

# Optiver Prove It: Episode 3

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The video containing the problem statement is available here: [https://www.youtube.com/watch?v=YKE5SRXK\\_nI](https://www.youtube.com/watch?v=YKE5SRXK_nI).

## 1. Expected number of rolls with an $n$ -sided die

Let  $n \geq 1$  denote the number of sides of the die, and let  $X_1, X_2, \dots$  be i.i.d. random variables uniformly distributed over  $\{1, \dots, n\}$ . For each integer  $p \geq 1$ , let  $S_p = \sum_{k=1}^p X_k$  and define the random variable  $N$  as

$$N = \inf\{p \geq 1 : S_p \geq n\}.$$

### 1.1. A first-step analysis

For each  $i \in \{1, \dots, n\}$ , let  $e_i$  denote the expected number of rolls needed to reach a sum  $\geq i$ . Clearly,  $e_1 = 1$  and we are looking for the value of  $e_n$ .

If on the first throw we obtain a value  $j < i$ , then the game starts over on the next turn, except that the target to be surpassed is now  $i - j$ . For convenience, let  $N_i = \inf\{p \geq 1 : S_p \geq i\}$ . Conditioning on the first throw, we obtain therefore

$$\begin{aligned} e_i &= \mathbb{E}[N_i] = \sum_{j=1}^{i-1} \mathbb{E}[N_i | X_1 = j] \mathbb{P}(X_1 = j) + \sum_{j=i}^n \mathbb{E}[N_i | X_1 = j] \mathbb{P}(X_1 = j) \\ &= \sum_{j=1}^{i-1} (1 + e_{i-j}) \frac{1}{n} + \sum_{j=i}^n 1 \cdot \frac{1}{n} \\ &= 1 + \frac{1}{n} \sum_{j=1}^{i-1} e_{i-j}. \end{aligned}$$

This recursion makes it possible, for a given value of  $n$ , to numerically compute  $e_2, e_3, \dots$  up to  $e_n$ . However, obtaining a closed form for  $e_n$  seems tedious and we will not proceed further with this approach.

### 1.2. A combinatorial approach

Note that the random variable  $N$  takes values in  $\{1, \dots, n\}$  and

$$\mathbb{E}[N] = \sum_{p=0}^{\infty} \mathbb{P}(N > p) = 1 + \sum_{p=1}^{n-1} \mathbb{P}(N > p) = 1 + \sum_{p=1}^{n-1} \mathbb{P}(S_p < n).$$

Next,

$$\begin{aligned}
\mathbb{P}(S_p < n) &= \mathbb{P}\left((X_1, \dots, X_p) \in \left\{(x_1, \dots, x_p) \in \{1, \dots, n\}^p : \sum_{k=1}^p x_k < n\right\}\right) \\
&\stackrel{(i)}{=} \frac{\text{card}\{(x_1, \dots, x_p) \in \{1, \dots, n\}^p : \sum_{k=1}^p x_k < n\}}{\text{card}(\{1, \dots, n\}^p)} \\
&= \frac{1}{n^p} \sum_{s=p}^{n-1} \text{card}\left\{(x_1, \dots, x_p) \in \{1, \dots, n\}^p : \sum_{k=1}^p x_k = s\right\} \\
&\stackrel{(ii)}{=} \frac{1}{n^p} \sum_{s=p}^{n-1} \binom{s-1}{p-1} \\
&\stackrel{(iii)}{=} \frac{1}{n^p} \binom{n-1}{p},
\end{aligned} \tag{1}$$

where (i) holds because the random vector  $(X_1, \dots, X_p)$  is uniformly distributed over the set  $\{1, \dots, n\}^p$ , (ii) follows from stars and bars, and (iii) from the hockeystick identity.

Thus

$$\mathbb{E}[N] = 1 + \sum_{p=1}^{n-1} \frac{1}{n^p} \binom{n-1}{p} = \sum_{p=0}^{n-1} \binom{n-1}{p} \frac{1}{n^p} = \left(1 + \frac{1}{n}\right)^{n-1}.$$

## 2. Bonus exercise

### 2.1. Theoretical solution

With the previous notation,  $n$  is equal to 6 and the first player wins if and only if the random variable  $N$  takes an odd value. Therefore

$$\begin{aligned}
\mathbb{P}(\text{First player wins}) &= \sum_{p=0}^2 \mathbb{P}(N = 2p + 1) \\
&= \sum_{p=0}^2 \mathbb{P}(N > 2p) - \mathbb{P}(N > 2p + 1) \\
&= \sum_{p=0}^2 \mathbb{P}(S_{2p} < 6) - \mathbb{P}(S_{2p+1} < 6) \\
&= \sum_{p=0}^2 \frac{1}{6^{2p}} \binom{5}{2p} - \frac{1}{6^{2p+1}} \binom{5}{2p+1} \quad \text{by (1)} \\
&= \sum_{p=0}^2 \frac{1}{6^{2p}} \left( \binom{5}{2p} - \frac{1}{6} \binom{5}{2p+1} \right) \\
&= \frac{3125}{7776} \\
&\approx 0.402.
\end{aligned}$$

Consequently  $\mathbb{P}(\text{Second player wins}) = \frac{4651}{7776} \approx 0.598$  and you should play second.

## 2.2. Numerical simulation

The following piece of code repeats the experiment 20 000 times. The output is 0.401 which confirms that our computations are correct.

```
import numpy as np

n = 6
n_reps = 20000
np.random.seed(2024)

X = np.random.randint(1, n+1, (n_reps, n))
X = np.cumsum(X, axis=1) >= 6
X = np.argmax(X, axis=1) - 1 #subtract 1 to convert zero-based into one-based indexing
print(np.mean(X % 2)) #proportion of wins by the first player
```