

# Markov Chain Monte Carlo

Galin L. Jones

School of Statistics

University of Minnesota

Draft: January 28, 2023

## Contents

<b>1</b>	<b>Introduction</b>	<b>1</b>
<b>2</b>	<b>Markov Chains</b>	<b>2</b>
2.1	Stability . . . . .	4
<b>3</b>	<b>Metropolis-Hastings</b>	<b>5</b>
	<b>Exercises</b>	<b>5</b>
	<b>Appendix</b>	<b>6</b>

## 1 Introduction

Later.

## 2 Markov Chains

Let  $(\mathbf{X}, \mathcal{B})$  be a measurable space. A sequence of  $\mathbf{X}$ -valued random variables  $\{X_1, X_2, X_3, \dots\}$  is a Markov chain if for all  $g$

$$E[g(X_{n+1}, X_{n+2}, \dots) \mid X_n, X_{n-1}, \dots, X_1] = E[g(X_{n+1}, X_{n+2}, \dots) \mid X_n].$$

Then  $P$  is a *Markov kernel* if  $P : \mathbf{X} \times \mathcal{B} \rightarrow \mathbb{R}$  satisfying (i) for each fixed  $x \in \mathbf{X}$ ,  $P(x, \cdot)$  is a probability measure and (ii) for each fixed  $B \in \mathcal{B}$ ,  $P(\cdot, B)$  is a measurable function.

When  $\mathbf{X}$  is a discrete set a Markov kernel can be represented as a square matrix whose entries are nonnegative and whose rows sum to 1.

*Example 2.1.* Suppose

$$P = \begin{pmatrix} 1/2 & 1/2 \\ 1/3 & 2/3 \end{pmatrix}$$

Then  $P$  is a Markov matrix on two states  $\{0, 1\}$ , say. The first row for example, is interpreted as the probability of moving in one step from state 0 to state 0 is  $1/2$  which is the same as the probability of moving in one step from state 0 to start 1.

*Example 2.2.* Let  $\mathbf{X} = \mathbb{Z}$  and let  $0 < \theta < 1$ . If  $x \geq 1$ , then a Markov kernel is defined by the matrix  $P$  with elements

$$P(x, x+1) = P(-x, -x-1) = \theta, \quad P(x, 0) = P(-x, 0) = 1 - \theta,$$

and  $P(0, 1) = P(1, 0) = 1/2$ .

Most often  $\mathbf{X}$  will be uncountable and  $\mathcal{B}$  will be countably generated. If  $\mathbf{X}$  is topological, then  $\mathcal{B}$  will be the Borel  $\sigma$ -algebra generated by  $\mathbf{X}$ .

*Example 2.3.* Let  $\mathbf{X} = (0, 1)$  and consider the Markov chain that evolves as follows. Draw  $U \sim \text{Uniform}(0, 1)$ . If  $u \leq 0.5$ ,  $X_{n+1} \sim \text{Uniform}(0, X_n)$ , but if  $u > 0.5$ ,  $X_{n+1} \sim \text{Uniform}(X_n, 1)$ . Then if  $X_n = x$  and  $B \in \mathcal{B}$

$$P(x, B) = \int_B \left[ \frac{1}{2} \frac{1}{x} I_y((0, x)) + \frac{1}{2} \frac{1}{1-x} I_y(x, 1) \right] dy.$$

In Example 2.3, the integrand in the Markov kernel is a conditional density on  $\mathbf{X}$ . This is a setting that will be encountered repeatedly throughout. If there is a conditional density  $k(y \mid x)$  such that the Markov kernel satisfies for  $B \in \mathcal{B}$

$$P(x, B) = \int_B k(y \mid x) dy$$

then say  $k$  is a *Markov transition density*.

*Example 2.4.* Suppose  $f(x, y)$  is a joint density with support  $\mathbb{R}^2$  and conditional densities  $f_{X|Y}(x \mid y)$  and  $f_{Y|X}(y \mid x)$ . Then

$$k(x', y' \mid x, y) = f_{X|Y}(x' \mid y) f_{Y|X}(y' \mid x')$$

is a Markov transition density. The Markov chain evolves from  $(X_k = x, Y_k = y)$  to  $(X_{k+1}, Y_{k+1})$  by drawing  $X_{k+1} \sim F_{X|Y}(\cdot \mid y)$  followed by  $Y_{k+1} \sim F_{Y|X}(\cdot \mid X_{k+1})$ . This is a special case of the so-called two-variable Gibbs sampler.

Suppose  $\lambda$  is a positive measure on  $(\mathbf{X}, \mathcal{B})$ , define

$$\lambda P(B) = \int_{\mathbf{X}} \lambda(dx) P(x, B). \quad (1)$$

When  $\lambda$  is a probability measure, the encouraged interpretation is that  $X_{n+1} \mid X_n \sim P(X_n, \cdot)$  and  $X_n \sim \lambda$ , the product  $\lambda(dx) P(x, \cdot)$  is the joint distribution of  $(X_n, X_{n+1})$  and  $\lambda P$  is the marginal distribution of  $X_{n+1}$ .

Since Markov kernels act to the left on measures (1),

$$P^2(x, B) = \int_{\mathbf{X}} P(x, dx_k) P(x_k, B).$$

Continuing in this fashion obtain for every  $n \geq 2$

$$P^n(x, B) \int_{\mathbf{X}} P(x, dx_k) P(x_k, dx_{k+1}) \cdots P(x_{k+n-2}, B).$$

More generally, the so-called Chapman-Kolmogorov equations hold for  $n \geq m \geq 0$

$$P^n(x, B) \int_{\mathbf{X}} P^m(x, dy) P^{n-m}(y, B).$$

If  $\lambda = \lambda P$ , then  $\lambda$  is *invariant* for  $P$ . Notice that if  $\lambda$  is invariant for  $P$  and  $X_n \sim \lambda$ , then  $X_{n+1} \sim \lambda$ . That is, the marginal distribution does not depend upon  $n$  in which case the Markov chain is *stationary*.

*Example 2.5.*

*Example 2.6.* Recall the Markov chain defined in Example (2.4)

One common way of establishing invariance of MCMC Markov chains is to verify a *detailed balance condition*; see Exercise 3.1. Detailed balance holds if

$$\lambda(dx)P(x, dy) = \lambda(dy)P(y, dx). \quad (2)$$

When  $\lambda$  is a probability measure, one interpretation is that the joint distribution of  $(X_k, X_{k+1})$  is the same as the distribution of  $(X_{k+1}, X_k)$  so that this is also often called the *reversibility condition*. Another name often encountered is that  $P$  is  $\lambda$ -*symmetric*.

## 2.1 Stability

Most MCMC applications are constructed so that a specific probability distribution  $F$  is invariant. However, in applications where MCMC is required it is typically difficult to simulate from the invariant distribution. The most that can be hoped for is that the simulation will eventually produce a representative sample from  $F$ . This long-run behavior is in no way guaranteed.

*Example 2.7.* Suppose  $F$  lives on  $\{1, 2\}$  with  $F(1) = 1 - F(2) = 1/4$  and

$$P = \begin{pmatrix} 0 & 1 \\ 1 & 0 \end{pmatrix}$$

Since the Markov chain moves deterministically between the two states, it will over represent state 1 and underrepresent state 2 no matter how many iterations there are.

### 3 Metropolis-Hastings

#### Exercises

*Exercise 3.1.* Prove that if Equation 2 holds, then  $\lambda$  is invariant for  $P$ .

*Exercise 3.2.* What is the invariant distribution of the Markov chain defined in Example 2.6?

**Appendix**

**References**