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EE5603:Concentration Inequalities

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1 Markov Inequality

1.1 Let $X \ge 0$ be a positive random integer. Show that

$$E[X] = \sum_{m=0}^{\infty} \Pr(X \ge m)$$
 (1.1)

Solution: By definition,

$$E[X] = \sum_{m=0}^{\infty} m \Pr(X = m)$$

$$= \Pr(X = 1) + 2 \Pr(X = 2) + 3 \Pr(X = 3)$$

$$+ \dots \qquad (1.3)$$

$$= \{\Pr(X = 1) + \Pr(X = 2) + \Pr(X = 3)$$

$$+ \dots \} \qquad (1.4)$$

$$+ \{\Pr(X = 2) + \Pr(X = 3) + \dots \} \qquad (1.5)$$

$$+ \{\Pr(X = 3) + \dots \} + \dots \qquad (1.6)$$

$$= \Pr(X \ge 1) + 2 \Pr(X \ge 2) + 3 \Pr(X \ge 3)$$

$$+ \dots \qquad (1.7)$$

resulting in (1.2).

1.2 For a continuous r.v $X \ge 0$, show that

$$E[X] = \int_0^\infty \Pr(x \ge t) dt \qquad (1.8)$$

1.3 For r.v $X \ge 0$ and $\varepsilon > 0$, show that

$$\Pr(X \ge \varepsilon) \le \frac{E[X]}{\varepsilon}$$
 (1.9)

Solution: $:: X \ge 0$,

$$E[X] = \int_0^\infty x p_X(x) dx$$
 (1.10)
=
$$\int_0^\varepsilon x p_X(x) dx + \int_\varepsilon^\infty x p_X(x) dx$$
 (1.11)

$$\geqslant \int_{\varepsilon}^{\infty} x p_X(x) \, dx \tag{1.12}$$

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which can be expressed as

$$E[X] \geqslant \int_{\varepsilon}^{\infty} \varepsilon p_X(x) \, dx \tag{1.13}$$

$$= \varepsilon \int_{\varepsilon}^{\infty} p_X(x) \, dx = \varepsilon \Pr(X \ge \varepsilon) \quad (1.14)$$

resulting in (1.9).

1.4 Chernoff Bound: For any r.v X with bounded variance, and for any t > 0, show using (1.9) that

$$\Pr(e^{tX} \geqslant e^{t\varepsilon}) \leqslant \frac{E\left(e^{tX}\right)}{e^{t\varepsilon}}$$
 (1.15)

1.5 Show that

$$Pr(X \ge \varepsilon) = Pr(e^{tX} \ge e^{t\varepsilon})$$
 (1.16)

Solution: This is true for any monotonic function.

2 Chebyschev inequality

2.1 Let

$$Y = (X - E[X])^2 (2.1)$$

and $\varepsilon > 0$. Show using (1.9) that

$$\Pr(Y \geqslant \varepsilon^2) \leqslant \frac{E(Y)}{\varepsilon^2}$$
 (2.2)

2.2 Show that

$$\Pr(Y \ge \varepsilon^2) = \Pr(\sqrt{Y} \ge \varepsilon) + \Pr(\sqrt{Y} \le -\varepsilon)$$
(2.3)

2.3 Show that

$$\Pr\left(\sqrt{Y} \leqslant -\varepsilon\right) = 0,\tag{2.4}$$

2.4 Show that

$$\Pr\left(\sqrt{Y} \ge \varepsilon\right) \le \frac{E(Y)}{\varepsilon^2}$$
 (2.5)

2.5 Show that

$$\Pr(|X - E[X]| \ge \varepsilon) \le \frac{\operatorname{Var}(X)}{\varepsilon^2}$$
 (2.6)

3 Law of Large Numbers (LLN)

3.1 Let

$$S_n = \frac{1}{n} \sum_{i=1}^n X_i \tag{3.1}$$

where X_i are i.i.d r.v. with mean μ and bounded variance σ^2 . Show that

$$E(S_n) = \mu \tag{3.2}$$

$$Var(S_n) = \frac{\sigma^2}{n}$$
 (3.3)

3.2 Using Chebyschev inequality in (2.6), show that

$$\Pr(|S_n - \mu| \ge \varepsilon) \le \frac{\sigma^2}{n\varepsilon^2}$$
 (3.4)

3.3 Show that

$$\lim_{n \to \infty} \Pr(|S_n - \mu| \ge \varepsilon) = 0 \tag{3.5}$$

4 Hoeffding's Lemma

4.1 Consider a r.v. X such that a < X < b. If $\theta = \frac{a}{a-b}$ and E[X] = 0, show that

$$0 < \theta < 1 \tag{4.1}$$

4.2 A convex function g is defined as

$$g(x) \le f(x), \quad x \in (a, b) \tag{4.2}$$

where f is the line joining the points (a, g(a)) and (b, g(b)). Show that

$$E[f(X)] \le \frac{bg(a) - ag(b)}{b - a} \tag{4.3}$$

4.3 Show that

$$g(x) = e^{sx}, s > 0$$
 (4.4)

is convex.

4.4 Using (4.3) and (4.4), show that the moment generating function (MGF) of X

$$M_X(s) = E\left[e^{sX}\right] \le e^{-u\theta} (1 - \theta + \theta e^u), \quad (4.5)$$

where u = s(b - a).

4.5 Let

$$e^{-u\theta} \left(1 - \theta + \theta e^u \right) = e^{\psi(u)} \tag{4.6}$$

Show that

$$\psi(0) = 0 \tag{4.7}$$

$$\psi'(0) = 0 \tag{4.8}$$

$$\psi''(v) = t(1-t), \quad t = \frac{\theta e^{v}}{1-\theta + \theta e^{v}}.$$
 (4.9)

4.6 Show that

$$\psi''(v) \le \frac{1}{4} \tag{4.10}$$

4.7 According to Taylor's theorem,

$$\psi(u) = \psi(0) + u\psi'(0) + \frac{u^2}{2}\psi''(v) \quad 0 < v < u.$$
(4.11)

4.8 Using (4.7), (4.10) and (4.11), show that

$$\psi(u) \le \frac{u^2}{8} \quad 0 < v < u. \tag{4.12}$$

4.9 From (4.5), (4.6) and (4.12), show that

$$M_X(s) \le e^{\frac{s^2(b-a)^2}{8}} \tag{4.13}$$

5 Hoeffding's Inequality

5.1 Let

$$S_n = \sum_{i=1}^n X_i, \quad X_i \in [a_i, b_i], E[X_i] = 0$$
 (5.1)

wherer X_i are independent and

$$Y = S_n - E[S_n]. (5.2)$$

5.2 Using (1.15) and (1.16), show that

$$\Pr(Y \ge t) \le e^{-st} M_Y(s). \tag{5.3}$$

5.3 Show that

$$M_Y(s) = \prod_{i=1}^n M_{X_i}(s)$$
 (5.4)

5.4 From (5.3), (5.5) and (4.13), show that

$$\Pr(Y \ge t) \le \exp\left(-st + \frac{1}{8}s^2||\mathbf{b} - \mathbf{a}||^2\right)$$
 (5.5)

5.5 Show that

$$\min_{s} \left(-st + \frac{1}{8}s^2 ||\mathbf{b} - \mathbf{a}||^2 \right) = -\frac{2t^2}{||\mathbf{b} - \mathbf{a}||^2} \quad (5.6)$$

5.6 Show that

$$\Pr(S_n - E[S_n] \ge t) \le \exp\left(-\frac{2t^2}{\|\mathbf{b} - \mathbf{a}\|^2}\right)$$
 (5.7)

(6.11)

6 Bennet's Inequality

where

for $X_i \leq b$.

6.8 From (6.6) and (6.8), show that

$$v = \sum_{i=1}^{n} E(X_i^2),$$
 (6.9)

$$\sum_{i=1}^{n} E(X_i) \ge 0, \tag{6.10}$$

 $\log M_S(\lambda) \le \frac{v}{h^2} \phi(\lambda b),$

6.1 A real valued r.v X is said to be σ^2 -sub Gaussian if there exists a σ such that

$$M_X(\lambda) < e^{\frac{\lambda^2 \sigma^2}{2}} \tag{6.1}$$

6.2 Let

$$\phi(x) = e^x - x - 1. \tag{6.2}$$

Show that $u^{-2}\phi(u)$ is non-decreasing.

6.3 Show that

$$(\lambda X_i)^{-2} \phi(\lambda X_i) \le \lambda^{-2} \phi(\lambda), \quad \lambda > 0, X_i < 1.$$
(6.3)

6.4 Show that

$$E[e^{\lambda X_i}] \le 1 + \lambda E[X_i] + E[X_i^2]\phi(\lambda)$$
 (6.4)

6.5 Let

$$S = \sum_{i=1}^{n} X_i - E[X_i]. \tag{6.5}$$

Show that

$$\log M_{S}(\lambda) \leq \sum_{i=1}^{n} \log \left[1 + \lambda E(X_{i}) + E(X_{i}^{2}) \phi(\lambda) \right]$$
$$-\lambda \sum_{i=1}^{n} E(X_{i}) \quad (6.6)$$

6.6 Given that log is convex, show that

$$\sum_{i=1}^{n} \log \left[1 + \lambda E(X_i) + E(X_i^2) \phi(\lambda) \right]$$

$$\leq n \log \left[\frac{1}{n} \sum_{i=1}^{n} \left(1 + \lambda E(X_i) + E(X_i^2) \phi(\lambda) \right) \right]$$
(6.7)

6.7 Using the fact that $\log(1 + x) \le x, x \ge 0$, show that

$$n \log \left[\frac{1}{n} \sum_{i=1}^{n} \left(1 + \lambda E(X_i) + E(X_i^2) \phi(\lambda) \right) \right]$$
$$-\lambda \sum_{i=1}^{n} E(X_i) \le v \phi(\lambda) \quad (6.8)$$