Conditional Expectation: Linear Regression as A Computational Tool

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CEF: Motivation, Definition, Properties, and Use

CEF: Linear Regression as a computational tool

CEF: Discrete Variable and Linearity

When is linear regression not working? And what to do?

CEF: Motivation, Definition, Properties, and Use

Motivation

Contrary to the **unconditional expectation** \mathbb{Y} , the conditional expectation is a **random variable** (since $f_{Y|X}(y|X)$ is a random function)!

We can think of Y as a function of X:

$$\mathbb{E}[Y|X] = g(X)$$
 for some function $g(\cdot)$

where $g(\cdot)$ is a function determined by the joint (and hence conditional) distribution of (X, Y).

Our goal: To figure out what is g(X)!

Conditional Expectation

$$\mathbb{E}[Y \mid X]$$

has a value for every value of X. If X is multi-dimensional, X_1, X_2, \ldots, X_k just for **every** combination of the values of X_1, X_2, \ldots, X_k .

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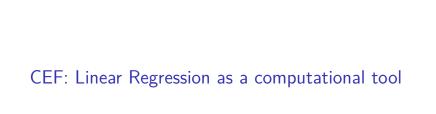
Wage for different race (black vs non-blacks) and marital status (married vs non-married) 3. Wage for different race (black vs non-blacks) and marital status (married vs non-married)

$$\mathbb{E}[\text{wages}|\text{married,black}] = \begin{cases} 841.9756 & \text{if unmarried non-blacks} \\ 1007.2797 & \text{if married non-blacks} \\ 600.1111 & \text{if unmarried blacks} \\ 759.7941 & \text{if married blacks} \end{cases}$$

Use of CEF

We focus on the CEF for the following two purposes

- 1. Prediction
- 2. Marginal Effects



As emphasized earlier, we focus our attention on discrete X first. It		
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As emphasized earlier, we focus our attention on discrete *X* first. It is treated as a **factor variable**. Every value of *X* should be considered as a separate category and coresponds to a new variable indicating whether or not the observation belonging to this category.

In a model, if you type factor(x), R will automatically create the dummy variables for each possible value of x and include them properly in estimations.

Dummy Variable or Indicator Variables

A **binary** (or **dummy** variable takes the value of one for some observations to indicate membership in a group and zero for the remaining observations.

For a discrete variable, x, that can take on two values, say $\{0,1\}$, there are two corresponding dummy variables:

$$x_1 = \mathbb{I}[x = 0]$$
$$x_2 = \mathbb{I}[x = 1]$$

Example

Χ	$x_1 = \mathbb{I}[x = 0]$	$x_2 = \mathbb{I}[x = 1]$
1	0	1
0	1	0
1	0	1
1	0	1
0	1	0

In this special case of binary independent variable, note that $x_2 = x$.

If X can take only four different values (say, 1, 2, 3, 4 or first, second, third, fourth seasons), then we need to create four **additional** dummy variables, denoted by x_1, x_2, x_3, x_4 :

 $x_4 = \mathbb{I}[X = 4]$

$$x_1 = \mathbb{I}[X = 1]$$

 $x_2 = \mathbb{I}[X = 2]$
 $x_3 = \mathbb{I}[X = 3]$

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2. With intercept

$$y = \beta_0 + \beta_2 x_2 + \beta_3 \cdot x_3 + \beta_4 \cdot x_4 + \epsilon$$

These two equivalent forms of linear regression will produce **identical** predictions, but the coefficients will have different interpretations.

Let's now examine it using the example of wages and marital status (married vs non-married) as an example

Example: 01_linear_regression_example01.do

Lessons Learned

- 1. Use of a discrete predictor, two approaches deliver the same predictions
 - (a) An indicator for all possible values: directly informing **predictions**
 - (b) Omitting one of them but include an intercept: directly informing **partial effects**
- 2. Linear regressions are equilvalent to find conditional means!
- 3. OLS are equivalent to finding conditional averages!

Question: Why?

Consider first how we solve the OLS problem

$$\sum_{i=1}^{N} \epsilon_i^2 = \sum_{i=1}^{N} (\mathsf{wage}_i - \beta_1 \cdot \mathsf{non\text{-}married}_i - \beta_2 \cdot \mathsf{married}_i)^2$$

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$$\begin{split} \sum_{i=1}^{N} \epsilon_{i}^{2} &= \sum_{i=1}^{N} \left(\mathsf{wage}_{i} - \beta_{1} \cdot \mathsf{non\text{-}married}_{i} - \beta_{2} \cdot \mathsf{married}_{i} \right)^{2} \\ &= \sum_{\mathsf{non\text{-}married}} \left(\mathsf{wage}_{i} - \beta_{1} \cdot \mathsf{non\text{-}married}_{i} - \beta_{2} \cdot \mathsf{married}_{i} \right)^{2} \\ &+ \sum_{\mathsf{married}} \left(\mathsf{wage}_{i} - \beta_{1} \cdot \mathsf{non\text{-}married}_{i} - \beta_{2} \cdot \mathsf{married}_{i} \right)^{2} \end{split}$$

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 (β_1, β_2) that minimize

$$\sum_{i=1}^{N} \epsilon_i^2 = \sum_{\text{non-married}} (\mathsf{wage}_i - \beta_1)^2 + \sum_{\text{married}} (\mathsf{wage}_i - \beta_2)^2$$

is the same as β_1 that minimizes

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In your homework, you will be asked to show that

$$v = \beta_0 + \beta_1 \cdot D + \epsilon$$

 $= \mathbb{E}[y \mid D = 1] - \mathbb{E}[y \mid D = 0]$

where D is a binary variable = $\{0,1\}$

$$eta^{\mathsf{linear regression}} = rac{\mathsf{cov}(y, D)}{\mathsf{var}(D)}$$

Quantile Regression Model and Conditional Quantile

Quantile regression estimates the following model

$$Q_{\tau}(y|x) = x'\beta(\tau)$$

 β is chosen to minimize the following objective function

$$\sum \rho(y_i - x_i'\beta(\tau))$$

where $\rho(\cdot) = \tau \cdot |y_i - x_i'\beta(\tau)|$ if $y_i - x_i'\beta(\tau) \ge 0$ (positive error) and $\rho(\cdot) = (1 - \tau) \cdot |y_i - x_i'\beta(\tau)|$ if $y_i - x_i'\beta(\tau) < 0$ (negative error)

Using the same technique above, you can easily show that the following model with a discrete variable x_1 :

$$Q_{\tau}(y|x) = \beta_0(\tau) + x_1\beta_1(\tau)$$

The quantile regression coefficients, $\beta_0(\tau)$, $\beta_1(\tau)$, are nothing but the sample quantile for the ommitted group and the difference in the sample quantiles between the included group and the ommitted group, respectively.

CEF: Discrete Variable and Linearity

Why does it work in general? (Intuition)

Suppose that

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_k x_k + \epsilon$$

Regression coefficients are those, $\beta_1, \beta_2, \dots, \beta_k$, that solve the following minimization problem.

min
$$\mathbb{E}[(y - (\beta_0 + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_k x_k))^2]$$

\$

Intuition: The linear regression coefficient is the best **linear** prediction in the sense that it minimizes mean squared error!

Note that when you choose $\beta_1, \beta_2, \dots, \beta_k$, you are choosing a function, m(x), but a linear one!

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Conditional Expectation, $\mathbb{E}[y|x]$ solves the following minimization problem.

min
$$\mathbb{E}[(y-m(x))^2]$$

If the conditional expectation function, $\mathbb{E}[y|x]$, is indeed linear, and of course, the regression coefficient is just the CEF!!

In other words, **if the CEF** is linear, then you can just use regression to obtain the CEF!

Now, let's go through the previous examples and see why the CEFs are linear.

Prior to continuing, it is worth pointing out that linearity refers to linearity in coefficients.

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Examples of Linear Models:

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- $1. \ \ y = \beta_0 + \beta_1 x_1 + \epsilon$
- 2. $y = \beta_1 x_1 + \beta_2 x_2 + \epsilon$

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Examples of Linear Models:

- 1. $\mathbf{v} = \beta_0 + \beta_1 \mathbf{x}_1 + \epsilon$
- $2. \ \ y = \beta_1 x_1 + \beta_2 x_2 + \epsilon$
- 3. $y = \beta_1 x_1 + \beta_2 \exp(x_2) + \beta_3 x_1^{10} + \epsilon$

As long as we can write it in the following form

$$y = \beta_0 + \beta_1 \cdot x_1 + \beta_2 \cdot x_2 + \dots + \beta_k \cdot x_k + \epsilon$$

Examples of non-linear regression models

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$$y = \beta_0 + \beta_1 \cdot \beta_0 \cdot x_1 + \epsilon$$

$$2. \ y = \exp(\beta_0 + \beta_1 \cdot x_1 + \epsilon)$$

$$\mathbb{E}[\mathsf{wages}|\mathsf{married}] = \begin{cases} 798.4400 & \mathsf{if married} = 0 \\ 977.0479 & \mathsf{if married} = 1 \end{cases}$$



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 $\mathbb{E}[\mathsf{wages}|\mathsf{married}] = 798.4400 \cdot (1-\mathsf{married}) + 977.0479 \cdot \mathsf{married}$

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$$\begin{split} \mathbb{E}[\mathsf{wages}|\mathsf{married}] &= 798.4400 \cdot (1-\mathsf{married}) + 977.0479 \cdot \mathsf{married} \\ &= \gamma_1 x_0 + \gamma_1 x_1 \end{split}$$

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$$\mathbb{E}[\mathsf{wages}|\mathsf{married}] = 798.4400 + (977.0479 - 798.4400) \cdot \mathsf{married}$$

$$\mathbb{E}[\mathsf{wages}|\mathsf{married}] = \begin{cases} 798.4400 & \mathsf{if married} = 0 \\ 977.0479 & \mathsf{if married} = 1 \end{cases}$$

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$$\mathbb{E}[\text{wages}|\text{married}] = 798.4400 + (977.0479 - 798.4400) \cdot \text{married}$$

= $798.4400 + 178.6079 \cdot \text{married}$

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These two models are equilvalent but the interpretation of the coefficients has changed.

- 1. Group averages from each sub-group
- Partial Effects: Everything relative to the base group. Group averages for married people are the sum of group averages for non-married people and the difference in group averages (i.e., marriage premium).

By Decomposition Property:

$$y = \mathbb{E}[y|x] + \epsilon$$

We can write our wage equation as follows

wage =
$$\gamma_1 \cdot \text{non-married} + \gamma_2 \cdot \text{married} + \epsilon$$

wage =
$$\beta_0 + \beta_1 \cdot \text{married} + \epsilon$$

Lets look at the education example

 $\mathbb{E}[\text{wage}|\text{education}] = \begin{cases} 774.2500 & \text{if educ} = \text{below high school} \\ 862.6718 & \text{if educ} = \text{high school} \\ 1076.0242 & \text{if educ} = \text{above high school} \end{cases}$

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We can write our model as follows:

$$\begin{split} \text{wage} &= 774.2500 \cdot \text{below high school} + 862.6718 \cdot \text{high school} \\ &+ 1076.0242 \cdot \text{above high school} + \epsilon \\ &= \beta_1 \cdot \text{below high school} + \beta_2 \cdot \text{high school} \\ &+ \beta_3 \cdot \text{above high school} + \epsilon \end{split}$$

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Or, you can similarly manipulate it into another form of conditional expectation function:

 $\mathsf{wage} = 774.2500 + 88.42 \cdot \mathsf{high} \ \mathsf{school} + 301.77 \cdot \mathsf{above} \ \mathsf{high} \ \mathsf{school} + \epsilon$

key result:

The conditional mean function is linear whenever all regressors are discrete (i.e., take on only a finite number of possible values).

Example: What is the relationship between marital status, races and wages?

When is linear regression not working? And what to do?

Motivating Example

Let's look at our example code first: 01_linear_regression_example02.do

Two Lessons:

- 1. The conditional mean is better. *Prediction Property*
- 2. The linear regression result is not too far off (We will have a name for this!).

Now, lets try to understand why our linear regression is not delivering the best predictions or our conditional means.

 $\mathbb{E}[\text{wages}|\text{married,black}] = \begin{cases} 841.9756 & \text{if unmarried non-blacks} \\ 1007.2797 & \text{if married non-blacks} \\ 600.1111 & \text{if unmarried blacks} \\ 759.7941 & \text{if married blacks} \end{cases}$

From last semester, If Q is a statement, such as "Last name is Wang", then $\mathbb{I}[Q]$ is set to 1 when the statement Q is true, and zero otherwise. For an event

$$A \in \mathcal{F}\mathbb{I}[\omega \in A] = egin{cases} 1 & ext{if } \omega \in A \ 0 & ext{otherwise} \end{cases}$$

Properties of Indicator Function: If $A_1, A_2, \dots A_N$, then

- 1. $\mathbb{I}[\cap A_N] = \prod_{n=1}^N \mathbb{I}[A_n]$
 - 2. $\mathbb{I}[\cup A_N] = \sum_{n=1}^{N} \mathbb{I}[A_n]$ whenever the sets are disjoint

unmarried, non-blacks:	$\mathbb{I}[married = 0 \ and \ black = 0]$
	$= (1 - married) \cdot (1 - black)$

unmarried, non-blacks: $\mathbb{I}[\mathsf{married} = 0 \text{ and black} = 0]$ = $(1 - \mathsf{married}) \cdot (1 - \mathsf{black})$

To see why,

1. If married=0 and black=0, $(1 - married) \cdot (1 - black) = 1$

unmarried, non-blacks: $\mathbb{I}[\text{married} = 0 \text{ and black} = 0]$ = $(1 - \text{married}) \cdot (1 - \text{black})$

To see why,

- 1. If married=0 and black=0, $(1 \text{married}) \cdot (1 \text{black}) = 1$
- 2. If either married $\neq 0$ or black $\neq 0$, $(1 \text{married}) \cdot (1 \text{black}) = 0$

Similarly, we can write all the possible values in the following forms

married,non-blacks = $\mathbb{I}[\text{married}=1,\text{black}=0]$ = married $\cdot (1 - \text{black})$ Similarly, we can write all the possible values in the following forms

$$\begin{aligned} \mathsf{married}, \mathsf{non\text{-}blacks} &= \mathbb{I}[\mathsf{married}\text{=}1, \mathsf{black}\text{=}0] \\ &= \mathsf{married} \cdot (1 - \mathsf{black}) \end{aligned}$$

non-married,blacks =
$$\mathbb{I}[\text{married}=0,\text{black}=1]$$

= $(1 - \text{married}) \cdot \text{black}$

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$$married, non-blacks = I[married=1, black=0]$$

= $married \cdot (1 - black)$

non-married,blacks =
$$\mathbb{I}[\text{married}=0,\text{black}=1]$$

= $(1 - \text{married}) \cdot \text{black}$

$$\begin{aligned} \mathsf{married}, \mathsf{blacks} &= \mathbb{I}[\mathsf{married}{=}1, \mathsf{black}{=}1] \\ &= \mathsf{married}{\cdot}\mathsf{black} \end{aligned}$$

$$\mathbb{E} \big[\text{wages} | \text{married,black} \big] = \begin{cases} \mu_{00} = 841.9756 & \text{if married} = 0 \text{ and black} = 0 \\ \mu_{10} = 1007.2797 & \text{if married} = 1 \text{ and black} = 0 \\ \mu_{01} = 600.1111 & \text{if married} = 0 \text{ and black} = 1 \\ \mu_{11} = 759.7941 & \text{if married} = 0 \text{ and black} = 1 \end{cases}$$

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$$\Longrightarrow$$

$$\begin{split} \mathbb{E}[\mathsf{wages}|\mathsf{married},\mathsf{black}] &= \mu_{00} \cdot (1 - \mathsf{married}) \cdot (1 - \mathsf{black}) \\ &+ \mu_{10} \cdot \mathsf{married} \cdot (1 \text{-black}) \\ &+ \mu_{01} \cdot (1 - \mathsf{married}) \cdot \mathsf{black} \\ &+ \mu_{11} \cdot \mathsf{married} \cdot \mathsf{black} \\ \\ &= \mu_{00} \cdot x_1 + \mu_{10} \cdot x_2 \\ &+ \mu_{01} \cdot x_3 + \mu_{11} \cdot x_4 \end{split}$$

where x_1 can be thought of as a new variable (1 - married)(1 - black), other x_1 are similarly defined based on the variables in the equation.

Let's use this insight to implement the correct linear regression
Lets look at our example: mv06_cond_expectation04.Rmd

Why do I want to go through all the troubles to manipulate my four
binary variables into the products between married and black?

We can show the missing part in the original regression model when

you include only two variables (married and black).

$$\begin{split} \mathbb{E}[\mathsf{wages}|\mathsf{married},\mathsf{black}] &= \mu_{00} \cdot (1-\mathsf{married}) \cdot (1-\mathsf{black}) \\ &+ \mu_{10} \cdot \mathsf{married} \cdot (1\!-\!\mathsf{black}) \\ &+ \mu_{01} \cdot (1-\mathsf{married}) \cdot \mathsf{black} \\ &+ \mu_{11} \cdot \mathsf{married} \cdot \mathsf{black} \end{split}$$

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$$\begin{split} \mathbb{E}[\mathsf{wages}|\mathsf{married},\mathsf{black}] &= \mu_{00} \cdot (1 - \mathsf{married}) \cdot (1 - \mathsf{black}) \\ &+ \mu_{10} \cdot \mathsf{married} \cdot (1 \text{-black}) \\ &+ \mu_{01} \cdot (1 - \mathsf{married}) \cdot \mathsf{black} \\ &+ \mu_{11} \cdot \mathsf{married} \cdot \mathsf{black} \\ &= \mu_{00}[1 - \mathsf{black} - \mathsf{married} + \mathsf{married} \cdot \mathsf{black}] \\ &+ \mu_{10}[\mathsf{married} - \mathsf{married} \cdot \mathsf{black}] \\ &+ \mu_{01}[\mathsf{black} - \mathsf{married} \cdot \mathsf{black}] + \mu_{11} \cdot \mathsf{married} \cdot \mathsf{black} \\ &= \mu_{00} + (-\mu_{00} + \mu_{01}) \cdot \mathsf{black} + (-\mu_{00} + \mu_{10}) \cdot \mathsf{married} \\ &+ (\mu_{00} - \mu_{10} - \mu_{01} + \mu_{11}) \mathsf{married} \cdot \mathsf{black} \end{split}$$

$$\begin{split} \mathbb{E}[\mathsf{wages}|\mathsf{married},\mathsf{black}] &= \mu_{00} \cdot (1 - \mathsf{married}) \cdot (1 - \mathsf{black}) \\ &+ \mu_{10} \cdot \mathsf{married} \cdot (1 \text{-black}) \\ &+ \mu_{01} \cdot (1 - \mathsf{married}) \cdot \mathsf{black} \\ &+ \mu_{11} \cdot \mathsf{married} \cdot \mathsf{black} \\ &= \mu_{00}[1 - \mathsf{black} - \mathsf{married} + \mathsf{married} \cdot \mathsf{black}] \\ &+ \mu_{10}[\mathsf{married} - \mathsf{married} \cdot \mathsf{black}] \\ &+ \mu_{01}[\mathsf{black} - \mathsf{married} \cdot \mathsf{black}] + \mu_{11} \cdot \mathsf{married} \cdot \mathsf{black} \\ &= \mu_{00} + (-\mu_{00} + \mu_{01}) \cdot \mathsf{black} + (-\mu_{00} + \mu_{10}) \cdot \mathsf{married} \\ &+ (\mu_{00} - \mu_{10} - \mu_{01} + \mu_{11}) \mathsf{married} \cdot \mathsf{black} \\ &= \beta_0 + \beta_1 \cdot \mathsf{black} + \beta_2 \cdot \mathsf{married} + \beta_3 \cdot \mathsf{married} \cdot \mathsf{black} \end{split}$$

By Decomposition Property of the CEF, the outcome variable should be written as follows

wage ==
$$\beta_0 + \beta_1 \cdot \text{black} + \beta_2 \cdot \text{married} + \beta_3 \cdot \text{married} \cdot \text{black} + \epsilon$$

Let's use this insight to implement the correct linear regression
Lets look at our example: mv06_cond_expectation04.Rmd

wages =
$$\gamma_0 + \gamma_1 \cdot \text{married} + \gamma_2 \cdot \text{black} + \epsilon$$

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1. For blacks: wage difference between married and non-married workers:

$$(\gamma_0 + \gamma_1 \cdot 1 + \gamma_2 \cdot 1 + \epsilon) - (\gamma_0 + \gamma_1 \cdot 0 + \gamma_2 \cdot 1 + \epsilon) = \gamma_1$$

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2. For non-blacks: wage difference between married and non-married workers:

$$(\gamma_0 + \gamma_1 \cdot 1 + \gamma_2 \cdot 0 + \epsilon) - (\gamma_0 + \gamma_1 \cdot 0 + \gamma_2 \cdot 0 + \epsilon) = \gamma_1$$

$$\mathsf{wages} = \gamma_0 + \gamma_1 \cdot \mathsf{married} + \gamma_2 \cdot \mathsf{black} + \epsilon$$

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Why does marriage premium have to be the same across races?

wages =
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For blacks: wage difference between married and non-married workers:

$$(\beta_0 + \beta_1 \cdot 1 + \beta_2 \cdot 1 + \beta_3 \cdot 1 \cdot 1 + \epsilon) - (\beta_0 + \beta_1 \cdot 0 + \beta_2 \cdot 1 + \beta_3 \cdot 0 \cdot 1 + \epsilon) = \beta_1 + \beta_3$$

$$\mathsf{wages} = \beta_0 + \beta_1 \cdot \mathsf{married} + \beta_2 \cdot \mathsf{black} + \beta_3 \cdot \mathsf{married} \cdot \mathsf{black} + \epsilon$$

For blacks: wage difference between married and non-married workers:

$$(\beta_0 + \beta_1 \cdot 1 + \beta_2 \cdot 1 + \beta_3 \cdot 1 \cdot 1 + \epsilon) - (\beta_0 + \beta_1 \cdot 0 + \beta_2 \cdot 1 + \beta_3 \cdot 0 \cdot 1 + \epsilon) = \beta_1 + \beta_3$$

For non-blacks: wage difference between married and non-married workers:

$$\mathsf{wages} = \beta_0 + \beta_1 \cdot \mathsf{married} + \beta_2 \cdot \mathsf{black} + \beta_3 \cdot \mathsf{married} \cdot \mathsf{black} + \epsilon$$

For blacks: wage difference between married and non-married workers:

$$(\beta_0+\beta_1\cdot 1+\beta_2\cdot 1+\beta_3\cdot 1\cdot 1+\epsilon)-(\beta_0+\beta_1\cdot 0+\beta_2\cdot 1+\beta_3\cdot 0\cdot 1+\epsilon)=\beta_1+\beta_3$$

$$(\beta_0 + \beta_1 \cdot 1 + \beta_2 \cdot 0 + \beta_3 \cdot 1 \cdot 0 + \epsilon) - (\beta_0 + \beta_1 \cdot 0 + \beta_2 \cdot 0 + \beta_3 \cdot 0 \cdot 0 + \epsilon) = \beta_1$$

Truth: (average) marriage premium varies across races!!

Heterogenous Partial Effects and Interactions

Including only married and black in your model artificially impose the assumption that the partial effect is the same across groups.

Heterogenous Partial Effects (across groups) are captured by the interactions between the main variable of interest and the group variable that you think the partial effects may vary with.

For example,

1. The effects of education on wages may depend on the education level itself (decreasing marginal returns)

wages=
$$\beta_0 + \beta_1 \cdot \text{educ} + \beta_2 \cdot \text{educ} \cdot \text{educ} + \epsilon$$

Synergy effects in marketing: The effects of spending on TV advertising on sales may depend on the level of spending on radio advertising, too!

 $\textit{textsales} = \beta_0 + \beta_1 \cdot \mathsf{TV} \text{ expenditures} + \beta_2 \cdot \mathsf{Radio} \text{ expenditures} + \beta_3 \cdot \mathsf{TV} \cdot \mathsf{Radio}$

Fact:

In general, if there are p dummy variables, x_1, x_2, \ldots, x_p , then the CEF $\mathbb{E}[y|x_1, x_2, \ldots, x_p]$ takes at most 2^p distinct values, and can be written as a linear function of the 2^p regressors including x_1, x_2, \ldots, x_p and all cross products. Cannot be practically!

$$\mathbb{E}[y|x_1, x_2, x_3] = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_3 + \beta_4 x_1 x_2 + \beta_5 x_1 x_3 + \beta_6 x_2 x_3 + \beta_7 x_1 x_2 x_3$$

Let's use this insight to implement the correct linear regression
Lets look at our example:
01_linear_regression_example02.do

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- 3. Two forms of the conditional mean
- (a) Prediction form: Include dummy variables for all possible levels (exclusive groups), but omit an intercept
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- (a) Prediction form: Include dummy variables for all possible levels (exclusive groups), but omit an intercept
- (b) Partial Effects form: Include an intercept and omit one of the dummy variables. The partial effects is the group difference.
 - 4. In the case of all discrete variables, linear regression delivers the Conditional Mean.
 - (a) Straightforward case: binary or multi-valued cases
 - (b) Less straightforward case: two or more than two binary variables (interactions)