

# Covariance Matrix Adaptation - Evolution Strategy:

## A summary

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### 1. Introduction

The *Covariance Matrix Adaptation - Evolution Strategy* (CMA-ES) is a global and black-box optimization algorithm. It is a randomized black-box search algorithm as it samples evaluation points from a distribution conditioned with the previous parameters. This kind of algorithm is detailed in Figure 1. CMA-ES uses a multivariate Gaussian as the sampling distribution  $\mathcal{N}(m, C)$ . The authors made this choice as, given the variances and covariances between components, the normal distribution has the largest entropy in  $\mathbb{R}^d$ . To goal is to find how update the mean and covariance matrix of this distribution to minimize the trade-off between finding a good approximation of the optimum and evaluate the objective function as few times as possible. In this small paper, we will present the ideas behind CMA-ES. For more details, see (Hansen 2023). Throughout this paper, we will suppose that the objective function is to be maximized.

For  $g$  in  $1 \dots k$ :

1. Let  $d_\theta$  a distribution on  $\mathcal{X}$  parametrized by  $\theta$ ;
2. Sample  $\lambda$  points:  $(x_i)_{1 \leq i \leq \lambda} \sim d_{\theta_i}$ ;
3. Evaluate the points:  $f((x_i)_{1 \leq i \leq \lambda})$ ;
4. Update the parameters  $\theta_{i+1} = F(\theta_i, (x_1, f(x_1)), \dots, (x_\lambda, f(x_\lambda)))$ .

Algorithm 1: Black-box search algorithm.

### 2. Update the mean

In the CMA Evolution Strategy, the  $\lambda$  points are sampled from a multivariate Gaussian distribution which writes:

$$(x_i^{(g)})_{1 \leq i \leq \lambda} \sim m^{(g)} + \sigma^{(g)} \mathcal{N}(0, C^{(g)}), \quad (1)$$

where  $g$  is the generation number,  $m^{(g)}$  is the mean vector,  $\sigma^{(g)}$  is the “overall” standard deviation and  $C^{(g)}$  is the covariance matrix. It is equivalent to say that  $x_i^{(g)} \sim \mathcal{N}(m^{(g)}, (\sigma^{(g)})^2 C^{(g)})$ .

To update the mean, we begin by selecting the  $\mu$  best points, i.e. :

$$f(x_1^{(g)}) \geq \dots \geq f(x_\mu^{(g)}) \geq f(x_{\mu+1}^{(g)}) \geq \dots f(x_\lambda^{(g)}). \quad (2)$$

We introduce the index notation  $i : \lambda$ , denoting the index of the  $i$ -th best point. The mean at generation  $g + 1$  becomes a weighted average of those points:

$$m^{(g+1)} = m^{(g)} + c_m \sum_{i=1}^{\mu} w_i (x_{i:\lambda} - m^{(g)}), \quad (3)$$

where:

$$\sum_{i=1}^{\mu} w_i = 1, \quad w_1 \geq \dots \geq w_{\mu} \geq 0, \quad (4)$$

and  $c_m$  is a learning rate, usually set to 1. In that case, Eq. 3 simply becomes:

$$m^{(g+1)} = \sum_{i=1}^{\mu} w_i x_{i:\lambda}. \quad (5)$$

The choice of the weights is crucial in CMA-ES as they represent the trade-off between exploration and exploitation. To do so, we define the quantity  $\mu_{\text{eff}}$  as:

$$\mu_{\text{eff}} = \left( \frac{\|w\|_1}{\|w\|_2} \right)^2 = \frac{\left( \sum_{i=1}^{\mu} |w_i| \right)^2}{\sum_{i=1}^{\mu} w_i^2} = \frac{1}{\sum_{i=1}^{\mu} \frac{w_i^2}{w_i}}. \quad (6)$$

From Eq. 4, one can easily derive  $1 \leq \mu_{\text{eff}} \leq \mu$ , the latter happens when all the weights are equal, i.e.  $\forall 1 \leq i \leq \mu, w_i = \frac{1}{\mu}$ .  $\mu_{\text{eff}}$  quantize the loss of variance due to the selection of the best points. According to the author,  $\mu_{\text{eff}} \approx \frac{\lambda}{4}$  indicates a reasonable choice of  $w_i$ . A simple and decent way to achieve that is to set  $w_i \propto \mu - i + 1$  (see Eq. 10). Choosing  $c_m < 1$  can work well on noisy function. However, the step-size  $\sigma$  is roughly proportional to  $\frac{1}{c_m}$  and thus, with a too small  $c_m$ , the search would moves away from the current region of relevance.

### 3. Update the covariance matrix

To update the covariance matrix, we need to estimate it using the points  $(x_i)_{1 \leq i \leq \lambda}$ . In this section, we assume  $\sigma = 1$  for simplicity. If  $\sigma \neq 1$ , one can simply rescale the covariance matrix by  $\frac{1}{\sigma^2}$ . If we have enough sample, one can use the empirical covariance matrix:

$$C_{\text{emp}}^{(g+1)} = \frac{1}{\lambda - 1} \sum_{i=1}^{\lambda} \left( x_i^{(g+1)} - \frac{1}{\lambda} \sum_{i=1}^{\lambda} x_i^{(g+1)} \right) \left( x_i^{(g+1)} - \frac{1}{\lambda} \sum_{i=1}^{\lambda} x_i^{(g+1)} \right)^{\top}. \quad (7)$$

A different would be to use the real mean  $m^{(g+1)}$  computed before instead of the empirical mean:

$$C_{\lambda}^{(g+1)} = \frac{1}{\lambda} \sum_{i=1}^{\lambda} \left( x_i^{(g+1)} - m^{(g+1)} \right) \left( x_i^{(g+1)} - m^{(g+1)} \right)^{\top}. \quad (8)$$

Both are unbiased estimators of the covariance matrix. However, they do not influence the search towards the direction of the  $\mu$  best points. To do so, one can use the same *weighted selection* as in Eq. 3 :

$$C_{\mu}^{(g+1)} = \sum_{i=1}^{\mu} w_i \left( x_i^{(g+1)} - m^{(g+1)} \right) \left( x_i^{(g+1)} - m^{(g+1)} \right)^{\top}. \quad (9)$$

This last estimator tends to reproduce the current best points and thus allows a faster convergence. However, this estimation method requires a lot of samples and  $\mu_{\text{eff}}$  must be large enough to be reliable. The author suggests another method to estimate  $C^{(g+1)}$  that tackles these two issues, the *rank- $\mu$*  method.

### 3.1. Rank- $\mu$ method

## Bibliography

N. Hansen, “The CMA Evolution Strategy: A Tutorial,” arXiv, 2023. Accessed: Apr. 4, 2023. [Online]. Available: <http://arxiv.org/abs/1604.00772> (Comment: ArXiv e-prints, arXiv:1604.00772, 2016, pp.1-39)

#### 4. Annex

Let  $w_i = \frac{\mu-i+1}{\sum_{i=1}^{\mu} \mu-i+1}$ . Then:

$$\begin{aligned}
 \mu_{\text{eff}} &= \frac{1}{\sum_{i=1}^{\mu} \left( \frac{\mu-i+1}{\sum_{i=1}^{\mu} \mu-i+1} \right)^2} = \frac{1}{\sum_{i=1}^{\mu} \left( \frac{i}{\sum_{i=1}^{\mu} i} \right)^2} \\
 &= \frac{1}{\sum_{i=1}^{\mu} \frac{i^2}{\left( \frac{\mu(\mu+1)}{2} \right)^2}} = \frac{1}{\frac{\mu(\mu+1)(2\mu+1)}{6 \left( \frac{\mu(\mu+1)}{2} \right)^2}} \\
 &= \frac{6 \left( \frac{\mu(\mu+1)}{2} \right)^2}{\mu(\mu+1)(2\mu+1)} = \frac{3\mu(\mu^3 + 2\mu^2 + \mu)}{2(2\mu^3 + 3\mu^2 + \mu)} \\
 &= \frac{3\mu(1 + \mu)}{2(1 + 2\mu)} \approx \frac{3\frac{\lambda}{2} \left( 1 + \frac{\lambda}{2} \right)}{2(1 + \lambda)} \\
 &= \frac{\frac{3\lambda^2 + 6\lambda}{2}}{4(1 + \lambda)} = \frac{3\lambda(2 + \lambda)}{8(1 + \lambda)} \\
 &\approx \frac{3\lambda}{8}.
 \end{aligned} \tag{10}$$