

4th year Review - Statistical Physics Perspectives on Learning in High Dimensions

Advisor: Surya Ganguli

Madhu Advani

Stanford University

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Current Research

Optimal Tractable High Dimensional M-estimation

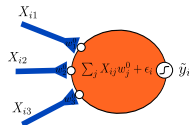
Future Directions

- ① LN and GLM extensions of M-estimation also Optimal Signal Processing (Structured Coefficients)
- ② Random Dimensionality Reduction
- ③ Phase transitions in clustering Behavior

Previous Research and Future Plan

- ① Review Paper
- ② Timeline

Problem Setup



$$y_i = \mathbf{X}_i \cdot \mathbf{w}^0 + \epsilon_i \quad i \in [1, \dots, N]$$

- Noise $\epsilon_i \sim f$ not necessarily gaussian
- (Easy) Classical Regime: $\kappa = P/N \rightarrow 0$
- (Hard) High Dimensional Regime: $\kappa = P/N \neq 0$

We want to find \mathbf{w}^0

$$\hat{\mathbf{w}} = \arg \min_{\mathbf{w}} \left[\sum_i \rho(y_i - \mathbf{X}_i \cdot \mathbf{w}) \right]$$

E.g. $\rho(x) = x^2, |x|, -\log f(x)$

Maximum Likelihood

P fixed, $N \rightarrow \infty$

$$\|\hat{\mathbf{w}} - \mathbf{w}^0\|^2 = \frac{(\int \psi(x) f(x) dx)^2}{\int \psi^2 f(x) dx}$$

$$\psi = \frac{\partial}{\partial \mathbf{w}} \rho$$

We find that $\rho_{\text{opt}} = -\log f$

$$\hat{\mathbf{w}}_{\text{ML}} = \arg \max_{\mathbf{w}^0} P(\mathbf{y}, \mathbf{X} | \mathbf{w}^0) = \arg \min_{\mathbf{w}} \left[\sum_i -\log f(y_i - \mathbf{X}_i \cdot \mathbf{w}) \right]$$

Classical vs High Dimensional Optimal M-estimation

Classical

- $N \rightarrow \infty, P/N = \kappa \rightarrow 0$
- $\rho_{\text{opt}} = -\log f$
- $\langle\langle (\hat{w}_i - w_i^0)^2 \rangle\rangle \geq \frac{\kappa}{\int \frac{\kappa}{f'^2}}$

High Dimensional

- $N, P \rightarrow \infty, P/N = \kappa \in [0, 1]$
- $\rho_{\text{opt}}(x) = -\inf_y \left[\ln(\zeta(y)) + \frac{(x-y)^2}{2\hat{q}_0} \right] \quad \zeta = f * \phi_{\hat{q}_0}$
- $\langle\langle (\hat{w}_i - w_i^0)^2 \rangle\rangle \geq \frac{\kappa}{\int \frac{\kappa}{\zeta'^2}}$

Adding a Regularizer

$$P(\mathbf{w}^0 | \mathbf{X}, \mathbf{y}) = \frac{P(\mathbf{X}, \mathbf{y} | \mathbf{w}^0) P(\mathbf{w}^0)}{P(\mathbf{X}, \mathbf{y})} \propto f(\mathbf{y} - \mathbf{X} \cdot \mathbf{w}^0) g(\mathbf{w}^0)$$

$$w_j^0 \sim g$$

Maximum a Priori

$$\hat{\mathbf{w}}_{\text{MAP}} = \arg \min_{\mathbf{w}} \left[\sum_i -\log f(y_i - \mathbf{X}_i \cdot \mathbf{w}) + \sum_j -\log g(w_j) \right]$$

Regularized M-estimation

$$\hat{\mathbf{w}} = \arg \min_{\mathbf{w}} \left[\sum_i \rho(y_i - \mathbf{X}_i \cdot \mathbf{w}) + \sum_j \sigma(w_j) \right]$$

Note separability. Solvable for convex σ, ρ

$$\hat{\mathbf{w}} = \arg \min_{\mathbf{w}} \sum_a \rho(y_a - \mathbf{X}_a \cdot \mathbf{w}) + \sum_i \sigma(w_i)$$

Applications

- Maximum Likelihood (ML) and MAP commonly applied to High Dimensional Bio-informatics problems, where we should expect poor performance
- Deriving/Understanding the Optimal M-estimator has potential applications for statistical inference and signal processing.
- Tractability of this form of optimization popular: Compressed Sensing, LASSO, Elastic Net

Statistical Physics Formulation

Define a spin glass system to solve M-estimator inference

Spin Glass System

Define continuous spins $\mathbf{u} = \mathbf{w}^0 - \mathbf{w}$. Let the Energy of the system be a function of these spins

$$E_{\Lambda}(\mathbf{u}) = \sum_i \rho(\mathbf{X}_i \cdot \mathbf{u} + \epsilon_i) + \sum_a \sigma(w_a^0 - u_a)$$

This in turn induces an equilibrium probability distribution on the state

$$P_G(\mathbf{u}) = \frac{e^{-\beta E_{\Lambda}(\mathbf{u})}}{Z_{\Lambda}} \quad Z_{\Lambda} = \int e^{-\beta E_{\Lambda}(\mathbf{u})} d\mathbf{u}$$

$$\lim_{\beta \rightarrow \infty} P_G(\mathbf{u}) = \delta(\mathbf{u} - \mathbf{w}^0 + \hat{\mathbf{w}})$$

The Unregularized Case

Coupled Equations Relating Order Parameters q, c

$$\left\langle\left\langle (\text{prox}_{c\rho}(\sqrt{q}z + \epsilon) - \sqrt{q}z - \epsilon)^2 \right\rangle\right\rangle_{z,\epsilon} = \kappa q$$

$$\left\langle\left\langle \text{prox}'_{c\rho}(\sqrt{q}z + \epsilon) \right\rangle\right\rangle_{z,\epsilon} = 1 - \kappa$$

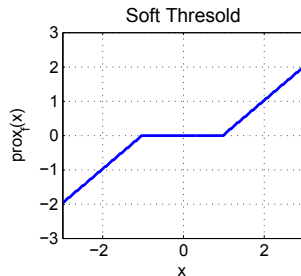
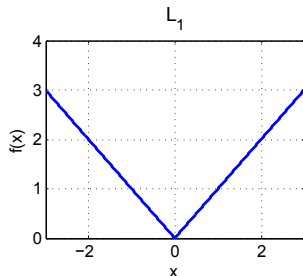
- $X_{ij} \in \mathcal{N}(0, 1/P)$
- ρ convex
- $\epsilon_i \sim f$ iid

Proximal Map

Definition

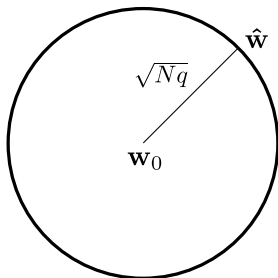
$$\begin{aligned} \text{prox}_f(x) &= \arg \min_y \left[\frac{(x - y)^2}{2} + f(y) \right] \\ &= [I + \partial f]^{-1}(x) \end{aligned}$$

Example:



Analytic Estimator Error

How to choose ρ to minimize q ?



$$\left\langle \left\langle (\text{prox}_{c\rho}(\sqrt{q}z + \epsilon) - \sqrt{q}z - \epsilon)^2 \right\rangle \right\rangle_{z, \epsilon} = \kappa q$$

$$\left\langle \left\langle \text{prox}'_{c\rho}(\sqrt{q}z + \epsilon) \right\rangle \right\rangle_{z, \epsilon} = 1 - \kappa$$

Optimal Unregularized M-estimator

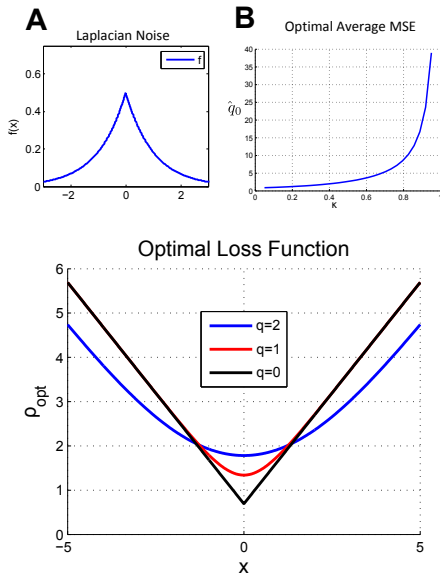
Optimal M-estimator

$$\rho_{\text{opt}}(x) = -\inf_y \left[\ln(\zeta_{\hat{q}}(y)) + \frac{(x-y)^2}{2\hat{q}} \right] \quad \zeta_{\hat{q}} = f * \phi_{\hat{q}}$$

$$\hat{q} = \min q \quad \text{s.t.} \quad q l_q = \kappa \quad l_q = \int_{-\infty}^{\infty} \frac{\zeta_q'^2(y)}{\zeta_q(y)} dy$$

- \hat{q} - best possible asymptotic MSE for a convex M-estimator
- ρ_{opt} is the optimal loss function (assuming log-concave noise)
- Note: Not maximum likelihood, and ρ varies with dimensionality κ

Optimal Unregularized M-estimator

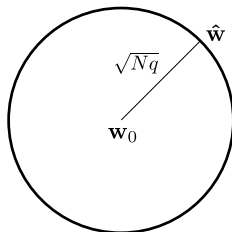


Regularized M-estimation

$$y_i = \mathbf{X}_i \cdot \mathbf{w}^0 + \epsilon_i$$

$$w_j^0 \sim g, \epsilon_i \sim f$$

$$\hat{\mathbf{w}} = \arg \min_{\mathbf{w}} \sum_a \rho(y_a - \mathbf{X}_a \cdot \mathbf{w}) + \sum_i \sigma(w_i)$$



Optimal Inference

$$\rho_{\text{opt}}^R(x) = - \inf_y \left[\ln(\zeta_{\tilde{q}_0}(y)) + \frac{(x - y)^2}{2\tilde{q}_0} \right]$$

$$\sigma_{\text{opt}}^R(x) = - \frac{\tilde{q}_0}{\tilde{a}} \inf_y \left[\ln(\xi_{\tilde{a}}(y)) + \frac{(x - y)^2}{2\tilde{a}} \right]$$

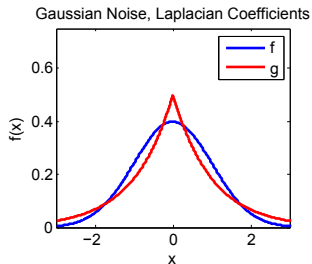
$$\tilde{q}_0, \tilde{a} = \arg \min_{q_0, a} q_0$$

$$\text{s.t.} \quad a I_{q_0} = \kappa, \quad a^2 J_a = a - q_0$$

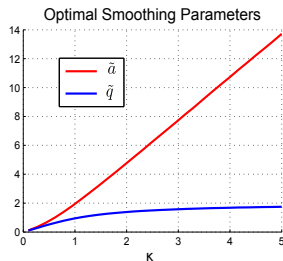
- $\rho_{\text{opt}}, \sigma_{\text{opt}}$ are optimal M-estimator (log concave f, g)
- \tilde{q}_0 is the asymptotic MSE
- \tilde{q}_0, \tilde{a} are smoothing parameters

Unregularized M-estimator

A

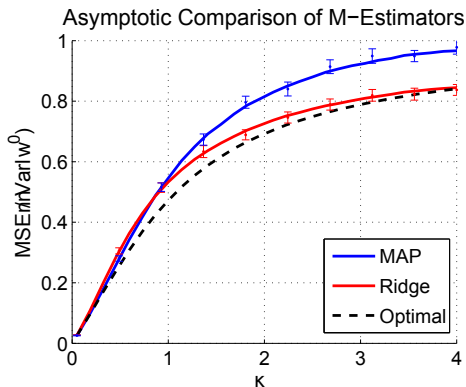


B



M-estimator Comparison

$$\begin{aligned}\rho_{\text{MAP}}(x) &= \frac{x^2}{2} & \sigma_{\text{MAP}}(x) &= |x| \\ \rho_{\text{Ridge}}(x) &= \frac{x^2}{2} & \sigma_{\text{Ridge}}(x) &= \frac{x^2}{2}\end{aligned}$$



Future Directions- LN models

$$y_i = \eta(X_i \cdot w^0) + \epsilon_i$$

Future Directions - Random Dimensionality Reduction

Future Directions - Phase Transitions in Clustering

Future Directions - Phase Transitions in Clustering