

## FraseR autoencoder gradients

Beta binomial density:

$$\begin{aligned} \text{BetaBin}(k|n, \alpha, \beta) &= \binom{n}{k} \cdot \frac{\text{Be}(\alpha + k, \beta + n - k)}{\text{Be}(\alpha, \beta)} \\ &= \frac{\Gamma(n+1)}{\Gamma(k+1)\Gamma(n-k+1)} \cdot \frac{\Gamma(\alpha+k)\Gamma(\beta+n-k)}{\Gamma(\alpha+\beta+n)} \cdot \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} \end{aligned}$$

Parametrization with  $\mu$  and  $\rho$ :

$$\begin{aligned} \mu &= \frac{\alpha}{\alpha + \beta} & \rho &= \frac{1}{1 + \alpha + \beta} \\ \Rightarrow \alpha &= \mu \left( \frac{1 - \rho}{\rho} \right) & \beta &= (\mu - 1) \left( \frac{\rho - 1}{\rho} \right) \end{aligned}$$

$\alpha + \beta$  is only dependent on  $\rho$ :

$$\begin{aligned} \alpha + \beta &= \mu \left( \frac{1}{\rho} - 1 \right) + \frac{(\mu - 1)(\rho - 1)}{\rho} = \frac{\mu(1 - \rho)}{\rho} + \frac{-(1 - \mu) - (1 - \rho)}{\rho} \\ &= \frac{\mu(1 - \rho) + (1 - \mu)(1 - \rho)}{\rho} = \frac{(1 - \rho)(\mu + (1 - \mu))}{\rho} \\ &= \frac{1 - \rho}{\rho} = \frac{1}{\rho} - 1 \end{aligned}$$

Definition of  $\mu_{ij}$  and  $x_{ij}$ :

$$\begin{aligned} Y &= XW_e W_d^T + b \\ \text{logit}(\mu_{ij}) &= y_{ij} \Rightarrow \mu_{ij} = \frac{e^{y_{ij}}}{1 + e^{y_{ij}}} \\ x_{ij} &= \text{logit} \left( \frac{k_{ij} + \frac{1}{2}}{n_{ij} + 1} \right) \end{aligned}$$

Negative log-likelihood:

$$\begin{aligned} nll &= -\log \left( \frac{\Gamma(n+1)}{\Gamma(k+1)\Gamma(n-k+1)} \cdot \frac{\Gamma(\alpha+k)\Gamma(\beta+n-k)}{\Gamma(\alpha+\beta+n)} \cdot \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} \right) \\ &= -\log \left( \frac{\Gamma(n+1)}{\Gamma(k+1)\Gamma(n-k+1)} \right) - \log \left( \frac{\Gamma(\alpha+k)\Gamma(\beta+n-k)}{\Gamma(\alpha+\beta+n)} \right) - \log \left( \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} \right) \\ &= -\log(\Gamma(n+1)) + \log(\Gamma(k+1)) + \log(\Gamma(n-k+1)) - \log(\Gamma(\alpha+k)) - \log(\Gamma(\beta+n-k)) \\ &\quad + \log(\Gamma(\alpha+\beta+n)) + \log(\Gamma(\alpha)) + \log(\Gamma(\beta)) - \log(\Gamma(\alpha+\beta)) \end{aligned}$$

$$nll_{\mathbf{W}} = \log(\Gamma(\alpha)) + \log(\Gamma(\beta)) - \log(\Gamma(\alpha+k)) - \log(\Gamma(\beta+n-k))$$

NLL with pseudocounts ( $k+1, n+2$ ):

$$nll_{\mathbf{W}} = \log(\Gamma(\alpha)) + \log(\Gamma(\beta)) - \log(\Gamma(\alpha+k+1)) - \log(\Gamma(\beta+n-k+1))$$

Full form of  $\log(\Gamma(\alpha))$ :

$$\begin{aligned}\log(\Gamma(\alpha)) &= \log\left(\Gamma\left(\mu \frac{1-\rho}{\rho}\right)\right) = \log\left(\Gamma\left(\frac{e^Y}{1+e^Y} \frac{1-\rho}{\rho}\right)\right) \\ &= \log\left(\Gamma\left(\left(\frac{e^{XW_e W_d^T + b}}{1+e^{XW_e W_d^T + b}}\right) \left(\frac{1-\rho}{\rho}\right)\right)\right)\end{aligned}$$

Derivatives:

$$\begin{aligned}\frac{d}{d\alpha} [\log(\Gamma(\alpha))] &= \psi^{(0)}(\alpha) \\ \frac{d}{d\mu} \left[ \mu \left( \frac{1-\rho}{\rho} \right) \right] &= \frac{1-\rho}{\rho} \\ \frac{d}{dY} \left[ \frac{e^Y}{1+e^Y} \right] &= \frac{e^Y}{(1+e^Y)^2} \\ \frac{d}{dW_d} [XW_e W_d^T + b] &= XW_e \\ \frac{d}{dW_e} [XW_e W_d^T + b] &= X^T W_d \\ \frac{d}{db} [XW_e W_d^T + b] &= 1\end{aligned}$$

Derivatives of the first term of the negative log-likelihood:

$$\begin{aligned}\frac{d}{dW_d} \log(\Gamma(\alpha)) &= \frac{d}{dW_d} \log\left(\Gamma\left(\left(\frac{e^Y}{1+e^Y}\right) \left(\frac{1-\rho}{\rho}\right)\right)\right) \\ &= \left(\psi\left(\left(\frac{e^Y}{1+e^Y}\right) \left(\frac{1-\rho}{\rho}\right)\right) \cdot \frac{1-\rho}{\rho} \cdot \frac{e^Y}{(1+e^Y)^2}\right)^T XW_e\end{aligned}$$

$$\begin{aligned}\frac{d}{dW_e} \log(\Gamma(\alpha)) &= \frac{d}{dW_e} \log\left(\Gamma\left(\left(\frac{e^Y}{1+e^Y}\right) \left(\frac{1-\rho}{\rho}\right)\right)\right) \\ &= X^T \left(\psi\left(\left(\frac{e^Y}{1+e^Y}\right) \left(\frac{1-\rho}{\rho}\right)\right) \cdot \frac{1-\rho}{\rho} \cdot \frac{e^Y}{(1+e^Y)^2}\right) W_d\end{aligned}$$

$$\begin{aligned}\frac{d}{db} \log(\Gamma(\alpha)) &= \frac{d}{db} \log\left(\Gamma\left(\left(\frac{e^Y}{1+e^Y}\right) \left(\frac{1-\rho}{\rho}\right)\right)\right) \\ &= \left(\psi\left(\left(\frac{e^Y}{1+e^Y}\right) \left(\frac{1-\rho}{\rho}\right)\right) \cdot \frac{1-\rho}{\rho} \cdot \frac{e^Y}{(1+e^Y)^2}\right)^T\end{aligned}$$

Derivatives of the second term of the negative log-likelihood:

$$\begin{aligned}\frac{d}{dW_d} \log(\Gamma(\beta)) &= \frac{d}{dW_d} \log \left( \Gamma \left( \left( \frac{e^Y}{1+e^Y} - 1 \right) \left( \frac{\rho-1}{\rho} \right) \right) \right) \\ &= \left( \psi \left( \left( \frac{e^Y}{1+e^Y} - 1 \right) \left( \frac{\rho-1}{\rho} \right) \right) \cdot \frac{\rho-1}{\rho} \cdot \frac{e^Y}{(1+e^Y)^2} \right)^T XW_e\end{aligned}$$

$$\begin{aligned}\frac{d}{dW_e} \log(\Gamma(\beta)) &= \frac{d}{dW_e} \log \left( \Gamma \left( \left( \frac{e^Y}{1+e^Y} - 1 \right) \left( \frac{\rho-1}{\rho} \right) \right) \right) \\ &= X^T \left( \psi \left( \left( \frac{e^Y}{1+e^Y} - 1 \right) \left( \frac{\rho-1}{\rho} \right) \right) \cdot \frac{\rho-1}{\rho} \cdot \frac{e^Y}{(1+e^Y)^2} \right) W_d\end{aligned}$$

$$\begin{aligned}\frac{d}{db} \log(\Gamma(\beta)) &= \frac{d}{db} \log \left( \Gamma \left( \left( \frac{e^Y}{1+e^Y} - 1 \right) \left( \frac{\rho-1}{\rho} \right) \right) \right) \\ &= \left( \psi \left( \left( \frac{e^Y}{1+e^Y} - 1 \right) \left( \frac{\rho-1}{\rho} \right) \right) \cdot \frac{\rho-1}{\rho} \cdot \frac{e^Y}{(1+e^Y)^2} \right)^T\end{aligned}$$

Derivatives of the third term of the negative log-likelihood:

$$\begin{aligned}\frac{d}{dW_d} \log(\Gamma(\alpha + k + 1)) &= \frac{d}{dW_d} \log \left( \Gamma \left( \left( \frac{e^Y}{1+e^Y} \right) \left( \frac{1-\rho}{\rho} \right) \right) + k + 1 \right) \\ &= \left( \psi \left( \left( \frac{e^Y}{1+e^Y} \right) \left( \frac{1-\rho}{\rho} \right) + k + 1 \right) \cdot \frac{1-\rho}{\rho} \cdot \frac{e^Y}{(1+e^Y)^2} \right)^T XW_e\end{aligned}$$

$$\begin{aligned}\frac{d}{dW_e} \log(\Gamma(\alpha + k + 1)) &= \frac{d}{dW_e} \log \left( \Gamma \left( \left( \frac{e^Y}{1+e^Y} \right) \left( \frac{1-\rho}{\rho} \right) \right) + k + 1 \right) \\ &= X^T \left( \psi \left( \left( \frac{e^Y}{1+e^Y} \right) \left( \frac{1-\rho}{\rho} \right) + k + 1 \right) \cdot \frac{1-\rho}{\rho} \cdot \frac{e^Y}{(1+e^Y)^2} \right) W_d\end{aligned}$$

$$\begin{aligned}\frac{d}{db} \log(\Gamma(\alpha + k + 1)) &= \frac{d}{db} \log \left( \Gamma \left( \left( \frac{e^Y}{1+e^Y} \right) \left( \frac{1-\rho}{\rho} \right) \right) + k + 1 \right) \\ &= \left( \psi \left( \left( \frac{e^Y}{1+e^Y} \right) \left( \frac{1-\rho}{\rho} \right) + k + 1 \right) \cdot \frac{1-\rho}{\rho} \cdot \frac{e^Y}{(1+e^Y)^2} \right)^T\end{aligned}$$

Derivatives of the fourth term of the negative log-likelihood:

$$\begin{aligned}\frac{d}{dW_d} \log(\Gamma(\beta + n - k + 1)) &= \frac{d}{dW_d} \log \left( \Gamma \left( \left( \frac{e^Y}{1 + e^Y} - 1 \right) \left( \frac{\rho - 1}{\rho} \right) + n - k + 1 \right) \right) \\ &= \left( \psi \left( \left( \frac{e^Y}{1 + e^Y} - 1 \right) \left( \frac{\rho - 1}{\rho} \right) + n - k + 1 \right) \cdot \frac{\rho - 1}{\rho} \cdot \frac{e^Y}{(1 + e^Y)^2} \right)^T X W_e\end{aligned}$$

$$\begin{aligned}\frac{d}{dW_e} \log(\Gamma(\beta + n - k + 1)) &= \frac{d}{dW_e} \log \left( \Gamma \left( \left( \frac{e^Y}{1 + e^Y} - 1 \right) \left( \frac{\rho - 1}{\rho} \right) + n - k + 1 \right) \right) \\ &= X^T \left( \psi \left( \left( \frac{e^Y}{1 + e^Y} - 1 \right) \left( \frac{\rho - 1}{\rho} \right) + n - k + 1 \right) \cdot \frac{\rho - 1}{\rho} \cdot \frac{e^Y}{(1 + e^Y)^2} \right) W_d\end{aligned}$$

$$\begin{aligned}\frac{d}{db} \log(\Gamma(\beta + n - k + 1)) &= \frac{d}{db} \log \left( \Gamma \left( \left( \frac{e^Y}{1 + e^Y} - 1 \right) \left( \frac{\rho - 1}{\rho} \right) + n - k + 1 \right) \right) \\ &= \left( \psi \left( \left( \frac{e^Y}{1 + e^Y} - 1 \right) \left( \frac{\rho - 1}{\rho} \right) + n - k + 1 \right) \cdot \frac{\rho - 1}{\rho} \cdot \frac{e^Y}{(1 + e^Y)^2} \right)^T\end{aligned}$$