

## 1 Axioms of Probability

# 1.1 Definitions

- 1. For any event  $A, 0 \le Pr(A) \le 1$ .
- 2.  $A \cup B \triangleq A + B$ .
- 3.  $A \cap B \triangleq AB$ .
- 4. The null and complete event are  $\phi = 0, S = 1$ .
- 5. If AB = 0, Pr(A + B) = Pr(A) + Pr(B).
- 6. (A + B)' = A'B'

### 1.2 Problems

Prove the following:

1.

$$A = AB + AB' \tag{1.2.1.1}$$

2.

$$Pr(A) = Pr(AB) + Pr(AB')$$
 (1.2.2.1)

3.

$$A + B = B + AB' \tag{1.2.3.1}$$

4.

$$Pr(A + B) = Pr(A) + Pr(B) - Pr(AB)$$
 (1.2.4.1)

### 2 Distribution of the sum of random variables

# 2.1 Definitions

1. The mean of X is defined as

$$E(X) = \sum_{k} k p_X(k)$$
 (2.1.1.1)

2. The Z transform of X is defined as

$$M_X(z) = E(z^{-X}) = \sum_{k=-\infty}^{\infty} z^{-k} p_X(k)$$
 (2.1.2.1)

- 3. There is a one to one relationship between the pmf and its Z transform.
- 4. If If  $X_1$  and  $X_2$  are independent,

$$E[f(X_1)g(X_2)] = E[f(X_1)]E[g(X_2)]$$
 (2.1.4.1)

5. For a Bernoulli random variable X, the pmf is

$$p_X(n) = \begin{cases} p & k = 1\\ 1 - p & k = 0\\ 0 & \text{otherwise} \end{cases}$$
 (2.1.5.1)

6.  $X_i$  are said to be i.i.d (independent and identically distributed) if they are independent and have the same pmf.

## 2.2 Problem

- 1. Find the Z-transform for X, given that X is a Bernoulli random variable with parameter p.
- 2. If  $X_1$  and  $X_2$  are independent, and

$$Y = X_1 + X_2, (2.2.2.1)$$

show that

$$M_Y(z) = M_{X_1}(z)M_{X_2}(z) (2.2.2.2)$$

- 3. Find the Z-transform of Y, given that  $X_i$  are i.i.d Bernoulli random variables with parameter p.
- 4. Find the pmf of *Y*.
- 5. Find the pmf of

$$Y = \sum_{i=1}^{N} X_i,$$
 (2.2.5.1)

where  $X_i$  are i.i.d.

### 3 Moments and variance

## 3.1 Definitions

1. The variance of *X* is defined as:

$$Var(X) = E(X - E(X))^{2}$$
 (3.1.1.1)

2. The Z transform of X is defined as

$$M_X(z) = E(z^{-X}) = \sum_{k=-\infty}^{\infty} z^{-k} p_X(k)$$
 (3.1.2.1)

3. Let X be a random variable with pmf.

$$p_X(k) = \begin{cases} 1/6 & 1 \le k \le 6\\ 0 & \text{otherwise} \end{cases}$$
 (3.1.3.1)

X is said to be Discrete Uniform Random Variable

4. The  $n^{th}$  moment of X is defined as:

$$E(X^{n}) = \sum_{k=-\infty}^{\infty} k^{n} p_{X}(k)$$
 (3.1.4.1)

### 3.2 Problems

- 1. Show that  $Var(X) = E(X^2) [E(X)]^2$
- 2. Find  $M_X(z)$
- 3. Show that  $E(X) = \frac{d}{dz} M_X(z^{-1}) |_{z=1}$
- 4. Find  $E(X^2)$
- 5. Find Var(X).

#### 4 Convolution

### 4.1 Definitions

1. The Z transform of X is defined as

$$M_X(z) = E(z^{-X}) = \sum_{k=-\infty}^{\infty} z^{-k} p_X(k)$$
 (4.1.1.1)

2. Let *X* be a random variable with pmf.

$$p_X(k) = \begin{cases} 1/6 & 1 \le k \le 6\\ 0 & \text{otherwise} \end{cases}$$
 (4.1.2.1)

X is said to be Discrete Uniform Random Variable

3. Convolution of two sequences using Toeplitz matrices

$$\mathbf{y} = \mathbf{x} \otimes \mathbf{h}$$

$$\mathbf{y} = \begin{pmatrix} h_1 & 0 & \dots & 0 \\ h_2 & h_1 & \dots & 0 \\ h_3 & h_2 & h_3 & \dots & 0 \\ h_{m-1} & \dots & h_2 & h_1 \\ h_m & h_{m-1} & \dots & h_2 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & h_m \end{pmatrix} \begin{pmatrix} x_1 \\ x_2 \\ \vdots \\ \vdots \\ x_n \end{pmatrix}$$

$$(4.1.3.1)$$

- 4.2 Problems
  - 1. If  $\mathbf{x} = \mathbf{h} = \frac{1}{2} \begin{pmatrix} 1 \\ 1 \end{pmatrix}$ , find  $\mathbf{y}$ .
  - 2. Find  $p_{X_1}(k) \circledast p_{X_2}(k)$  using toeplitz matrices.
  - 3. Find  $M_Y(z)$ , such that  $Y = X_1 + X_2$
  - 4. Find  $p_Y(k)$
- 5 Z-Transform Applications
- 5.1 Definitions

1.

$$u(n) = \begin{cases} 1 & n \ge 0 \\ 0 & n < 0 \end{cases}$$
 (5.1.1.1)

2. The Z transform of X is defined as

$$M_X(z) = E(z^{-X}) = \sum_{k=-\infty}^{\infty} z^{-k} p_X(k)$$
 (5.1.2.1)

- 5.2 Problems
  - 1. If

$$p_Y(n) \stackrel{\mathcal{Z}}{\longleftrightarrow} M_Y(z),$$
 (5.2.1.1)

show that

$$p_Y(n-k) \stackrel{\mathcal{Z}}{\longleftrightarrow} M_Y(z)z^{-k},$$
 (5.2.1.2)

2. Show that

$$u(n) \stackrel{\mathcal{Z}}{\longleftrightarrow} \frac{1}{(1-z^{-1})}, \quad |z| > 1 \tag{5.2.2.1}$$

3. Show that

$$nu(n) \stackrel{\mathcal{Z}}{\longleftrightarrow} \frac{z^{-1}}{(1-z^{-1})^2}, \quad |z| > 1$$
 (5.2.3.1)

4. Let

$$M_Y(z) = \left\{ \frac{z^{-1} \left( 1 - z^{-6} \right)}{6 \left( 1 - z^{-1} \right)} \right\}^2, \quad |z| > 1$$
 (5.2.4.1)

Show that

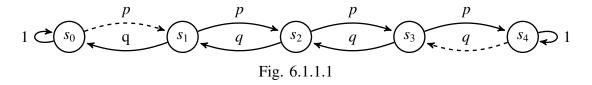
$$p_Y(n) = \frac{(n-1)u(n-1) - 2(n-7)u(n-7) + (n-13)u(n-13)}{36}$$
 (5.2.4.2)

## 6 Markov Chain

## 6.1 Definitions

1. Fig. 6.1.1.1 shows a Markovs chain with 5 states. Transition from one state to another happens over time.  $s_0$  and  $s_4$  are absorbing states.

$$p + q = 1 \tag{6.1.1.1}$$



2. At time instant n,

$$\mathbf{p}^{(n)} = \begin{pmatrix} P_0^{(n)} \\ P_1^{(n)} \\ P_2^{(n)} \\ P_3^{(n)} \\ P_4^{(n)} \end{pmatrix}$$
(6.1.2.1)

where  $P_i^{(n)}$  are defined to be the *stationary* probabilities.

3.  $P_{i|j}$  is defined as the *transition* probability of going to state i from state j.

4. For a matrix A, let

$$\mathbf{A}\mathbf{x} = \lambda \mathbf{x}.\tag{6.1.4.1}$$

Then  $\lambda$  is a scalar defined to be the eigenvalue of **A** and **x** is the corresponding eigenvector.

#### 6.2 Problems

1. Let

$$P_0^{(n+1)} = P_0^{(n)} (6.2.1.1)$$

$$P_1^{(n+1)} = pP_0^{(n)} + qP_2^{(n)} (6.2.1.2)$$

$$P_2^{(n+1)} = pP_1^{(n)} + qP_3^{(n)} (6.2.1.3)$$

$$P_3^{(n+1)} = pP_2^{(n)} + qP_4^{(n)} (6.2.1.4)$$

$$P_0^{(n+1)} = P_0^{(n)}$$

$$P_1^{(n+1)} = pP_0^{(n)} + qP_2^{(n)}$$

$$P_2^{(n+1)} = pP_1^{(n)} + qP_3^{(n)}$$

$$P_3^{(n+1)} = pP_2^{(n)} + qP_4^{(n)}$$

$$P_4^{(n+1)} = P_4^{(n)}$$

$$(6.2.1.4)$$

$$P_4^{(n+1)} = P_4^{(n)}$$

$$(6.2.1.5)$$

Find the matrix **P** such that  $\mathbf{p}^{(n+1)} = \mathbf{P}\mathbf{p}^{(n)}$ 

- 2. Show that 1 is an eigen value of **P**.
- 3. Show that

$$P_2^{(n+1)} = pP_1^{(n)} + qP_3^{(n)} (6.2.3.1)$$

4. If  $P_0 = 1$ ,  $P_N = 0$  and

$$P_i = pP_{i-1} + qP_{i+1} (6.2.4.1)$$

show that

$$P_{i} = \frac{\left(\frac{p}{q}\right)^{i} - \left(\frac{p}{q}\right)^{N}}{1 - \left(\frac{p}{q}\right)^{N}}, 0 \le i \le N$$

$$(6.2.4.2)$$

### 7 Gaussian Distribution

## 7.1 Definitions

1. The CDF of X is defined as,

$$F_X(x) = \Pr(X \le x)$$
 (7.1.1.1)

2. The PDF of X is defined as,

$$p_X(x) = \frac{d}{dx} F_X(x) \tag{7.1.2.1}$$

3. Let  $X \sim \mathcal{N}(0, 1)$ . Then the Q function is defined as,

$$Q(x) = \Pr(X > x), \quad x \ge 0$$
 (7.1.3.1)

# 7.2 Problems

1. Find

$$\Pr\left(|X - \mu| \le k\sigma\right) \tag{7.2.1.1}$$

in terms of Q function.

2. Find

$$\Pr\left(X \le x, |X - \mu| \le k\sigma\right) \tag{7.2.2.1}$$

in terms of  $F_X(x)$ 

3. Find

$$F_X(x||X - \mu| \le k\sigma) \tag{7.2.3.1}$$

4. Find

$$p_X(x||X-\mu| \le k\sigma) \tag{7.2.4.1}$$

8 BIVARIATE GAUSSIAN

# 8.1 Definitions

1. Let

$$\mathbf{X} = \begin{pmatrix} x_1 \\ x_2 \end{pmatrix}, \boldsymbol{\mu} = E\left(\mathbf{X}\right) \tag{8.1.1.1}$$

$$\Sigma_{\mathbf{x}} = E\left[ (\mathbf{x} - \boldsymbol{\mu}) (\mathbf{x} - \boldsymbol{\mu})^{\mathsf{T}} \right]$$
 (8.1.1.2)

Then  $\Sigma_x$  is defined to be the *covariance* matrix of x.

2. For  $\mathbf{x} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma}_{\mathbf{x}})$ ,

$$f_{\mathbf{X}}(\mathbf{x}) = \frac{1}{2\pi\sqrt{|\Sigma_{\mathbf{x}}|}} \exp{-\frac{1}{2}\left((\mathbf{x} - \boldsymbol{\mu})^{\top} \Sigma_{\mathbf{x}}^{-1} (\mathbf{x} - \boldsymbol{\mu})\right)}$$
(8.1.2.1)

3. The correlation coefficient is defined as

$$\rho = \frac{E[(x_1 - \mu_1)(x_2 - \mu_2)]}{\sigma_1 \sigma_2}$$
 (8.1.3.1)

where  $\mu_i$ ,  $\sigma_i^2$  are the mean and variance of  $x_i$ .

1. Show that

$$E\left[\left(\mathbf{x} - \boldsymbol{\mu}\right)\left(\mathbf{x} - \boldsymbol{\mu}\right)^{\mathsf{T}}\right] = \begin{pmatrix} \sigma_1^2 & \rho\sigma_1\sigma_2\\ \rho\sigma_1\sigma_2 & \sigma_2^2 \end{pmatrix}$$
(8.2.1.1)

- 2. Prove that  $\Sigma_x$  is a diagonal matrix when  $x_1$  and  $x_2$  are independent.
- 3. Let

$$z_1 = x_1 + x_2 \tag{8.2.3.1}$$

$$z_2 = x_1 - x_2 \tag{8.2.3.2}$$

Find **P** such that

$$\mathbf{z} = \begin{pmatrix} z_1 \\ z_2 \end{pmatrix} = \mathbf{P}\mathbf{x} \tag{8.2.3.3}$$

4. Show that

$$\Sigma_{\mathbf{z}} = \mathbf{P} \Sigma_{\mathbf{x}} \mathbf{P}^{\mathsf{T}} \tag{8.2.4.1}$$

- 5. Check the independence of  $z_1$  and  $z_2$  given that  $\sigma_1 = \sigma_2$ .
- 6. Show that columns of **P** are eigenvectors of  $\Sigma_z$ .
- 7. Show that the eigenvectors of  $\Sigma_z$  are orthogonal to each other.
- 8. Summarize your conclusion in one line.

## 9 Transformation of Random Variables

## 9.1 Definitions

1. The pdf of  $X \sim \mathcal{N}(\mu, \sigma^2)$  is defined as

$$f_X(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp^{-\frac{(x-\mu)^2}{2\sigma^2}} - \infty < x < \infty$$
 (9.1.1.1)

2. The Jacobian matrix transforming  $R, \Theta$  to  $X_1, X_2$  is defined as

$$\mathbf{J} = \begin{pmatrix} \frac{\partial X_1}{\partial R} & \frac{\partial X_1}{\partial \Theta} \\ \frac{\partial X_2}{\partial R} & \frac{\partial X_2}{\partial \Theta} \end{pmatrix}$$
(9.1.2.1)

3.

$$p_{R,\Theta}(r,\theta) = p_{X_1,X_2}(x_1,x_2) |\mathbf{J}|$$
 (9.1.3.1)

4. The marginal distribution

$$p_R(r) = \int_0^{2\pi} p_{R,\Theta}(r,\theta) \ d\theta$$
 (9.1.4.1)

5. The Laplace transform of  $p_Y(y)$  is given by

$$M_Y(s) = E\left(e^{-sY}\right) \tag{9.1.5.1}$$

6. The unit step function is defined as

$$u(y) = \begin{cases} 1 & y \ge 0 \\ 0 & \text{otherwise} \end{cases}$$
 (9.1.6.1)

### 9.2 Problems

- 1. Let  $X_1, X_2 \in \mathbb{N} \sim (0, 1)$  be i.i.d. Find  $p_{X_1, X_2}(x_1, x_2)$ .
- 2. Let

$$X_1 = R\cos\Theta \tag{9.2.2.1}$$

$$X_2 = R\sin\Theta \tag{9.2.2.2}$$

Find  $p_{R,\theta}(r,\theta)$ .

- 3. Find  $p_R(r)$ .
- 4. Find  $p_{\Theta}(\theta)$ .
- 5. Find the distribution of

$$Y = X_1^2 + X_2^2 (9.2.5.1)$$

- 6. Find the Laplace transform of  $e^{-y}u(y)$
- 7. Find the Laplace transform of  $p_{X_1^2}(x_1)$ .
- 8. Find the Laplace transform of  $p_Y(y)$  using (9.2.5.1).
- 9. Find the distribution of Y.

## 10 Order Statistics

### 10.1 Definitions

1. The pdf of an exponential distribution is given by

$$p_X(x) = e^{-x}u(x) (10.1.1.1)$$

where  $u(\cdot)$  is the unit step function.

- 1. Find  $F_X(x)$ .
- 2. Let X and Y be iid exponential. Find  $F_{XY}(z,z)$
- 3. Show that

$$\Pr(X \le z, X > Y) = \int_{x = -\infty}^{z} \int_{y = -\infty}^{x} p_{X,Y}(x, y) \, dx dy$$
 (10.2.3.1)

$$=\frac{e^{-2z}}{2}-e^{-z} \tag{10.2.3.2}$$

4. Find

$$\Pr(Y \le z, X \le Y)$$
 (10.2.4.1)

5. Let

$$Z = \max(X, Y)$$
 (10.2.5.1)

Show that

$$F_Z(z) = \Pr(X \le z, X > Y) + \Pr(Y \le z, X \le Y)$$
 (10.2.5.2)

- 6. Find the pdf of Z.
- 7. Find the pdf of  $W = \min(X, Y)$ .
- 8. Find  $F_W(vZ)$ , where v is a constant.
- 9. Find  $E[F_W(vZ)]$ .
- 10. Find the pdf of  $V = \frac{W}{Z}$ .

## 11 Some distributions

## 11.1 Definitions

1. If Y = f(Z) be monotonic,

$$p_Y(y) dy = p_Z(z) dz$$
 (11.1.1.1)

2. The pdf of the  $\chi^2(k)$  distribution is given by

$$p_X(x) = \frac{x^{\frac{k}{2} - 1} e^{-\frac{x}{2}}}{2^{\frac{k}{2}} \Gamma\left(\frac{k}{2}\right)} u(x)$$
 (11.1.2.1)

3. The Beta function is defined as

$$B(x,y) = \frac{\Gamma(x)\Gamma(y)}{\Gamma(x+y)}$$
(11.1.3.1)

1. Let  $X_1 \sim \chi^2(m)$  and  $X_2 \sim \chi^2(n)$  be independent. For

$$Z = \frac{X_1/m}{X_2/n},\tag{11.2.1.1}$$

show that

$$F_Z(z) = E\left[F_{X_1}\left(\frac{mzX_2}{n}\right)\right]$$
 (11.2.1.2)

2. Show that

$$p_Z(z) = E\left[\frac{mX_2}{n}p_{X_1}\left(\frac{mzX_2}{n}\right)\right]$$
 (11.2.2.1)

3. Show that

$$p_Z(z) = \frac{\left(\frac{m}{n}\right)^{m/2}}{B\left(\frac{m}{2}, \frac{n}{2}\right)} x^{\frac{m}{2} - 1} \left(1 + \frac{m}{n}z\right)^{-\frac{m+n}{2}} u(z)$$
 (11.2.3.1)

Z has an F distribution with (m, n) degrees of freedom.

- 4. Show that  $Y = \frac{1}{Z}$  is monotonic.
- 5. Show that Y also has an F distribution with (n, m) degrees of freedom. 6. Find the pdf of  $\frac{mZ}{mZ+n}$ .

#### 12 Inequalities

## 12.1 Definitions

1. The mean of Y is defined as

$$E(Y) = \sum_{k} k p_{Y}(k)$$
 (12.1.1.1)

2. The MGF of X is defined as

$$M_X(s) = E(e^{-sX})$$
 (12.1.2.1)

3. Let  $X \sim \mathcal{N}(0, 1)$ . Then the Q function is defined as,

$$Q(x) = \Pr(X > x), \quad x \ge 0$$
 (12.1.3.1)

## 12.2 Problems

1. Show that

$$E(Y) \ge \sum_{k=m}^{\infty} k p_Y(k), \quad m > 0$$
 (12.2.1.1)

2. Show that

$$\Pr(Y > m) \le \frac{E(Y)}{m}, \quad m > 0$$
 (12.2.2.1)

3. Using (12.2.2.1), show that Show that

$$\Pr([Y - E(Y)]^2 > b^2) \le \frac{var(Y)}{b^2}, \quad b > 0$$
 (12.2.3.1)

4. Show that

$$\Pr(|Y - E(Y)| > b) \le \frac{var(Y)}{b^2}, \quad b > 0$$
 (12.2.4.1)

5. Show that

$$\Pr(X > a) = \Pr(e^{-sX} > e^{-sa}), \quad s < 0$$
 (12.2.5.1)

6. Using (12.2.2.1), show that

$$\Pr(X > a) \le e^{as} M_X(s), \quad s < 0$$
 (12.2.6.1)

7. Show that the MGF of X is

$$M_X(s) = e^{\frac{1}{2}s^2} \tag{12.2.7.1}$$

8. Using (12.2.6.1) show that

$$Q(x) \le e^{-\frac{x^2}{2}} \tag{12.2.8.1}$$

### 13 Jensen's Inequality

## 13.1 Definitions

1. A function g is said to be convex if

$$g(\lambda x_1 + (1 - \lambda) x_2) \le \lambda g(x_1) + (1 - \lambda) g(x_2)$$
 (13.1.1.1)

2. g is convex if and only if

$$g''(x) \ge 0 \tag{13.1.2.1}$$

3. The *information* associated with an event E is defined as

$$I(E) = -\log_2(p(E)) \tag{13.1.3.1}$$

4. The *entropy* of a random variable X is given by

$$H(X) = -E[\log_2(p(X))]$$
 (13.1.4.1)

5. For a *bernoulli* random variable  $Y \in \{0, 1\}$ , the p.m.f is given by

$$p_Y(k) = \begin{cases} p & k = 0\\ 1 - p & k = 1 \end{cases}$$
 (13.1.5.1)

### 13.2 Problems

- 1. Show that  $\log(\frac{1}{x})$  is convex. 2. Show that

$$H(Y) = -p\log_2 p - (1-p)\log_2 (1-p)$$
 (13.2.2.1)

and find the maximum value of H(Y).

3. Let  $X \in \{x_1, x_2\}$  and

$$\lambda = q = p_X(X = x_1) \tag{13.2.3.1}$$

in (13.1.1.1). Show that

$$E[g(X)] \ge g[E(X)].$$
 (13.2.3.2)

4. Using (13.2.3.2), show that

$$H(Y) \le 1 \tag{13.2.4.1}$$

### 14 MAXIMUM LIKELIHOOD CONDITION

## 14.1 Definitions

1. The pmf for a Binomial Random variable  $Y \sim (n, p)$  is given by

$$p_Y(Y=k) = \binom{n}{k} p^k (1-p)^{n-k}, \quad 0 \le k \le n$$
 (14.1)

2. The pdf of an Exponential Random variable X is given by

$$p_X(x) = ce^{-cx}u(x) \tag{14.1}$$

#### 14.2 Problems

1. Show that the solution of

$$\max_{p} p_Y(Y = k) \tag{14.2.1.1}$$

is

$$\hat{p} = \frac{k}{n} \tag{14.2.1.2}$$

2. Find

$$\Pr(X > T)$$
 (14.2.2.1)

where T is a constant.

3. If X represents the lifetime of a bulb, show that the value of c that maximizes the probability of k out of n bulbs working after T hours is

$$c = \frac{1}{T} \log \left( \frac{n}{k} \right) \tag{14.2.3.1}$$

## 15 Maximum Likelihood Condition

### 15.1 Definitions

1. The pmf for a Bernoulli r.v.  $X_i$  is given by

$$p_{X_i}(k) = \begin{cases} p_0 & k = 1\\ 1 - p_0 & k = 0\\ 0 & \text{otherwise} \end{cases}$$
 (15.1)

2. The pmf for a Binomial Random variable

$$X = \sum_{i=1}^{n} X_i \tag{15.1}$$

is given by

$$p_X(X=k) = \binom{n}{k} p_0^k (1-p_0)^{n-k}, \quad 0 \le k \le n$$
 (15.2)

where  $X_i$  are i.i.d.

3. Let  $Z \sim \mathcal{N}(0, 1)$ . Then the Q function is defined as,

$$Q(z) = \Pr(Z > z), \quad z \ge 0$$
 (15.1)

4. The MGF of Y is defined as

$$M_Y(s) = E\left(e^{-sY}\right) \tag{15.1}$$

5. The MGF of Z is

$$M_Z(s) = \frac{1}{2}e^{\frac{1}{2}s^2} \tag{15.1}$$

6.

$$\lim_{n \to \infty} \left( 1 + \frac{1}{n} \right)^n = e \tag{15.1}$$

- 1. Show that the mean and variance of  $X_i$  are  $\mu = p_0$  and  $\sigma^2 = p(1 p)$ .
- 2. Show that the mean and variance of

$$Y_i = \frac{X_i - \mu}{\sigma} \tag{15.1}$$

are 0 and 1 respectively.

3. Show that the mean and variance of

$$Y = \frac{1}{\sqrt{n}} \sum_{i=1}^{n} Y_i \tag{15.1}$$

are 0 and 1 respectively.

4. Show that

$$M_Y(s) = \left[ E\left(e^{-\frac{sY_i}{\sqrt{n}}}\right) \right]^n \tag{15.1}$$

5. Show that

$$E\left(e^{-\frac{sY_i}{\sqrt{n}}}\right) = 1 + \frac{1}{2n} + \frac{1}{n}R(s,n)$$
 (15.1)

where R(s, n) is an infinite series.

6. Show that

$$\lim_{n \to \infty} Y = Z \tag{15.1}$$

This is known as the Central Limit Theorem.

7. Let

$$p = \frac{1}{n} \sum_{i=1}^{n} X_i \tag{15.1}$$

Show that

$$E(p) = \mu = p_0 \tag{15.2}$$

$$var(p) = \frac{\sigma^2}{n} = \frac{p_0 (1 - p_0)}{n}$$
 (15.3)

8. Show that

$$\lim_{n \to \infty} \frac{p - p_0}{\sqrt{\frac{p_0(1 - p_0)}{n}}} = Z \tag{15.1}$$

9. Find z if

$$\Pr(|Z| < z) = 1 - \alpha \tag{15.1}$$

10. Show that

$$p_0 - \sqrt{\frac{p_0 (1 - p_0)}{n}} Q^{-1} \left(\frac{\alpha}{2}\right) (15.1)$$

11. Among 4000 newborns, 2080 are male. Find the  $1 - \alpha = 0.99$  confidence interval of the probability that a male child is born.