

# **Stats 519: HW 5 - EFA (CH 5)**

Due on March 27, 2009

*Dr. Stephen Lee 1:30*

**Monte J. Shaffer**

## Problem 0

Can the PC analysis result generalize well outside the original sample data? To answer this question, let us consider bootstrapping the original data to assess the validity of the PC analysis as follows: Resample 50 observations with replacement and form the linear combination of the bootstrapped data using the 1st PC direction of the original data set (i.e., `scale(gsp.share)`). Compute the variance of this linear combination. Compare this variance with the variance of the 1st PC from the bootstrapped data sample. Repeat this process 1000 times to produce Figure 4.18 (pg 120) Lattin et al. and comment on the validation of the PCA result on the data set `gsp.share`.

---

### Program 1 PCA for 0

---

```
setwd("C:/latex/statsMultiVariate/datasets");
gsp.share = read.table("GSP_SHARE.txt");
myColumns = c("AGRICULTURE", "MINING", "CONSTRUC", "MFR_DUR", "MFR_NON", "TRANSPORT", "COMMUN", "UTILITIES", "W
myStates = gsp.share[,1];

Xs= scale(gsp.share[,2:14]); rownames(Xs) = myStates; colnames(Xs) = myColumns;
Xs.PCA=princomp(Xs);          rownames(Xs.PCA$loadings)=myColumns; rownames(Xs.PCA$scores)=myStates;
summary(Xs.PCA);

# PCA Variable Factor Map
library(FactoMineR);
result = PCA(Xs); # graphs generated automatically

biplot(Xs.PCA);

## nFactors
n=dim(Xs)[1];p=dim(Xs)[2];
library(nFactors);
nResults=nScree(eig = as.numeric(Xs.PCA$sdev),aparallel = parallel(subject=n,var=p)$eigen$qevpea);
plotnScree(nResults, main="Component Retention Analysis");

# r_1 is x*u_1/x*u_1*          # r_2 is xu_1*/xu_1

nsim = 1000;
Xorig = as.matrix(Xs);
u_1 = as.matrix(Xs.PCA$loadings[,1]);
r_1 = r_2 = c();
for(i in 1:nsim)
{
  Xboot = Xs[sample(1:50,50,replace=T),];
  Xboot.PCA = princomp(Xboot);
  u_1_star = as.matrix(Xboot.PCA$loadings[,1]);

  r_1[i] = var(Xboot%*%u_1)          /    var(Xboot%*%u_1_star);
  r_2[i] = var(Xorig%*%u_1_star)    /    var(Xorig%*%u_1);
}
hist(r_1,breaks=33,main=nsim);
hist(r_2,breaks=33,main=nsim);
```

---

*Proof.* Proof. Two ratios can be used to try and explain what is captured in this exercise:

$$r_1 = \frac{\text{var}(X_{boot}^* \cdot u_1)}{\text{var}(X_{boot}^* \cdot u_1^*)} \text{ and } r_2 = \frac{\text{var}(X_S \cdot u_1^*)}{\text{var}(X_S \cdot u_1)}$$

where  $u_1 = \text{as.matrix}(X_S.PCA\$loadings[, 1])$  and  $u_1^* = \text{as.matrix}(X_{boot}.PCA\$loadings[, 1])$ , the first column of eigen values. FIGURE 1 shows summary of the PCA of the original scales ( $X_S$ ) data.

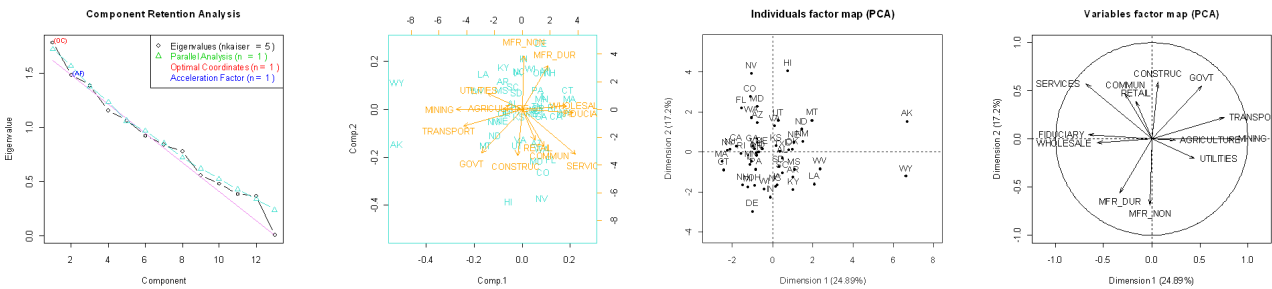


Figure 1: Basic PCA Explanation: Scree, BiPlot, Factor Maps using princomp

Results are reported for different bootstrap simulations (from 100 to 100,000).

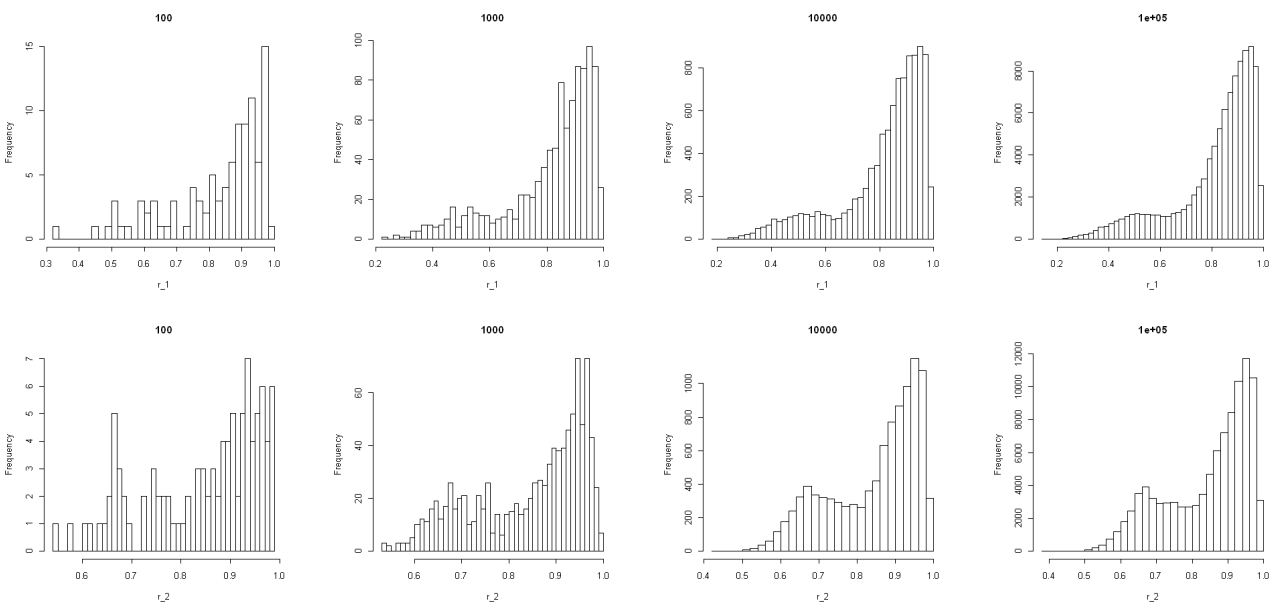


Figure 2: Comparing  $r_1$  and  $r_2$  for different number of bootstraps

From pg. 119–120, bootstrapping will work if we assume the data is representative of the true distribution. Is the variation explained by the first PC due to chance? To answer this we use one of the ratios above. If the ratio is close to 1 – we can conclude that the variation captured is systematic to the sample (and boot), thereby concluding generalizability. As seen above, the mass of each ratio is close to one, so if the assumptions underlying bootstrapping hold, we can establish generalizability. As noted in the text, the values for Alaska and Wyoming, if not included in the bootstrap draw, change the variance of PC1.

□

## Problem 1

The MBA\_CAR dataset may be found on the website. This has ratings on 16 attributes of 10 different car models 1 through 10, from MBA students. There are some missing values, so use `read.table` with the `na.strings='.'` option, and `na.omit` to remove the incomplete observations. [See problem 5.7, pg 169 of Lattin] Save the car variable separately, and remove it from your data matrix. Conduct a principal components analysis on this data. How many components would you keep? Can you interpret the loadings on those components meaningfully? Does rotation of the loadings for the components you are keeping improve interpretability?

---

### Program 2 PCA for 0

---

```
setwd("C:/latex/statsMultiVariate/datasets");
mba.cars = na.omit(read.table("MBA_CAR_ATTRIB.txt",na.strings='.'));
  myCars = c("BMW 328i","Ford Explorer","Infiniti J30","Jeep Grand Cherokee","Lexus ES300","Chrysler
  myAttributes = c("Exciting","Dependable","Luxurious","Outdoorsy","Powerful","Stylish","Comfortable

cars = mba.cars[,3:18];          colnames(cars) = myAttributes;

Cs= scale(cars);                colnames(Cs) = myAttributes;
Cs.PCA=princomp(Cs);            rownames(Cs.PCA$loadings)=myAttributes;
  summary(Cs.PCA);

plot(Cs.PCA$scores[,1],Cs.PCA$scores[,2]);
  library(rgl);
plot3d(Cs.PCA$scores[,1:3], type="s", radius=.1, col=rainbow(10));

# PCA Variable Factor Map
library(FactoMineR)
result = PCA(Cs) # graphs generated automatically

## nFactors
n=dim(Cs)[1];p=dim(Cs)[2];
library(nFactors);
  nResults = nScree(eig = as.numeric(Cs.PCA$sdev),aparallel = parallel(subject = n, var = p)$eigen$q
plotuScree(as.numeric(Cs.PCA$sdev)); ## basic scree
plotnScree(nResults, main="Component Retention Analysis");

biplot(Cs.PCA);

Cs.PCA$loadings;
Cs.rotatedLoadingsVarimax = varimax(Cs.PCA$loadings); Cs.rotatedLoadingsVarimax$rot[,1:3];
Cs.rotatedLoadingsPromax = promax(Cs.PCA$loadings); Cs.rotatedLoadingsPromax$rot[,1:3];

Cs.rotated = Cs.rotatedLoadingsVarimax$rot[,1:3];
  rownames(Cs.rotated)=myAttributes; colnames(Cs.rotated)=c("Comp.1","Comp.2","Comp.3");
```

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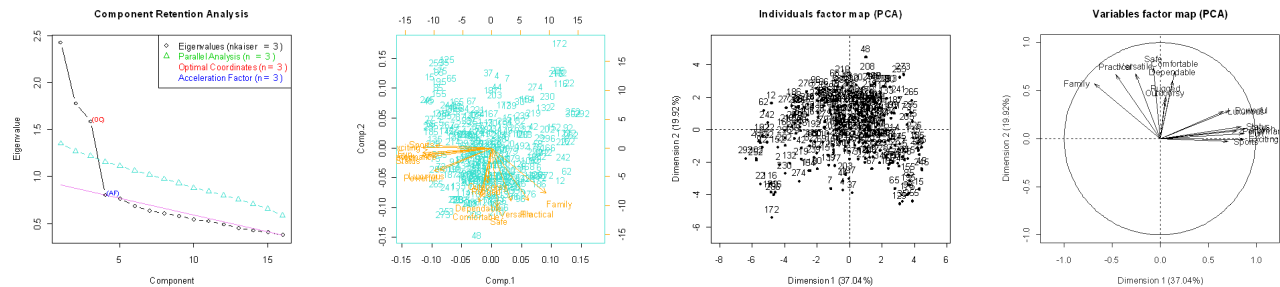


Figure 3: Basic PCA Explanation: Scree, BiPlot, Factor Maps using princomp

I would keep 3 PCs; since PCA only gives one solution, a rotation of the loadings should be rigid (in this case in 3-d space), and the rotation would help to find meaning in the 3 PCs; I could use varimax or promax rotation, which in this case are equivalent. Based on the rotations of the loadings, I would interpret the following components: PC1 is correlated with Exciting and Practical and Sporty. PC2 is correlated with Dependable, Stylish, Comfortable, Safe. PC3 is correlated Luxurious, Versatile, Status Symbol.

## Loadings:

	Comp.1	Comp.2	Comp.3	Comp.1	Comp.2	Comp.3
Exciting	-0.359		0.131	0.359	-0.030	-0.019
Dependable		-0.341	-0.268	0.031	0.417	-0.221
Luxurious	-0.270	-0.159	-0.291	-0.097	0.224	0.506
Outdoorsy		-0.221	0.506	0.100	-0.293	0.280
Powerful	-0.299	-0.166	0.143	-0.034	-0.027	-0.114
Stylish	-0.364			0.140	-0.387	-0.030
Comfortable		-0.391	-0.235	0.143	0.530	0.149
Rugged		-0.248	0.489	0.346	0.051	-0.028
Fun	-0.359			-0.318	0.046	-0.277
Safe		-0.417	-0.224	-0.236	-0.444	-0.047
Performance	-0.329		-0.146	0.169	-0.074	-0.255
Family	0.281	-0.320		0.121	-0.149	-0.053
Versatile	0.103	-0.376	0.240	-0.091	0.018	0.444
Sports	-0.292		0.310	-0.416	0.039	-0.027
Status	-0.347		-0.139	-0.224	0.163	-0.447
Practical	0.190	-0.373		0.508	0.006	-0.190

## Rotated Loadings:

## Problem 2

Now examine this same data with an exploratory factor analysis. How many factors should you use? How do you interpret them?

There are two techniques I could use to determine the number of factors to use. From the PCA analysis, I would conclude 3 factors. This is because the form of PCA ( $R = FF'$  to  $R \approx F_c F_c'$  where I keep  $c$  principal components) is similar to the model specification of EFA ( $R = \Lambda_c \Lambda_c' + \Theta$ ). If the common factors are large and  $\Theta$  as the specific factors are small (or close to zero), then the PCA and EFA are close to the same. Using the statistics from `factanal` we get conflicting results, since this is a  $\chi^2$  test of model fit.

Factors	d.f.	$\chi^2$	p-value	Conclusion
1	104	1634.44	2.84e-273	Reject $H_0$ : 1 solution is sufficient
2	89	994.86	5.19e-153	Reject $H_0$ : 2 solutions are sufficient
3	75	228.56	1.98e-17	Reject $H_0$ : 3 solutions are sufficient
4	62	144.33	1.63e-08	Reject $H_0$ : 4 solutions are sufficient
5	50	90.91	0.00036	Reject $H_0$ : 5 solutions are sufficient
6	39	51.97	0.0799	Fail to Reject $H_0$ : 6 solutions are sufficient

Figure 4: Basic EFA Model Fit Using  $\chi^2$

The model fit requires 6 factors, but the PCA tests suggest 3. Model fit will be assumed going forward, but this demonstrates that maybe the data should not be reduced; commonalities and uniqueness might explain part of the problem:

```
> Cs.EFA.none = factanal(Cs,factors=6,rotation='none');
> Cs.EFA.varimax = factanal(Cs,factors=6,rotation='varimax');

> round(1-Cs.EFA$uniquenesses,digits=2);
```

Exciting	Dependable	Luxurious	Outdoorsy	Powerful	Stylish	Comfortable	Rugged
0.79	0.44	0.72	0.77	0.60	0.86	0.60	0.83
Family	Versatile	Sports	Status	Practical			
0.80	0.60	0.69	0.76	0.70			

Since EFA may have multiple solutions, rotations should be used consistently to find meaning in the data reduction. F1 is Exciting, Luxurious, Power, Stylish, Fun, Performance, NOT family, Sporty, Status Symbol. F2 is Dependable, Luxurious, Comfortable, Safe. F3 is Outdoorsy, Rugged. F4 is Family, Versatile, Practical. F5 is Performance (cross-loaded with F1). F6 is Fun (also cross-loaded with F1). The common factor of F1 is an overall bias, in my opinion, to the fact that the ten models were mostly high-end with a few family cars. Also, two obvious SUVs are a factor. In the future, I would suggest doing either analysis on a class of cars: e.g., all SUVs, all sedans, etc. In marketing, this is how perceptual maps are generally created with attributes.

Varimax Rotated

Loadings:

	Factor1	Factor2	Factor3	Factor4	Factor5	Factor6
Exciting	0.855		0.142	-0.164		
Dependable	0.104	0.640		0.132		
Luxurious	0.579	0.530	-0.209	-0.169		-0.179
Outdoorsy			0.858	0.157		
Powerful	0.662	0.175	0.307		0.183	
Stylish	0.889	0.135		-0.103		-0.182
Comfortable	0.149	0.739		0.156		
Rugged			0.894	0.148		
Fun	0.903			-0.123		0.390
Safe		0.762		0.198	0.102	
Performance	0.724	0.203	-0.172	-0.137	0.617	
Family	-0.555	0.321	0.143	0.604		
Versatile		0.193	0.432	0.606		
Sports	0.699	-0.235	0.357	-0.105		
Status	0.788	0.258	-0.120	-0.172	0.107	-0.145
Practical	-0.284	0.322	0.115	0.709		

## Problem 3

Rerun your final EFA model with the option `scores = 'regression'`. Construct a scatterplot or pairs plot of the scores, setting `color` or `pch` to the car ID variable. Use `legend`, `text`, or `identify` to label the cars. What do your plots tell you about the similarities and differences between the 10 car models?

I am going to do “paired plots” in addition to the pairs plots (of the varimax rotation with option `scores = 'regression'`) so that I can interpret the color schema.

```
> Cs.EFA.varimax.regression = factanal(Cs,factors=6,rotation='varimax',scores='regression');

> legendV = c(1:10);
> plot(legendV,legendV,pch=myScores[,1],col=myScores[,1],xlab="",ylab="",main="LEGEND");
> text(legendV,legendV,pch=myScores[,1],col=myScores[,1],xlab="",ylab="",main="",labels=myCars);

> myScores=cbind(mba.cars[,2],Cs.EFA.varimax.regression$scores);
> pairs(myScores[,2:7],pch=myScores[,1],col=myScores[,1],cex=.1);
> library(MASS);
> parcoord(myScores[,2:7],pch=myScores[,1],col=myScores[,1]);

> myMeans = read.table('clipboard');
> colnames(myMeans)=c("CarID","n","Factor1","Factor2","Factor3","Factor4","Factor5","Factor6");
## clustering activity
> parcoord(myMeans[,3:8],pch=myScores[,1],col=myScores[,1]);
```

```
## paired plots
for(j in 2:7)
{
  for(i in 2:7)
  {
    plot(myScores[,i],myScores[,j],pch=myScores[,1],col=myScores[,1],xlab=i-1,ylab=j-1);
  }
}
```

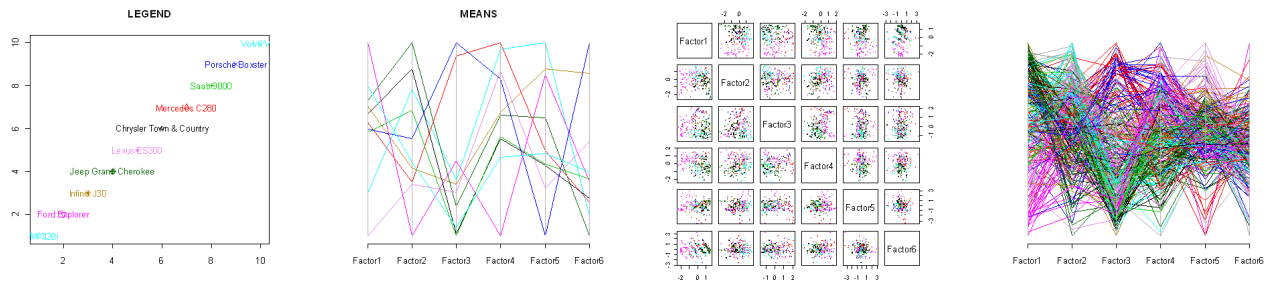


Figure 5: EFA Factors: Legend, Pairs, ParCoord, ParCoord with means

I also compared the means of the car models, which is a classification (clustering) technique based on the common factors. The means for each model can help us identify a given car, please see FIGURE 6 and MEANS of FIGURE 5.

Car ID	Car Name	N	Factor 1 Exciting, Luxurious, Power, Stylish, Fun, Per- formance, NOT family, Sporty, Status Symbol	Factor 2 Dependable, Luxurious, Com- fortable, Safe	Factor 3 Outdoorsy, Rugged	Factor 4 Family, Ver- satile, Prac- tical	Factor 5 Performance	Factor 6 Fun
1	BMW 328i	29	0.60	-0.22	-0.63	-0.36	-0.12	-0.03
2	Ford Explorer	29	-0.03	0.01	1.46	0.28	-0.70	0.47
3	Infiniti J30	28	-0.09	0.28	-0.73	-0.19	-0.20	-0.06
4	Jeep Grand Cherokee	30	0.07	-0.41	1.30	0.57	-0.11	-0.15
5	Lexus ES300	30	0.20	0.68	-0.70	-0.21	-0.21	-0.14
6	Chrysler Town and Country	31	-1.61	-0.43	-0.24	0.35	-0.37	0.09
7	Mercedes C280	27	0.40	0.93	-0.40	-0.01	0.13	-0.28
8	Saab 9000	30	0.30	-0.28	-0.14	0.01	0.47	0.35
9	Porsche Boxster	30	1.25	-0.93	0.12	-1.00	0.43	-0.07
10	Volvo V90	30	-0.98	0.48	-0.10	0.52	0.65	-0.20

Figure 6: Vehical Model Types and Factor Comparison



## Problem 4

The PSYCH\_TESTS dataset is a triangular correlation matrix for nine tests on 145 children, an expansion of the five-test data we looked at in class and in the text. Use `read.tri` to import this matrix and use it to answer question 5.1.

Analyze these data using factor analysis. How many factors are there? How would you interpret them? How do the results differ from the results based on five tests presented in the chapter?

```
setwd("C:/latex/statsMultiVariate/datasets");
read.tri = function(file)
{
  x = scan(file);
  lx = length(x);
  d = (sqrt(8*lx+1)-1)/2;
  m = matrix(0, d, d);
  m[upper.tri(m, T)] = x;
  m = m + t(m) - diag(diag(m));
  return(m);
}

R=read.tri("PSYCH_TESTS.txt");

psychQuestions = c("Visual perception","Cubes","Lozenges","Paragraph Comprehension","Sentence completion");
psychQuestionShort = c("VP","Cub","Loz","PC","SC","WM","Add","CD","SCC");

rownames(R)=psychQuestions;
colnames(R)=psychQuestionShort;

n=145;p=dim(R)[2];
R.e = eigen(R);
Lambda = R.e$values;
U = R.e$vectors;
## nFactors
library(nFactors);
nResults = nScree(eig = R.e$values,aparallel = parallel(subject = n, var = p)$eigen$vevpea);
plotuScree(R.e$values);
plotnScree(nResults, main="Component Retention Analysis");

psych.EFA.none = factanal(cov=R,factors=5,rotation='none');
psych.EFA.varimax = factanal(cov=R,factors=5,rotation='varimax');
```

Since the data is not available, the goodness of fit  $\chi^2$  and the scores are not available. The degrees of freedom are a determining variable to determine model fit. [<http://tolstoy.newcastle.edu.au/R/e2/help/07/05/16135.html>] The `criteria` variable is assumed to be related to some fit measure (AIC, BIC, LL, etc.); I will choose 5, even though PCA would suggest 3.

```
factanal(factors = 5, covmat = R, rotation = "varimax")
```

Uniquenesses:

VP	Cub	Loz	PC	SC	WM	Add	CD	SCC
0.523	0.595	0.509	0.005	0.392	0.011	0.005	0.464	0.367

Loadings:

	Factor1	Factor2	Factor3	Factor4	Factor5
Visual perception	0.213		0.409	0.510	
Cubes			0.622		
Lozenges	0.187		0.605	0.281	0.105
Paragraph Comprehension	0.959		0.218		-0.130
Sentence completion	0.720	0.121	0.113	0.207	0.142
Word meaning	0.769		0.193		0.590
Addition	0.147	0.980		0.114	
Counting dots			0.546	0.137	0.467
Straight-curved capitals	0.190	0.315	0.230	0.667	

	Factor1	Factor2	Factor3	Factor4	Factor5
SS loadings	2.175	1.381	1.089	1.083	0.401
Proportion Var	0.242	0.153	0.121	0.120	0.045
Cumulative Var	0.242	0.395	0.516	0.636	0.681

The degrees of freedom for the model is 1 and the fit was 0.0019

Factor 1 is Verbal (Paragraph Comprehension, Sentence completion, word Meaning); Factor 2 is Addition; Factor 3 is Visual (Visual perception, Cubes, Lozenges); Factor 4 is Spatial (Visual perception, Counting dots, Straight-curved capitals); Factor 5 is Language (Word Meaning). These 5 factors explain 68% of the variance in the data. PCA keeps each component consistent when one is added or removed, EFA is not like that. Since more factors are chosen than in the book examples, we can pull out more of the meaning from the data collected. There are still Math and Language components, but the abilities are further refined and explained using more data and more factors.

## FROM <http://tolstoy.newcastle.edu.au/R/e2/help/07/05/16135.html>

```
multifactanal = function(factors=1:3, ...)
{
  names(factors) = factors;
  ret = lapply(factors, function(factors)
  {
    try(factanal(factors=factors, ...))
  });
  class(ret) = "multifactanal";
  ret;
}
```

```

summary.multifactanal = function(object,...)
{
  do.call("rbind", lapply(object, summary.factanal));
}

print.multifactanal = function(x,...)
{
  ret = summary.multifactanal(x);
  print(ret, ...);
  invisible(ret);
}

summary.factanal =function(object, ...)
{
  if (inherits(object, "try-error"))
  {
    c(n=NA, items=NA, factors=NA, total.df=NA, rest.df=NA, model.df=NA, LL=NA, AIC=NA, AICc=NA, BIC=NA)
  }
  else
  {
    n = object$n.obs;
    p = length(object$uniquenesses);
    m = object$factors;

    model.df = (p*m) + (m*(m+1))/2 + p - m^2;
    total.df = p*(p+1)/2;
    rest.df = total.df - model.df; # = object$dof
    LL = -as.vector(object$criteria["objective"]);
    k = model.df;
    aic = 2*k - 2*LL;
    aicc = aic + (2*k*(k+1))/(n-k-1);
    bic = k*log(n) - 2*LL;
    c(n=n, items=p, factors=m, total.df=total.df, rest.df=rest.df, model.df=model.df, LL=LL, AIC=aic, AICc=aicc, BIC=bic)
  }
}

multifactanal(factors=1:5, covmat=R, n.obs=145);

```

	n	items	factors	total.df	rest.df	model.df	LL	AIC	AICc	BIC
1	145	9	1	45	27	18	-1.234756121	38.46951	43.89808	92.05072
2	145	9	2	45	19	26	-0.440387442	52.88077	64.77908	130.27585
3	145	9	3	45	12	33	-0.069084915	66.13817	86.35439	164.37038
4	145	9	4	45	6	39	-0.018929229	78.03786	107.75214	194.13047
5	145	9	5	45	1	44	-0.001949608	88.00390	127.60390	218.98018