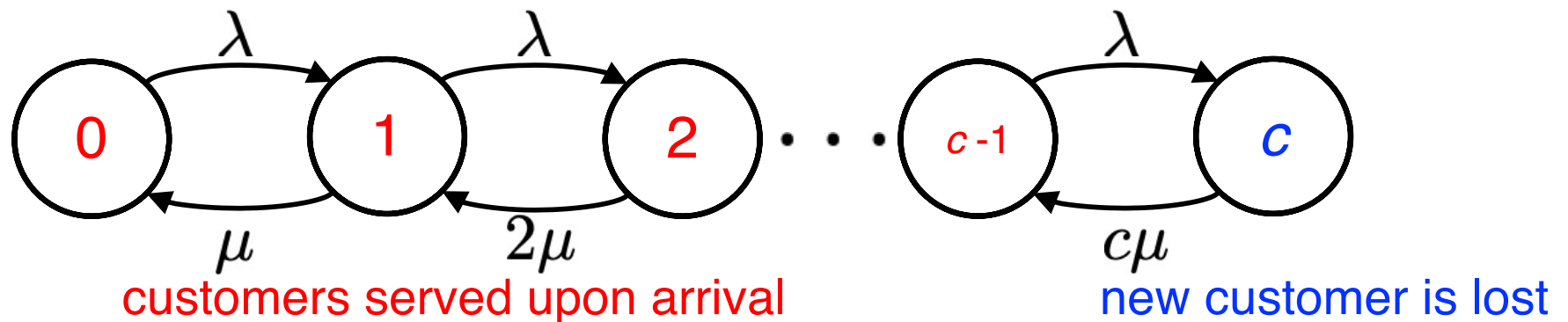
$$\rho = \frac{\lambda}{c\mu}$$

$$P_{\text{wait}} = \sum_{i \geq c} \pi_i = \frac{\pi_0 (\lambda/\mu)^c}{c!} \sum_{i \geq 0} \left(\frac{\lambda}{c\mu}\right)^i = \frac{\pi_0 (c\rho)^c}{c!(1-\rho)}$$

15 minutes break

$M / M / c / c$ Queue

- Multi-server queue with no waiting room
- Pure loss system (call center without music)
- Finite system \rightarrow no stability condition
- Queue size is a birth-and-death process
 - ▶ $i \rightarrow i + 1$ birth rate $\lambda_i = \lambda \quad i = 0, 1, \dots, c - 1$
 - ▶ $i \rightarrow i - 1$ death rate $\mu_i = i\mu, \quad i = 1, 2, \dots, c$
- Transition diagram $\mathcal{E} = \{0, 1, \dots, c\}$



$M / M / c / c$ Queue

- Stationary distribution

$$\pi_i = \pi_0 \left(\frac{\lambda}{\mu} \right)^i \frac{1}{i!} \quad i = 0, 1, \dots, c$$

$$\pi_0 = \left[\sum_{i=0}^c \left(\frac{\lambda}{\mu} \right)^i \frac{1}{i!} \right]^{-1}$$

- Loss probability

$$P_{\text{loss}} = \pi_c = \frac{\left(\frac{\lambda}{\mu} \right)^c \frac{1}{c!}}{\sum_{i=0}^c \left(\frac{\lambda}{\mu} \right)^i \frac{1}{i!}}$$

- Erlang's loss formula (1917)

- Major historical role in dimensioning phone systems

- Insensitivity property: holds for any service time distribution

Example: Repair Person Model

- K machines
- One repair person
- Each machine breaks after time $\text{Exp}(\alpha)$ (break rate is α)
- Upon a breakdown repair request is sent
- Repair person spends time $\text{Exp}(\mu)$ to repair one machine
- Questions
 - ▶ Probability that i machines are working normally
 - ▶ Overall failure rate?
- Define $X(t)$ number of functional machines at time t
- State space $\mathcal{E} = \{0, 1, \dots, K\}$

Example: Repair Person Model

- Possible transitions?

- ▶ $i \rightarrow i + 1$ repair is over, time $\text{Exp}(\mu)$

- ▶ $i \rightarrow i - 1$ break occurs, time $\text{Exp}(i\alpha)$

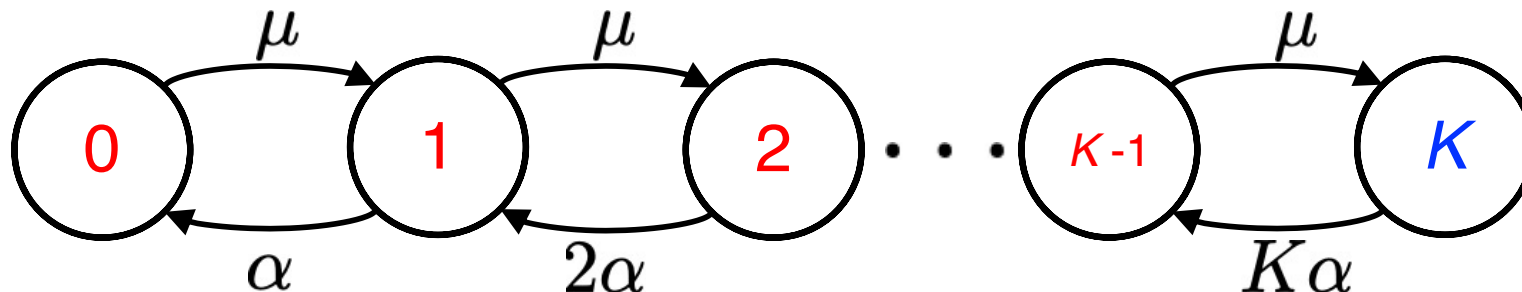
- Using construction rule #2, process is a (homogeneous) CTMC

- It is also a birth-and-death-process

- ▶ Birth rate (of functional machines) $\lambda_i = \mu$

- ▶ Death rate $\mu_i = i\alpha$

- Transition diagram



Example: Repair Person Model

- Number of functional machines

= Queue size of $M / M / K / K$

- Probability that i machines are working normally

$$\pi_i = \frac{\left(\frac{\mu}{\alpha}\right)^i \frac{1}{i!}}{\sum_{j=0}^K \left(\frac{\mu}{\alpha}\right)^j \frac{1}{j!}}$$

- Overall failure rate

$$\sum_{i=1}^K (i \alpha) \pi_i = \alpha E[X]$$

- Overall repair rate

$$\sum_{i=0}^{K-1} \mu \pi_i = \mu (1 - \pi_K)$$

- In steady-state overall failure rate = overall repair rate

Little's Formula

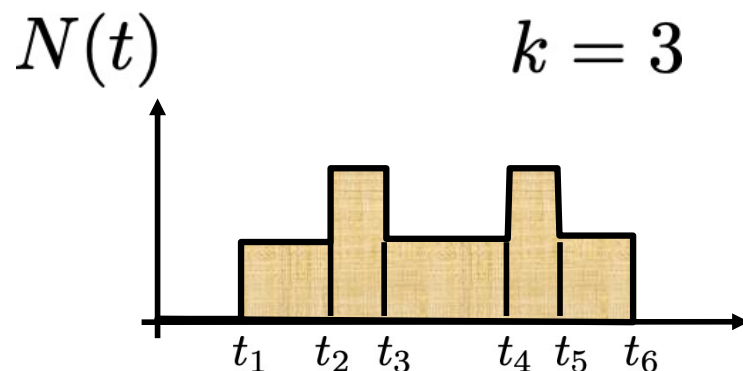
- Relates three quantities in **steady-state**
 - ▶ **Occupancy** in a system \bar{N}
 - ▶ **Entrance rate** to the system λ
 - ▶ **Sojourn** time in system \bar{T}

$$\bar{N} = \lambda \bar{T}$$

- Valid for **work-conserving** systems
 - system may not be idle if customers waiting
- No assumption on any distribution
- No assumption on service discipline
- Independence assumption

Proof of Little's Formula

- Steady-state \rightarrow system empties infinitely often
- Let 0 and C be two times when system is empty
- Let k be number of customers served in $(0, C)$
- $\{a_i\}_{i=1,\dots,k}$ arrival instants
- $\{d_i\}_{i=1,\dots,k}$ departure instants
- $\{t_n\}_{n=1,\dots,2k}$ all instants
- Number of customers over time $N(t)$

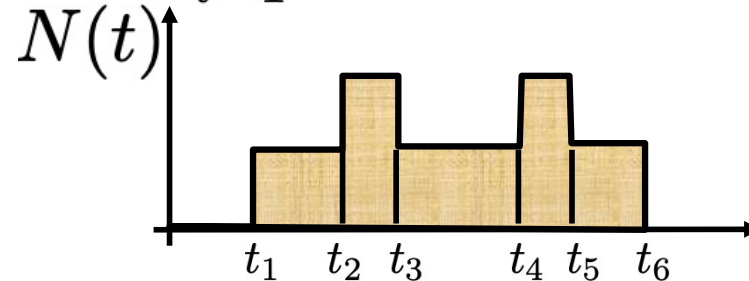
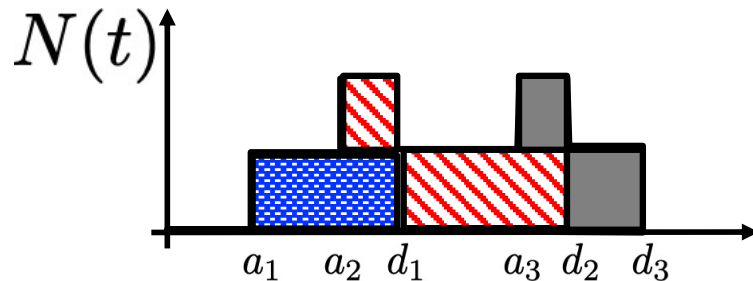


$$\begin{aligned}\bar{N} &= \frac{1}{C} \int_0^C N(t) dt \\ &= \frac{1}{C} \sum_{i=1}^{2k-1} N(t_i^+) (t_{i+1} - t_i)\end{aligned}$$

Proof of Little's Formula

- Mean sojourn time

$$\bar{T} = \frac{1}{k} \sum_{i=1}^k (d_i - a_i)$$



sum of horizontal boxes = sum of vertical boxes

$$\sum_{i=1}^k (d_i - a_i) = \sum_{i=1}^{2k-1} N(t_i^+) (t_{i+1} - t_i)$$

$$\Rightarrow \bar{T} k = \bar{N} C$$

- When $C \rightarrow \infty$, $\frac{k}{C} \rightarrow \lambda \Rightarrow \boxed{\bar{T} \lambda = \bar{N}}$

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- Consider $M / M / 1$, arrival rate λ , service rate μ
- If $\rho = \frac{\lambda}{\mu} < 1$ (queue is stable)
- Mean number of customers $\bar{N} = E[X] = \frac{\rho}{1 - \rho} > 0$
- Entrance rate = arrival rate (no losses)
- Expected sojourn time

By Little's formula $\bar{T} = \frac{\bar{N}}{\lambda} = \frac{\rho}{\lambda(1 - \rho)} \Rightarrow \boxed{\bar{T} = \frac{1}{\mu - \lambda}}$

- Expected waiting time

$$\bar{W} = \bar{T} - \frac{1}{\mu} = \frac{\rho}{\mu(1 - \rho)}$$

For next week

- Lesson 4 to revise
- Homework 4 to return on Tuesday 8 October before 9 am
- Lesson 5 to read before Lecture 5