

# Homework Assignment 1

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**Problem 3.7.** Suppose  $p(x, y, z)$ , the joint probability mass function of the random variables  $X$ ,  $Y$ , and  $Z$ , is given by

$$p(1, 1, 1) = \frac{1}{8}, \quad p(2, 1, 1) = \frac{1}{4},$$

$$p(1, 1, 2) = \frac{1}{8}, \quad p(2, 1, 2) = \frac{3}{16},$$

$$p(1, 2, 1) = \frac{1}{16}, \quad p(2, 2, 1) = 0,$$

$$p(1, 2, 2) = 0, \quad p(2, 2, 2) = \frac{1}{4}.$$

What is  $E[X|Y = 2]$ ? What is  $E[X|Y = 2, Z = 1]$ ?

*Solution.* Recall that the conditional probability mass function of  $X$  given that  $Y = y$  is given by

$$p_{X|Y}(x|y) = P\{X = x|Y = y\} = \frac{P\{X = x, Y = y\}}{P\{Y = y\}}.$$

As a natural extension, we have that the conditional expectation of  $X$  given that  $Y = y$  is given by

$$E[X|Y = y] = \sum_x xP\{X = x|Y = y\} = \sum_x xp_{X|Y}(x|y).$$

Thus, in order to find the conditional expectation of  $X$  given that  $Y = 2$ , i.e.  $E[X|Y = 2]$ , we first need to determine  $p_{X|Y}(x|2)$ . We note from the above joint probability mass function that

$$P\{Y = 2\} = \sum_{x,z} p(x, 2, z) = p(1, 2, 1) + p(2, 2, 1) + p(1, 2, 2) + p(2, 2, 2) = \frac{5}{16}.$$

Similarly, we have from the above joint probability mass function that

$$P\{X = x, Y = 2\} = \sum_z p(x, 2, z) = p(x, 2, 1) + p(x, 2, 2).$$

Thus, the conditional probability mass function of  $X$  given that  $Y = 2$  is given by

$$p_{X|Y}(x|2) = \frac{P\{X = x, Y = 2\}}{P\{Y = 2\}} = \begin{cases} \frac{p(1,2,1)+p(1,2,2)}{5/16} = \frac{1}{5} & \text{if } x = 1 \\ \frac{p(1,2,1)+p(1,2,2)}{5/16} = \frac{4}{5} & \text{if } x = 2. \end{cases}$$

Using  $p_{X|Y}(x|2)$ , we readily see that

$$E[X|Y = 2] = \sum_x x p_{X|Y}(x|2) = 1 \cdot p_{X|Y}(1|2) + 2 \cdot p_{X|Y}(2|2) = \frac{9}{5}.$$

In order to find the conditional expectation of  $X$  given that  $Y = 2$  and  $Z = 1$ , i.e.  $E[X|Y = 2, Z = 1]$ , we proceed in a similar manner as previously by first finding  $p_{X|Y,Z}(x|2, 1)$ . We note from the above joint probability mass function that

$$P\{Y = 2, Z = 1\} = \sum_x p(x, 2, 1) = p(1, 2, 1) + p(2, 2, 1) = \frac{1}{16}$$

Similarly, we have from the above joint probability mass function that

$$P\{X = x, Y = 2, Z = 1\} = p(x, 2, 1).$$

Thus, the conditional probability mass function of  $X$  given that  $Y = 2$  and  $Z = 1$  is given by

$$p_{X|Y,Z}(x|2, 1) = \frac{P\{X = x, Y = 2, Z = 1\}}{P\{Y = 2, Z = 1\}} = \begin{cases} \frac{p(1,2,1)}{1/16} = 1 & \text{if } x = 1 \\ \frac{p(2,2,1)}{1/16} = 0 & \text{if } x = 2. \end{cases}$$

Using  $p_{X|Y,Z}(x|2, 1)$ , we readily see that

$$E[X|Y = 2, Z = 1] = \sum_x x p_{X|Y,Z}(x|2, 1) = 1 \cdot p_{X|Y,Z}(1|2, 1) + 2 \cdot p_{X|Y,Z}(2|2, 1) = 1.$$

□

**Problem 3.8.** An unbiased die is successively rolled. Let  $X$  and  $Y$  denote, respectively, the number of rolls necessary to obtain a six and a five. Find:

- a.  $E[X]$ ,
- b.  $E[X|Y = 1]$ ,
- c.  $E[X|Y = 5]$ .

*Solution.* The experiment of rolling a die, assuming the die is six-sided, has six possible outcomes: the die lands oriented such that the side with 1, 2, 3, 4, 5, or 6 pips is face-up. Assuming the die is unbiased, each outcome occurs with probability  $p = 1/6$  and each trial of rolling the die is independent of any other trial. If  $X$  and  $Y$  denote, respectively, the number of rolls necessary to obtain a six and a five, then under the given assumptions,  $X$  and  $Y$  are both geometric random variables with parameter  $p = 1/6$ . The probability mass function for these random variables is given by  $p(n) = (1 - p)^{n-1}p = (5/6)^{n-1}(1/6)$ .

- a. Let  $Z$  be the random variable defined as  $Z = 1$  if the result of the first roll is a six and  $Z = 0$  if the result of the first roll is not a six. We may compute  $E[X]$  by conditioning on the variable  $Z$ . Note that, by conditioning, we obtain

$$\begin{aligned} E[X] &= \sum_z E[X|Z = z]P\{Z = z\} \\ &= \left[\frac{1}{6}\right] E[X|Z = 1] + \left[\frac{5}{6}\right] E[X|Z = 0]. \end{aligned}$$

If  $Z = 1$ , then the result of the first roll is a six, so the number of rolls to obtain a six is clearly 1 and  $E[X|Z = 1] = 1$ . Likewise, if the result of the first roll is not a six, then the expected number of rolls to obtain a six given that the first roll is not a six is 1 more than the expected number of rolls to obtain a six so that  $E[X|Z = 0] = 1 + E[X]$ . Therefore,

$$\begin{aligned} E[X] &= \left[\frac{1}{6}\right] E[X|Z = 1] + \left[\frac{5}{6}\right] E[X|Z = 0] \\ &= \frac{1}{6} + \left[\frac{1}{6}\right] (1 + E[X]) \end{aligned}$$

which implies that  $E[X] = 6$ .

- b. We wish to find  $E[X|Y = 1]$ , i.e. the expected number of rolls to obtain a six given that the first roll is a five. Using the same reasoning as in part a, we know that the expected number of rolls to obtain a six given that the first roll is not a six (it's a five) is 1 more than the expected number of rolls to obtain a six. Therefore,

$$E[X|Y = 1] = 1 + E[X] = 7$$

where we used the result previously obtained that  $E[X] = 6$ .

- c. In order to calculate  $E[X|Y = y]$  for some  $y > 1$ , we first compute  $p_{X|Y}(x|y)$ . Suppose that  $Y = y$  for some  $y > 1$ . From this we gather that the first  $y - 1$  trials result in not rolling a five while the  $y$ -th trial results in rolling a five.

As a consequence, if  $X = x$  where  $x < y$  then the first  $x$  trials have only five possible outcomes with the  $x$ -th trial resulting in a success out of those five outcomes. Thus,

$$P\{X = x|Y = y\} = \frac{1}{5} \left[ \frac{4}{5} \right]^{x-1},$$

i.e. for  $x < y$  the conditional probability that  $X = x$  given that  $Y = y$  is the probability mass function of a geometric random variable with parameter  $p = 1/5$ .

Note that if  $X = x$  where  $x = y$ , then

$$P\{X = x|Y = y\} = 0$$

since it cannot happen that on the  $y$ -th trial the outcome of the trial is that both a five and a six were rolled.

Finally, if  $X = x$  where  $x > y$ , then as mentioned, the first  $y - 1$  trials do not result in a five, but after the  $y$ -th trial the result obtained can in fact be a five. Thus, the first  $y - 1$  failures each occur with probability  $4/5$  while the  $y$ -th failure occurs with probability 1. However, after that, the failures of the trials  $y + 1$  through  $x - 1$  all occur with probability  $5/6$  since it is possible for the die to roll a five during these trials. On the  $x$ -th trial the trial succeeds with probability  $1/6$ . Thus, if  $x > y$ , then

$$P\{X = x|Y = y\} = \frac{1}{6} \left[ \frac{4}{5} \right]^{y-1} \left[ \frac{5}{6} \right]^{x-y-1}.$$

Combining the above statements, we see that the conditional probability mass function that  $X = x$  given that  $Y = y$  with  $y > 1$  is given by

$$p_{X|Y}(x|y) = \begin{cases} \frac{1}{5} \left[ \frac{4}{5} \right]^{x-1} & \text{if } x < y \\ 0 & \text{if } x = y \\ \frac{1}{6} \left[ \frac{4}{5} \right]^{y-1} \left[ \frac{5}{6} \right]^{x-y-1} & \text{if } x > y \end{cases}.$$

Therefore, we have that the expected value of  $X$  given that  $Y = 5$  is

$$\begin{aligned} E[X|Y = 5] &= \sum_{x=1}^{\infty} x p_{X|Y}(x|5) \\ &= \frac{1}{5} \sum_{x=1}^4 x \left[ \frac{4}{5} \right]^{x-1} + \frac{1}{6} \left[ \frac{4}{5} \right]^4 \sum_{x=6}^{\infty} x \left[ \frac{5}{6} \right]^{x-6} \\ &= \frac{3637}{625} \approx 5.82. \end{aligned}$$

□

**Problem 3.9.** Show in the discrete case that if  $X$  and  $Y$  are independent, then

$$E[X|Y = y] = E[X] \text{ for all } y.$$

*Solution.* Suppose that  $X$  and  $Y$  are discrete, independent random variables. Then, due to the independence of the random variables, we know that

$$\begin{aligned} P\{X = x|Y = y\} &= \frac{P\{X = x, Y = y\}}{P\{Y = y\}} \\ &= \frac{P\{X = x\}P\{Y = y\}}{P\{Y = y\}} \\ &= P\{X = x\}. \end{aligned} \tag{1}$$

Therefore, combining the definition of conditional expectation for discrete random variables and result (1), we have that for any  $y$ ,

$$E[X|Y = y] = \sum_x xP\{X = x|Y = y\} = \sum_x xP\{X = x\} = E[X]$$

and we are done. □

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**Problem 3.10.** Suppose  $X$  and  $Y$  are independent continuous random variables. Show that

$$E[X|Y = y] = E[X] \text{ for all } y.$$

*Solution.* Suppose that  $X$  and  $Y$  are continuous, independent random variables with probability density functions  $f_X(x)$  and  $f_Y(y)$ , respectively. The conditional probability density function of  $X$  given that  $Y = y$  is given by

$$f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)}$$

where  $f(x, y)$  is the joint probability density function of  $X$  and  $Y$ . Due to the independence of the random variables  $X$  and  $Y$ , we have that  $f(x, y) = f_X(x)f_Y(y)$ . Thus, if  $X$  and  $Y$  are independent, then the conditional probability density function of  $X$  given that  $Y = y$  is

$$f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)} = \frac{f_X(x)f_Y(y)}{f_Y(y)} = f_X(x). \quad (2)$$

Therefore, combining the definition of conditional expectation for continuous random variables and result (2), we have that for any  $y$ ,

$$E[X|Y = y] = \int_{-\infty}^{\infty} x f_{X|Y}(x|y) dx = \int_{-\infty}^{\infty} x f_X(x) dx = E[X]$$

and we are done. □

**Problem 3.13.** Let  $X$  be exponential with mean  $1/\lambda$ ; that is,

$$f_X(x) = \lambda e^{-\lambda x}, \quad 0 < x < \infty.$$

Find  $E[X|X > 1]$ .

*Solution.* Suppose that  $X$  is an exponential random variable with mean  $1/\lambda$ . Let  $Y$  be the discrete random variable defined as

$$Y = \begin{cases} 1 & \text{if } X > 1 \\ 0 & \text{if } 0 < X \leq 1 \end{cases}$$

with probability mass function

$$p_Y(y) = \begin{cases} P\{X > 1\} & \text{if } y = 1 \\ P\{X \leq 1\} & \text{if } y = 0. \end{cases}$$

From these definitions of  $X$  and  $Y$ , we see that the conditional density function of  $X$  given  $Y = 1$  is given by

$$f_{X|Y}(x|1) = \frac{f(x, 1)}{p_Y(1)} = \begin{cases} \frac{f_X(x)}{P\{Y=1\}} & \text{if } x > 1 \\ 0 & \text{if } 0 < x \leq 1 \end{cases}$$

where we know that if  $x > 1$  then

$$\begin{aligned} \frac{f_X(x)}{P\{Y = 1\}} &= \frac{\lambda e^{-\lambda x}}{P\{X > 1\}} \\ &= \frac{\lambda e^{-\lambda x}}{1 - P\{X \leq 1\}} \\ &= \frac{\lambda e^{-\lambda x}}{1 - \int_0^1 \lambda e^{-\lambda x} dx} \\ &= \frac{\lambda e^{-\lambda x}}{e^{-\lambda}}. \end{aligned}$$

Thus, by definition, we now have that

$$\begin{aligned} E[X|X > 1] &= E[X|Y = 1] \\ &= \int_{-\infty}^{\infty} x f_{X|Y}(x|1) dx \\ &= \int_1^{\infty} x \frac{\lambda e^{-\lambda x}}{e^{-\lambda}} dx \\ &= \lambda e^{\lambda} \int_1^{\infty} x e^{-\lambda x} dx \\ &= \lambda e^{\lambda} \left[ \frac{e^{-\lambda}(1 + \lambda)}{\lambda^2} \right] \\ &= \frac{1 + \lambda}{\lambda}. \end{aligned}$$

□

**Problem 3.14.** Let  $X$  be uniform over  $(0, 1)$ . Find  $E[X|X < 1/2]$ .

*Solution.* Suppose that  $X$  is a uniform random variable over  $(0, 1)$ . Let  $Y$  be the discrete random variable defined as

$$Y = \begin{cases} 1 & \text{if } 0 < X < 1/2 \\ 0 & \text{if } 1/2 \leq X < 1 \end{cases}$$

with probability mass function

$$p_Y(y) = \begin{cases} P\{0 < X < 1/2\} & \text{if } y = 1 \\ P\{1/2 \leq X < 1\} & \text{if } y = 0 \end{cases}.$$

From these definitions of  $X$  and  $Y$ , we see that the conditional density function of  $X$  given  $Y = 1$  is given by

$$f_{X|Y}(x|1) = \frac{f(x, 1)}{p_Y(1)} = \begin{cases} \frac{f_X(x)}{P\{Y=1\}} & \text{if } 0 < x < 1/2 \\ 0 & \text{if } 1/2 \leq x < 1 \end{cases}$$

where we know that if  $0 < x < 1/2$  then

$$\begin{aligned} \frac{f_X(x)}{P\{Y = 1\}} &= \frac{1}{P\{0 < X < 1/2\}} \\ &= \frac{1}{\int_0^{1/2} dx} \\ &= \frac{1}{1/2} = 2. \end{aligned}$$

Thus, by definition, we now have that

$$\begin{aligned} E[X|X < 1/2] &= E[X|Y = 1] \\ &= \int_{-\infty}^{\infty} x f_{X|Y}(x|1) dx \\ &= \int_0^{1/2} 2x dx = \frac{1}{4}. \end{aligned}$$

□