

Movement matrix goodness of fit

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Variables used in this example (three datasets/runs)

- East Atlantic
- West Atlantic large fish
- West Atlantic small fish

```
eatl.ct = get.trans.prob(eatl.datbox, perc = F)
watl.big.ct = get.trans.prob(watl.big.datbox, perc = F)
watl.small.ct = get.trans.prob(watl.small.datbox, perc = F)
eatl = get.trans.prob(eatl.datbox, perc = T)
watl.big = get.trans.prob(watl.big.datbox, perc = T)
watl.small = get.trans.prob(watl.small.datbox, perc = T)
```

functions for multinomial variance.

The expected number of times the outcome i observed over n trials is

$$E(X_i) = np_i \quad (1)$$

$$v(X_i) = np_i(1 - p_i) \quad (2)$$

where n is the number of trials and p is the probability. Here, I consider each row of a movement matrix a multinomial. This is what we discussed in the proposal defense. code modified from [here](#)

Question:

I'm not sure how to represent n , but I assumed the transition between seasons was a single trial for each row.

```
# function to calculate variance of a multinomial row
get.mvar = function(x){
  p = drop(x) # use the percentages
  v = p*(1-p)
  # c = outer(p, 1-p)
  # vcv = diag(c)
  v
}

# function to sample from a posterior multinomial distribution
get.mvar.samp = function(vmu, vsd) {
  o = rtruncnorm(1, mean = vmu, sd = vsd, a = 0, 1)
  o[is.nan(o)] = 0
  o
}
```

```
# example for seasonal transition matrices
vmat = lapply(watl.big, function(x) apply(x, 1, get.mvar))
```

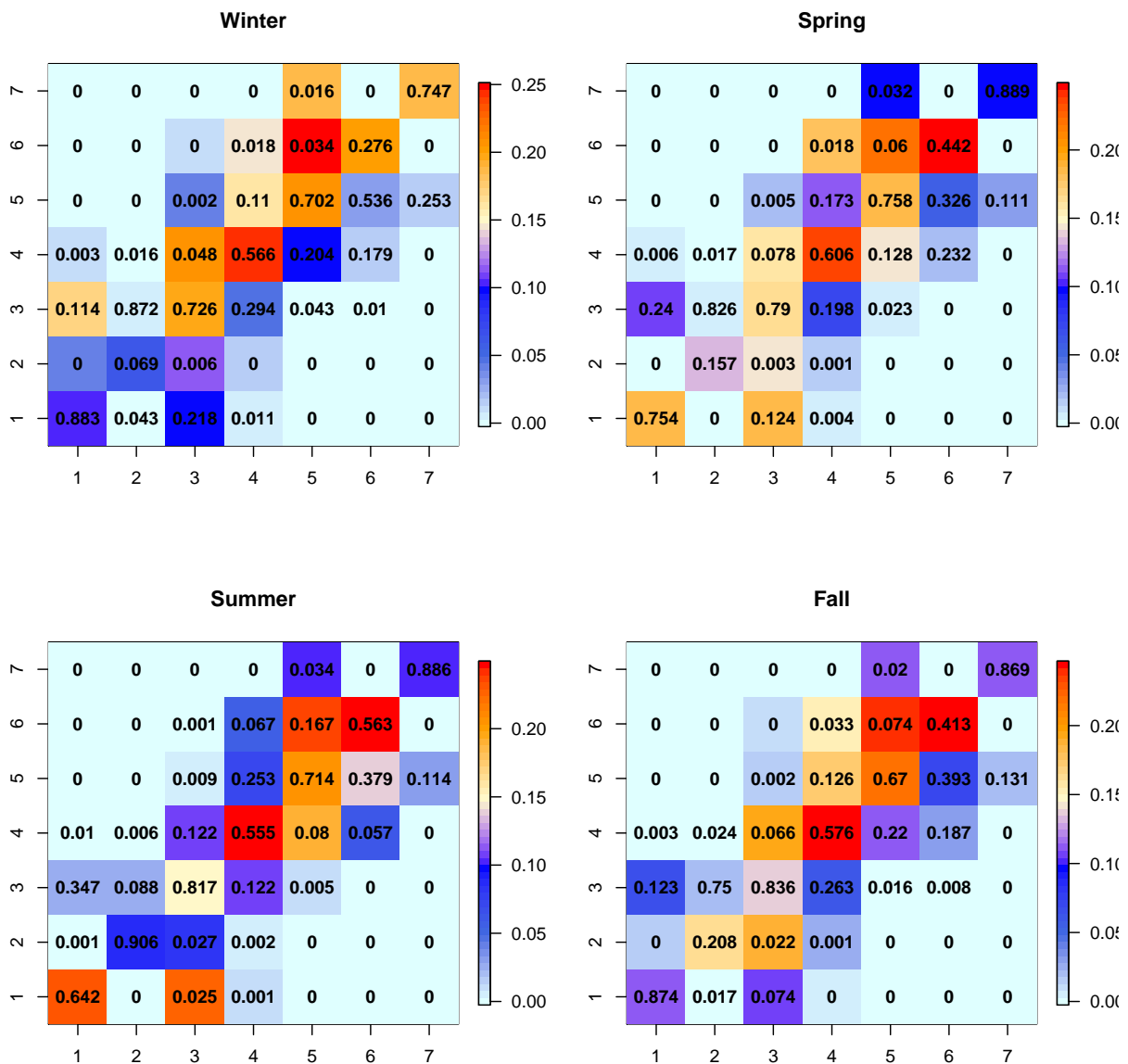
Table 1: Multinomial variance for Western bluefin >185cm, in Winter

	1	2	3	4	5	6	7
1	0.1034	0.0409	0.1704	0.0109	0.0000	0.0000	0.0000
2	0.0000	0.0643	0.0060	0.0000	0.0000	0.0000	0.0000
3	0.1010	0.1117	0.1990	0.2077	0.0412	0.0101	0.0000
4	0.0032	0.0160	0.0461	0.2456	0.1626	0.1467	0.0000
5	0.0000	0.0000	0.0017	0.0983	0.2092	0.2487	0.1889
6	0.0000	0.0000	0.0001	0.0176	0.0333	0.1996	0.0000
7	0.0000	0.0000	0.0000	0.0000	0.0158	0.0000	0.1889

Plot the variance for each cell, each season (colorbar)

Overlay with original movement matrix values

Example using Western bluefin >185cm



sample from posterior

```
samp = data.frame(sapply(1:7, function(x) get.mvar.samp(watl.big[[1]][x,], vmat[[1]][x,])))
names(samp) = 1:7
kable(samp, row.names = 1:7, caption = 'Posterior sample: Western Atlantic bluefin > 185cm, Winter season')
```

Table 2: Posterior sample: Western Atlantic bluefin > 185cm, Winter season

	1	2	3	4	5	6	7
1	0.8952	0.0427	0.1925	0.0092	0.0000	0.0000	0.0000
2	0.0766	0.1335	0.2093	0.0060	0.0000	0.0000	0.0000
3	0.3652	0.8779	0.6493	0.2873	0.0440	0.0101	0.0000
4	0.0006	0.0163	0.0988	0.6622	0.1492	0.1808	0.0000
5	0.0000	0.0000	0.0607	0.0861	0.8687	0.5023	0.2628
6	0.0000	0.0000	0.0055	0.1078	0.1534	0.5465	0.0000
7	0.0000	0.0000	0.0000	0.0000	0.2236	0.0000	0.6696

Goodness of Fit

function: compare pairs of rows of two matrices, using loglinear model

returns the pchisq value

χ^2 multinomial goodness of fit test from [here](#)

$$\chi^2 = \sum_i \frac{(f_i - e_i)^2}{e_i} \quad (3)$$

*# first argument is counts, second argument is probability. In this way, row-wise comparisons of two matrices
p-values of the log-linear (chisq test) are returned*

```
get.loglin <- function(mat1, mat2){  
  fm = NULL  
  for(i in 1:nrow(mat1)){  
    fm[[i]] = loglin(rbind(mat1[i,], mat2[i,]+1e-30),c(1), print = F, fit = F)  
  }  
  unlist(lapply(fm, function(x) 1-pchisq(x$lrt, x$df)))  
}  
  
get.mgof <- function(mat1, mat2){  
  fm = NULL  
  for(i in 1:nrow(mat1)){  
    fm[i] = chisq.test(mat1[i,], p = mat2[i,]+1e-30)$p.value  
  }  
  fm  
}
```

I tried several versions of log-linear and χ^2 tests. Log linear models, and their GLM versions, are essentially wrappers on χ^2 tests for ease of use and flexibility in using contingency tables in various formats.

Since each row of the movements matrices is dependent on its associated region in the previous time step, the matrix as a whole cannot be considered a contingency table. Each row, however, may be considered its own 1-dimensional contingency table.

To compare any two rows, within a matrix or between matrices, the best method appeared to be the χ^2 test. To set this up, counts within a row are compared to the probability in another row. The probability from the second matrix row is simply $\frac{x_i}{\sum x_i}$, and is the markovian result we have used to date.

To do goodness of fit comparisons, I used the `get.mgof` function. If any probabilities are 0, the χ^2 has issues. I added $1e - 30$ to alleviate this problem. Perhaps there is a better way?

Compare between seasons of same run/stock (e.g. Western Atlantic > 185cm)

```
bw.bs = get.mgof(watl.big.ct[[1]], watl.big[[2]])
print(bw.bs)
```

```
FALSE [1] 2.370314e-100 0.000000e+00 1.512262e-285 9.761739e-37 4.226017e-13
FALSE [6] 0.000000e+00 7.020061e-03
```

Compare between runs, same season

```
# same row, same season, large and small WABFT
# (bg.sm = get.loglin(watl.big.ct[[1]], watl.small[[1]]))

bg.sm = data.frame(sapply(1:4, function(x) get.mgof(watl.big.ct[[x]], watl.small[[x]])))
names(bg.sm) = names(watl.big)
kable(bg.sm, row.names = 1:7, caption = "Western Atlantic bluefin > 185 cm vs < 185 cm", digits = 4)
```

Table 3: Western Atlantic bluefin > 185 cm vs < 185 cm

	Winter	Spring	Summer	Fall
1	0.0000	0.0000	0.0000	0.0000
2	0.0209	0.8376	0.0000	0.0000
3	0.0000	0.0000	0.0000	0.0000
4	0.0000	0.0000	0.0000	0.0000
5	0.0144	0.6251	0.0000	0.0000
6	0.0000	0.0078	0.0096	0.0000
7	1.0000	0.2848	0.0000	0.1587

Compare stocks and season

```
e.w.s = data.frame(sapply(1:4, function(x) get.mgof(watl.big.ct[[x]], eatl[[x]])))
names(e.w.s) = names(watl.big)

# library(xtable)

# print(xtable(e.w.s, include.rownames = T), type = 'html')
kable(e.w.s, row.names = 1:7, caption = "Western Atlantic bluefin > 185 cm vs Eastern Atlantic bluefin")
```

Table 4: Western Atlantic bluefin > 185 cm vs Eastern Atlantic bluefin

	Winter	Spring	Summer	Fall
1	0.0000	0.0000	0.0000	0.0000
2	0.0000	0.0000	0.0000	0.0000
3	0.0000	0.0000	0.0000	0.0000
4	0.0000	0.0000	0.0000	0.0000
5	0.0000	0.0000	0.0000	0.0000
6	0.0000	0.0000	0.6999	0.0000
7	0.9228	0.9756	0.7166	0.8458

Compare between stocks, non-seasonal

this doesn't work... need a different test if this comparison is necessary

```
# get overall transition rate; multiply the seasonal matrices
b1 = watl.big
b2 = eatl

ball.w = b1[[1]]%*%b1[[2]]%*%b1[[3]]%*%b1[[4]]
ball.e = b2[[1]]%*%b2[[2]]%*%b2[[3]]%*%b2[[4]]

(e.w = get.loglin(ball.e, ball.w))
```

Tradeoff between degrees of freedom and likelihood ratio within the χ^2 test

this helped answer a question I had earlier

