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# Granular Asset Pricing

Junxiong Gao

October 20, 2022

# Granularity in macroeconomics

Granular Asset Pricing

Introduction

Granularity: fat tail distribution of firm size. Large firms.

- Firm size is measured by fundamentals (value of product, number of employee, etc).
- For example, Nokia in 2000. 1.6 % of Finland's GDP growth.
- Implication: Idiosyncratic shocks of large firms
  - Impact on the aggregate output. (Gabaix (2011), Acemoglu et al. (2012))
  - Help identify price elasticity of aggregate demand/supply. (Gabaix and Koijen (2020)).

# Granularity in stock market

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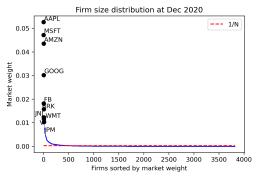
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 Nokia in 2000: 60% of market cap in the Finnish market index.

Market weight of US firms in 2020:



The 10 largest firms account for over 25 percents of the total market value (about 4,000 firms).

• Similar finding over time. • evidence



#### Theoretical motivation

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- Classical view in asset pricing:
  - Common risk factors in asset returns.
  - Idiosyncratic risks are defined relative to factors.
  - Only common factors explain expected returns. Idiosyncratic risks do not.
- Key assumption: diversification in market.
  - No large firms in a market with sufficiently many assets.
  - Idiosyncratic risks are diversified away.
- This paper: a granular asset pricing model
  - Granularity/Fat tail distribution⇒ Large firms. Failure of diversification.
  - Idiosyncratic risk explains expected returns.

# Theoretical framework: APT+granularity

Granular Asset Pricing

Introduction

 APT: two independent components in shocks of asset returns:

$$r_i - E[r_i] = \sum_{s=1}^k \beta_{i,s} f_s + \epsilon_i$$

Factors  $f_{s=1...k} \perp \text{Idiosyncratic shocks } \epsilon_i \longrightarrow \text{discussion}$ 



- Definition of f and  $\epsilon$  is independent of firm size distribution.
  - Statistical criteria based on covariance among shocks.
  - f (strong correlation) v.s.  $\epsilon_i$  (weak/no correlation)
- Add regulating condition from firm size distribution perspective.
  - Thin tail. Diversification. Factor model results.
  - With granularity. Variance of  $\epsilon_i$  explains the expected returns.

# Model Implication

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- In the cross-section, the impact of  $\epsilon_i$  depends on firm size.
  - Large firms have their idiosyncratic risk explain expected returns.
  - Small firms do not.
- On aggregate, Idiosyncratic risk of large firms also explains market variation.
- Quantify these implications of granularity by fitting Pareto distribution of firm size
  - Widely used in macroeconomic literature to fit fat tail distribution. Gabaix (1999), Gabaix (2011), Acemoglu, Ozdaglar, and Tahbaz-Salehi (2015), etc.

#### Related literature

Granular Asset Pricing

Introduction

• Granularity in macroeconomic: Gabaix (2011), Acemoglu et al. (2012), Acemoglu, Akcigit, and Kerr (2016), Gabaix and Koijen (2020)

#### Asset pricing theory/test

- APT models Chamberlain and Rothschild (1983), Chamberlain (1983), Dybvig (1983), Connor and Korajczyk (1993), Connor and Korajczyk (1995), Huberman (2005)
- Test of factors: Feng, Giglio, and Xiu (2020), Kelly, Pruitt, and Su (2020), Giglio, Xiu, and Zhang (2021), Giglio and Xiu (2021), Giglio, Kelly, and Xiu (2022)
- Idiosyncratic risk and expected returns: Ang et al. (2006), Ang et al. (2009), Hou and Loh (2016) Campbell et al. (2001), Xu and Malkiel (2003), Goyal and Santa-Clara (2003) and Herskovic et al. (2016).
- This paper: granularity's asset pricing implication.

# Key result with granularity

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Expected returns implied by model:

$$E[r_i] - r_f = \text{factor terms}_i + \gamma w_i \theta_i$$

- ullet  $\gamma$  risk aversion coefficient.
- w<sub>i</sub> market weight.
- $\theta_i$  variance of idiosyncratic shocks  $\epsilon_i$ .
- $w_i\theta_i$  size-adjusted idiosyncratic risk explains the expected returns.
  - Large firms. High  $w_i$ , high impact of idiosyncratic risk.
  - Small firms. Negligible  $w_i$ .  $w_i\theta_i \to 0$ . No impact.

# Implication: Granularity and idiosyncratic risk puzzle

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Granularity and idiosyncratic risk

 Idiosyncratic risk puzzle in Ang et al. (2006) and Ang et al. (2009) (IRP hereafter). Estimate:

$$E[r_i] - r_f = \text{factor terms} + \eta \sqrt{\theta_i}$$

- $\hat{\eta} < 0$ . Negative relation between idiosyncratic risk and return.
  - No satisfying enough explanation in literature (Hou and Loh (2016)).
- This paper:
  - Explain the puzzle. Misspecification if ignoring the size difference among firms.
  - Identify a positive risk-return relation using  $w_i\theta_i$ . Risk aversion coefficient  $\hat{\gamma} > 0$ .

# Implication: Granularity and aggregate market variation

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• Market expected return.  $E[r_m] = \sum_i w_i E[r_i]$ 

$$E[r_m] - r_f = \sum w_i(w_i\theta_i) + \text{factor terms}$$

- Relate to whether idiosyncratic risk matters for market returns in literature.
  - $\theta \uparrow$  overall, more aggregate risk, more expected returns.
  - Time-series implication in Campbell et al. (2001), Goyal and Santa-Clara (2003), Bali et al. (2005).
- This paper:
  - Granularity ↑. Idiosyncratic risks are less diversified on aggregate.
  - Time-series implication. Controlling for the magnitude of idiosyncratic risk, the level of granularity explains market expected returns.

# Expected returns: APT+granularity

Granular Asset Pricing

#### Theoretical framework

- n assets. w<sub>i</sub> market weight for each. A representative agent maximizes expected utility based on portfolio return  $r_m = \sum_i w_i r_i$ .
- CARA utility for simplicity. Shocks of the pricing kernel:

$$-\gamma \sum w_i \left(\beta_i f + \epsilon_i\right) = -\gamma \sum_i^n w_i \beta_i f - \gamma \sum_i^n w_i \epsilon_i$$

 If firm size is granular in equilibrium, what's the expected returns?

$$E[r_i] - r_f = \sum_{s=1}^{k} \beta_{i,s} \mu_s + \gamma COV(\epsilon_i, \sum_{i}^{n} w_i \epsilon_i)$$
(1)
APT granularity

 $\mu_s$  risk premium for factor  $f_s$ ,  $\gamma$  risk aversion coefficient.

### Diversification and factor models

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Diversification assumed by APT models:

$$\lim_{n \to \infty} \sum w_i^2 = 0 \tag{2}$$

- ullet No granularity. All firms have market weight  $w_i 
  ightarrow 0$
- The granularity terms  $COV(\epsilon_i, \sum_i^n w_i \epsilon_i)$  converge to zero
- Classical implication. A multi-factor model for expected returns?

$$E[r_i] - r_f = \sum_{s=1}^k \beta_{i,s} \mu_s$$

• The diversification  $\lim_{n\to\infty} \sum w_i^2$  depends on firm size distribution.

### Thin tail size distribution and diversification

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• Work on the market value  $X_i$ , assume i.i.d for simplicity.  $w_i = X_i / \sum_{i=1}^n X_i$ 

• The convergence of  $\sum w_i^2$  depends on the first and second moments of  $X_i$ :

$$\lim_{n\to\infty} \sum w_i^2 = \lim_{n\to\infty} \sum \frac{(X_i)^2}{(\sum X_i)^2} = \lim_{n\to\infty} \left(\frac{1}{n}\right) \frac{1/n\sum (X_i)^2}{(1/n\sum X_i)^2}$$
(3)

ullet A thin-tail distribution of  $X_i$ . Finite mean and variance

$$\lim_{n\to\infty} \sum w_i^2 = \lim_{n\to\infty} \left(\frac{1}{n}\right) \frac{E[X^2]}{E[X]^2} = 0$$

•  $\sum w_i^2$  scales as 1/n

# Granularity and failure of diversification

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- Quantify granularity by the Pareto distribution (Gabaix (2011)). Two implications:
  - With granularity. Infinite moments of  $X_i$ . Failure of diversification.

$$\lim_{n\to\infty}\sum w_i^2=\lim_{n\to\infty}\frac{1}{n}\frac{1/n\sum(X_i)^2}{(1/n\sum X_i)^2}\neq 0$$

• Large firms have non-negligible market weight.

$$\lim_{n\to\infty}w_i\neq 0$$

 Asset pricing result: Idiosyncratic risk explains the expected returns in cross-section and on aggregate.

# Quantify the granularity by the Pareto distribution

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Pareto distribution with survival probability:

$$P(X_i > x) = \left(\frac{x}{X_m}\right)^{-\xi}, x > X_m \tag{4}$$

fit for large firms over a threshold  $x_m$ 

- The Pareto coefficient  $\xi$  determines thickness of the tails (granularity). High  $\xi \to \text{low}$  level of granularity.
  - ullet  $\zeta$  is estimated to be around 1 in literature (Zipf's law)
  - $\bullet$   $\zeta$  < 2, granularity, failure of diversification.
  - $\bullet$   $\zeta >$  2, finite first and second moments, diversification.
- A linear relation between i (rank) and  $X_i$  (size)

$$\log(i/n) \approx \log(X_i/X_m)^{-\zeta} = -\zeta(\log X_i - \log X_m)$$

# Fit of the Pareto distribution: log-log plot

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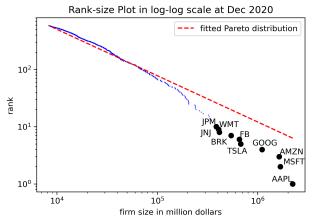
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As an example, fit the largest 20% firms at Dec 2020 with a Pareto distribution. Check the linear relation between rank and size.



### Pareto distribution and failure of diversification

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#### Theoretical framework

• Pareto distribution  $\zeta < 2$ 

$$\lim_{n \to \infty} \sum w_i^2 = c_1 \frac{Y_2 + 1}{(Y_1)^2}, \zeta < 1$$
 (5)

$$= c_2 n^{2/\zeta - 2} \frac{Y_2 + 1}{E[X]^2}, \zeta > 1$$
 (6)

where  $c_1$ ,  $c_2$  are constants under different range of  $\zeta$ .

- $Y_1$ ,  $Y_2$  are the non-degenerate terms due to infinite moments
- Failure of diversification. For example, let  $n = 10^5$  and  $\zeta = 1.1.$ 
  - $\sum w_i^2$  scales as  $n^{2/\zeta-2} = 1/10$
  - Instead of 1/n = 1/10000 in thin tail case.

# Pareto distribution and large firms

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•  $Y_1$  comes from the extreme values. Leads to infinite first moment, such that  $X_i > a_n$ 

$$a_n = \inf\{x : P(X_i > x) \le n^{-1}\} = n^{1/\xi}$$

- Stable law: need to adjust for these extreme values to regulate the convergence.
- A firm with extreme size  $a_n$  have weights in market portfolio equals to  $\frac{a_n}{\sum X_i}$

$$\lim_{n \to \infty} \frac{a_n}{\sum X_i} = 1/(1+Y_1), \xi < 1$$
 (7)

$$= n^{1/\zeta - 1} (1 - 1/\zeta), \, \xi > 1 \tag{8}$$

- Same example.  $n=10^5$  and  $\zeta=1.1$ .
  - The market weight of a large firm scales as  $a_n$  is roughly 0.03.

# Asset pricing implications

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• For a firm large enough  $\lim_{n\to\infty} w_i \neq 0$ . Assume  $\epsilon_i$  independent simplifies:

$$\lim_{n\to\infty} COV(\epsilon_i, \sum_{i=1}^{n} w_i \epsilon_i) \underset{\epsilon_i \text{independent } n\to\infty}{\approx} \lim_{n\to\infty} w_i \theta_i \neq 0$$

- $\bullet$   $\theta_i$  variance of idiosyncratic shocks.  $w_i$  market weight
- A "granular alpha", abnormal return relative to factors:

$$\alpha_i = \gamma w_i \theta_i$$

• On aggregate, if diversification fails  $\lim_{n\to\infty} \sum w_i^2 \neq 0$  $\sum w_i(w_i\theta_i)$  matters for market expected return.

# 5 portfolios sorted by idiosyncratic risk

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Portfolio level

 Replicate Ang et al. (2006) to form five portfolios sorted by  $\theta$ 

- FF 3 factors as benchmark. Measure  $\theta_i$  by daily returns in each month.
- Compute w<sub>i</sub> total market weight of assets in each portfolio.

Panel A: alpha relative to FF3								
•	L	2	3	4	Н	L-H		
$\sqrt{\theta}_{FF3}$	2.82	3.97	5.80	8.96	13.85			
Wi	0.60	0.23	0.11	0.05	0.02			
$\alpha_{FF3}$	1.18	-0.20	-0.44	-5.29	-11.42	12.60		
T-stat	2.91	-0.38	-0.53	-3.88	-6.54	6.34		
$\alpha_{FF3}/\theta_{FF3}$	14.85	-1.29	-1.31	-6.60	-5.95			

From the lowest  $\theta$  to highest:

- Decreasing size  $w_i$ , decreasing  $\alpha_i$  (as found in literature)
- Model implies  $\alpha_i = \gamma w_i \theta_i$
- Decreasing  $\alpha_i/\theta_i$  as  $w_i \downarrow$  since  $\alpha_i/\theta_i = \gamma w_i$

# 5 portfolios sorted by idiosyncratic risk (Cont'd)

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Similar pattern. Measure  $\alpha$  relative to CAPM/PCA factors.

Pa	nel A:	alpha re	elative t	o CAP	М	
	L	2	3	4	Н	L-H
$\sqrt{\theta}_{CAPM}$	3.67	4.00	7.16	12.29	18.38	
$\alpha_{CAPM}$	1.34	-0.02	-0.50	-5.39	-10.59	11.92
T-stat	2.52	-0.04	-0.48	-2.98	-4.35	4.20
$\alpha_{CAPM}/\theta_{CAPM}$	9.94	-0.14	-0.97	-3.57	-3.13	
Panel C:	alpha	relative	to thre	ee PCA	factors	
$\sqrt{\theta}_{PC}$	12.83	15.28	17.24	19.31	20.11	
$\alpha_{PC}$	5.90	5.29	5.06	0.11	-5.88	11.79
T-stat	3.45	2.66	2.33	0.04	-2.16	5.29
$\alpha_{PC}/\theta_{PC}$	3.59	2.27	1.70	0.03	-1.45	

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# Evidence in 100 portfolios sorted

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Portfolio level

- Cross-sectional test of the relation between  $\alpha$ ,  $\theta$ , w
  - Requires a big enough cross-section for regression.
  - Same measure and portfolio construction. Extend to 100 portfolios.
  - Estimate  $\alpha_i$ ,  $\theta_i$  of the 100 portfolios.  $w_i$  total market weight of assets in each portfolio.
- Same as in IRP.

$$\alpha_i = constant + \eta \sqrt{\theta_i}$$

$$\hat{\eta} = -0.74$$
, std $(\hat{\eta}) = 0.039$ .

Compare to the granular channel of alpha

$$\alpha_i = constant + \gamma w_i \theta_i$$

$$\hat{\gamma} = 5.17$$
, std $(\hat{\gamma}) = 0.59$ .

# A granular explanation for IRP

Granular

If the data generating process is as model implied.

$$\alpha_i = \gamma w_i \theta_i$$

• The negative relation between risk  $\sqrt{\theta_i}$  and return  $\alpha_i$  is because of:

$$corr(\sqrt{\theta_i}, w_i\theta_i) = -0.72$$

driven by

- Negative relation between size and risk. (feature in data)  $corr(\sqrt{\theta_i}, w_i) = -0.68$
- Granularity  $\Rightarrow$  a few large firms. Other firms have negligible w<sub>i</sub>.
- High  $\theta_i$  firms. Low  $w_i$ . Low  $w_i\theta_i$ . Low  $\alpha_i$
- Low  $\theta_i$  firms. High  $w_i$ . High  $w_i\theta_i$ . High  $\alpha_i$ .

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#### Individual asset level test

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• Replicate Ang et al. (2009).  $\hat{\eta} < 0$  in:

$$r_{i,t} - r_f = controls + \eta \sqrt{\theta_{i,t-1}} + \epsilon_{i,t}$$

ullet Use the size-adjusted idiosyncratic risk instead,  $\hat{\gamma}>0$ 

$$r_{i,t} - r_f = controls + \gamma w_{i,t-1}\theta_{i,t} + \epsilon_{i,t}$$

- As in Ang et al. (2009). Fama-Macbeth approach.
  - Estimate the Fama French 3 factors model using daily returns to measure  $\theta_{i,t}$  and factor exposures in each month.
  - Cross-sectional regression in each month. Take the average estimates as  $\hat{\eta}$  and  $\hat{\gamma}$ .

#### Results of individual asset level test

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	Cross-sectional Regression, Stock Level								
	r <sub>i,t</sub> controls	r <sub>i,t</sub> controls	r <sub>i,t</sub> controls	r <sub>i,t</sub> controls	r <sub>i,t</sub> controls				
$\sqrt{\hat{\theta}_{i,t-1}}$	-0.01	-0.01			-0.01				
·	-1.98	-2.10			-2.16				
$w_{i,t-1}$		-0.10	-0.11	-1.86	-1.78				
		-0.63	-0.59	-5.05	-5.26				
$w_{i,t-1}\hat{\theta}_{i,t}$				9.15	8.95				
				8.99	8.87				

- Same controls in each column: factor exposures to FF3, lagged book to market value ratio, size, momentum (past six month returns) as in Ang et al. (2009).
- In column 1,  $\hat{\eta} < 0$ . Simply controlling for size  $w_i$  does not change the sign in columns 2.
- Functional form is important. Use size-adjusted risk gives  $\hat{\gamma} > 0$ .

#### Robustness checks for cross-sectional results

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 Benchmark results. Portfolio level and individual asset level.

- Take FF3 as factor model.
- Measure  $\theta$  using daily returns in the past month.
- Robust for longer measurement window, using daily returns in the past 3,6,12 months.
- Robust for using PCA factor models.
- The tests relies on cross-sectional difference of  $\theta$  among firms. Should be insensitive to factor model selection. For example:
  - Firm1  $\theta_1$  > Firm2  $\theta_2$
  - Different shocks/residuals from fitting different factor models. (omitted factors in Giglio and Xiu (2021), weakly tested factors in Giglio, Xiu, and Zhang (2021), etc.)
  - But similar  $\theta$ .  $\theta_1 > \theta_2$  for various factor models.

#### Time-series tests

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• The market expected return  $E[r_m] = \sum w_i E[r_i]$ 

$$\lim_{n \to \infty} E[r_m] - rf = \text{factor terms} + \gamma \lim_{n \to \infty} \sum w_i^2 \theta_i$$

- Time-series implication.
  - Two components matter for market expected returns.

$$\sum w_i^2 \theta_i = \sum w_i^2 \left( \sum \frac{w_i^2}{\sum w_i^2} \theta_i \right)$$

- Control the time-variation of idiosyncratic risk. Does level of granularity matter?
- Pareto coefficient measures the level of granularity

$$\zeta \downarrow \Rightarrow \lim_{n \to \infty} \sum w_i^2 \uparrow \Rightarrow \lim_{n \to \infty} E[r_m] - rf \uparrow$$
?

A time-series test:

$$\log r_{m,t+1} = controls + A \log \zeta_t$$

$$H_0: A = 0, H_1: A < 0$$

# Estimate of the Pareto coefficient $\zeta$

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Time-series results

 Hill estimator (see Hill (1975)). Sort all the n firm sizes in a descending order. Use the largest k firms  $X_{1,2}$  k to estimate:

$$\zeta = \left\{ 1/k \sum_{i=1}^{k} (\log X_i - \log X_k) \right\}^{-1}$$
 (9)

- Maximum likelihood estimator of  $\zeta$  conditioning on a known minimum threshold  $X_k$  (simple inference)
- Pick cutoff k proportional to number of total assets n.

$$k/n = 1\%, 5\%, 10\%, 20\%...$$

 Select k/n faces a trade-off between bias and variance. More observations. Less variance. Yet, more bias (more small firms included). Adjusting small firms included.

# Example: Hill estimator use the largest 20 % firms

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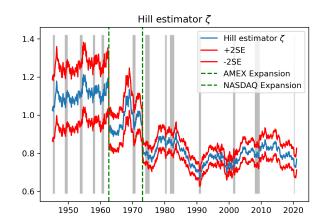
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- Dataset expands in June 1962 (AMEX) and December 1973 (NASDAQ). Over time, n increases such that:
  - Variance ↓. the estimation error decreases.
  - Bias ↑. downward bias increases due to more small firms.

# De-bias by co-integration

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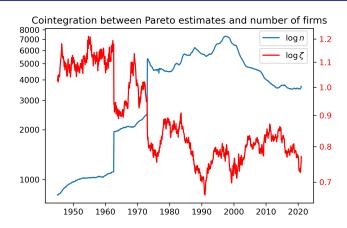
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#### For time-series test purpose:

- Form a "de-biased" predictor by subtracting the non-stationary trend due to increasing  $n_t$
- Select optimal cutoff ratio k/n for best out-of-sample results (10-fold cross-validation)

# Single variable prediction results

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 $r_{m,t+1} - r_f = constant + A \log \zeta_t(debias)$ 

 $H_0: A = 0, H_1: A < 0$ 

Panel A: Single variable prediction, multiple-horizon results						
	$\log r_{m,t  o t+1}$	$\log r_{m,t \to t+12}$	$\log r_{m,t \to t+60}$			
$\log \zeta_t$	-0.28	-2.03	-10.81			
T-stat	-2.11	-1.70	-3.42			
$R^2(\%)$	0.43	1.67	9.61			
$\log \zeta_t$ (de Stambaugh-bias)	-0.27	-2.04	-10.78			
T-stat	-1.87	-3.61	-8.77			
$Out-of-sampleR^2(\%)$	-0.17	1.50	13.34			

Panel B: Single variable prediction, sub-sample results

	Whole Sample	NBER Recession	Non-NBER Recession
$\log \zeta_t$	-0.28	-1.05	-0.10
T-stat	-2.11	-2.59	-0.76
$R^{2}(\%)$	0.43	5.00	0.05
$\log \zeta_t$ (de Stambaugh-bias)	-0.27	-1.05	-0.09
T-stat	-1.88	-4.18	-0.59

- Panel A: Significance at all predictive horizons. Positive out-of-sample  $R^2$  at the long horizon (12, 60 months).
- Panel B: More predictive power during bad times.

# Results controlling for idiosyncratic risk

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• Three measures using Fama French 3 factors, PCA and Campbell et al. (2001). •• plot

Panel	A: control ∑	$w_i\theta_i$ (FF3)	
	$\log r_{m,t \to t+1}$	$\log r_{m,t \to t+12}$	$\log r_{m,t \to t+60}$
$\log \zeta_t$	-0.27	-1.70	-6.17
T-stat	-1.91	-1.31	-1.91
$\sum w_i \theta_i(FF3)$	-0.20	-1.69	0.80
T-stat	-0.99	-0.91	0.18
$R^{2}(\%)$	0.48	1.79	3.34
$Out-of-sample R^2(\%)$	-2.45	-8.12	5.85
Panel	B: control ∑ v	$v_i\theta_i$ (PCA)	
	$\log r_{m,t \to t+1}$	$\log r_{m,t\to t+12}$	$\log r_{m,t\to t+60}$
$\log \zeta_t$	-0.26	-1.64	-5.83
T-stat	-1.81	-1.23	-1.78
$\sum w_i \theta_i(PCA)$	-0.10	-1.11	1.53
T-stat	-0.47	-0.56	0.33
$R^2(\%)$	0.33	1.12	3.47
$Out-of-sample R^2(\%)$	-2.85	-10.20	4.72
Panel C: co	ntrol $\sum w_i \theta_i$ (C	Campbell et al)	
	$\log r_{m,t \to t+1}$	$\log r_{m,t \to t+12}$	$\log r_{m,t \to t+60}$
$\log \zeta_t$	-0.28	-1.78	-6.52
T-stat	-1.94	-1.38	-2.06
$\sum w_i \theta_i$ (Campbell et al)	-0.23	-2.18	-0.21
T-stat	-1.11	-1.20	-0.05
$R^2(\%)$	0.55	2.55 ∢ □ ▶ ∢	<u>-</u> 3.29∢ <u>≥</u> ▶ ∢ <u>≥</u>
$Out - of - sampleR^2(\%)$	-2.69	-7.11	3.04

# Results controlling for other predictors

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T-stat

T-stat

 $R^{2}(\%)$ 

predictor

-3.56

6.32

1.97

12.92

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		Panel A:	Predict	tors Cont	rolled, 1	Month I	Horizon		
	bm	dspr	dp	ер	ltr	ntis	svar	tspr	corpr
$\log \zeta_t$	-0.29	-0.34	-0.25	-0.29	-0.28	-0.27	-0.28	-0.30	-0.29
T-stat	-2.13	-2.38	-1.86	-2.13	-2.22	-2.09	-2.14	-2.31	-2.30
predictor	0.13	0.20	0.26	0.22	0.37	-0.07	-0.16	0.26	0.52
T-stat	0.85	0.83	1.82	1.14	2.68	-0.36	-0.48	1.73	3.42
$R^{2}(\%)$	0.53	0.62	0.80	0.70	1.21	0.46	0.57	0.82	1.92
Panel B: Predictors Controlled, 12 Month Horizon									
	bm	dspr	dp	ер	ltr	ntis	svar	tspr	corpr
$\log \zeta_t$	-2.15	-2.69	-1.65	-2.11	-2.07	-2.03	-2.02	-2.21	-2.08
T-stat	-1.74	-2.12	-1.34	-1.78	-1.76	-1.71	-1.69	-1.90	-1.78
predictor	2.01	1.96	3.49	2.82	1.63	-0.43	0.68	3.04	1.96
T-stat	1.40	1.48	2.59	1.77	3.44	-0.23	0.94	2.45	4.00
$R^{2}(\%)$	3.31	3.07	6.57	4.89	2.74	1.74	1.80	5.46	3.22
	Panel C: Predictors Controlled, 60 Month Horizon								
	bm	dspr	dp	ер	ltr	ntis	svar	tspr	corpr
$\log \zeta_t$	-10.78	-13.31	-8.27	-10.63	-10.85	-10.95	-10.75	-11.21	-10.88

-3.85

8.86

2.03

16.22

-3.44

1.84

1.56

9.89

-3.39

-1.56

-0.48

9.79

Stable coefficient controlling for other predictors.

-3.06

14.24

25.99

5.79

-3.80

7.11

2.47

13.56



-3.46

2.43

1.96

10.10

-3.62

10.52

19.23

3.56

-3.40

2.56

1.39

10.03

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- APT+granularity framework
- Treat granularity as a feature of DGP, study the asset pricing implication.
- Granularity implies
  - Failure of diversification
  - Large firms in cross-section
- Asset pricing implications:
  - Size-adjusted idiosyncratic risk explains the expected returns in cross-section.
  - Granularity explains a puzzling risk-return relation in the cross-section.
  - Level of granularity explains the time-variation of market expected returns.

#### Future Research

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- Future research question: reasons for an under-diversified stock market.
- An asset pricing model to generate endogenous granularity
- Endogenous granularity in firm size: Champernowne (1953), Wold and Whittle (1957), Gabaix (1999), Beare and Toda (2022)
- Combine with asset pricing features:
  - negative relation between size and variance
  - factor structure in covariance.
  - joint process of size and return

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# Summary of the 10 largest firms, 40s-70s

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	Panel A: Sur	mmary of the 10 lar	gest firms, 40s-70s	
	1940	1950	1960	1970
1	GENERAL MOTORS (0.05)	STANDARD OIL NJ(0.05)	IBM (0.05)	IBM (0.05)
2	STANDARD OIL NJ(0.04)	GENERAL MOTORS (0.05)	GENERAL MOTORS (0.04)	STANDARD OIL NJ(0.03)
3	DUPONT (0.04)	DUPONT (0.04)	STANDARD OIL NJ(0.04)	GENERAL MOTORS (0.02)
4	GENERAL ELECTRIC (0.03)	GENERAL ELEC- TRIC(0.03)	TEXACO INC(0.02)	EASTMAN KO- DAK(0.02)
5	TEXASCO(0.02)	TEXASCO(0.02)	GENERAL ELEC- TRIC(0.02)	GENERAL ELEC- TRIC(0.02)
6	STANDARD OIL IND(0.01)	STANDARD OIL CAL(0.02)	DUPONT (0.02)	TEXACO(0.01)
7	STANDARD OIL CAL(0.01)	GULF OIL (0.02)	EASTMAN KO- DAK(0.01)	PROCTER & GAM- BLE(0.01)
8	COCA COLA(0.01)	IBM (0.01)	GULF OIL (0.01)	MINNESOTA MINING & MFG(0.01)
9	GULF OIL (0.01)	SOCONY VACUUM OIL(0.01)	STANDARD OIL CAL(0.01)	DUPONT (0.01)
10	KENNECOTT COP- PER (0.01)	STANDARD OIL IND(0.01)	MINNESOTA MINING & MFG(0.01)	STANDARD OIL CO IND(0.01)
Total weight	0.24	0.26	0.24	0.19
Number of assets	1019	1215	2995	6718

# Summary of the 10 largest firms, 80s-2020

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Panel B:Summary of the 10 largest firms, 80s-2020							
	1980	1990	2000	2010	2020		
1	IBM(0.04)	GE(0.02)	XOM(0.03)	AAPL(0.03)	AAPL(0.05)		
2	XON(0.02)	XON(0.02)	GE(0.03)	GOOG(0.02)	MSFT(0.05		
3	GE(0.02)	KO(0.02)	MSFT(0.02)	MSFT(0.02)	AMZN(0.04		
4	SUO(0.01)	WMT(0.01)	WMT(0.02)	XOM(0.02)	GOOG(0.03		
5	SN(0.01)	IBM(0.01)	C(0.02)	BRK(0.02)	FB(0.02)		
6	GM(0.01)	MSFT(0.01)	PFE(0.02)	BRK(0.02)	BRK(0.02)		
7	MOB(0.01)	MRK(0.01)	JNJ(0.01)	AMZN(0.01)	JNJ(0.01)		
8	SD(0.01)	PG(0.01)	INTC(0.01)	JNJ(0.01)	WMT(0.01)		
9	BLS(0.01)	BMY(0.01)	CSCO(0.01)	WMT(0.01)	V(0.01)		
10	DD(0.01)	JNJ(0.01)	IBM(0.01)	JPM(0.01)	JPM(0.01)		
Summed weight	0.15	0.14	0.17	0.18	0.25		
Number of assets	10428	12477	9040	6060	3823		

- Existence of large firms over decades.
- In each decade, commonality of large firms
- Can be merged as one larger firm. Granularity is not overestimated.

# Granularity increases over time

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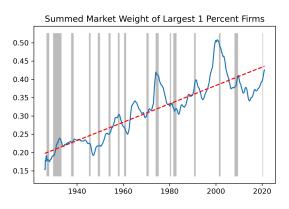
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- Fat tail persists over time.
- Measure the market weight of the largest 1% firms.
- Implication: The market portfolio is more "concentrated" over time → intro

# Granularity and other factor models

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 For example, CAPM states a relation between expected returns by assuming multiple conditions

$$E[r_i] - r_f = \beta_i E[r_m - r_f]$$

- CAPM+APT risk structure:  $w_i\theta_i$  in  $\beta_i$
- APT assumes  $f \perp \epsilon$ . Simplicity for empirical test.
- Econometric issue: bias in measuring factor exposure  $\beta_i$ . OLS:

$$r_i - rf = \beta_i (r_m - rf) + \varepsilon_i$$

 $r_m - r_f$  includes  $\sum w_i \varepsilon_i$ . Endogenous when  $w_i$  is non-negligible.

- More complicated to identify shocks propagation in an equilibrium. Loop feedback effect:  $\varepsilon_i \to r_m \to \varepsilon_i$
- Different instrumental variables to use depending on what equilibrium mechanism to identify.

# Bias v.s. Variance in the $\zeta$ estimation

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- A trade-off between bias and variance. More observations.
   Less variance. Yet, more concave deviations and more bias (more small firms included).
- Vast literature to correct the bias by assuming a distribution to adjust for the concavity.
- For example, Feuerverger and Hall (1999) assumes a survival function as:

$$P(X_i > x) = ax^{-\zeta}(1 + bx^{-\rho} + o(bx^{-\rho}))$$

- Derive the bias by the distribution assumed.
- Generally, the bias depends on number of observations n
  and level of the concavity.

### Time-variation of idiosyncratic risk

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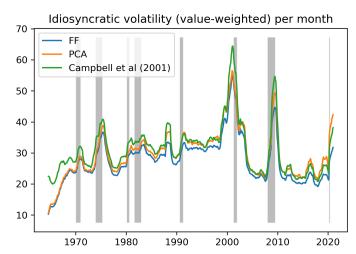
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# Definition of predictors controlled

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#### **Summary of Predictors** Description AR1 Corr with $\xi_t$ $\xi_t$ granularity measure 0.97 1.00 book to market ratio 0.07 bm 0.99default spread 0.27dspr 0 97 -0.11dp dividend price ratio 0.990.990.04 earning price ratio ep ltr long term government bond return 0.050.01 ntis net equity expansion ratio 0.980.08 stock variance 0.40-0.07svar tspr term spread 0.960.11corporate bond return 0.110.01 corpr