

Probability

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Abstract—This book provides solved examples on Probability

1 AXIOMS

1.1. The probability that a given positive integer lying between 1 and 100 (both inclusive) is NOT divisible by 2,3 or 5 is ...

Solution: Table 1.1.1 summarizes the given information.

Event	Definition	Probability
A	$n \equiv 0 \pmod{2}$	$\frac{50}{100}$
B	$n \equiv 0 \pmod{3}$	$\frac{33}{100}$
C	$n \equiv 0 \pmod{5}$	$\frac{20}{100}$
AB	$n \equiv 0 \pmod{6}$	$\frac{16}{100}$
BC	$n \equiv 0 \pmod{15}$	$\frac{6}{100}$
AC	$n \equiv 0 \pmod{10}$	$\frac{10}{100}$
ABC	$n \equiv 0 \pmod{30}$	$\frac{3}{100}$

TABLE 1.1.1: $1 \leq n \leq 100$

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$$\begin{aligned}
 \therefore \Pr(A + B + C) &= \Pr(A) + \Pr(B) + \Pr(C) \\
 &\quad - \Pr(AB) - \Pr(BC) \\
 &\quad - \Pr(AC) + \Pr(ABC) \quad (1.1.1)
 \end{aligned}$$

Substituting from Table 1.1.1 in (1.1.1),

$$\begin{aligned}\Pr(A + B + C) &= \frac{50}{100} + \frac{33}{100} + \frac{20}{100} \\ &\quad - \frac{16}{100} - \frac{6}{100} - \frac{10}{100} + \frac{3}{100} \\ &= \frac{74}{100} \quad (1.1.2)\end{aligned}$$

Thus, the required probability is

$$1 - \Pr(A + B + C) = \frac{26}{100} \quad (1.1.3)$$

1.2. P and Q are considering to apply for a job. The probability that P applies for the job is $\frac{1}{4}$, the probability that P applies for the job given that Q applies for the job is $\frac{1}{2}$, and the probability that Q applies for the job given that P applies for the job is $\frac{1}{3}$. Then the probability that P does not apply for the job given that Q does not apply for the job is

- a) $\frac{4}{5}$ b) $\frac{5}{6}$ c) $\frac{7}{8}$ d) $\frac{11}{12}$

Solution: The given information can be expressed as

$$\Pr(P) = \frac{1}{4} \quad (1.2.1)$$

$$\Pr(P|Q) = \frac{1}{2} = \frac{\Pr(PQ)}{\Pr(Q)} \quad (1.2.2)$$

$$\Pr(Q|P) = \frac{1}{3} = \frac{\Pr(PQ)}{\Pr(P)} \quad (1.2.3)$$

which yields

$$\Pr(PQ) = \frac{1}{3} \times \frac{1}{4} = \frac{1}{12} \quad (1.2.4)$$

$$\Pr(Q) = \frac{\frac{1}{12}}{\frac{1}{3}} = \frac{1}{6}$$

The objective is to find

$$\Pr(P'|Q') \quad (1.2.5)$$

(1.2.1) can be expressed as

$$\Pr(P'|Q') = \frac{\Pr(P'Q')}{\Pr(Q')} \quad (1.2.6)$$

$$= \frac{\Pr(1 - (P + Q)')}{\Pr(Q')} \quad (1.2.7)$$

$$= \frac{1 - \Pr(P) - \Pr(Q) + \Pr(PQ)}{1 - \Pr(Q)} \quad (1.2.8)$$

Substituting from (1.2.4) and (1.2.1) in (1.2.8),

$$\Pr(P'|Q') = \frac{4}{5} \quad (1.2.9)$$

1.3. An experiment consists of two papers. paper1 and paper2. The probability of failing in paper 1 is .3 and that in paper 2 is .2. Given that a student has failed in paper 2, the probability of failing in paper 1 is .6. The probability of student failing in both is

a) .5

b) .18

c) .12

d) .06

Solution: Table 1.3.1 summarises the given

	Description	Probability
0	failure	$\Pr(X = 0) = 0.3$
1	success	$\Pr(Y = 0) = 0.2$
X	Paper 1	$\Pr(X = 0 Y = 0) = 0.6$
Y	Paper 2	

TABLE 1.3.1: Description

information. The desired probability is

$$\Pr(X = 0, Y = 0) = \Pr(X = 0|Y = 0) \Pr(Y = 0) \quad (1.3.1)$$

$$= .12 \quad (1.3.2)$$

2 GEOMETRIC DISTRIBUTION

2.1. Suppose X has density

$$f(x|\theta) = \frac{1}{\theta} e^{-x/\theta}, x > 0 \quad (2.1.1)$$

Define

$$Y = k, \quad k \leq X < k + 1, \quad k = 0, 1, 2, \dots \quad (2.1.2)$$

Then the distribution of Y is

- a) Normal c) Poisson
b) Binomial d) Geometric

Solution:

$$\Pr(Y = k) = \Pr(k \leq X < k + 1) \quad (2.1.3)$$

$$= \int_k^{k+1} f(x|\theta) dx \quad (2.1.4)$$

$$= \int_k^{k+1} \frac{1}{\theta} e^{-\frac{x}{\theta}} dx \quad (2.1.5)$$

$$= \left[-e^{-\frac{x}{\theta}} \right]_k^{k+1} \quad (2.1.6)$$

$$= e^{-\frac{k}{\theta}} \left(1 - e^{-\frac{1}{\theta}} \right) \quad (2.1.7)$$

$$\Rightarrow \Pr(Y = k) = (1 - p)^k p \quad k = 0, 1, 2, \dots \quad (2.1.8)$$

where

$$p = 1 - e^{-\frac{1}{\theta}} \quad (2.1.9)$$

Therefore, the distribution of Y is 4) Geometric.

b) By definition,

$$f_Y(y) = \int_{-\infty}^{\infty} f(x, y) dx \quad (3.1.2)$$

$$= \int_0^y cxy dx \quad (3.1.3)$$

$$= cy \left(\frac{x^2}{2} \right) \Big|_0^y \quad (3.1.4)$$

$$= \frac{cy^3}{2} \quad (3.1.5)$$

$$\Rightarrow f_Y(y) = \begin{cases} \frac{cy^3}{2}, & 0 < y < 1 \\ 0 & \text{otherwise.} \end{cases} \quad (3.1.6)$$

\therefore the area under the pdf is 1, from (3.1.6),

$$\Rightarrow \int_{-\infty}^{\infty} f_Y(y) dy = 1 \quad (3.1.7)$$

$$\Rightarrow \int_0^1 c \frac{y^3}{2} dy = 1 \quad (3.1.8)$$

$$\Rightarrow \frac{c}{8} = 1 \quad (3.1.9)$$

$$\text{or, } c = 8 \quad (3.1.10)$$

Also,

$$f_Y(y) = \begin{cases} 4y^3, & \text{if } 0 < y < 1 \\ 0, & \text{otherwise} \end{cases} \quad (3.1.11)$$

c)

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy \quad (3.1.12)$$

$$= \int_x^1 cxy dy \quad (3.1.13)$$

$$= cx \left(\frac{y^2}{2} \right) \Big|_x^1 \quad (3.1.14)$$

$$= cx \left(\frac{1 - x^2}{2} \right) \quad (3.1.15)$$

$$\Rightarrow f_X(x) = \begin{cases} 4x(1 - x^2), & 0 < x < 1 \\ 0 & \text{otherwise} \end{cases} \quad (3.1.16)$$

3 TWO DIMENSIONS

3.1. Let $c \in \mathbb{R}$ be a constant. Let X, Y be random variables with joint probability density function

$$f(x, y) = \begin{cases} cxy & 0 < x < y < 1, \\ 0 & \text{otherwise} \end{cases} \quad (3.1.1)$$

Which of the following statements are correct ?

- a) $c = \frac{1}{8}$
b) $c = 8$
c) X and Y are independent
d) $\Pr(X = Y) = 0$

Solution:

- a) False

From (3.1.16) and (3.1.11)

$$f_X(x) \times f_Y(y) = \begin{cases} 16xy^3(1-x^2) & , \text{ if } 0 < x, y < 1 \\ 0 & , \text{ otherwise} \end{cases} \quad (3.1.17)$$

$$\neq f(x, y) \quad (3.1.18)$$

Hence, X and Y are not independent.

d)

$$F_X(x) = \int_{-\infty}^x f_X(x) dx \quad (3.1.19)$$

$$= \int_0^x 4x(1-x^2) dx \quad (3.1.20)$$

$$= \int_0^x 4x - 4x^3 dx \quad (3.1.21)$$

$$= 2x^2 - 4x^4 \text{ for } 0 < x < 1 \quad (3.1.22)$$

yielding

$$F_X(x) = \begin{cases} 0 & x \leq 0 \\ 2x^2 - 4x^4 & 0 < x < 1 \\ 1 & x \geq 1 \end{cases} \quad (3.1.23)$$

From (3.1.23),

$$\begin{aligned} \Pr(Y - \epsilon < X < Y + \epsilon) \\ = F_X(Y + \epsilon) - F_X(Y - \epsilon) \\ = 8\epsilon Y(1 - Y^2 - \epsilon^2) \end{aligned} \quad (3.1.24)$$

upon simplification. Letting

$$g(Y) = 8\epsilon Y(1 - Y^2 - \epsilon^2), \quad (3.1.25)$$

$$E[g(Y)] = \int_{-\infty}^{\infty} g(y)f_Y(y) dy \quad (3.1.26)$$

$$= \int_0^1 (4y^3)(8\epsilon y)(1 - y^2 - \epsilon^2) dy \quad (3.1.27)$$

$$\begin{aligned} \implies \Pr(Y - \epsilon < X < Y + \epsilon) \\ = 32\epsilon \left(\frac{2 - 7\epsilon^2}{35} \right) \end{aligned} \quad (3.1.28)$$

Now substituting $\epsilon = 0$ in the above,

$$\Pr(X = Y) = 0 \quad (3.1.29)$$

3.2. Let X and Y be random variables having the

joining probability density function

$$f_{XY}(x, y) = \begin{cases} \frac{1}{\sqrt{2\pi y}} e^{\frac{-1}{2y}(x-y)^2} & x \in (-\infty, \infty), \\ & y \in (0, 1) \\ 0 & \text{otherwise} \end{cases} \quad (3.2.1)$$

The covariance between the random variables X and Y is

Solution:

3.3. Let a random variable X follow exponential distribution with mean 2. Define $Y = [X-2|X > 2]$. The value of $\Pr(Y \geq t)$ is ...

Solution: From the given information,

$$\Pr(Y \geq t) = \frac{\Pr(X - 2 \geq t, X > 2)}{\Pr(X > 2)} \quad (3.3.1)$$

$$= \frac{\Pr(X \geq t + 2, X > 2)}{\Pr(X > 2)} \quad (3.3.2)$$

$\therefore X$ has an exponential distribution with parameter $\lambda = \frac{1}{2}$,

$$F_X(x) = \begin{cases} 1 - e^{-\lambda x}, & \text{if } 0 < x < \infty \\ 0, & \text{otherwise} \end{cases} \quad (3.3.3)$$

and

$$\Pr(X > 2) = 1 - F_X(2) = e^{-2\lambda} \quad (3.3.4)$$

Also,

$$\Pr(X \geq t + 2, X > 2) = \begin{cases} \Pr(X \geq t + 2) & t \geq 0 \\ \Pr(X > 2) & t < 0 \end{cases} \quad (3.3.5)$$

Substituting (3.3.5) in (3.3.2), using (3.3.4) and simplifying,

$$\Pr(Y \geq t) = \begin{cases} e^{-\frac{t}{2}} & t \geq 0 \\ 1 & t < 0 \end{cases} \quad (3.3.6)$$

4 MARKOV CHAIN

4.1. **Step 1.** Flip a coin twice.

Step 2. If the outcomes are (TAILS, HEADS) then output Y and stop.

Step 3. If the outcomes are either (HEADS, HEADS) or (HEADS, TAILS), then output N and stop.

Step 4. If the outcomes are (TAILS, TAILS), then go to Step 1.

The probability that the output of the experiment is Y is (upto two decimal places).....

Solution: The given problem can be repre-

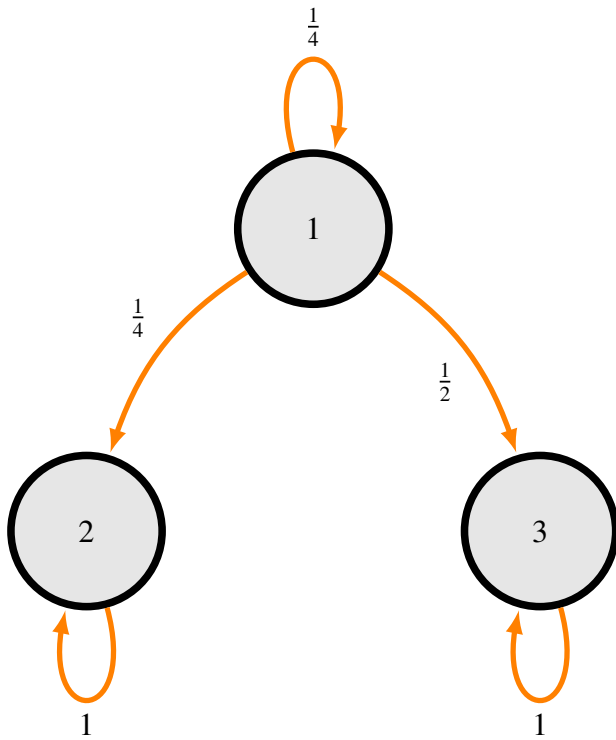


Fig. 4.1.1

sented using Table 4.1.1 and the Markov chain in Fig. 4.1.1. The state transition matrix for the

State	Description
1	$\{T, T\}$
2	$Y = \{T, H\}$
3	$N = \{\{H, H\}, \{H, T\}\}$

TABLE 4.1.1: States and their notations

Markov chain can be expressed as

$$P = \begin{matrix} & \begin{matrix} 2 & 3 & 1 \end{matrix} \\ \begin{matrix} 2 \\ 3 \\ 1 \end{matrix} & \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0.25 & 0.5 & 0.25 \end{bmatrix} \end{matrix} \quad (4.1.1)$$

Clearly, the state 1 is transient, while 2, 3 are absorbing. Comparing (4.1.1) with the standard form of the state transition matrix

$$P = \begin{matrix} & \begin{matrix} A & N \end{matrix} \\ \begin{matrix} A \\ N \end{matrix} & \begin{bmatrix} I & O \\ R & Q \end{bmatrix} \end{matrix} \quad (4.1.2)$$

TABLE 4.1.2: Notations and their meanings

Notation	Meaning
A	All absorbing states
N	All non-absorbing states
I	Identity matrix
O	Zero matrix
R, Q	Other submatrices

where, From (4.1.1) and (4.1.2),

$$R = \begin{pmatrix} 0.25 & 0.5 \end{pmatrix}, Q = \begin{pmatrix} 0.25 \end{pmatrix} \quad (4.1.3)$$

The limiting matrix for absorbing Markov chain is

$$\bar{P} = \begin{pmatrix} I & O \\ FR & O \end{pmatrix} \quad (4.1.4)$$

where

$$F = (I - Q)^{-1} = (1 - 0.25)^{-1} = \frac{4}{3} \quad (4.1.5)$$

is called the fundamental matrix of P . Upon substituting from (4.1.3) in (4.1.5),

$$F = (1 - 0.25)^{-1} = \frac{4}{3} \quad (4.1.6)$$

and

$$FR = \begin{pmatrix} \frac{1}{3} & \frac{2}{3} \end{pmatrix} \quad (4.1.7)$$

which, upon substituting in (4.1.4) yields

$$\bar{P} = \begin{matrix} & \begin{matrix} 2 & 3 & 1 \end{matrix} \\ \begin{matrix} 2 \\ 3 \\ 1 \end{matrix} & \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ \frac{1}{3} & \frac{2}{3} & 0 \end{bmatrix} \end{matrix} \quad (4.1.8)$$

$$\therefore \bar{p}_{12} = \frac{1}{3} \quad (4.1.9)$$