

Selected Topics of Mathematical Statistics: Quiz 4

Josephine Kraft

Ladislaus von Bortkiewicz Chair of Statistics
Humboldt-Universität zu Berlin
<http://lvb.wiwi.hu-berlin.de>



Quiz 4

**Show several characteristic functions under
different CDFs!**



Binomial Distribution

Let $X \sim \text{Bin}(n, p)$. Then:

$$\begin{aligned}\varphi_X(t) &= \mathbb{E}[e^{itX}] = \sum_{k=0}^n e^{itk} \cdot \mathbb{P}(X = k) \\&= \sum_{k=0}^n e^{itk} \cdot \binom{n}{k} \cdot p^k \cdot (1-p)^{n-k} \\&= \sum_{k=0}^n \binom{n}{k} \cdot (e^{it} \cdot p)^k \cdot (1-p)^{n-k} \\&= [e^{it} \cdot p + (1-p)]^n\end{aligned}$$

because $(a+b)^n = \sum_{k=0}^n \binom{n}{k} a^k b^{n-k}$ for $n \in \mathbb{N}$.



Poisson Distribution

Let $X \sim \text{Pois}(\lambda)$. Then:

$$\begin{aligned}\varphi_X(t) &= \mathbb{E}[e^{itX}] = \sum_{k=0}^{\infty} e^{itk} \cdot \mathbb{P}(X = k) \\ &= \sum_{k=0}^{\infty} e^{itk} \cdot e^{-\lambda} \cdot \frac{\lambda^k}{k!} \\ &= e^{-\lambda} \cdot \sum_{k=0}^{\infty} \frac{(\lambda \cdot e^{it})^k}{k!} \\ &= e^{-\lambda} \cdot e^{\lambda e^{it}} = e^{\lambda(e^{it}-1)}\end{aligned}$$

because $e^x = \sum_{k=0}^{\infty} \frac{x^k}{k!}$ for $x \in \mathbb{R}$.



Standard Normal Distribution

Let $X \sim N(0, 1)$. Then:

$$\begin{aligned}\varphi_X(t) &= \mathbb{E}[e^{itX}] = \int_{-\infty}^{\infty} e^{itx} \cdot f(x) \, dx \\&= \int_{-\infty}^{\infty} e^{itx} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \, dx \\&= \frac{1}{\sqrt{2\pi}} \left[\int_{-\infty}^0 e^{itx} \cdot e^{-\frac{x^2}{2}} \, dx + \int_0^{\infty} e^{itx} \cdot e^{-\frac{x^2}{2}} \, dx \right] \\&= \frac{1}{\sqrt{2\pi}} \left[\int_0^{\infty} e^{-itx} \cdot e^{-\frac{x^2}{2}} \, dx + \int_0^{\infty} e^{itx} \cdot e^{-\frac{x^2}{2}} \, dx \right]\end{aligned}$$



Standard Normal Distribution

$$\begin{aligned}\varphi_X(t) &= \frac{1}{\sqrt{2\pi}} \cdot \int_0^\infty e^{-\frac{x^2}{2}} \cdot e^{itx} + e^{-\frac{x^2}{2}} \cdot e^{-itx} \, dx \\&= \frac{1}{\sqrt{2\pi}} \cdot \int_0^\infty e^{-\frac{x^2}{2}} \cdot [e^{itx} + e^{-itx}] \, dx \\&= \frac{1}{\sqrt{2\pi}} \cdot \int_0^\infty e^{-\frac{x^2}{2}} \cdot 2 \cdot \cos(tx) \, dx \\&= \frac{2}{\sqrt{2\pi}} \cdot \int_0^\infty e^{-\frac{x^2}{2}} \cdot \cos(tx) \, dx \\&\stackrel{(1)}{=} e^{-\frac{t^2}{2}}\end{aligned}$$



Standard Normal Distribution

To prove (1) let

$$F(t) := \varphi_X(t) = \frac{2}{\sqrt{2\pi}} \cdot \int_0^{\infty} e^{-\frac{x^2}{2}} \cdot \cos(tx) \, dx$$

We show that

$$F'(t) = -t \cdot F(t) \quad (*)$$

This ordinary differential equation has the solution $F(t) = c \cdot e^{-\frac{t^2}{2}}$, where c is a constant. Since we have

$$F(0) = \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} \cdot e^{-\frac{x^2}{2}} \, dx = 1$$

we get $c = 1$. Hence, $\varphi_X(t) = F(t) = e^{-\frac{t^2}{2}}$.



Standard Normal Distribution

$$\begin{aligned}F'(t) &= \frac{d}{dt} \frac{2}{\sqrt{2\pi}} \cdot \int_0^{\infty} e^{-\frac{x^2}{2}} \cdot \cos(tx) dx \\&= \frac{2}{\sqrt{2\pi}} \cdot \int_0^{\infty} e^{-\frac{x^2}{2}} \cdot \frac{d}{dt} \cos(tx) dx \\&= \frac{2}{\sqrt{2\pi}} \cdot \int_0^{\infty} -e^{-\frac{x^2}{2}} \cdot x \cdot \sin(tx) dx\end{aligned}$$

Integration by parts:

$$\begin{aligned}u(x) &= e^{-\frac{x^2}{2}}, \quad u'(x) = -x \cdot e^{-\frac{x^2}{2}} \quad \text{and} \\v(x) &= \sin(tx), \quad v'(x) = t \cos(tx)\end{aligned}$$



Standard Normal Distribution

$$\begin{aligned} F'(t) &= \frac{2}{\sqrt{2\pi}} \cdot [(e^{-\frac{x^2}{2}} \cdot \sin(tx))|_0^\infty - \int_0^\infty e^{-\frac{x^2}{2}} \cdot t \cdot \cos(tx) \, dx] \\ &= -\frac{2}{\sqrt{2\pi}} \cdot \int_0^\infty e^{-\frac{x^2}{2}} \cdot t \cdot \cos(tx) \, dx \\ &= -t \cdot \int_0^\infty \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \cdot 2 \cdot \cos(tx) \, dx \\ &= -t \cdot F(t) \end{aligned}$$

Hence, (*) was proved and we get $\varphi_X(t) = e^{-\frac{t^2}{2}}$.

