rdlocrand: Local Randomization Methods for RD Designs*

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Abstract

The regression discontinuity (RD) design is a popular quasi-experimental design for causal inference and policy evaluation. Under the local randomization approach, RD designs can be interpreted as randomized experiments inside a window around the cutoff. The rdlocrand package provides tools to analyze RD designs under local randomization: rdrandinf to perform hypothesis testing using randomization inference, rdwinselect to select a window around the cutoff in which randomization is likely to hold, rdsensitivity to assess the sensitivity of the results to different window lengths and null hypotheses and rdrbounds to construct Rosenbaum bounds for sensitivity to unobserved confounders. We illustrate the implementation of these four functions, which have the same syntax and capabilities of the Stata commands described in Cattaneo, Titiunik, and Vazquez-Bare [2016]. For more details, and related Stata and R packages useful for analysis of RD designs, visit https://rdpackages.github.io/.

Keywords: regression discontinuity designs, quasi-experimental techniques, causal inference, randomization inference, finite-sample methods, Fisher's exact p-values, Neyman's repeated sampling approach.

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1 Introduction

This article illustrates the R package rdlocrand, which provides tools to analyze RD designs under a local randomization approach. The functions included in this package have the same syntax, and offer the same functionalities, as our companion Stata commands described in Cattaneo, Titiunik, and Vazquez-Bare [2016].

For brevity, we focus exclusively on software implementation issues. Extensive discussion and details on methodological and practical aspects can be found in Cattaneo, Frandsen, and Titiunik [2015], Cattaneo, Titiunik, and Vazquez-Bare [2016] and Cattaneo, Titiunik, and Vazquez-Bare [2017]. For help on the functions' syntax and related issues, please refer to reference manual.

For related Stata and R packages useful for analysis of RD designs, visit:

https://rdpackages.github.io/

2 Illustration of Methods

We illustrate how to implement the four functions described above using the dataset from Cattaneo et al. [2015]. This section replicates, as close as possible, section 7 of Cattaneo et al. [2016].

First, to install the rdlocrand package, type:

```
install.packages("rdlocrand")
    Next, load the data:
> data <- read.csv("rdlocrand_senate.csv")</pre>
> dim(data)
[1] 1390
> names(data)
                                                                                          "presdemvoteshlag1" "demmv"
 [1] "state"
                          "year"
                                               "dopen"
                                                                     "population"
[10] "demvoteshfor2"
                          "demwinprv1"
                                               "demwinprv2"
                                                                     "dmidterm"
                                                                                          "dpresdem"
 # Select predetermined covariates to be used for window selector
    <- cbind(data$presdemvoteshlag1,</pre>
              data$population/1000000,
              data$demvoteshlag1,
              data$demvoteshlag2,
              data$demwinprv1,
              data$demwinprv2,
              data$dopen,
              data$dmidterm,
              data$dpresdem)
 # Assign names to the covariates
```

The most basic syntax for rdwinselect is the following:

```
> tmp <- rdwinselect(R,X)
Mass points detected in running variable</pre>
```

Window selection for RD under local randomization

Number of obs = 1390
Order of poly = 0
Kernel type = uniform
Reps = 1000
Testing method = rdrandinf
Balance test = diffmeans

Cutoff $c = 0.000$	Left of c	Right of c
Number of obs	640	750
1st percentile	7	7
5th percentile	32	37
10th percentile	64	75
20th percentile	127	149

Obs>=c	Obs <c< th=""><th>Bin.test</th><th>Var. name</th><th>p-value</th><th>Window length / 2</th></c<>	Bin.test	Var. name	p-value	Window length / 2
16	10	0.327	DemSen.Vote.t.2	0.206	0.529
23	15	0.256	DemSen.Win.t.1	0.339	0.731
31	19	0.119	Open	0.034	1.080
35	25	0.245	Open	0.107	1.307
39	35	0.728	Midterm	0.106	1.539
44	38	0.581	Midterm	0.074	1.676
49	43	0.602	Midterm	0.040	1.904
54	48	0.621	Midterm	0.071	2.185
59	53	0.637	Open	0.136	2.366

Because in this particular application the cutoff is zero, which is the default value, the cutoff option can be omitted. For this reason, this and all the remaining examples will not specify this option. In practice, when the cutoff is not zero, the user can simply specify cutoff = c. Alternatively, it may be easier to simply redefine the running variable by recentering it at the cutoff. By default, rdwinselect uses the difference-in-means statistic to perform hypothesis tests—but this can be changed with the statistic option.

The output of rdwinselect is divided in three panels. The upper panel indicates the total sample size, the degree of the polynomial used by rdrandinf, the type of kernel used for the weighting scheme (uniform, triangular or epan), the number of replications in the permutation test (whenever this test is performed), the method used to perform the covariate balance tests (approximate or rdrandinf), the test statistic used (test, ksmirnov or ranksum).

The middle panel provides information on sample sizes. The first row gives the total number of observations to the left and to the right of the cutoff, and also the total sample size. The following four rows provide the same information but around small neighborhoods around the cutoffs defined by the first, fifth, tenth and twentieth percentile of the running variable.

Finally, the main panel gives the result of the two balance tests performed at each of the windows considered. The first column provides the window length of each window considered, divided by two. For example, a value of 0.529 in this column refers to the window $[\bar{r}-0.529\,;\,\bar{r}+0.529]$, where \bar{r} is the cutoff (equal to zero in this case) and the window length is $\bar{r}+0.529-(\bar{r}-0.529)=1.058$. The second column, labeled "p-value", provides the minimum p-value of the difference-in-means test, and the name of the corresponding variable associated to this p-value is given in column 3, "Var. name". The p-value is obtained by either permutation testing or a t-test, depending on the option specified. The fourth column gives the p-value from a Binomial probability test of the hypothesis that the probability of treatment is 0.5 using the binom.test command. Columns 5 and 6 give the number of observations to the left and right of the cutoff inside each window. As indicated in the last line of the output, the largest recommended window (the largest window for which the second column is equal to or above 0.15) in this case is [-0.731; 0.731] and contains 15 observations below the cutoff and 23 observations above.

Notice that rdwinselect gives the following message in the first line of the output:

Mass points detected in running variable

This message indicates that the running variable has repeated values, that is, multiple observations sharing the same value of the running variable. The presence of mass points in the running variable requires careful examination; see Cattaneo, Titiunik, and Vazquez-Bare [2017], Cattaneo, Idrobo, and Titiunik [2020] for details. In this application, the only mass points occur at the minimum and maximum of the running variable (which correspond to uncontested elections). Since

these observations are not used for estimation and inference, this warning can be ignored in this case.

By default, rdwinselect starts with a window that contains at least 10 observations at each side of the cutoff, and increases the length ensuring that at least two observations are added in each successive window. The user can choose these two values using the obsmin and obsstep options, respectively, or can define the windows in terms of their length instead of the number of observations. For instance, Cattaneo et al. [2015] start from the window [-0.5; 0.5] and increase the width by 0.125 using 10,000 replications in the permutation test. To replicate their results, we can type:

> tmp <- rdwinselect(R,X,wmin=.5,wstep=.125,reps=10000)
Mass points detected in running variable</pre>

Window selection for RD under local randomization

Number of obs = 1390
Order of poly = 0
Kernel type = uniform
Reps = 10000
Testing method = rdrandinf
Balance test = diffmeans

Cutoff c =	0.000	Left of c	Right of c
Number	of obs	640	750
1st per	centile	7	7
5th per	centile	32	37
10th per	centile	64	75
20th perc	centile	127	149

Window length / 2	p-value	Var. name	Bin.test	Obs <c< th=""><th>0bs>=c</th></c<>	0bs>=c
0.500	0.265	DemSen.Vote.t.2	0.230	9	16
0.625	0.427	Open	0.377	13	19
0.750	0.257	Open	0.200	15	24
0.875	0.150	Open	0.211	16	25
1.000	0.074	Open	0.135	17	28
1.125	0.038	Open	0.119	19	31
1.250	0.059	Open	0.105	21	34
1.375	0.137	Midterm	0.539	30	36
1.500	0.090	Midterm	0.640	34	39
1.625	0.112	Midterm	0.734	37	41

Recommended window is [-0.75;0.75] with 39 observations (15 below, 24 above).

We see that the command selects the window [-0.75; 0.75], as in Cattaneo et al. [2015]. How-

ever, it is important to note that these results can vary slightly because of the randomization process behind the selection procedure. Additionally, observe that the minimum p-value is not necessarily monotonic on the length of the window. The plot option allows the user to depict graphically how these values change for different lengths. We will set the number of windows to 80 to have more observations in the plot, and we will specify the approx option to speed up the calculations. By specifying this option, the command uses the large-sample approximation instead of randomization inference. It is useful for illustration purposes as it is much faster, but it can be misleading since the approximation may be poor when the sample is small. The output from rdwinselect with 80 windows is a long table and will be omitted. The resulting graph is shown in Figure 1.

> tmp <- rdwinselect(R,X,wmin=.5,wstep=.125,approx=TRUE,nwin=80,quietly=TRUE,plot=TRUE)
Mass points detected in running variable</pre>

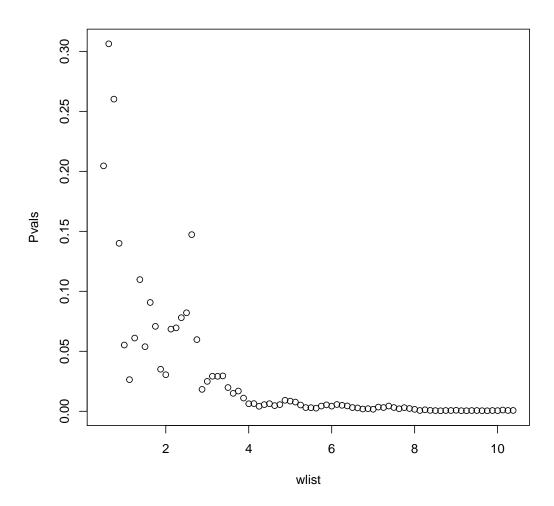


Figure 1. Plot of p-values.

The figure shows that the p-values vary widely for very short windows, but the sequence stabilizes once the window length is large enough (around the value 3 in this case).

Once the window has been selected, randomization inference to test the sharp null hypothesis of no treatment effect can be performed using rdrandinf. For example, take the window [-0.75; 0.75], which is the one selected by Cattaneo et al. [2015] and replicated above. The basic syntax for rdrandinf is:

```
> tmp <- rdrandinf(Y,R,wl=-.75,wr=.75)</pre>
Selected window = [-0.75; 0.75]
Running randomization-based test...
Randomization-based test complete.
Number of obs
                              1297
Order of poly
                                 0
Kernel type
                           uniform
                              1000
Reps
Window
                       set by user
HO:
             tau
                  = fixed margins
Randomization
Cutoff c =
              0.000
                       Left of c Right of c
      Number of obs
                             595
                                          702
Eff. number of obs
                              15
    Mean of outcome
                          42.808
                                       52.497
    S.d. of outcome
                           7.042
                                       7.742
```

Window

-0.750

		Finite sample	Large sample		
Statistic	T	P> T	P> T	Power vs d =	3.521
Diff. in means	9.689	0.000	0.000		0.300

0.750

22

Like rdwinselect, the output of rdrandinf is divided in three panels. The upper panel gives the total sample size, the order of the polynomial, the type of kernel used for the weighting scheme (uniform, triangular or epan), the number of replications in the randomization test, and whether the window was specified by the user by setting wl and wr or calculated using rdwinselect as will be illustrated shortly. The middle panel provides the number of observations at each side of the cutoff, sample size below and above the cutoff inside the specified window, some descriptive statistics for the outcome inside the window, and the selected window. Note that the first line in this panel displays the number of observations with non-missing values of the outcome and running variable, so the sample sizes shown can differ from the total sample size.

Finally, the main panel gives the results from the randomization test. The first column, labeled "Statistic", indicates the statistic used in the randomization test. The second column gives the observed value of the selected statistic and the third column shows its finite-sample p-value obtained from the randomization test. The fourth column gives the asymptotic p-value, that is, the p-value obtained from the corresponding asymptotic distribution of the chosen statistic. Finally the fifth column gives the asymptotic power against an alternative value that can be specified using the options d() or dscale(). The default is dscale(.5), that is, an effect size equal to half the standard deviation of the outcome for the control group inside the window (the critical value for the power calculation is set to 1.96).

As mentioned above, rdrandinf uses the difference in means as the default statistic, but it can also use the Kolmogorov-Smirnov and the rank sum statistics. By adding statistic(all) as an option we can obtain the result for all three statistics. The output is:

```
> tmp <- rdrandinf(Y,R,wl=-.75,wr=.75,statistic='all')</pre>
```

Selected window = [-0.75; 0.75]

Running randomization-based test...

Randomization-based test complete.

 Number of obs
 =
 1297

 Order of poly
 =
 0

 Kernel type
 =
 uniform

 Reps
 =
 1000

 Window
 =
 set by user

 H0:
 tau
 =
 0

 Randomization
 =
 fixed margins

Cutoff c =	0.000	Left of	С	Right	of c
Number o	of obs	59	95		702
Eff. number o	of obs	1	.5		22
Mean of ou	tcome	42.80	8	52	2.497
S.d. of ou	tcome	7.04	2	7	7.742
V	lindow	-0.75	0	C	750

	_	Finite sample	ple Large sample		
Statistic	Т	P> T	P> T	Power vs d =	3.521
Diff. in means	9.689	0.000	0.000		0.300
Kolmogorov-Smirnov	0.552	0.008	0.005		NA

Rank sum z-stat -3.217 0.000 0.001 0.209

We can see that the three statistics provide basically the same result in terms of inference; the randomization test rejects the sharp null of no treatment effect at one percent significance level in all three cases. Also note that the rdrandinf command does not provide the asymptotic power for the Kolmogorov-Smirnov statistic.

The window in which to perform the randomization-based tests can be set manually using wl and wr. This options specify the lower and upper limits of the chosen window. Importantly, these are window limits and not lengths, so for instance, if the cutoff is 100 and the user wants a window of ± 5 , the correct syntax is wl = 95 and wr = 105. We advise the user to always normalize the cutoff to zero by centering the running variable to avoid confusion.

Alternatively, the user can specify the list of covariates to have rdrandinf select the window automatically using rdwinselect. All the options allowed in rdwinselect can be passed through rdrandinf. For example:

```
> tmp <- rdrandinf(Y,R,statistic='all',covariates=X,wmin=.5,wstep=.125,rdwreps=10000)</pre>
Running rdwinselect...
Mass points detected in running variable
rdwinselect complete.
Selected window = [-0.75; 0.75]
Running randomization-based test...
Randomization-based test complete.
Number of obs
                              1297
Order of poly
Kernel type
                          uniform
Reps
                              1000
Window
                      rdwinselect
HO:
             tau =
Randomization
                  = fixed margins
Cutoff c =
              0.000
                      Left of c Right of c
                            595
                                         702
      Number of obs
Eff. number of obs
                             15
                                          22
    Mean of outcome
                         42.808
                                      52.497
    S.d. of outcome
                                       7.742
                          7.042
             Window
                         -0.750
                                       0.750
```

		Finite sample	Large sample		
Statistic	- Т	P> T	P> T	Power vs d =	3.521
Diff. in means	9.689	0.000	0.000		0.300
Kolmogorov-Smirnov	0.552	0.006	0.005		NA
Rank sum z-stat	-3.217	0.000	0.001		0.209

The rdrandinf command allows the user to specify a polynomial transformation model for the outcomes using the option p. By default, the command sets p=0, which means no transformation. When the p is set to an integer larger than zero, the slopes (and possibly higher order terms) are subtracted from the outcomes, leaving a residualized version of the outcome that only differs above and below the cutoff in the intercept. For instance, to perform a linear transformation, the syntax is:

```
> tmp <- rdrandinf(Y,R,wl=-.75,wr=.75,statistic='all',p=1)</pre>
```

Selected window = [-0.75; 0.75]

Running randomization-based test...

Randomization-based test complete.

Number of obs = 1297

Order of poly = 1

Kernel type = uniform

Reps = 1000

Window = set by user

HO: tau = 0

Randomization = fixed margins

0.000 Cutoff c = Left of c Right of c 595 702 Number of obs Eff. number of obs 15 22 Mean of outcome 42.808 52.497 S.d. of outcome 7.042 7.742 Window -0.750 0.750

		Finite sample		Large sample	
Statistic	T	P> T	P> T	Power vs d =	3.521
Diff. in means	15.297	0.000	0.066		0.071
Kolmogorov-Smirnov	0.797	0.000	NA		NA
Rank sum z-stat	-4.455	0.000	NA		NA

When a model for the outcomes is specified—that is, when p is set to a number greater than zero—with the option statistic="ttest" fits a regression of the outcome on the treatment dummy interacted with a polynomial of the running variable, and uses the difference in intercepts as the test-statistic. The other test-statistics use as outcomes the residuals described above. Note that the command does not provide the asymptotic p-value nor the asymptotic power of the Kolmogorov-Smirnov and rank sum statistics, as the asymptotic distribution does not account for the model transformation and hence can be misleading.

In the presence of arbitrary interference, a confidence interval for a particular measure of the effects of the program can be obtained with the interfci option. For example, to obtain a 95 percent confidence interval, we type:

```
> tmp <- rdrandinf(Y,R,wl=-.75,wr=.75,interfci=.05)</pre>
Selected window = [-0.75; 0.75]
Running randomization-based test...
Randomization-based test complete.
                              1297
Number of obs
Order of poly
Kernel type
                           uniform
Reps
                              1000
Window
                      set by user
HO:
             tau =
Randomization
                  = fixed margins
Cutoff c =
              0.000
                      Left of c Right of c
                             595
                                         702
      Number of obs
Eff. number of obs
                                          22
                              15
    Mean of outcome
                         42.808
                                      52.497
    S.d. of outcome
                          7.042
                                       7.742
             Window
                         -0.750
                                       0.750
```

		Finite sample	Large sample		
Statistic	T	P> T	P> T	Power vs d =	3.521
Diff. in means	9.689	0.000	0.000		0.300

95% confidence interval under interference: [3.795;15.824]

In terms of interpretation, it is important to keep in mind that the confidence interval under interference is not a confidence interval for the point estimate (and in fact, it may even not contain the point estimate). The interference confidence interval is constructed based on the difference between the observed statistic and the statistic that would be observed if the treatment was withheld from all units. In our application, allowing for arbitrary interference we can say with 95 percent confidence that the "excess" benefit of the treated group compared to the control group is between 3.795 and 15.824. Again, in this particular example the point estimate under SUTVA happens to be contained in the confidence interval under interference, but this need not be the case and has no clear interpretation.

The rdlocrand package provides two types of sensitivity analyses to assess how p-values change with window length. The first one, rdsensitivity, calculates and plots a matrix of p-values over a range of values for the treatment effect under the null hypothesis (rows) and window lengths (columns). For instance, we can see how the p-values change by starting from the selected window, increasing the window length by 0.25 and over a range of treatment effects that is roughly the point estimate plus and minus 10:

```
> tmp <- rdsensitivity(Y,R,wlist=seq(.75,10,by=.25),tlist=seq(0,20,by=1))
Running sensitivity analysis...
Sensitivity analysis complete.</pre>
```

Note that the rdsensitivity command does not display any output, but its results can be called using tmp\$results. In addition to the p-values, the rdsensitivity command returns the plot shown in Figure 2. The plot depicts the grid of window lengths in the horizontal axis and the grid of treatment effects under the null. The color represents the p-value for each pair of window length and treatment effect, where white corresponds to zero and black corresponds to one. This is simply a graphical display of the results given by rdsensitivity. The plot can be replicated (or modified) with the following code:

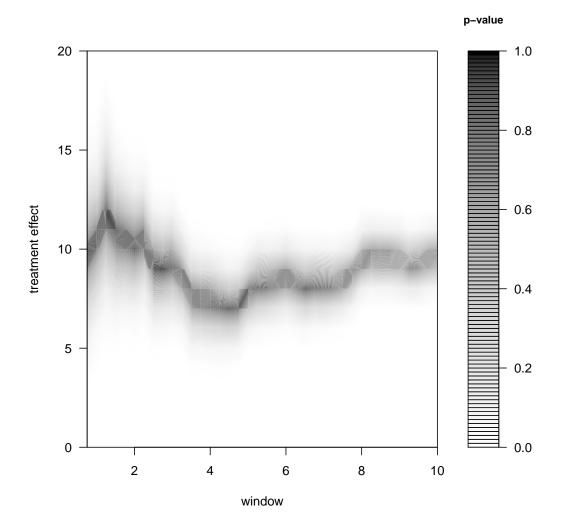


Figure 2. Sensitivity analysis.

One way to interpret these results is to see the range of values for which the p-value is above, say, .05, as a 95 percent confidence interval for the point estimate (assuming a constant additive treatment effect). Thus, the above table shows how the confidence interval for the treatment effect changes as the window length increases. For instance, the 95 percent confidence interval for the window [-.75; .75] is roughly [5; 14], whereas for the window [-2; 2] it becomes [7; 14]. In this case, the point estimate seems to be relatively stable over the range of windows considered. The confidence interval for the window [-.75; .75] can be obtained using the ci option:

```
> tmp <- rdsensitivity(Y,R,wlist=seq(.75,2,by=.25),tlist=seq(0,20,by=1),ci=0.75)
Running sensitivity analysis...
Sensitivity analysis complete.
> tmp$ci
```

[1] 5 14

Additionally, rdsensitivity can be called from within rdrandinf to obtain confidence intervals for the point estimates obtained using the ci option. The syntax is the following:

```
> tmp <- rdrandinf(Y,R,wl=-.75,wr=.75,ci=c(.05,seq(3,20,by=1)))</pre>
Selected window = [-0.75; 0.75]
Running randomization-based test...
Randomization-based test complete.
Running sensitivity analysis...
Sensitivity analysis complete.
Number of obs
                              1297
Order of poly
Kernel type
                           uniform
Reps
                              1000
Window
                      set by user
             tau =
Randomization
                  = fixed margins
Cutoff c =
              0.000
                      Left of c Right of c
      Number of obs
                            595
                                         702
Eff. number of obs
                             15
                                          22
   Mean of outcome
                         42.808
                                      52.497
    S.d. of outcome
                                       7.742
                          7.042
             Window
                         -0.750
                                       0.750
                                   Finite sample
                                                             Large sample
          Statistic
                                       P>|T|
                                                     P>|T|
                                                              Power vs d =
                                                                             3.521
```

0.000

95% confidence interval: [5,14]

9.689

Diff. in means

The second type of sensitivity analysis is performed with rdrbounds, which calculates Rosenbaum bounds [Rosenbaum, 2002]. As explained above, this command calculates upper and lower bounds for the randomization p-value under Bernoulli trials for a range of values of a parameter $\Gamma \equiv \exp(\gamma)$ that captures the strength with which an unobservable binary variable U_i affects the probability of selection into treatment. The basic syntax is:

0.000

0.300

```
> tmp <- rdrbounds(Y,R,expgamma=c(1.5,2,3),wlist=c(.5,.75,1),reps=1000)</pre>
Calculating randomization p-value...
Bernoulli p-value (w = 0.5) = 0.01
Bernoulli p-value (w = 0.75) = 0
Bernoulli p-value (w = 1) = 0
Running sensitivity analysis...
Sensitivity analysis complete.
> tmp$lower.bound
            [,1] [,2] [,3]
[1,] 0.006000000 0.000
[2,] 0.002000000 0.001
[3,] 0.002004008 0.000
> tmp$upper.bound
      [,1] [,2] [,3]
[1,] 0.027 0.013 0.003
[2,] 0.088 0.058 0.016
[3,] 0.262 0.280 0.162
```

The output from rdrbounds is divided in several parts. The first one shows the randomization p-value based on Bernoulli trials for each window. The matrices lower.bound and lower.bound present the lower and upper bounds for the p-values for different values of γ and windows. The wider the distance between the lower and upper bounds, the more sensitive the inference to deviations from a randomized experiment. The remaining elements give the list of values used to calculate the bounds.

The fmpval option adds the fixed margins randomization p-value to the first panel of the output. This allows the user to compare the p-values obtained using each method.

```
[1,] 0.006000000 0.000 0
[2,] 0.002000000 0.001 0
[3,] 0.002004008 0.000 0
> tmp$upper.bound
       [,1] [,2] [,3]
[1,] 0.027 0.013 0.003
[2,] 0.088 0.058 0.016
[3,] 0.262 0.280 0.162
```

We can see that the p-values obtained by both methods are very similar, which we found to be usually true in applications as long as the number of replications is large enough.

Finally, when using outcome transformation, the rdrandinf command allows the user to choose in which point to evaluate the transformed outcomes. By default, the evaluation point is the cutoff, which emulates the idea used in the local polynomial approach of estimating the effect at the cutoff. However, whenever the local randomization assumption is plausible, the cutoff need not be the point of interest. For instance, to set the evaluation points at the means of the running variable inside the window below and above the cutoff, we can type:

```
> ii <- (R>=-.75) & (R<=.75) & !is.na(Y) & !is.na(R)</pre>
> m0 <- mean(R[ii & D==0],na.rm=TRUE)</pre>
> m1 <- mean(R[ii & D==1],na.rm=TRUE)</pre>
> tmp <- rdrandinf(Y,R,wl=-.75,wr=.75,p=1,evall=m0,evalr=m1)</pre>
Selected window = [-0.75; 0.75]
Running randomization-based test...
Randomization-based test complete.
Number of obs
                              1297
Order of poly
Kernel type
                           uniform
                              1000
Reps
Window
                       set by user
             tau =
Randomization
                   = fixed margins
Cutoff c =
              0.000
                       Left of c Right of c
      Number of obs
                             595
                                          702
Eff. number of obs
                                           22
                              15
    Mean of outcome
                          42.808
                                       52.497
    S.d. of outcome
                          7.042
                                        7.742
             Window
                          -0.750
                                        0.750
```

Finite sample

Large sample

Statistic	Т	P> T	P> T	Power vs d =	3.521
Diff. in means	9.689	0.000	0.000		0.283

The user can verify that the point estimate in this case is the same as when no transformation is used, which is due to the fact that the transformation comes from a linear regression which by construction passes through the means. The p-values, however, can differ. Incidentally, note that the means are taken over the sample inside the window with non-missing values for the outcome and the running variable. The reason is that rdrandinf drops the observations inside the window with missing outcomes of running variable. Similarly, rdwinselect drops, at each evaluated window, the observations with missing values of the covariates and running variable.

3 Auxiliary Functions

This section introduces the auxiliary functions used by the four functions described above. All the auxiliary functions are included in the file rdlocrand_fun. We do not recommend to use these functions directly, as they typically do not contain error handling and incorrectly using them may lead to unexpected mistakes.

3.1 Statistics for randomization inference

```
rdrandinf.model(Y,D,statistic,pvalue=FALSE,kweights,endogtr,delta="")
```

This function uses the outcome and treatment variable $D = \mathbf{1}(R \ge c)$ to calculate the observed statistics, asymptotic p-values and asymptotic power for all the statistics used by rdrandinf, namely, "ttest", "ksmirnov", "ranksum", "all", "ar" and "wald".

3.2 Hotelling's T^2 statistic

```
hotelT2(X,D)
```

This function computes Hotelling's T^2 statistic and its asymptotic p-value for the matrix of covariates X for rdwinselect.

3.3 Default window length

wlength(R,D,num)

This function finds the window containing at least num observations at each side of the cutoff. It is used to find the default initial window in rdwinselect.

3.4 Default window increment

findstep(R,D,obsmin,obsstep,times)

This function finds a list of increments of length times, starting from window with obsmin observations and adding obsstep observations at each side in each step. It is used to find the default window increments in rdwinselect.

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