

11-2015-10-28-gail-trying-to-match-the-table-for-09-kernelPvals092415

I began by following [gitboxes/b-blueb/trunk/docum/2015/grits/b-2015-05-11-tracking-down-independent-independent/19-NL-and-GBPvals051715.pdf](#). Now I am following [38-NL&GBPvals062415.pdf](#) and [38-NL&GBPvals062415Tbl1.pdf](#)

After moving to “/Users/ggong/aaa-ebony/43-gritsr2/gritsr2/gails-stuff/b-examples/02-v-9010/a-docum” I am following [/Volumes/2015a-stanford/aaa/gitboxes/b-blueb/trunk/docum/2015/grits/e-2015-10-13-grits-competitor/03-kernelPvals091815.pdf](#)

/Users/ggong/aaa-ebony/43-gritsr2/gritsr2/gails-stuff/b-examples/02-v-9010/a-docum

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1 davies_fn

Given the J dimensional random vector (of functions of genotypes) $Z \sim N(\mu_Z, V_Z)$ and the matrix A , the p-value for the statistic $Q = Z^T A Z$ can be gotten from `davies_fn(zzz, mu_z, V_z, AAA)`

This function performs the following calculations

$$\tilde{Z} = V_Z^{-1/2} Z \sim N(\mu_{\tilde{Z}} = V_Z^{-1/2} \mu_Z, I).$$

$$Q = Z^T A Z = Z^T V^{-1/2} V^{1/2} A V^{1/2} V^{-1/2} Z = \tilde{Z}^T \tilde{A} \tilde{Z}$$

$$\tilde{A} = V^{1/2} A V^{1/2}$$

$\tilde{A} = U^T \Lambda U$. The spectral decomposition of \tilde{A} , so Λ is a diagonal matrix containing the eigenvalues of \tilde{A} , and $U^T U = U U^T = I$.

$$Q = \tilde{Z}^T \tilde{A} \tilde{Z} = \tilde{Z}^T U^T \Lambda U \tilde{Z} = X^T \Lambda X$$

$$X = U \tilde{Z} \sim N(\mu_X = U \mu_{\tilde{Z}}, U I U^T = I)$$

$Q = X^T \Lambda X = \sum_{j=1}^J \lambda_j X_j^2$ is a sum of independent χ^2 random variables with noncentrality parameters μ_X^2 .

The R package `CompQuadForm` provides the function `davies` which computes $P(Q > q)$ where $Q = \sum_{j=1}^J \lambda_j X_j + \sigma X_0$ where X_j are independent random variables having a non-central χ^2 distribution with n_j degrees of freedom and noncentrality parameter δ_j^2 , and X_0 having a standard normal distribution.

In our case, the degrees of freedom is $n_j = 1$, and the noncentrality parameter is $\delta_j^2 = \mu_X^2$.

2 The genotype matrix

G is the genotype matrix.

$$E(G_{nm}) = 2\pi_m$$

$\Gamma = \text{cov}(G_n)$, a covariance matrix of dimension $M \times M$. I can calculate the empirical covariance matrix by `cov(genotype)`.

Ψ is 2 times the kinship matrix. $\Psi_{n_1 n_2} = \text{Cor}(G_{n_1 m}, G_{n_2 m})$ while holding m equal to any number in $1, \dots, M$.

$$\text{Cov}(G_{n_1 m_1}, G_{n_2 m_2}) = \Gamma_{m_1 m_2} \Psi_{n_1 n_2}$$

$W = W_{M \times M}$ is the diagonal matrix of weights.

3 Y_1 , Y_2 , N_{cases} and e

Y_1 is a vector of dimension N , the n -th element being the indicator for disease of the n -th person.

N_{cases} is the number of cases.

Y_2 is a vector of outcome predictors. By design, each element of this vector is equal to N_{cases}/N .

$e = Y_1 - Y_2$ is the vector of residuals.

4 Couch everything in terms if $Z = Z^{(2)}$

$$Z = \begin{pmatrix} Z_1 \\ Z_2 \end{pmatrix} = \begin{pmatrix} WGy \\ WGp \end{pmatrix}$$

5 Lines 1, 2, and 5 of Alice's Table 1

$$\begin{aligned}
Z^{(1)} &= Z_1 - Z_2 \\
A_{\text{burden}}^{(1)} &= J_M \\
A_{\text{skat}}^{(1)} &= I_M \\
A_{\text{lee}}^{(1)} &= (1 - \rho)A_{\text{burden}}^{(1)} + \rho A_{\text{skat}}^{(1)} = (1 - \rho)J_M + \rho I_M
\end{aligned}$$

6 Lines 3, 4, and 6 of Alice's Table 1

$$\begin{aligned}
Z^{(2)} &= Z \\
A_{\text{burden}}^{(2)} &= \begin{pmatrix} J_M & 0 \\ 0 & -J_M \end{pmatrix} = \Delta(J_M, -J_M) \\
A_{\text{skat}}^{(2)} &= \begin{pmatrix} I_M & 0 \\ 0 & -I_M \end{pmatrix} = \Delta(I_M, -I_M) \\
A_{\text{lee}}^{(2)} &= (1 - \rho)A_{\text{burden}}^{(2)} + \rho A_{\text{skat}}^{(2)} = (1 - \rho)\Delta(J_M, -J_M) + \rho\Delta(I_M, -I_M)
\end{aligned}$$

7 Lines 7, 8, and 9 of Alice's Table 1

$$\begin{aligned}
Z^{(3)} &= \begin{pmatrix} Z_1 - Z_2 \\ Z_1 \\ Z_2 \end{pmatrix} \\
A_{\text{burden}}^{(3)} &= \begin{pmatrix} (1 - \tau_b)A_{\text{burden}}^{(1)} & 0 \\ 0 & \tau_b A_{\text{burden}}^{(2)} \end{pmatrix} = \Delta\{(1 - \tau_b)J_M, \tau_b \Delta(J_M, -J_M)\} \\
A_{\text{skat}}^{(3)} &= \begin{pmatrix} (1 - \tau_s)A_{\text{skat}}^{(1)} & 0 \\ 0 & \tau_s A_{\text{skat}}^{(2)} \end{pmatrix} = \Delta\{(1 - \tau_s)I_M, \tau_s \Delta(I_M, -I_M)\} \\
A_{\text{lee}}^{(3)} &= (1 - \rho)A_{\text{burden}}^{(3)} + \rho A_{\text{skat}}^{(3)} \\
&= \Delta\{(1 - \rho)(1 - \tau_b)J_M + \rho(1 - \tau_s)I_M, (1 - \rho)\tau_b \Delta(J_M, -J_M) + \rho\tau_s \Delta(I_M, -I_M)\} \\
&= \Delta\{\beta_1 J_M + \beta_2 I_M, \beta_3 \Delta(J_M, -J_M) + \beta_4 \Delta(I_M, -I_M)\}
\end{aligned}$$

Another way to generalize is like this

$$\begin{aligned}
A_{\text{another}}^{(3)} &= \begin{pmatrix} (1 - \tau)A_{\text{lee}}^{(1)}(\rho^{(1)}) & 0 \\ 0 & \tau A_{\text{lee}}^{(2)}(\rho^{(2)}) \end{pmatrix} = \Delta\{(1 - \tau)A_{\text{lee}}^{(1)}(\rho^{(1)}), \tau A_{\text{lee}}^{(2)}(\rho^{(2)})\} \\
&= \Delta\{(1 - \tau)(1 - \rho^{(1)})J_M + (1 - \tau)\rho^{(1)}I_M, \tau(1 - \rho^{(2)})\Delta(J_M, -J_M) + \tau\rho^{(2)}\Delta(I_M, -I_M)\} \\
&= \Delta\{\beta_1 J_M + \beta_2 I_M, \beta_3 \Delta(J_M, -J_M) + \beta_4 \Delta(I_M, -I_M)\}
\end{aligned}$$

Record the following in case I need it

$$\begin{aligned}
A_{\text{lee}}^{(3)} &= \Delta \{A_1^{(3)}, A_2^{(3)}, A_3^{(3)}\} \\
A_1^{(3)} &= (1 - \rho)(1 - \tau_b)J_M + \rho(1 - \tau_s)I_M \\
A_2^{(3)} &= (1 - \rho)\tau_b J_M + \rho\tau_s I_M \\
A_3^{(3)} &= -A_2^{(3)}
\end{aligned}$$

$$\begin{aligned}
A_{\text{another}}^{(3)} &= \Delta \{A_1^{(4)}, A_2^{(4)}, A_3^{(4)}\} \\
A_1^{(4)} &= (1 - \tau)(1 - \rho^{(1)})J_M + (1 - \tau)\rho^{(1)}I_M \\
A_2^{(4)} &= \tau(1 - \rho^{(2)})J_M + \tau\rho^{(2)}I_M \\
A_3^{(4)} &= -A_2^{(4)}
\end{aligned}$$

8 To get the distribution of $Q = Z^T A Z$.

For lines 7, 8, and 9 of Alice's Table 1, I note that Z^3 has covariance that is singular, but I can write

$$\begin{aligned}
Z^{(3)} &= \begin{pmatrix} Z_1 - Z_2 \\ Z_1 \\ Z_2 \end{pmatrix} = BZ \\
B &= \begin{pmatrix} I_M & -I_M \\ I_M & 0 \\ 0 & I_M \end{pmatrix} \\
Q &= Z^{(3)T} A^{(3)} Z^{(3)} = Z^T B^T A^{(3)} B Z = Z^T A Z \\
A &= B^T A^{(3)} B = \begin{pmatrix} A_1 + A_2 & -A_1 \\ -A_1 & A_1 - A_2 \end{pmatrix} \\
A^{(3)} &= \Delta \{A_1, A_2, -A_2\} \\
A_1 &= \beta_1 J_M + \beta_2 I_M \\
A_2 &= \beta_3 J_M + \beta_4 I_M
\end{aligned}$$