Homework Assignment 4

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Problem 2.3. Find the ACVF of the time series $X_t = Z_t + aZ_{t-1} + bZ_{t-2}$ where $Z_t \sim WN(0, \sigma^2)$ when:

a.
$$a = 0.3$$
, $b = -0.4$, and $\sigma^2 = 1$.

b.
$$a = -1.2$$
, $b = -1.6$, and $\sigma^2 = 0.25$.

Solution. The ACVF of the time series $\{X_t\}$, $\gamma_X(h)$, is by definition:

$$\gamma_X(h) = \operatorname{Cov}(X_{t+h}, X_t)
= \operatorname{Cov}(Z_{t+h} + aZ_{t+h-1} + bZ_{t+h-2}, Z_t + aZ_{t-1} + bZ_{t-2})
= \operatorname{Cov}(Z_{t+h}, Z_t) + a\operatorname{Cov}(Z_{t+h}, Z_{t-1}) + b\operatorname{Cov}(Z_{t+h}, Z_{t-2})
+ a\operatorname{Cov}(Z_{t+h-1}, Z_t) + a^2\operatorname{Cov}(Z_{t+h-1}, Z_{t-1}) + ab\operatorname{Cov}(Z_{t+h-1}, Z_{t-2})
+ b\operatorname{Cov}(Z_{t+h-2}, Z_t) + ab\operatorname{Cov}(Z_{t+h-2}, Z_{t-1}) + b^2\operatorname{Cov}(Z_{t+h-2}, Z_{t-2}).$$
(1)

Using (1), we can see that since $Z_t \sim WN(0, \sigma^2)$

$$\gamma_X(h) = \begin{cases} (1+a^2+b^2)\sigma^2 & \text{if } h = 0\\ a(1+b)\sigma^2 & \text{if } h = \pm 1\\ b\sigma^2 & \text{if } h = \pm 2\\ 0 & \text{otherwise} \end{cases}.$$

Therefore, when

a. a = 0.3, b = -0.4, and $\sigma^2 = 1$, the ACVF of $\{X_t\}$ is:

$$\begin{cases} 1.25 & \text{if } h = 0 \\ 0.18 & \text{if } h = \pm 1 \\ -0.4 & \text{if } h = \pm 2 \\ 0 & \text{otherwise} \end{cases}$$

b. a = -1.2, b = -1.6, and $\sigma^2 = 0.25$, the ACVF of $\{X_t\}$ is:

$$\begin{cases} 1.25 & \text{if } h = 0 \\ 0.18 & \text{if } h = \pm 1 \\ -0.4 & \text{if } h = \pm 2 \\ 0 & \text{otherwise} \end{cases}$$

Problem 2.5. Suppose that $\{X_t, t = 0, \pm 1, \dots\}$ is stationary and that $|\theta| < 1$. Show that for each fixed n the sequence

$$S_m = \sum_{j=1}^m \theta^j X_{n-j}$$

is convergent absolutely and in mean square as $m \to \infty$.

Solution. Let $a_j = \theta^j X_{n-j}$. Since each X_i is a random variable, each X_i maps to a real, non-infinite value so let $X = \max\{|X_i|\}$. Then to see that S_m is convergent absolutely as $m \to \infty$, notice that

$$\sum_{j=1}^{m} |a_{j}| = \sum_{j=1}^{m} |\theta^{j} X_{n-j}|$$

$$= \sum_{j=1}^{m} |\theta|^{j} |X_{n-j}|$$

$$\leq \sum_{j=1}^{m} X |\theta|^{j} = \sum_{j=1}^{m} b_{j} = T_{m}$$

Since $|\theta| < 1$, we know that as $m \to \infty$, the partial sum $\sum_{j=1}^m X |\theta|^j \to 0$ and it must hold that $T_m \to 0$. Thus, we know that as $m \to \infty$, $\sum_{j=1}^m |a_j|$ converges to some L since $|a_j| \le b_j$ and T_m is convergent. Therefore, S_m is convergent absolutely.

To see that S_m is convergent in the mean square, it suffices to show that $\mathrm{E}(S_m-S_l)^2\to 0$ as $m,l\to\infty$.

Without loss of generality, assume that m > l > 0. Notice that $S_m - S_l = \sum_{j=1}^m a_j - \sum_{j=1}^n a_j = \sum_{j=l+1}^m a_j$. Thus,

$$E(S_m - S_l) = E(\sum_{j=l+1}^m a_j) = \sum_{j=l+1}^m E(a_j).$$

It is clear that $E(a_j) = E(\theta^j X_{n-j}) = \theta^j E(X_{n-j})$. Since $\{X_t\}$ is a stationary time series, its expectation does not depend on t, so say $E(X_{n-j}) = \mu_X$. Then

$$E(S_m - S_l) = \sum_{j=l+1}^m \theta^j E(X_{n-j})$$
$$= \mu_X \sum_{j=l+1}^m \theta^j$$
$$= \frac{\mu_X \theta^{l+1} (1 - \theta^{m-l-1})}{1 - \theta}$$

Since $|\theta| < 1$, it is clear then that $\mathrm{E}(S_m - S_l)^2 \to 0$ as $m, l \to \infty$ showing that S_m is convergent in mean square for any n.

Problem 2.11. Suppose that in a sample of size 100 from an AR(1) process with mean μ , $\phi = 0.6$, and $\sigma^2 = 2$ we obtain $x_{100}^- = 0.271$. Construct an approximate 95% confidence interval for μ . Are the data compatible with the hypothesis that $\mu = 0$.

Solution. Note that since AR(1) is a linear model, \bar{X}_n is approximately normal with mean μ for large n and an approximate 95% confidence interval for μ is

$$\left(\bar{X}_n - \frac{1.96\nu^{1/2}}{\sqrt{n}}, \bar{X}_n + \frac{1.96\nu^{1/2}}{\sqrt{n}}\right)$$

where $\nu = \sum_{|h| < \infty} \gamma_X(h)$.

Since $\{X_t\}$ is an AR(1) process, we know that $\gamma_X(h) = \gamma_X(0)\phi^{|h|}$ where $\gamma_X(0) = \sigma^2/(1-\phi^2)$. Thus

$$\nu = \sum_{|h| < \infty} \gamma_X(h) = \sum_{|h| < \infty} \frac{\sigma^2 \phi^{|h|}}{1 - \phi^2}$$

$$= \frac{\sigma^2}{1 - \phi^2} \left(1 + 2 \sum_{h=1}^{\infty} \phi^h \right)$$

$$= \frac{\sigma^2}{1 - \phi^2} \left(1 + \frac{2\phi}{1 - \phi} \right)$$

$$= \frac{\sigma^2(1 + \phi)}{(1 - \phi)(1 - \phi^2)} = \frac{\sigma^2}{(1 - \phi)^2}$$

If $\phi = 0.6$ and $\sigma^2 = 2$, then $\nu = 2/(1-0.6)^2 = 12.5$. Since n = 100, $\bar{x_n} = \bar{x_{100}} = 0.271$, and an approximate 95% confidence interval for μ is

$$\left(0.271 - \frac{1.96(12.5)^{1/2}}{\sqrt{100}}, 0.271 + \frac{1.96(12.5)^{1/2}}{\sqrt{100}}\right)$$

or (-0.42197, 0.96397). Given this confidence interval, it is plausible that $\mu = 0$.

Problem 2.12. Suppose that in a sample of size 100 from an MA(1) process with mean μ , $\theta = -0.6$, and $\sigma^2 = 1$ we obtain $x_{100}^- = 0.157$. Construct an approximate 95% confidence interval for μ . Are the data compatible with the hypothesis that $\mu = 0$.

Solution. Note that since MA(1) is a linear model, \bar{X}_n is approximately normal with mean μ for large n and an approximate 95% confidence interval for μ is

$$\left(\bar{X}_n - \frac{1.96\nu^{1/2}}{\sqrt{n}}, \bar{X}_n + \frac{1.96\nu^{1/2}}{\sqrt{n}}\right)$$

where $\nu = \sum_{|h| < \infty} \gamma_X(h)$.

Since $\{X_t\}$ is an MA(1) process, we know that

$$\gamma_X(h) = \begin{cases} \sigma^2(1+\theta^2) & \text{if } h = 0\\ \sigma^2\theta & \text{if } h = \pm 1\\ 0 & \text{otherwise} \end{cases}$$

Thus

$$\nu = \sum_{|h| < \infty} \gamma_X(h) = \sigma^2(1 + \theta^2) + 2\sigma^2\theta = \sigma^2(1 + \theta)^2$$

If $\theta = -0.6$ and $\sigma^2 = 1$, then $\nu = (1 - 0.6))^2 = 0.16$. Since n = 100, $\bar{x_n} = \bar{x_{100}} = 0.157$, and an approximate 95% confidence interval for μ is

$$\left(0.157 - \frac{1.96(0.16)^{1/2}}{\sqrt{100}}, 0.157 + \frac{1.96(0.16)^{1/2}}{\sqrt{100}}\right)$$

or (0.15198, 0.16202). Given this confidence interval, it is not plausible that $\mu = 0$.