

CFRM 541 Homework 3

Greg Damico

October 28, 2016

Problem #1

```
library(mpo)
```

```
## Loading required package: boot
```

```
## Loading required package: xts
```

```
## Loading required package: zoo
```

```
##
```

```
## Attaching package: 'zoo'
```

```
## The following objects are masked from 'package:base':
```

```
##
```

```
##      as.Date, as.Date.numeric
```

```
## Loading required package: lattice
```

```
##
```

```
## Attaching package: 'lattice'
```

```
## The following object is masked from 'package:boot':
```

```
##
```

```
##      melanoma
```

```
## Loading required package: mvtnorm
```

```
## Loading required package: rrcov
```

```
## Loading required package: robustbase
```

```
##
```

```
## Attaching package: 'robustbase'
```

```
## The following object is masked from 'package:boot':
```

```
##
```

```
##      salinity
```

```
## Scalable Robust Estimators with High Breakdown Point (version 1.4-3)
```

```
## Loading required package: robust
```

```

## Loading required package: fit.models

## Loading required package: MASS

## Loading required package: PortfolioAnalytics

## Loading required package: foreach

## Loading required package: PerformanceAnalytics

##
## Attaching package: 'PerformanceAnalytics'

## The following object is masked from 'package:graphics':
##
##     legend

## Loading required package: corpcor

## Loading required package: shiny

##
## Attaching package: 'mpo'

## The following object is masked from 'package:shiny':
##
##     runExample

## The following objects are masked from 'package:PerformanceAnalytics':
##
##     SFM.beta, StdDev.annualized

library(sn)

## Loading required package: stats4

##
## Attaching package: 'sn'

## The following object is masked from 'package:stats':
##
##     sd

library(quantmod)

## Loading required package: TTR

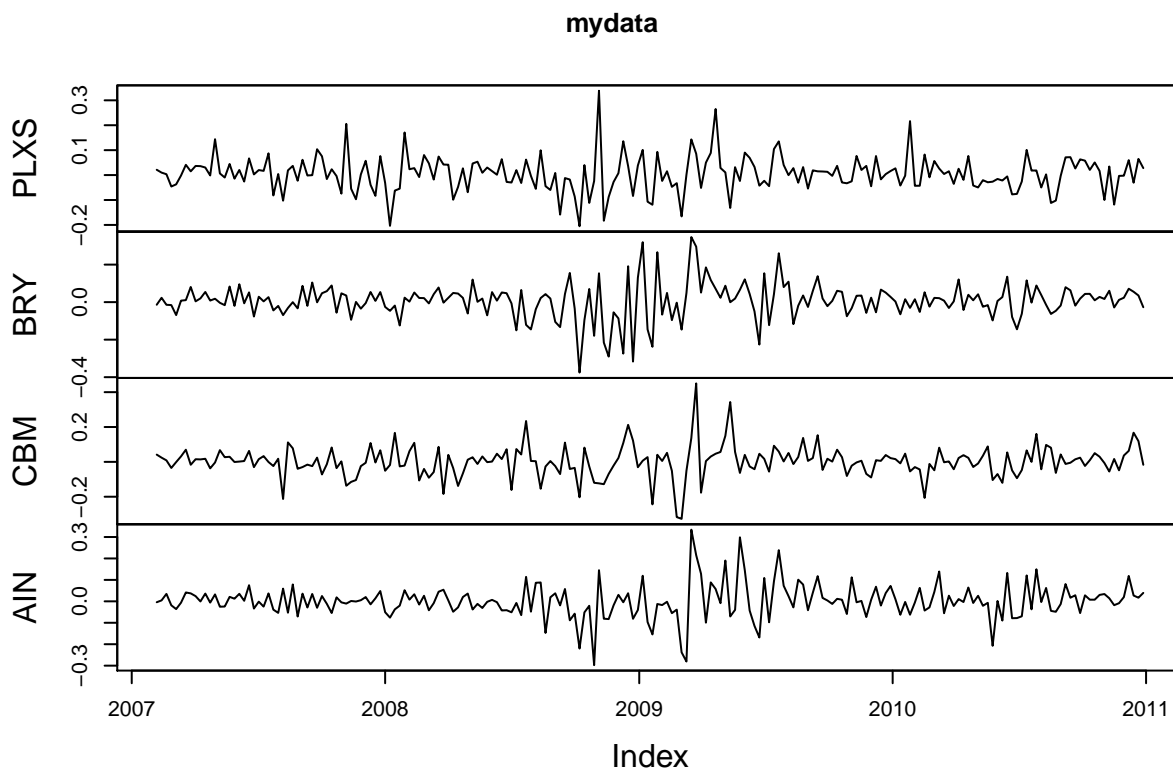
## Version 0.4-0 included new data defaults. See ?getSymbols.

```

```
library(TTR)
library(tseries)
library(norlmix)
```

*# There seems to be something funny going on here. I'm trying to call
 # "chart.QQPlot.norMix" from the mpo package, but when I ran the code below I would get an
 # error telling me that the compiler could not find the "chart.QQPlot.norMixEM" function.
 # A student on Piazza helpfully pointed out that this other function was in the norlmix
 # package. Hence the call above to load the norlmix package. But it still seems to me that
 # what I really want is not the EM version. Why the code below has been being taken to call
 # this other version remains mysterious to me.*

```
retSW4 <- smallcapW["2007-01-31/"]
tickers <- c("PLXS", "BRY", "CBM", "AIN")
mydata <- retSW4[, tickers]
plot.zoo(mydata)
```



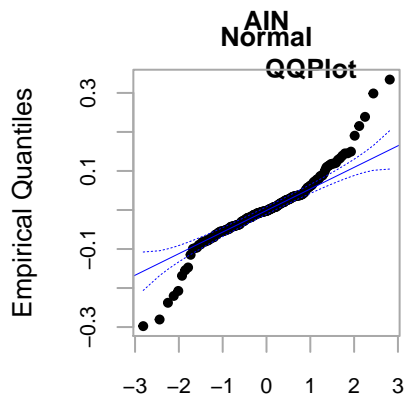
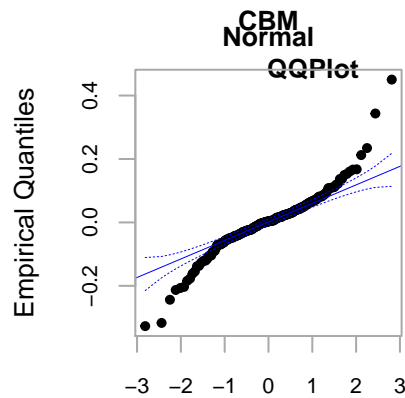
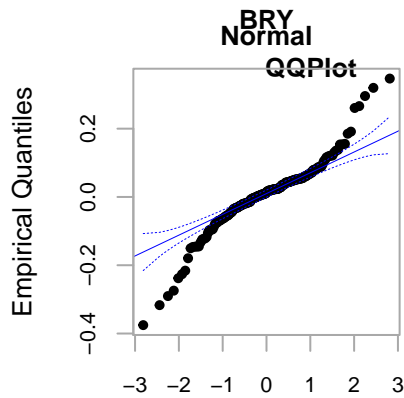
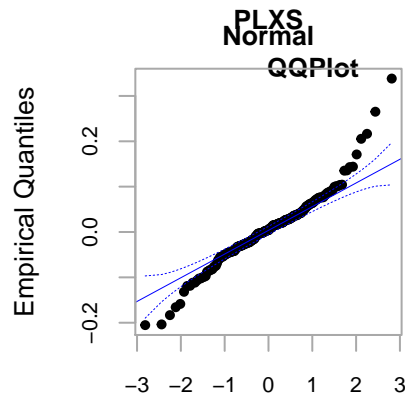
*# It does seem as though these returns are fairly non-normal. The value of the returns
 # seems generally unpredictable.*

```
nrow = 2
ncol = 2
par(mfrow = c(nrow, ncol))
par(pty = "s")
par(mar = c(3.1, 4.1, 2.1, 1.1))
n = 4
for (i in 1:n) {
  chart.QQPlot(mydata[, i], xlab = "Normal Quantiles", main = c(tickers[i], "Normal
```

```

    QQPlot"), envelope = 0.95, pch = 20, line = c("quartiles"), lwd = 0.5)
}

```



*# The normal qqplot leaves a lot to be desired here. There are many outliers outside even
the 0.95 (~ 2 sd) envelope.*

```

nrow = 2
ncol = 2
par(mfrow = c(nrow, ncol))
par(pty = "s")
par(mar = c(3.1, 4.1, 2.1, 1.1))
n = 4
for (i in 1:n) {
  main = c(tickers[i], "NorMix-2 QQPlot")
  chart.QQPlot.norMix(mydata[, i], xlab = "Two-Component NorMix Quantiles", main = main,
    na.rm = TRUE, envelope = 0.95, pch = ".", ylab = "", line =
    c("quartiles"), para = list(m = 2), distribution = "mixnormal", lwd = 0.5)
}

```

```

## Warning in norMixEM(x, para$m, trace = 0): EM did not converge in 100
## iterations (with 'tol'=1.49012e-08)

```

```

## [1] "fitted model:"
## 'Normal Mixture' object  ``norMixEM(x, m = para$m)[1:para$m,]''
##           mu sigma      w
## [1,] 0.00442 0.0476 0.766

```

```

## [2,] 0.00880 0.1207 0.234
## attr("class")
## [1] "fitEM" "nMfit"
## [1] "using fitted model as theoretical distribution"

## Warning in norMixEM(x, para$m, trace = 0): EM did not converge in 100
## iterations (with 'tol'=1.49012e-08)

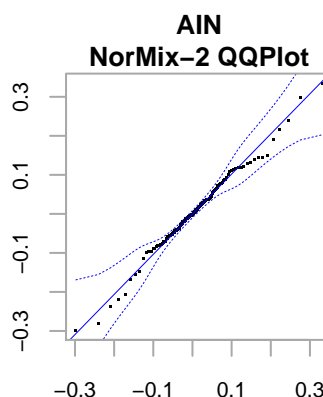
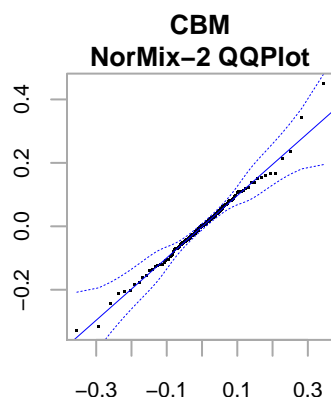
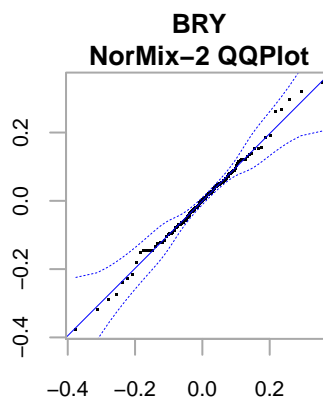
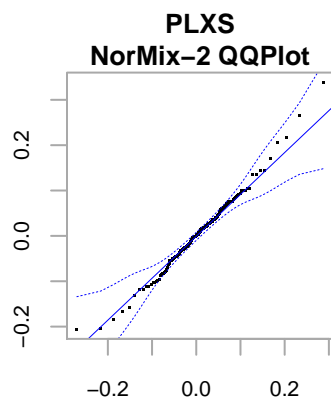
## Warning in norMixEM(x, para$m, trace = 0): EM did not converge in 100
## iterations (with 'tol'=1.49012e-08)

## [1] "fitted model:"
## 'Normal Mixture' object  ``norMixEM(x, m = para$m)[1:para$m,]''
##          mu  sigma      w
## [1,] -0.00948 0.1495 0.344
## [2,]  0.01514 0.0445 0.656
## attr("class")
## [1] "fitEM" "nMfit"
## [1] "using fitted model as theoretical distribution"

## [1] "fitted model:"
## 'Normal Mixture' object  ``norMixEM(x, m = para$m)[1:para$m,]''
##          mu  sigma      w
## [1,]  0.00464 0.0444 0.679
## [2,] -0.00533 0.1441 0.321
## attr("class")
## [1] "fitEM" "nMfit"
## [1] "using fitted model as theoretical distribution"

## [1] "fitted model:"
## 'Normal Mixture' object  ``norMixEM(x, m = para$m)[1:para$m,]''
##          mu  sigma      w
## [1,] -0.0038 0.0415 0.709
## [2,]  0.0172 0.1317 0.291
## attr("class")
## [1] "fitEM" "nMfit"
## [1] "using fitted model as theoretical distribution"

```



This normal-mixture plot seems to do much better. Now most of the values are inside the envelope.

```
nrow = 2
ncol = 2
par(mfrow = c(nrow, ncol))
par(pty = "s")
par(mar = c(3.1, 4.1, 2.1, 1.1))
n = 4
for (i in 1:n) {
  main = c(tickers[i], "NorMix-3 QQPlot")
  chart.QQPlot.norMix(mydata[, i], xlab = "Three-Component NorMix Quantiles", main = main,
    na.rm = TRUE, envelope = 0.95, pch = ".", ylab = "", line =
      c("quartiles"), para = list(m = 3), distribution = "mixnormal", lwd = 0.5)
}
```

```
## Warning in norMixEM(x, para$m, trace = 0): EM did not converge in 100
## iterations (with 'tol'=1.49012e-08)
```

```
## [1] "fitted model:"
## 'Normal Mixture' object    ``norMixEM(x, m = para$m)[1:para$m,]``
##           mu  sigma    w
## [1,] -0.0229 0.0742 0.337
## [2,]  0.0116 0.0407 0.557
## [3,]  0.0633 0.1277 0.106
## attr(,"class")
## [1] "fitEM" "nMfit"
```

```

## [1] "using fitted model as theoretical distribution"

## Warning in norMixEM(x, para$m, trace = 0): EM did not converge in 100
## iterations (with 'tol'=1.49012e-08)

## Warning in norMixEM(x, para$m, trace = 0): EM did not converge in 100
## iterations (with 'tol'=1.49012e-08)

## Warning in norMixEM(x, para$m, trace = 0): EM did not converge in 100
## iterations (with 'tol'=1.49012e-08)

## [1] "fitted model:"
## 'Normal Mixture' object  ``norMixEM(x, m = para$m)[1:para$m,]''
##           mu sigma      w
## [1,] -0.0577 0.1290 0.167
## [2,]  0.0159 0.0448 0.663
## [3,]  0.0341 0.1557 0.170
## attr(,"class")
## [1] "fitEM" "nMfit"
## [1] "using fitted model as theoretical distribution"

## Warning in norMixEM(x, para$m, trace = 0): EM did not converge in 100
## iterations (with 'tol'=1.49012e-08)

## Warning in norMixEM(x, para$m, trace = 0): EM did not converge in 100
## iterations (with 'tol'=1.49012e-08)

## Warning in norMixEM(x, para$m, trace = 0): EM did not converge in 100
## iterations (with 'tol'=1.49012e-08)

## [1] "fitted model:"
## 'Normal Mixture' object  ``norMixEM(x, m = para$m)[1:para$m,]''
##           mu sigma      w
## [1,] -0.000863 0.0880 0.387
## [2,]  0.004349 0.0371 0.490
## [3,] -0.002841 0.1874 0.124
## attr(,"class")
## [1] "fitEM" "nMfit"
## [1] "using fitted model as theoretical distribution"

## Warning in norMixEM(x, para$m, trace = 0): EM did not converge in 100
## iterations (with 'tol'=1.49012e-08)

## Warning in norMixEM(x, para$m, trace = 0): EM did not converge in 100
## iterations (with 'tol'=1.49012e-08)

## Warning in norMixEM(x, para$m, trace = 0): EM did not converge in 100
## iterations (with 'tol'=1.49012e-08)

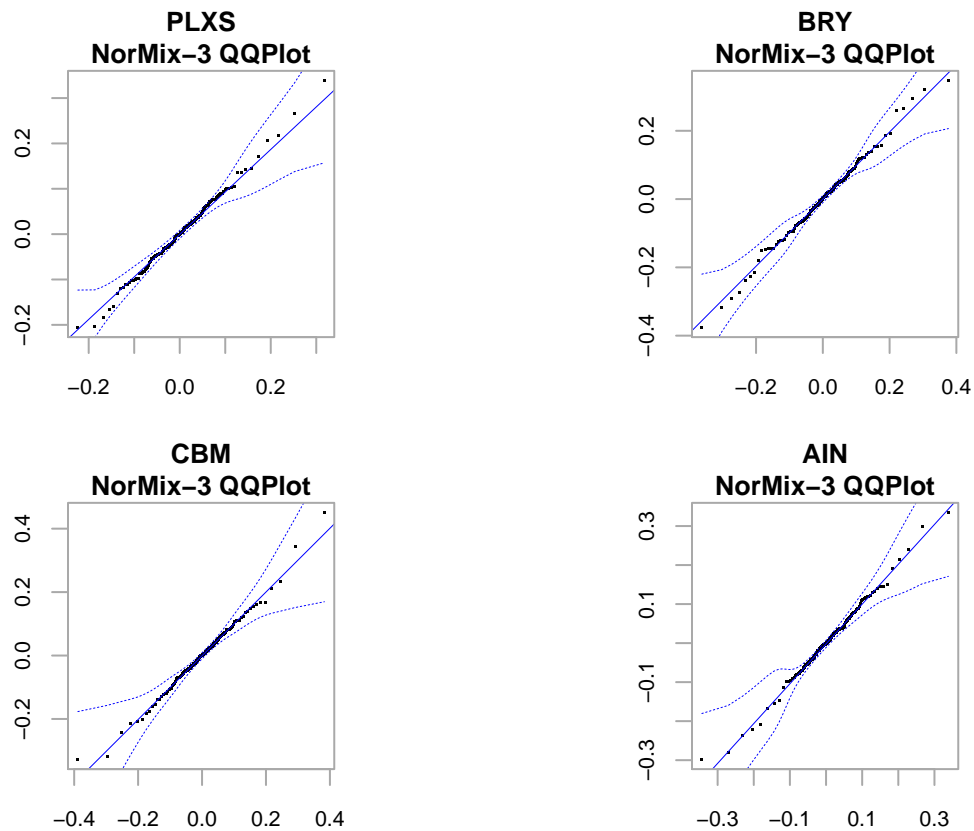
## [1] "fitted model:"
## 'Normal Mixture' object  ``norMixEM(x, m = para$m)[1:para$m,]''
##           mu sigma      w

```

```
## [1,] 0.0462 0.0587 0.212
## [2,] -0.0115 0.0381 0.614
## [3,] -0.0025 0.1554 0.174
## attr("class")
## [1] "fitEM" "nMfit"
## [1] "using fitted model as theoretical distribution"

## Warning in norMixEM(x, para$m, trace = 0): EM did not converge in 100
## iterations (with 'tol'=1.49012e-08)

## Warning in norMixEM(x, para$m, trace = 0): EM did not converge in 100
## iterations (with 'tol'=1.49012e-08)
```



I don't see a tremendous improvement here.

Problem #2

Consider a fully-invested portfolio with single period return $r_P = w_1 r_1 + w_2 r_2$ where returns r_1, r_2 have means μ_1, μ_2 , volatilities σ_1, σ_2 , and correlation coefficient ρ . Let μ_P and σ_P be the portfolio mean return and volatility.

- For the case $\rho = 1$ derive the straight-line relationship between μ_P and σ_P .

Now:

$$\mu_P = E[r_P] \quad (1)$$

$$= E[w_1 r_1 + w_2 r_2] \quad (2)$$

$$= E[w_1 r_1] + E[w_2 r_2] \quad (3)$$

$$= w_1 \mu_1 + w_2 \mu_2. \quad (4)$$

Also:

$$\sigma_P^2 = E[(r_P - \mu_P)^2] \quad (5)$$

$$= E[(w_1(r_1 - \mu_1) + w_2(r_2 - \mu_2))^2] \quad (6)$$

$$= w^T \Sigma w \quad (7)$$

$$(8)$$

for

$$\Sigma = \begin{bmatrix} \sigma_1^2 & \rho \sigma_1 \sigma_2 \\ \rho \sigma_1 \sigma_2 & \sigma_2^2 \end{bmatrix}$$

Thus $\sigma_P^2 = w_1^2 \sigma_1^2 + 2w_1 w_2 \rho \sigma_1 \sigma_2 + w_2^2 \sigma_2^2$.

Since the portfolio is fully invested, we know that $w_1 + w_2 = 1$.

For $\rho = 1$, $\sigma_P^2 = (w_1 \sigma_1 + w_2 \sigma_2)^2$.

Thus $\mu_P = w_1 \mu_1 + w_2 \mu_2 = w_1 \mu_1 + (1 - w_1) \mu_2$ and $\sigma_P = w_1 \sigma_1 + w_2 \sigma_2 = w_1 \sigma_1 + (1 - w_1) \sigma_2$.

Solving for w_1 from the σ - equation, we have:

$$w_1(\sigma_1 - \sigma_2) = \sigma_P - \sigma_2 \quad (9)$$

$$w_1 = \frac{\sigma_P - \sigma_2}{\sigma_1 - \sigma_2}. \quad (10)$$

Plugging this into the μ - equation, we have:

$$\mu_P = \frac{\sigma_P - \sigma_2}{\sigma_1 - \sigma_2} \mu_1 + \frac{\sigma_1 - \sigma_2 - \sigma_P + \sigma_2}{\sigma_1 - \sigma_2} \mu_2 \quad (11)$$

$$= \frac{\mu_1 - \mu_2}{\sigma_1 - \sigma_2} \sigma_P + \frac{\mu_2 \sigma_1 - \mu_1 \sigma_2}{\sigma_1 - \sigma_2}. \quad (12)$$

Thus there is a linear relationship between μ_P and σ_P .

b. For the case $\rho = -1$ derive the two straight-line relationships between μ_P and σ_P .

For $\rho = -1$, $\sigma_P^2 = (w_1 \sigma_1 - w_2 \sigma_2)^2$.

Thus:

$$\sigma_P = \begin{cases} w_1 \sigma_1 - w_2 \sigma_2, & w_1 \sigma_1 \geq w_2 \sigma_2 \\ w_2 \sigma_2 - w_1 \sigma_1, & w_1 \sigma_1 < w_2 \sigma_2 \end{cases}$$

Suppose $w_1\sigma_1 \geq w_2\sigma_2$. Then $\sigma_P = w_1\sigma_1 - (1 - w_1)\sigma_2$.

Hence:

$$w_1(\sigma_1 + \sigma_2) = \sigma_P + \sigma_2 \quad (13)$$

$$w_1 = \frac{\sigma_P + \sigma_2}{\sigma_1 + \sigma_2}. \quad (14)$$

Thus:

$$\mu_P = \frac{\sigma_P + \sigma_2}{\sigma_1 + \sigma_2} \mu_1 + \frac{\sigma_1 + \sigma_2 - \sigma_P - \sigma_2}{\sigma_1 + \sigma_2} \mu_2 \quad (15)$$

$$= \frac{\mu_1 + \mu_2}{\sigma_1 + \sigma_2} \sigma_P + \frac{\mu_2\sigma_1 + \mu_1\sigma_2}{\sigma_1 + \sigma_2}. \quad (16)$$

Otherwise: $\sigma_P = -w_1\sigma_1 + (1 - w_1)\sigma_2$.

Hence:

$$w_1(\sigma_1 + \sigma_2) = -\sigma_P + \sigma_2 \quad (17)$$

$$w_1 = \frac{\sigma_2 - \sigma_P}{\sigma_1 + \sigma_2}. \quad (18)$$

Thus:

$$\mu_P = \frac{\sigma_2 - \sigma_P}{\sigma_1 + \sigma_2} \mu_1 + \frac{\sigma_1 + \sigma_2 - \sigma_2 + \sigma_P}{\sigma_1 + \sigma_2} \mu_2 \quad (19)$$

$$= \frac{\mu_2 - \mu_1}{\sigma_1 + \sigma_2} \sigma_P + \frac{\mu_2\sigma_1 + \mu_1\sigma_2}{\sigma_1 + \sigma_2}. \quad (20)$$

Thus there is a linear relationship between μ_P and σ_P , since:

$$\mu_P = \begin{cases} \frac{\mu_1 + \mu_2}{\sigma_1 + \sigma_2} \sigma_P + \frac{\mu_2\sigma_1 + \mu_1\sigma_2}{\sigma_1 + \sigma_2}, & w_1\sigma_1 \geq w_2\sigma_2 \\ \frac{\mu_2 - \mu_1}{\sigma_1 + \sigma_2} \sigma_P + \frac{\mu_2\sigma_1 + \mu_1\sigma_2}{\sigma_1 + \sigma_2}, & w_1\sigma_1 < w_2\sigma_2 \end{cases}$$

Problem #3

Consider the quadratic utility optimal weights $\mathbf{w} = \lambda^{-1} \mathbf{\Sigma}^{-1} \boldsymbol{\mu}$ for the case of two assets with returns means, variances, and correlation $\mu_1, \mu_2, \sigma_1^2, \sigma_2^2$, and ρ respectively. First derive the expression for the two-by-two inverse covariance matrix $\mathbf{\Sigma}^{-1}$ in terms of $\mu_1, \mu_2, \sigma_1^2, \sigma_2^2$, and ρ . Then derive the expression for the derivative of w_1 with respect to μ_1 and explain how small changes in μ_1 affect w_1 when ρ is close to 1. Also derive an expression for the derivative of w_2 with respect to μ_1 and explain how changes in μ_1 affect changes in w_2 .

Now:

$$\mathbf{\Sigma} = \begin{bmatrix} \sigma_1^2 & \rho\sigma_1\sigma_2 \\ \rho\sigma_1\sigma_2 & \sigma_2^2 \end{bmatrix}$$

Thus:

$$\Sigma^{-1} = \frac{1}{\sigma_1^2 \sigma_2^2 - \rho \sigma_1^2 \sigma_2^2} \begin{bmatrix} \sigma_2^2 & -\rho \sigma_1 \sigma_2 \\ -\rho \sigma_1 \sigma_2 & \sigma_1^2 \end{bmatrix} \quad (21)$$

$$= \begin{bmatrix} \frac{1}{\sigma_1^2(1-\rho^2)} & \frac{-\rho}{\sigma_1 \sigma_2(1-\rho^2)} \\ \frac{-\rho}{\sigma_1 \sigma_2(1-\rho^2)} & \frac{1}{\sigma_2^2(1-\rho^2)} \end{bmatrix} \quad (22)$$

For quadratic utility, we have:

$U_q = w^T \mu - \frac{1}{2} \lambda w^T \Sigma w$. Hence $\mu - \lambda \Sigma w = 0$ and $w = \lambda^{-1} \Sigma^{-1} \mu$.

Thus:

$$\begin{bmatrix} w_1 \\ w_2 \end{bmatrix} = \lambda^{-1} \begin{bmatrix} \frac{1}{\sigma_1^2(1-\rho^2)} & \frac{-\rho}{\sigma_1 \sigma_2(1-\rho^2)} \\ \frac{-\rho}{\sigma_1 \sigma_2(1-\rho^2)} & \frac{1}{\sigma_2^2(1-\rho^2)} \end{bmatrix} \begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix} \quad (23)$$

$$w_1 = \frac{\mu_1}{\lambda \sigma_1^2(1-\rho^2)} - \frac{\mu_2 \rho}{\lambda \sigma_1 \sigma_2(1-\rho^2)}. \quad (24)$$

Thus:

$$\frac{\partial w_1}{\partial \mu_1} = \frac{1}{\lambda \sigma_1^2(1-\rho^2)}.$$

Thus when $\rho \approx 1$, i.e. for a highly correlated pair of assets, $\frac{\partial w_1}{\partial \mu_1}$ is huge and so w_1 is incredibly sensitive to small changes in μ_1 .

Similarly:

$$w_2 = \frac{-\rho \mu_1}{\lambda \sigma_1 \sigma_2(1-\rho^2)} + \frac{\mu_2}{\lambda \sigma_2^2(1-\rho^2)},$$

$$\text{and so } \frac{\partial w_2}{\partial \mu_1} = \frac{-\rho}{\lambda \sigma_1 \sigma_2(1-\rho^2)}.$$

Again, when $\rho \approx 1$, $|\frac{\partial w_2}{\partial \mu_1}|$ is huge and so w_2 is incredibly sensitive to small changes in μ_1 . In this case the differential is negative.

Problem #4

Show that the covariance of a global minimum variance (GMV) portfolio with that of any other frontier portfolio is constant, i.e. does not depend on any portfolio mean returns.

To express the covariance of a GMV portfolio with any other, we write:

$$\text{cov}(r_{GMV}, r_P) = E[(r_{GMV} - \mu_{GMV})(r_P - \mu_P)] \quad (25)$$

$$= E[w_{GMV}^T (r - \mu)^2 w_P] \quad (26)$$

$$= w_{GMV}^T \Sigma w_P. \quad (27)$$

Now:

$$w_{GMV}^T = \frac{1}{1^T \Sigma^{-1} 1} 1^T (\Sigma^{-1})^T \quad (28)$$

$$= \frac{1}{1^T \Sigma^{-1} 1} 1^T \Sigma^{-1}. \quad (29)$$

Thus:

$$\text{cov}(r_{GMV}, r_P) = \frac{1^T \Sigma^{-1} \Sigma w_P}{1^T \Sigma^{-1} 1} \quad (30)$$

$$= \frac{1^T w_P}{1^T \Sigma^{-1} 1}. \quad (31)$$

Thus the covariance is independent of μ_P .

Problem #5

Complete the steps to derive the expression (3.77) for the QU optimal portfolio weights w_{QU} with risky assets only.

Expression (3.77) tells us that $w_q = w_{GMV} + \gamma_{RA}^{-1} \Sigma^{-1} (\mu - \mu_{GMV} \cdot 1)$. Let us start with an expression for quadratic utility.

$$U_q = w^T \mu - \frac{1}{2} \gamma_{RA} w^T \Sigma w + \lambda (1 - w^T 1).$$

Thus, taking derivatives, $\mu - \gamma_{RA} \Sigma w - \lambda \cdot 1 = 0$.

So $w = \gamma_{RA}^{-1} \Sigma^{-1} (\mu - \lambda \cdot 1)$.

But now $1^T w = 1$, so $1 = \gamma_{RA}^{-1} 1^T \Sigma^{-1} (\mu - \lambda \cdot 1)$.

$$\text{Thus } \lambda = -\frac{1 - \gamma_{RA}^{-1} 1^T \Sigma^{-1} \mu}{\gamma_{RA}^{-1} 1^T \Sigma^{-1} 1} = \frac{\gamma_{RA}^{-1} 1^T \Sigma^{-1} \mu - 1}{\gamma_{RA}^{-1} 1^T \Sigma^{-1} 1}.$$

Plugging this equation for λ into the above expression for w , we have:

$$w = \gamma_{RA}^{-1} \Sigma^{-1} \left(\mu - \frac{\gamma_{RA}^{-1} 1^T \Sigma^{-1} \mu - 1}{\gamma_{RA}^{-1} 1^T \Sigma^{-1} 1} \cdot 1 \right) \quad (32)$$

$$= \frac{1}{\gamma} \Sigma^{-1} \left(\mu - \frac{1^T \Sigma^{-1} \mu - \gamma}{1^T \Sigma^{-1} 1} \cdot 1 \right) \quad (33)$$

$$= \frac{1}{\gamma} \Sigma^{-1} \left(\mu - \frac{1^T \Sigma^{-1} \mu}{1^T \Sigma^{-1} 1} \cdot 1 + \frac{\gamma}{1^T \Sigma^{-1} 1} \cdot 1 \right) \quad (34)$$

$$= \frac{1}{\gamma} \Sigma^{-1} (\mu - \mu_{GMV} \cdot 1) + \frac{\gamma \Sigma^{-1}}{\gamma 1^T \Sigma^{-1} 1} \cdot 1 \quad (35)$$

$$= w_{GMV} + \frac{1}{\gamma} \Sigma^{-1} (\mu - \mu_{GMV} \cdot 1). \quad (36)$$