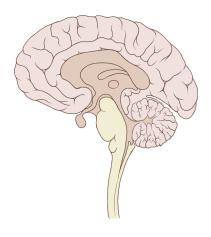
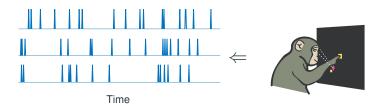
Statistical Machine Learning Methods for High-dimensional Neural Population Data Analysis

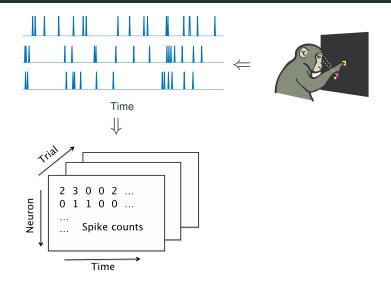
Yuanjun Gao

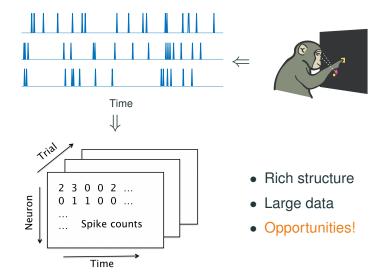
Department of Statistics Columbia University











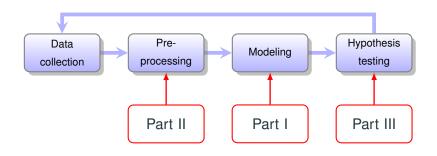


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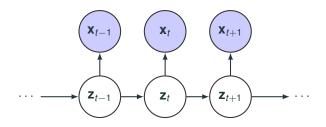
- I. Neural Population Data Analysis with Latent Variable Models
 - Generalized count linear dynamical system
 - Linear dynamical neural population models through nonlinear embeddings
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I. Neural Population Data Analysis with Latent Variable Models

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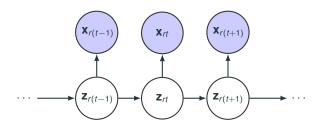
State space models



- $\mathbf{x}_t \in \mathbb{N}^n$: spike counts; $\mathbf{z}_t \in \mathbb{R}^m$: latent variables
- Joint distribution

$$p(\mathbf{x}, \mathbf{z}) = \underbrace{p(\mathbf{z}_1)}_{\text{Initial distribution}} \underbrace{\prod_{t=1}^{T-1} p(\mathbf{z}_{t+1} | \mathbf{z}_t)}_{\text{Transition model}} \underbrace{\prod_{t=1}^{T} p(\mathbf{x}_t | \mathbf{z}_t)}_{\text{Observation model}}$$

State space models: multiple trials



- r = 1, ..., R: trial number
- $\mathbf{x}_{rt} \in \mathbb{N}^n$: spike counts; $\mathbf{z}_{rt} \in \mathbb{R}^m$: latent variables
- Joint distribution

$$p(\mathbf{x}, \mathbf{z}) = \prod_{r=1}^{R} \left[\underbrace{p(\mathbf{z}_{r1})}_{\text{Initial distribution}} \underbrace{\prod_{t=1}^{T-1} p(\mathbf{z}_{r(t+1)} | \mathbf{z}_{rt})}_{\text{Transition model}} \underbrace{\prod_{t=1}^{T} p(\mathbf{x}_{rt} | \mathbf{z}_{rt})}_{\text{Observation model}} \right]$$

Common parameterization and our extensions

 Common assumptions for latent dynamics: Gaussian linear dynamical system (LDS)

$$\mathbf{z}_{r1} \sim \mathcal{N}(\mu_1, Q_1)$$
 $\mathbf{z}_{r(t+1)} | \mathbf{z}_{rt} \sim \mathcal{N}(A\mathbf{z}_{rt}, Q)$

Common observation models:

$$\mathbf{x}_{rt} | \mathbf{z}_{rt} \sim \underbrace{\mathcal{N}(C\mathbf{z}_{rt} + d, \Sigma)}_{ ext{model mismatch}} ext{ or } \underbrace{ ext{Poisson}\left(\exp(C\mathbf{z}_{rt} + d)\right)}_{ ext{equal dispersion}}$$

- Our extensions for observation model:
 - Generalized count distribution (GCLDS) (Gao et al. 2015)
 - Flexible nonlinear observation (fLDS) (Gao et al. 2016)

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Motivation

• Doubly stochastic Poisson model implies overdispersion

$$\left. egin{array}{ll} \mathbf{z} & \sim p(\mathbf{z}) \\ \mathbf{x} | \mathbf{z} & \sim \mathsf{Poisson}(f(\mathbf{z})) \end{array}
ight.
ight.$$

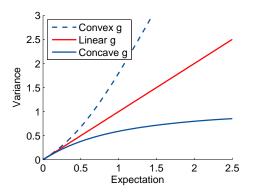
Generalized count distribution family

Generalized count (GC) distribution family

where $\theta \in \mathbb{R}$, $g(\cdot) : \mathbb{N} \to \mathbb{R}$.

- Parameterizes all the count distributions redundantly.
- \bullet θ acts as "rate" parameter.
- $g(\cdot)$ acts as "shape" parameter.

Generalized count distribution family



Probability mass function for GC distribution family

$$p_{\mathcal{GC}}(x;\theta,g(\cdot)) \propto \frac{\exp(\theta \cdot x + g(x))}{x!}, \quad x \in \mathbb{N}$$

Model formulation

 Linear dynamical systems with generalized count observation

$$egin{aligned} \mathbf{z}_{r1} &\sim \mathcal{N}(\mu_1, Q_1) \ \mathbf{z}_{r(t+1)} | \mathbf{z}_{rt} &\sim \mathcal{N}(A\mathbf{z}_{rt}, Q) \ X_{rti} | \mathbf{z}_{rt} &\sim \mathcal{GC}(c_i^T \mathbf{z}_{rt}, g_i(\cdot)) \end{aligned}$$

- Practical considerations
 - Set $g_i(k) = -\infty$ for k > K to facilitate computation.
 - Ridge penalty on the 2nd difference of g_i(·) to avoid overfitting;

penalty =
$$\lambda \sum_{k=1}^{K-1} (g_i(k-1) - 2g_i(k) + g_i(k+1))^2$$
.

Variational Bayes Expectation Maximization (VBEM)

- \mathbf{x} : data, \mathbf{z} : latent variables, θ : model parameters,
- Often hard to compute $p_{\theta}(\mathbf{x}) = \int p_{\theta}(\mathbf{x}, \mathbf{z}) d\mathbf{z}$ and $p_{\theta}(\mathbf{z}|\mathbf{x})$.
- Approximate the posterior by a tractable distribution family.

$$p_{\theta}(\mathbf{z}|\mathbf{x}) pprox q(\mathbf{z}) \in \mathcal{Q}$$

Optimize a lower bound of log likelihood, or ELBO

$$\mathsf{ELBO}(heta,q) = \int \left[\log rac{
ho_{ heta}(\mathbf{x},\mathbf{z})}{q(\mathbf{z})}
ight] q(\mathbf{z}) d\mathbf{z} \leq \log
ho_{ heta}(\mathbf{x})$$

Variational Bayes Expectation Maximization (VBEM)

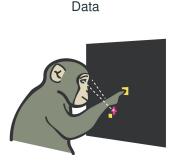
- VBEM: Optimize ELBO(θ, q) $\leq \log p_{\theta}(\mathbf{x})$ iteratively
 - E-step: For a fixed θ , optimize q
 - M-step: For a fixed q, optimize θ
- VBEM for GCLDS
 - We set q to be multivariate Gaussian
 - We derive a looser but tractable ELBO
 - E-step: fast Laplace approximation initialization + dual optimization (Emtiyaz Khan et al. 2013)
 - M-step: convex optimization + analytical solution

Experiments

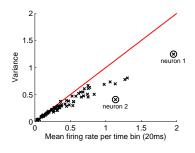
 For both simulated and real dataset, we compare GCLDS with PLDS (Poisson observation model)

	Mean	Variance	Likelihood
PLDS	✓	Х	Х
GCLDS	✓	\checkmark	✓

Real data analysis: data



Variance and mean of spike counts

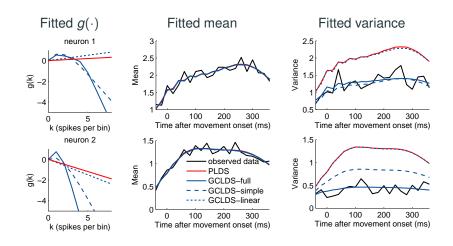


- Center-out reaching experiments
- Multi-electrode array recording
- Strong under-dispersion

Real data analysis: algorithms

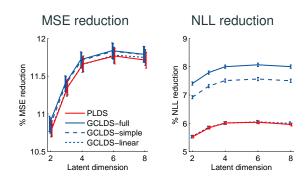
- Main algorithms to be compared
 - PLDS: Poisson observation (Macke, Buesing, and Sahani 2015)
 - GCLDS-full: Generalized count observation, individual $g(\cdot)$ across neurons
- Two control cases for GCLDS
 - GCLDS-linear: truncated linear $g(\cdot)$ (truncated Poisson)
 - GCLDS-simple: $g(\cdot)$ shared across neurons (up to a linear function)

Real data analysis: single neuron fit



Real data analysis: population fit

• Leave-one-neuron-out prediction



Conclusion and discussion

- Summary
 - Incorporated generalized count family into state space models.
 - Developed VBEM algorithm.
 - Observed superior fitted results on real neural data.
- Gao Y, Buesing L, Shenoy KV, Cunningham JP (2015)
 High-dimensional neural spike train analysis with generalized count linear dynamical systems. NIPS 2015.

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Motivation

- Neural activities lie in a low-dimensional nonlinear manifold rather than a linear subspace
- Flexible observation model makes the state space model more expressive

Model formulation: fLDS

Linear dynamical systems with nonlinear link and count observation

$$egin{aligned} \mathbf{z}_{r1} \sim & \mathcal{N}(\mu_1, Q_1) \ \mathbf{z}_{r(t+1)} | \mathbf{z}_{rt} \sim & \mathcal{N}(A\mathbf{z}_{rt}, Q) \ & x_{rti} \sim & \operatorname{Poisson}(\mathbf{f}_i(\mathbf{z}_{rt})) \text{ (PfLDS)} \ & \operatorname{or } \mathcal{GC}(\mathbf{f}_i(\mathbf{z}_{rt}), g_i(\cdot)) \text{ (GCfLDS)} \end{aligned}$$

where f_i is a nonlinear function parameterized by a neural network.

Inference algorithm: AEVB (high level idea)

- Auto-encoding Variational Bayes (AEVB)
- Learn a mapping (recognition model) from data to the approximate posterior distribution of latent variable.
- Jointly optimize the generative model parameters and recognition model parameters.
- Naturally incorporate stochastic optimization to handle large datasets.
- Tractable for a large class of graphical models

Inference algorithm: AEVB (algorithm)

Decompose ELBO by trials

$$\mathsf{ELBO}(\theta, q) = \sum_{r=1}^{R} \int \left[\log \frac{p_{\theta}(\mathbf{x}_r, \mathbf{z}_r)}{q(\mathbf{z}_r)} \right] q(\mathbf{z}_r) d\mathbf{z}_r$$

• Map data \mathbf{x}_r to $q(\mathbf{z}_r)$ by a parameterized function

$$q(\mathbf{z}_r) = q_{\phi}(\mathbf{z}_r|\mathbf{x}_r) = \mathcal{N}\left(\mu_{\phi}(\mathbf{x}_r), \Sigma_{\phi}(\mathbf{x}_r)\right)$$

• Learn both θ and ϕ by stochastic gradient descent on ELBO

$$abla \mathsf{ELBO}(heta, q_{\phi}) pprox R imes
abla \int \left[\log rac{p_{ heta}(\mathbf{x}_r, \mathbf{z}_r)}{q_{\phi}(\mathbf{z}_r | \mathbf{x}_r)}
ight] q_{\phi}(\mathbf{z}_r | \mathbf{x}_r) d\mathbf{z}_r$$
 $pprox R imes ext{an unbiased estimator of gradient}$

Structure of the recognition model

Generative model:

$$p_{\theta}(\mathbf{z}_r|\mathbf{x}_r) \propto \underbrace{p_{\theta}(\mathbf{z}_{r1}) \prod_{t=1}^{T-1} p_{\theta}(\mathbf{z}_{t(t+1)}|\mathbf{z}_{rt})}_{Gaussian} \underbrace{\prod_{t=1}^{T} \underbrace{p_{\theta}(\mathbf{x}_{rt}|\mathbf{z}_{rt})}_{Complicated}$$

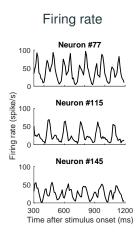
• Recognition model, product-of-Gaussian form:

$$q_{\phi}(\mathbf{z}_r|\mathbf{x}_r) \propto q_{\phi}(\mathbf{z}_{r1}) \prod_{t=1}^{T-1} q_{\phi}(\mathbf{z}_{r(t+1)}|\mathbf{z}_{rt}) \prod_{t=1}^{T} \underbrace{q_{\phi}(\mathbf{z}_{rt}|\mathbf{x}_{rt})}_{Gaussian}$$

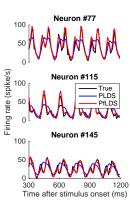
 Multivariate Gaussian distribution with block tri-diagonal precision matrix.

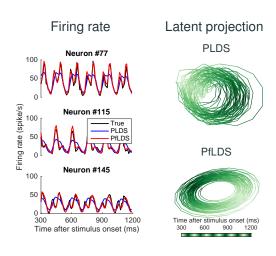
Experiments

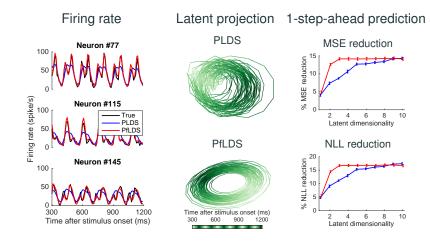
	Mean	Variance	Likelihood	Concise representation
PLDS	✓	Х	Х	Х
GCLDS	✓	✓	✓	X
PfLDS	✓	X	X	\checkmark
GCfLDS	✓	✓	✓	✓



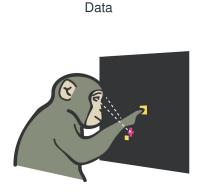








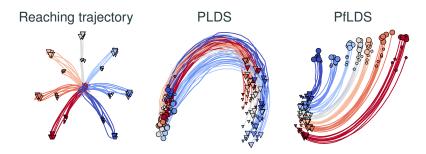
Real data analysis: Primate motor cortex





Real data analysis: Primate motor cortex

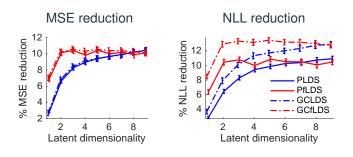
Latent projection with 2 latent dimensions



- Recovering the behavior structure from neural data in a unsupervised fashion.
- Concise and informative latent dimensions.

Real data analysis: Primate motor cortex

One-step-ahead predictive performance



Conclusion and discussion

- Summary
 - Incorporated nonlinear observation into state space models.
 - Developed AEVB algorithm (flexible and scalable).
 - Obtained concise latent representations.
- Gao Y*, Archer E*, Paninski L, Cunningham JP (2016)
 Linear dynamical neural population models through
 nonlinear embeddings. NIPS 2016. (* = equal contribution)

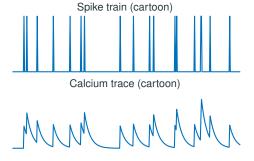
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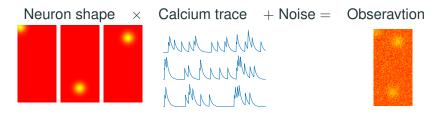
Introduction: calcium imaging data

 Basic principle: the spiking activity of a neuron induces a transient increase in calcium concentration, which can be indirectly observed by recording the fluorescent properties of certain calcium indicators.



Introduction: calcium imaging data

 Calcium imaging enables simultaneous recording of many neurons.



 Goal: recover the neuron shape and calcium trace given the observation.

Model formulation

- $X \in \mathbb{R}^{N \times T}$ represents the calcium imaging data, where each column is a (vectorized) frame that contains N pixels
- Decompose X into a product of K spatial component and temporal component

$$X = DA^T +$$
noise

- $D = [D_1, ..., D_K] \in \mathbb{R}^{N \times K}$ represents the neuron shapes
- $A = [A_1, ..., A_K] \in \mathbb{R}^{T \times K}$ is the neural activities
- Further exploit structure of the components (localized neuron shapes)

Model formulation: objective

Structured matrix factorization

minimize
$$\|X - DA^T\|_2^2 + f_D(D)$$
, subject to $D_k \in \mathcal{D}_w^+; k = 1, \dots, K$, $\|A_k\|_2 \le c_k$,

- \mathcal{D}_w^+ : non-negative vectors whose nonzero values is within a $w \times w$ window
- $f_D(D)$ regularizes the neuron shape (will discuss later)
- $||A_k||_2 \le c_k$ avoids degenerate solution

Greedy algorithm

- Identify ROIs one at a time, using the residual un-explained by existing signals.
- At iteration k, given the current residue (unexplained by existing ROIs):
 - Greedy identification: Identify the location p_k where the Gaussian kernel explains most of the data (across time).
 - Shape fine tuning: Locally optimize the spatial and temporal component.
 - Residue update: Subtract the newly identified signal.

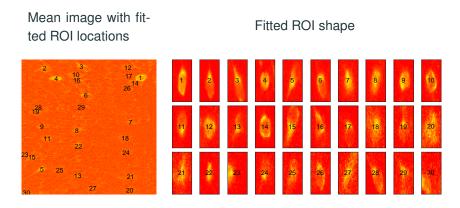
Shape fine tuning

 Given current residue R, an identified center pixel p_k, denote S_k as a w × w window centered at p_k

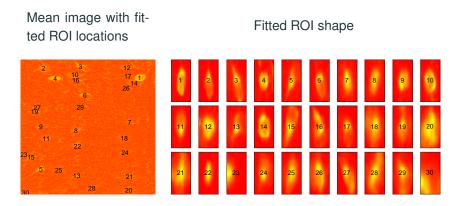
$$\begin{aligned} & \underset{D_k,A_k}{\text{minimize}} & & \|R - D_k A_k^T\|^2 + f(D_k), \\ & \text{subject to} & & D_{kp} \geq 0, p \in \mathcal{S}_k, \\ & & & D_{kp} = 0, p \notin \mathcal{S}_k, \\ & & & & \|A_k\|_2 \leq c_k, \end{aligned}$$

- $f(D_k) = \sum_{i=1}^3 \lambda_i f_i(D_k)$
 - $f_1(D_k) = \sum_{p} \tau_{(p,p_k)} |D_{kp}|$ encourages sparsity
 - $f_2(D_k) = \sum_{p} (D_{kp} G_{p_k})^2$ encourages Gaussian shape
 - $f_3(D_k) = \sum_{p_1 \text{ and } p_2 \text{ are neighbors}} (D_{kp_1} D_{kp_2})^2$ encourages smoothness
- Optimize by block coordinate descent

Real data analysis: sample patch, no shape regularization



Real data analysis: sample patch, with shape regularization



Conclusion and discussion

Summary

- Formulated calcium imaging ROI detection as a structured matrix factorization problem.
- Developed a fast greedy algorithm.
- Pnevmatikakis EA, Soudry D, Gao Y, Machado TA, Merel J, Pfau D, Reardon T, Mu Y, Lacefield C, Yang W, Ahrens M, Bruno R, Jessell TM, Peterka DS, Yuste R, Paninski L (2016) Simultaneous denoising, deconvolution, and demixing of calcium imaging data. Neuron, 89(2), 285-299.

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Maximum entropy principle

• Entropy: for a continuous distribution with density $p(\mathbf{z})$ where $\mathbf{z} \in \mathbb{R}^d$, the entropy is defined as

$$H(p) = -\int p(\mathbf{z}) \log p(\mathbf{z}) d\mathbf{z} = E_{\mathbf{Z} \sim p} \left[-\log p(\mathbf{Z}) \right].$$

- A popular measure of diversity or information content.
- Maximum entropy (ME) principle: Subject to some given prior knowledge (usually in the form of expectation or support constraints), the distribution that makes minimal additional assumptions is that which has the largest entropy of any distribution obeying those constraints

Applications of maximum entropy

- Neuroscience: generate a distribution of neural activity by specifying a set of features (pairwise correlation etc.) for hypothesis testing.
- Texture modeling: generate an image with a certain texture by specifying expected value of features relevant to texture.
- Ecology, natural language processing, finance...

Maximum entropy problem

• Maximum entropy (ME) problem

$$p^* = \max_{\rho} H(\rho)$$
 subject to $E_{\mathbf{Z} \sim \rho}[T(\mathbf{Z})] = 0$ $\sup_{\mathbf{Z} \sim \rho} [p] = \mathcal{Z},$

where $T(\mathbf{z}) = (T_1(\mathbf{z}), ..., T_m(\mathbf{z})) : \mathcal{Z} \to \mathbb{R}^m$ is the vector of known statistics, and \mathcal{Z} is the given support.

Gibbs distribution

 Lagrange multipliers argument ⇒ an exponential family form for ME distribution (Gibbs distribution):

$$p^*(\mathbf{z}) \propto e^{<\eta, T(\mathbf{z})>} \mathbb{1}(\mathbf{z} \in \mathcal{Z})$$

For some $\eta \in \mathbb{R}^m$ such that $E_{\mathbf{Z} \sim p^*}[T(\mathbf{Z})] = 0$.

- Hard to Identify η .
- Hard to sample from the distribution.
- Question: is there a better way to do this?

Maximum entropy flow network (MEFN)

• Identify a transformation $f_{\phi^*} \in \mathcal{F} = \{f_{\phi} : \mathbb{R}^d \to \mathbb{R}^d, \phi \in \mathbb{R}^q\}$ that transforms a simple distribution $\mathbf{Z}_0 \sim p_0$ (Gaussian) to $\mathbf{Z} = f_{\phi^*}(\mathbf{Z}_0) \sim p_{\phi^*}$ that approximates the ME distribution.

$$p^* = \max_{p} H(p)$$
subject to $E_{\mathbf{Z} \sim p}[T(\mathbf{Z})] = 0$
 $\sup_{\phi}(p) = \mathcal{Z},$
 ψ
 $\phi^* = \max_{\phi} H(p_{\phi})$
subject to $R(\phi) = E_{\mathbf{Z}_0 \sim p_0}[T(f_{\phi}(\mathbf{Z}_0))] = 0$
 $\sup_{\phi}(p_{\phi}) = \mathcal{Z}.$

Normalizing flow

- Normalizing flow: A smooth and invertible transformation.
- Explicit probability density form for $Z = f_{\phi}(Z_0) \sim p_{\phi}$ by change-of-variable theorem

$$p_{\phi}(\mathbf{z}) = p_0\left(f_{\phi}^{-1}(\mathbf{z})\right) |\det\left(J_{\phi}(\mathbf{z})\right)|^{-1}.$$

 Constructing a flexible family of normalizing flows by composing simple normalizing flows (deep learning)

$$\mathbf{Z}=f_k\circ f_{k-1}\circ\cdots\circ f_1(\mathbf{Z}_0).$$

Augmented Lagrangian method

Augmented Lagrangian method minimizes the objective

$$L(\phi; \lambda, c) = -H(p_{\phi}) + \lambda^{\top} R(\phi) + \frac{c}{2} ||R(\phi)||^2.$$

for a sequence of $\lambda \in \mathbb{R}^m$ and $c \geq 0$.

- Update rule: at iteration k, given λ_k and c_k .
 - Find ϕ_k that optimizes $L(\phi; \lambda_k, c_k)$.
 - Update λ and c by

$$\begin{split} \lambda_{k+1} = & \lambda_k + c_k R(\phi_k) \\ c_{k+1} = \begin{cases} \beta c_k & ||R(\phi_k)|| > \gamma ||R(\phi_{k-1})|| \\ c_k & \text{otherwise} \end{cases} \end{split}$$

for some
$$\gamma \in (0, 1)$$
, $\beta > 1$

Augmented Lagrangian method in stochastic setting

• Both $H(p_{\phi})$ and $R(\phi)$ are intractable, but we can approximate with an unbiased estimation,

$$H(p_{\phi}) pprox rac{1}{n} \sum_{i=1}^{n} -\log p_{\phi}(\mathbf{z}^{(i)}),$$
 $R(\phi) pprox rac{1}{n} \sum_{i=1}^{n} T(f_{\phi}(\mathbf{z}^{(i)})),$

where $\mathbf{z}^{(i)} \sim p_0$. We can then optimize the objective by stochastic gradient descent.

Texture modeling (overview)

- Goal: construct a distribution on images that mimics a given texture (from a training image).
 - Authenticity: the samples from the distribution should mimic the given texture (constraints).
 - Sample diversity: the distribution should generate a variety of images with certain texture (entropy).
- Natural application of maximum entropy modeling.

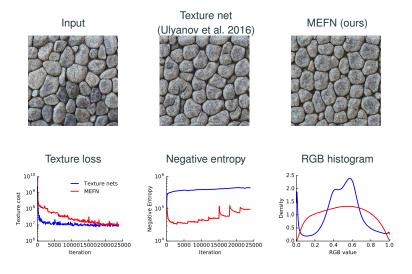
Texture modeling (formulation)

- Sample from the space of images $\mathbf{z} \in [0, 1]^{d=w \times h \times c}$ where $w \times h$ is the image size, c is the number of channels (c = 3 for RGB representation).
- Texture loss: $T(\mathbf{z}): [0,1]^{w \times h \times c} \to \mathbb{R}$, a complicated form proposed in Ulyanov et al. (2016).
- Ulyanov et al. (2016) proposes texture net, which solves the problem without considering entropy term.

$$\min E_{\mathbf{Z} \sim p_0} \left[T(f_{\phi}(\mathbf{Z})) \right]$$

• We build MEFN using $T(\mathbf{z})$ as the expectation constraint.

Experiment: Texture modeling (result)



Experiment: Texture modeling (diversity measure)

Method	d_{L^2}	SST	SSW	SSB
Texture net	0.077	0.043	0.036	0.006
MEFN	0.113	0.058	0.054	0.005

- For n = 20 randomly sampled image $\mathbf{z}^{(1)}, ..., \mathbf{z}^{(20)}$
- $d_{L^2} = \text{mean}_{i \neq j} \frac{1}{d} \|\mathbf{z}^{(i)} \mathbf{z}^{(j)}\|_2^2$: Mean Euclidean distance.
- ANOVA: SST = SSW + SSB, $\bar{\mathbf{z}} = \frac{1}{n} \sum_{i} \mathbf{z}^{(i)}$.
 - SST = $\frac{1}{nd}\sum_{i,k}(z_k^{(i)}-\bar{z})^2$: Total var.
 - SSW = $\frac{1}{nd} \sum_{i,k} (z_k^{(i)} \bar{z}_k)^2$: residue var, larger \Rightarrow better.
 - SSB = $\frac{1}{n}\sum_{k}(\bar{z}_{k}-\bar{z})^{2}$: mean image var, smaller \Rightarrow better.

Conclusion and discussion

- Summary
 - Solved maximum entropy problem by optimizing a normalizing flow.
 - Combined augmented Lagrangian optimization with stochastic optimization.
 - Obtained promising texture modeling result.
- Loaiza G*, Gao Y*, Cunningham JP (2017) Maximum entropy flow networks. ICLR 2017. (*=equal contribution)

Acknowledgement

Thank you!

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End of the talk. Questions?

Backup slides

GCLDS: Supervised case, GCGLM

• Given count data $x_i \in \mathbb{N}$, and associated covariates $z_i \in \mathbb{R}^p$, build Generalized Count Generalized Linear Model (GCGLM).

$$x_i \sim \mathcal{GC}(\theta(z_i), g(\cdot)), \quad \text{where} \quad \theta(z_i) = z_i \beta.$$

Reminder: generalized count distribution

$$p_{\mathcal{GC}}(x;\theta,g(\cdot)) \propto \frac{\exp(\theta \cdot x + g(x))}{x!}, \quad x \in \mathbb{N}$$

GCLDS: special cases of GCGLM

Model Name	Typical Parameterization	GCGLM Parametrization
Logistic regres- sion	$P(x = k) = \frac{\exp(k(\alpha + z\beta))}{1 + \exp(\alpha + x\beta)}$ $P(x = k) = \frac{\lambda^k}{k!} \exp(-\lambda);$	$g(k) = \alpha k; k = 0, 1$
Poisson regression	$P(x = k) = \frac{\lambda^k}{k!} \exp(-\lambda);$ $\lambda = \exp(\alpha + z\beta)$	$g(k) = \alpha k$
Adjacent cate- gory regression	$\frac{P(x=k+1)}{P(x=k)} = \exp(\alpha_k + z\beta)$	$g(k) = \sum_{i=1}^{k} (\alpha_{i-1} + \log i);$ k = 0, 1,, K
Negative bino- mial regression	$P(x = k) = \frac{(k+r-1)!}{k!(r-1)!} (1-\rho)^r \rho^k$ $\rho = \exp(\alpha + z\beta)$	$g(k) = \alpha k + \log(k + r - 1)!$
COM-Poisson regression	$P(x = k) = \frac{\lambda^k}{(k!)^{\nu}} / \sum_{j=1}^{+\infty} \frac{\lambda^j}{(j!)^{\nu}}$ $\lambda = \exp(\alpha + z\beta)$	$g(k) = \alpha k + (1 - \nu) \log k!$

GCLDS: VBEM detail

MEFN: Normalizing flow structures

- Rezende and Mohamed (2015) Proposes two specific families of transformations for variational inference
- Planar flow

$$f_i(\mathbf{z}) = \mathbf{z} + \mathbf{u}_i h(\mathbf{w}_i^T \mathbf{z} + b_i),$$

where $b_i \in \mathbb{R}$, \mathbf{u}_i , $\mathbf{w}_i \in \mathbb{R}^d$ and h is an activation function.

Circular flow

$$f_i(\mathbf{z}) = \mathbf{z} + \beta_i h(\alpha_i, r_i)(\mathbf{z} - \mathbf{z}_i'),$$

$$eta_i \in \mathbb{R}$$
, $lpha_i > 0$, $\mathbf{z}_i' \in \mathbb{R}^d$, $h(lpha, r) = 1/(lpha + r)$ and $r_i = ||\mathbf{z} - \mathbf{z}_i'||$.

MEFN: real-nvp normalizing flow structure

- Used in Dinh, Sohl-Dickstein, and Bengio (2016) for image denstiy estimation.
- Affine coupling layer: split variable $\mathbf{z} \in \mathbb{R}^D$ into $\mathbf{z}_1 \in \mathbb{R}^d$ and $\mathbf{z}_2 \in \mathbb{R}^{D-d}$. linearly transform \mathbf{z}_2 given \mathbf{z}_1 .

$$f\left(egin{array}{c} \mathbf{z}_1 \\ \mathbf{z}_2 \end{array}
ight) = \left(egin{array}{c} \mathbf{z}_1 \\ \mathbf{z}_2 \odot \exp\left(s(\mathbf{z}_1)\right) + t(\mathbf{z}_1) \end{array}
ight)$$

where \odot is element-wise product.

$$s(\mathbf{z}_1), t(\mathbf{z}_1) : \mathbb{R}^d \to \mathbb{R}^{D-d}.$$

- The transformation family is flexible because
 - s and t can be complex while maintaining tractability (inversion and Jacobian computation).
 - Partitioning z into (z₁, z₂) can be chosen arbitrarily.

Experiment: Dirichlet

- Dirichlet distribution is the ME distribution on a simplex $S = \{\mathbf{z} = (z_1, \dots, z_{d-1}) : z_i \ge 0 \text{ and } \sum_{k=1}^{d-1} z_k \le 1\}$ with expetation on the log of each coordinate $E[\log Z_k] = \kappa_k (k = 1, \dots, d)$, where $Z_d = 1 \sum_{k=1}^{d-1} Z_k$.
- 10 layers of planar flow (Rezende and Mohamed 2015).



Initial distribution p_0



MEFN result p_{ϕ^*} (Control case: moment matching)



Ground truth p*