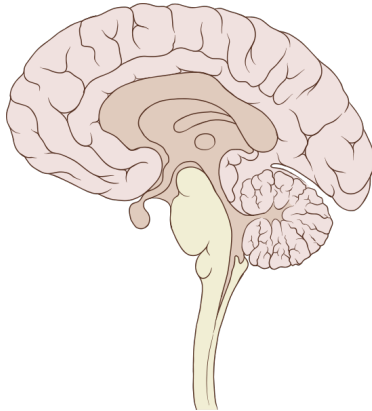**Statistical Machine Learning Methods for**

High-dimensional Neural Population Data Analysis

Yuanjun Gao

Department of Statistics
Columbia University
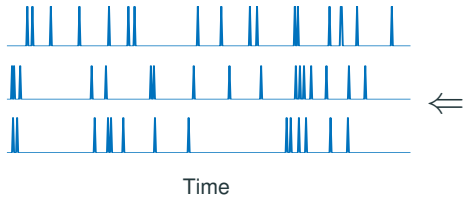
Time

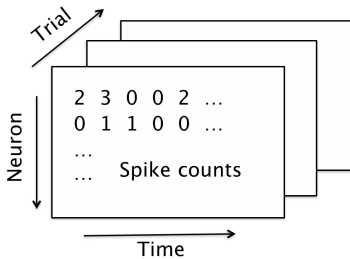Time

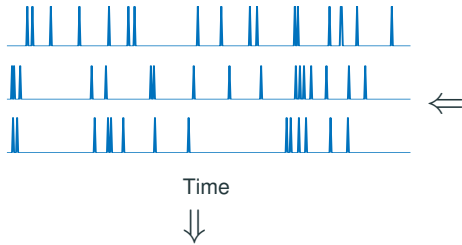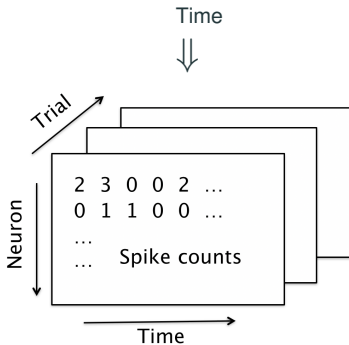Time

Trial

Neuron

2 3 0 0 2 ...
0 1 1 0 0 ...
...
...   Spike counts

Time

Time

Trial

Neuron

2 3 0 0 2 ...
0 1 1 0 0 ...
...
... Spike counts

Time

- Rich structure
- Large data
- Opportunities!

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State space models

- $\mathbf{x}_t \in \mathbb{N}^n$: spike counts; $\mathbf{z}_t \in \mathbb{R}^m$: latent variables
- Joint distribution

$$p(\mathbf{x}, \mathbf{z}) = \underbrace{p(\mathbf{z}_1)}_{\text{Initial distribution}} \underbrace{\prod_{t=1}^{T-1} p(\mathbf{z}_{t+1}|\mathbf{z}_t)}_{\text{Transition model}} \underbrace{\prod_{t=1}^{T} p(\mathbf{x}_t|\mathbf{z}_t)}_{\text{Observation model}}$$

State space models: multiple trials
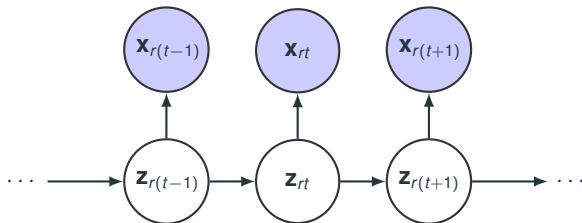
- $r = 1, ..., R$: trial number
- $\mathbf{x}_{rt} \in \mathbb{N}^n$: spike counts; $\mathbf{z}_{rt} \in \mathbb{R}^m$: latent variables
- Joint distribution

$$
p(\mathbf{x}, \mathbf{z}) = \prod_{r=1}^{R} \left[\underbrace{p(\mathbf{z}_{r1})}_{\text{Initial distribution}} \underbrace{\prod_{t=1}^{T-1} p(\mathbf{z}_{r(t+1)}|\mathbf{z}_{rt})}_{\text{Transition model}} \underbrace{\prod_{t=1}^{T} p(\mathbf{x}_{rt}|\mathbf{z}_{rt})}_{\text{Observation model}} \right]
$$

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Common parameterization and our extensions

- Common assumptions for latent dynamics: linear Gaussian dynamical system (LDS)

$$\mathbf{z}_1 \sim \mathcal{N}(\mu_1, Q_1)$$
$$\mathbf{z}_{t+1}|\mathbf{z}_t \sim \mathcal{N}(A\mathbf{z}_t, Q)$$

- Common observation models:

$$\mathbf{x}_t|\mathbf{z}_t \sim \underbrace{\mathcal{N}(C\mathbf{z}_t + d, \Sigma)}_{\text{model mismatch}} \text{ or } \underbrace{\text{Poisson}\left(\exp(C\mathbf{z}_t + d)\right)}_{\text{equal dispersion}}$$

$$\underbrace{\phantom{\mathcal{N}(C\mathbf{z}_t + d, \Sigma) \text{ or } \text{Poisson}\left(\exp(C\mathbf{z}_t + d)\right)}}_{\text{stringent assumptions}}$$

- Our extensions for observation model:
  - Generalized count distribution (GCLDS) (Gao et al. 2015)
  - Flexible nonlinear observation (fLDS) (Gao et al. 2016)

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- Doubly stochastic Poisson model implies overdispersion

$$\left.\begin{array}{ll} \mathbf{z} & \sim p(\mathbf{z}) \\ \mathbf{x} & \sim \text{Poisson}(f(\mathbf{z})) \end{array}\right\} \Rightarrow \begin{array}{ll} \text{var}(\mathbf{x}|\mathbf{z}) & = E(\mathbf{x}|\mathbf{z}) \\ \text{var}(\mathbf{x}) & \geq E(\mathbf{x}) \end{array}$$

- Need a more flexible distribution to separate firing rate variability with noise variability.

$$\text{var}(\mathbf{x}) = \underbrace{\text{var}\left(E(\mathbf{x}|\mathbf{z})\right)}_{\text{firing rate variability}} + \underbrace{E\left(\text{var}(\mathbf{x}|\mathbf{z})\right)}_{\text{noise variability}}$$

Generalized count distribution family

- Generalized count (GC) distribution family

$$p_{\text{Poisson}}(x; \lambda) \propto \frac{\exp\{\log \lambda \cdot x\}}{x!}, \quad x \in \mathbb{N}$$

$$\Downarrow$$

$$p_{\mathcal{GC}}(x; \theta, g(\cdot)) \propto \frac{\exp(\theta \cdot x + g(x))}{x!}, \quad x \in \mathbb{N}$$

where $\theta \in \mathbb{R}$, $g(\cdot) : \mathbb{N} \to \mathbb{R}$.

- Parameterizes all the count distributions redundantly.
- Given $g(\cdot)$, θ controls the expectation.
- $g(\cdot)$ controls the "shape" of the distribution.
  Convex/concave/linear $g(\cdot)$ implies
  overdispersed/underdispered/Poisson distribution.

Generalized count distribution family
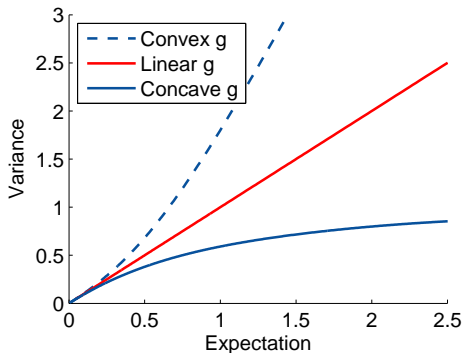
- Probability mass function for GC distribution family

$$p_{\mathcal{GC}}(x; \theta, g(\cdot)) \propto \frac{\exp(\theta \cdot x + g(x))}{x!}, \quad x \in \mathbb{N}$$

Model formulation

- Linear dynamical systems with generalized count observation

$$\mathbf{z}_{r1} \sim \mathcal{N}(\mu_1, Q_1)$$
$$\mathbf{z}_{r(t+1)}|\mathbf{z}_{rt} \sim \mathcal{N}(A\mathbf{z}_{rt}, Q)$$
$$x_{rti} \sim \mathcal{GC}(c_i^T \mathbf{z}_{rt}, g_i(\cdot)), i = 1, ..., n$$

- Practical considerations
    - Set $g_i(k) = -\infty$ for $k > K$ to facilitate computation;
    - Ridge penalty on the 2nd difference of $g_i(\cdot)$ to avoid overfitting;
    penalty $= \lambda \sum_{k=1}^{K-1}(g_i(k-1) - 2g_i(k) + g_i(k+1))^2$.

Variational Bayes Expectation Maximization (VBEM)

- **x**: data, **z**: latent variables, θ: model parameters,
- Often hard to compute $p_\theta(\mathbf{x}) = \int p_\theta(\mathbf{x}, \mathbf{z}) d\mathbf{z}$ and $p_\theta(\mathbf{z}|\mathbf{x})$.
- Approximate the posterior by a tractable distribution family.

$$p_\theta(\mathbf{z}|\mathbf{x}) \approx q(\mathbf{z}) \in \mathcal{Q}$$

- Optimize a lower bound of log likelihood, or ELBO

$$\text{ELBO}(\theta, q) = \int \left[\log p_\theta(\mathbf{x}, \mathbf{z}) - \log q(\mathbf{z}) \right] q(\mathbf{z}) d\mathbf{z}$$
$$= \log p_\theta(\mathbf{x}) - \text{KL}(q(\mathbf{z}) || p_\theta(\mathbf{z}|\mathbf{x})) \leq \log p_\theta(\mathbf{x})$$

Variational Bayes Expectation Maximization (VBEM)

- VBEM: Optimize $\text{ELBO}(\theta, q) \leq \log p_\theta(\mathbf{x})$ iteratively
  - E-step: For a fixed θ, optimize q
  - M-step: For a fixed q, optimize θ
- VBEM for GCLDS
  - We set q to be multivariate Gaussian
  - We derive a looser but tractable ELBO
  - E-step: fast Laplace approximation initialization + dual optimization (Emtiyaz Khan et al. 2013)
  - M-step: convex optimization + analytical solution
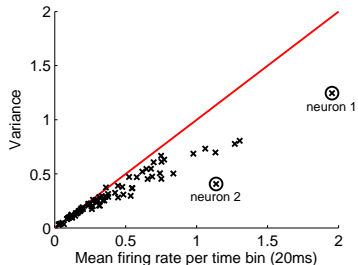
- For both simulated and real dataset, we compare GCLDS with PLDS (Poisson observation model)

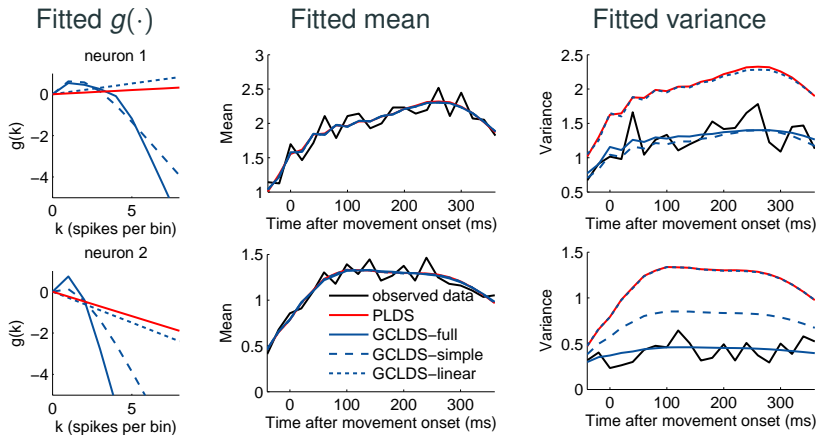|	Mean	Variance	Likelihood
PLDS	✓	✗	✗
GCLDS	✓	✓	✓
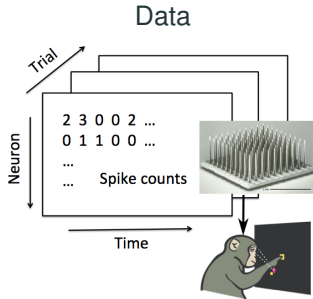
Data

Variance and mean of spike counts

- Center-out reaching experiments
- Multi-electrode array recording
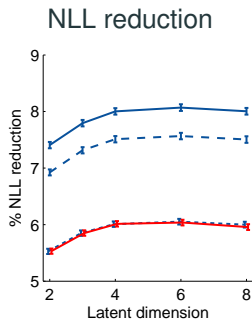- Strong under-dispersion

- Main algorithms to be compared
  - PLDS: Poisson observation (Macke, Buesing, and Sahani 2015)
  - GCLDS-full: Generalized count observation, individual $g(\cdot)$ across neurons
- Two control cases for GCLDS
  - GCLDS-linear: truncated linear $g(\cdot)$ (truncated Poisson)
  - GCLDS-simple: $g(\cdot)$ shared across neurons (up to a linear function)

Real data analysis: single neuron fit

Fitted $g(\cdot)$

Fitted mean

Fitted variance

Legend:
- observed data
- PLDS
- GCLDS-full
- GCLDS-simple
- GCLDS-linear

- Leave-one-neuron-out prediction
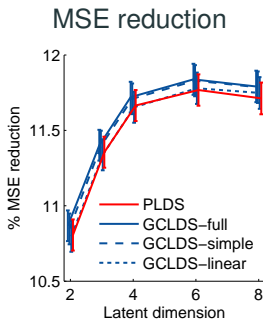
Conclusion and discussion

- Summary
    - Incorporated generalized count family into state space models.
    - Developed VBEM algorithm.
    - Observed superior fitted results on real neural data.
- Gao Y, Buesing L, Shenoy KV, Cunningham JP (2015) High-dimensional neural spike train analysis with generalized count linear dynamical systems. NIPS 2015.

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- Neural activities lie in a low-dimensional nonlinear manifold rather than a linear subspace
- Flexible observation model makes the state space model more expressive

Model formulation: fLDS

- Linear dynamical systems with nonlinear link and count observation

$$\mathbf{z}_{r1} \sim \mathcal{N}(\mu_1, Q_1)$$
$$\mathbf{z}_{r(t+1)} | \mathbf{z}_{rt} \sim \mathcal{N}(A\mathbf{z}_{rt}, Q)$$
$$x_{rti} \sim \text{Poisson}(f_i(\mathbf{z}_{rt})) \text{ (PfLDS)}$$
$$\text{or } \mathcal{GC}(f_i(\mathbf{z}_{rt}), g_i(\cdot)) \text{ (GCfLDS)}$$

  where f_i is a nonlinear function parameterized by a neural network

- Linear latent dynamics: simple, tractable, interpretable
- Nonlinear observation: flexible

Inference algorithm: AEVB (high level idea)

- Auto-encoding Variational Bayes (AEVB)
- Learn a mapping (recognition model) from data to the approximate posterior distribution of latent variable.
- Jointly optimize the generative model parameters and recognition model parameters.
- Naturally incorporate stochastic optimization to handle large datasets.
- Tractable for a large class of graphical models

Inference algorithm: AEVB (algorithm)

- Decompose ELBO by trials

$$\text{ELBO}(\theta, q) = \sum_{r=1}^{R} \int \left[\log \frac{p_\theta(\mathbf{x}_r, \mathbf{z}_r)}{q(\mathbf{z}_r)} \right] q(\mathbf{z}_r) d\mathbf{z}_r$$

- Map data $\mathbf{x}_r$ to $q(\mathbf{z}_r)$ by a parameterized function

$$q(\mathbf{z}_r) = q_\phi(\mathbf{z}_r | \mathbf{x}_r) = \mathcal{N}\left(\mu_\phi(\mathbf{x}_r), \Sigma_\phi(\mathbf{x}_r) \right)$$

- Learn both θ and ϕ by stochastic gradient descent on ELBO

$$\nabla \text{ELBO}(\theta, q_\phi) \approx R \times \nabla \int \left[\log \frac{p_\theta(\mathbf{x}_r, \mathbf{z}_r)}{q_\phi(\mathbf{z}_r | \mathbf{x}_r)} \right] q_\phi(\mathbf{z}_r | \mathbf{x}_r) d\mathbf{z}_r$$

$$\approx R \times \text{an unbiased estimator of gradient}$$

Structure of the recognition model

- Generative model:

$$p_\theta(\mathbf{z}_r | \mathbf{x}_r) \propto \underbrace{p_\theta(\mathbf{z}_{r1}) \prod_{t=1}^{T-1} p_\theta(\mathbf{z}_{t(t+1)} | \mathbf{z}_{rt})}_{Gaussian} \prod_{t=1}^{T} \underbrace{p_\theta(\mathbf{x}_{rt} | \mathbf{z}_{rt})}_{Complicated}$$

- Recognition model, product-of-Gaussian form:

$$q_\phi(\mathbf{z}_r | \mathbf{x}_r) \propto \underbrace{q_\phi(\mathbf{z}_{r1}) \prod_{t=1}^{T-1} q_\phi(\mathbf{z}_{r(t+1)} | \mathbf{z}_{rt})}_{Gaussian} \prod_{t=1}^{T} \underbrace{q_\phi(\mathbf{z}_{rt} | \mathbf{x}_{rt})}_{Gaussian}$$

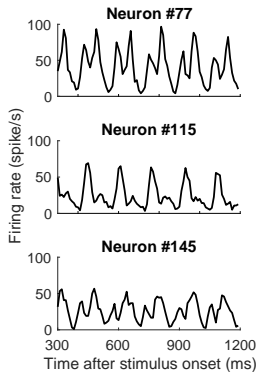- Multivariate Gaussian distribution with block tri-diagonal precision matrix.

	Mean	Variance	Likelihood	Concise representation
PLDS	✓	✗	✗	✗
GCLDS	✓	✓	✓	✗
PfLDS	✓	✗	✗	✓
GCfLDS	✓	✓	✓	✓

Real data analysis: primate visual cortex

Firing rate

Firing rate

Data

Reaching trajectory

- Latent projection with 2 latent dimensions
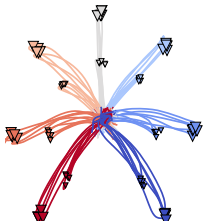
Reaching trajectory   PLDS   PfLDS
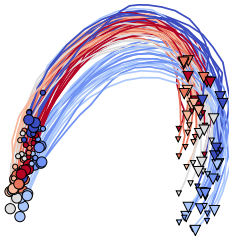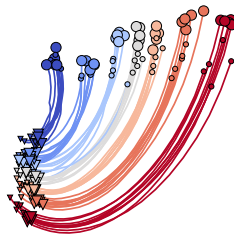
- Recovering the behavior structure from neural data in a unsupervised fashion.
- Concise and informative latent dimensions.

- One-step-ahead predictive performance

MSE reduction

NLL reduction
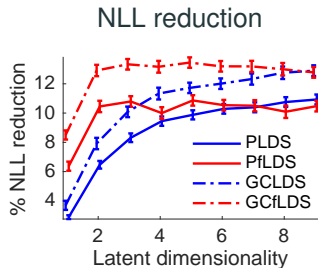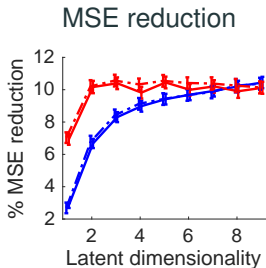
- Summary
  - Incorporated nonlinear observation into state space models.
  - Developed AEVB algorithm (flexible and scalable).
  - Obtained concise latent representations.
- Gao Y*, Archer E*, Paninski L, Cunningham JP (2016) Linear dynamical neural population models through nonlinear embeddings. NIPS 2016. (* = equal contribution)

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- Basic principle: the spiking activity of a neuron induces a transient increase in calcium concentration, which can be indirectly observed by recording the fluorescent properties of certain calcium indicators.

Spike train (cartoon)
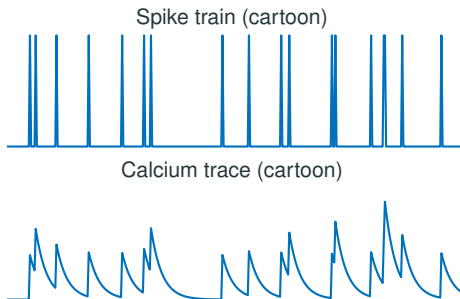
Calcium trace (cartoon)

- Calcium imaging enables simultaneous recording of many neurons.

Neuron shape $\times$ Calcium trace $+$ Noise $=$ Obseravtion
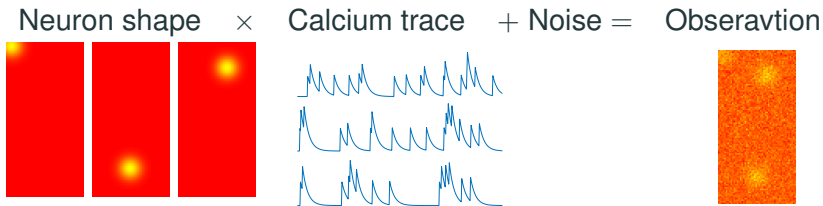
- Goal: recover the neuron shape and calcium trace given the observation.

Model formulation

- $X \in \mathbb{R}^{N \times T}$ represents the calcium imaging data, where each column is a (vectorized) frame that contains N pixels

- Decompose X into a product of K spatial component and temporal component

$$X = DA^T + \text{noise}$$

  - $D = [D_1, ..., D_K] \in \mathbb{R}^{N \times K}$ represents the neuron shapes
  - $A = [A_1, ..., A_K] \in \mathbb{R}^{T \times K}$ is the neural activities

- Further exploit structure of the components (localized neuron shapes)

Model formulation: objective

- Structured matrix factorization

$$\underset{D,A}{\text{minimize}} \quad \|X - DA^T\|_2^2 + f_D(D),$$

$$\text{subject to} \quad D_k \in \mathcal{D}_w^+; k = 1, \ldots, K,$$

$$\|A_k\|_2 \leq c_k,$$

- $\mathcal{D}_w^+$: non-negative vectors whose nonzero values is within a $w \times w$ window

- $f_D(D)$ regularizes the neuron shape (will discuss later)

- $\|A_k\|_2 \leq c_k$ avoids degenerate solution

Greedy algorithm

- Identify ROIs one at a time, using the residual un-explained by existing signals.
- At iteration k, given the current residue (unexplained by existing ROIs)
  - Greedy identification: Identify the location p_k where the Gaussian kernel explains most of the data (across time)
  - Shape fine tuning: Locally optimize the spatial and temporal component
  - Residue update: Subtract the newly identified signal

Shape fine tuning

- Given current residue R, an identified center pixel p_k, denote S_k as a $w \times w$ window centered at p_k

$$\underset{D_k, A_k}{\text{minimize}} \quad \|R - D_k A_k^T\|^2 + f(D_k),$$

$$\text{subject to} \quad D_{kp} \geq 0, p \in S_k,$$

$$D_{kp} = 0, p \notin S_k,$$

$$\|A_k\|_2 \leq c_k,$$

- $f(D_k) = \sum_{i=1}^{3} \lambda_i f_i(D_k)$
  - $f_1(D_k) = \sum_p \tau_{(p,p_k)} |D_{kp}|$ encourages sparsity
  - $f_2(D_k) = \sum_p (D_{kp} - G_{p_k})^2$ encourages Gaussian shape
  - $f_3(D_k) = \sum_{p_1 \text{ and } p_2 \text{ are neighbors}} (D_{kp_1} - D_{kp_2})^2$ encourages smoothness
- Optimize by block coordinate descent

Mean image with fitted ROI locations

Fitted ROI shape

Mean image with fitted ROI locations

Fitted ROI shape

Conclusion and discussion
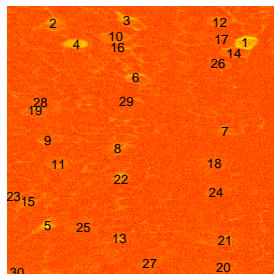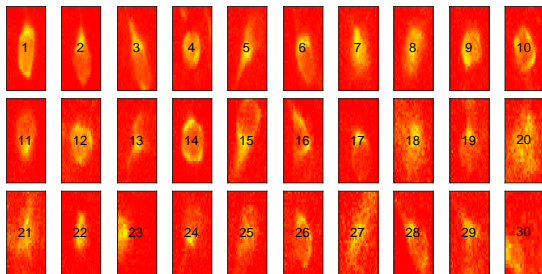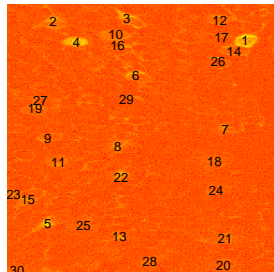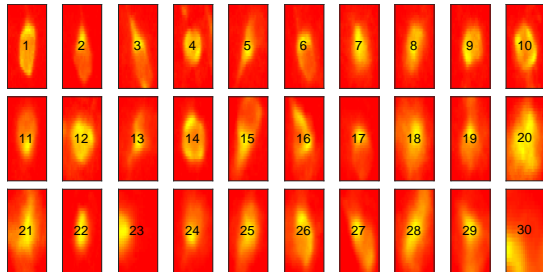
- Summary
  - Formulated calcium imaging ROI detection as a structured matrix factorization problem.
  - Developed a fast greedy algorithm.

- Pnevmatikakis EA, Soudry D, Gao Y, Machado TA, Merel J, Pfau D, Reardon T, Mu Y, Lacefield C, Yang W, Ahrens M, Bruno R, Jessell TM, Peterka DS, Yuste R, Paninski L (2016) Simultaneous denoising, deconvolution, and demixing of calcium imaging data. Neuron, 89(2), 285-299.

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Maximum entropy principle

- Entropy: for a continuous distribution with density $p(\mathbf{z})$ where $\mathbf{z} \in \mathbb{R}^d$, the entropy is defined as

$$H(p) = -\int p(\mathbf{z}) \log p(\mathbf{z}) d\mathbf{z} = E_{\mathbf{Z} \sim p} \left[-\log p(\mathbf{Z}) \right].$$

- A popular measure of diversity or information content.

- Maximum entropy (ME) principle: Subject to some given prior knowledge, the distribution that makes minimal additional assumptions is that which has the largest entropy of any distribution obeying those constraints

Applications of maximum entropy

- Neuroscience: generate a distribution of neural activity by specifying a set of features (pairwise correlation etc.) for hypothesis testing.
- Texture modeling: generate an image with a certain texture by specifying expected value of features relevant to texture.
- Ecology, natural language processing, finance...

Maximum entropy problem

- Maximum entropy (ME) problem

$$p^* = \underset{p}{\text{maximize}} \quad H(p)$$
$$\text{subject to} \quad E_{\mathbf{Z} \sim p}[T(\mathbf{Z})] = 0$$
$$\text{supp}(p) = \mathcal{Z},$$

where $T(\mathbf{z}) = (T_1(\mathbf{z}), ..., T_m(\mathbf{z})) : \mathcal{Z} \to \mathbb{R}^m$ is the vector of known statistics, and $\mathcal{Z}$ is the given support.

Gibbs distribution

- Lagrange multipliers argument $\Rightarrow$ an exponential family form for ME distribution (Gibbs distribution):

$$p^*(\mathbf{z}) \propto e^{<\eta, T(\mathbf{z})>} \mathbb{1}(\mathbf{z} \in \mathcal{Z})$$

- Hard to Identify $\eta \in \mathbb{R}^m$.
- Hard to sample from the distribution.
- Question: is there a better way to do this?

Maximum entropy flow network (MEFN)

- Identify a transformation $f_{\phi^*} \in \mathcal{F} = \{f_\phi : \mathbb{R}^d \to \mathbb{R}^d, \phi \in \mathbb{R}^q\}$ that transforms a simple distribution $\mathbf{Z}_0 \sim p_0$ (Gaussian) to $\mathbf{Z} = f_{\phi^*}(\mathbf{Z}_0) \sim p_{\phi^*}$ that approximates the ME distribution.

$$
\begin{aligned}
p^* = \underset{p}{\text{maximize}} \quad & H(p) \\
\text{subject to} \quad & E_{\mathbf{Z} \sim p}[T(\mathbf{Z})] = 0 \\
& \text{supp}(p) = \mathcal{Z},
\end{aligned}
$$

$$\Downarrow$$

$$
\begin{aligned}
\phi^* = \underset{\phi}{\text{maximize}} \quad & H(p_\phi) \\
\text{subject to} \quad & R(\phi) = E_{\mathbf{Z}_0 \sim p_0}[T(f_\phi(\mathbf{Z}_0))] = 0 \\
& \text{supp}(p_\phi) = \mathcal{Z}.
\end{aligned}
$$

Normalizing flow

- Normalizing flow: A parameterized family of smooth and invertible transformation.

- Explicit probability density form for $Z = f_\phi(Z_0) \sim p_\phi$ by change-of-variable theorem

$$p_\phi(\mathbf{z}) = p_0\left(f_\phi^{-1}(\mathbf{z})\right) |\det(J_\phi(\mathbf{z}))|^{-1}.$$

- Constructing a flexible family of normalizing flows by composing simple normalizing flows (deep learning)

$$\mathbf{Z} = f_k \circ f_{k-1} \circ \cdots \circ f_1(\mathbf{Z}_0).$$

Augmented Lagrangian method

- Augmented Lagrangian method minimizes the objective

$$L(\phi; \lambda, c) = -H(p_\phi) + \lambda^\top R(\phi) + \frac{c}{2}||R(\phi)||^2.$$

  for a sequence of $\lambda \in \mathbb{R}^m$ and $c \geq 0$.

- Update rule: at iteration k, given λ_k and c_k.
    - Find ϕ_k that optimizes $L(\phi; \lambda_k, c_k)$.
    - Update λ and c by

$$\lambda_{k+1} = \lambda_k + c_k R(\phi_k)$$

$$c_{k+1} = \begin{cases} \beta c_k & ||R(\phi_k)|| > \gamma||R(\phi_{k-1})|| \\ c_k & \text{otherwise} \end{cases}$$

  for some $\gamma \in (0, 1)$, $\beta > 1$

Augmented Lagrangian method in stochastic setting

- Both $H(p_\phi)$ and $R(\phi)$ are intractable, but we can approximate with an unbiased estimation,

$$H(p_\phi) \approx \frac{1}{n} \sum_{i=1}^{n} -\log p_\phi(\mathbf{z}^{(i)}),$$

$$R(\phi) \approx \frac{1}{n} \sum_{i=1}^{n} T(f_\phi(\mathbf{z}^{(i)})),$$

where $\mathbf{z}^{(i)} \sim p_0$. We can then optimize the objective by stochastic gradient descent.

Texture modeling (overview)

- Goal: construct a distribution on images that mimics a given texture (from a training image).
    - Authenticity: the samples from the distribution should mimic the given texture (constraints).
    - Sample diversity: the distribution should generate a variety of images with certain texture (entropy).
- Natural application of maximum entropy modeling.

Texture modeling (formulation)

- Sample from the space of images $\mathbf{z} \in [0, 1]^{d=w \times h \times c}$ where $w \times h$ is the image size, c is the number of channels ($c = 3$ for RGB representation).
- Texture loss: $T(\mathbf{z}) : [0, 1]^{w \times h \times c} \to \mathbb{R}$, a complicated form proposed in Ulyanov et al. (2016).
- Ulyanov et al. (2016) proposes texture net, which solves the problem without considering entropy term.

$$\min E_{\mathbf{z} \sim p_0} \left[T(f_\phi(\mathbf{Z})) \right]$$

- We build MEFN using $T(\mathbf{z})$ as the expectation constraint.
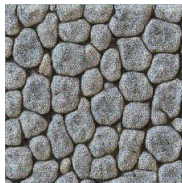- We use real-nvp (Dinh, Sohl-Dickstein, and Bengio 2016) as the normalizing flow structure.
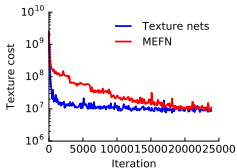
Input
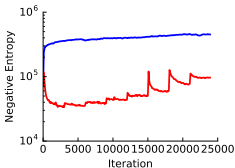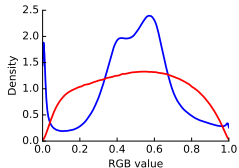
Texture net
(Ulyanov et al. 2016)

MEFN (ours)

Texture loss

Negative entropy

RGB histogram

Experiment: Texture modeling (diversity measure)

Method	d_{L^2}	SST	SSW	SSB
Texture net	0.077	0.043	0.036	0.006
MEFN	0.113	0.058	0.054	0.005

- For $n = 20$ randomly sampled image $\mathbf{z}^{(1)}, ..., \mathbf{z}^{(20)}$
- $d_{L^2} = \text{mean}_{i \neq j} \frac{1}{d} \|\mathbf{z}^{(i)} - \mathbf{z}^{(j)}\|_2^2$: Mean Euclidean distance.
- ANOVA: SST = SSW + SSB, $\bar{\mathbf{z}} = \frac{1}{n} \sum_i \mathbf{z}^{(i)}$.
  - SST $= \frac{1}{nd} \sum_{i,k} (z_k^{(i)} - \bar{z})^2$: Total var.
  - SSW $= \frac{1}{nd} \sum_{i,k} (z_k^{(i)} - \bar{z}_k)^2$: residue var, larger $\Rightarrow$ better.
  - SSB $= \frac{1}{n} \sum_k (\bar{z}_k - \bar{z})^2$: mean image var, smaller $\Rightarrow$ better.

Conclusion and discussion

- Summary
  - Solved maximum entropy problem by optimizing a normalizing flow.
  - Combined augmented Lagrangian optimization with stochastic optimization.
  - Obtained promising texture modeling result.
- Loaiza G*, Gao Y*, Cunningham JP (2017) Maximum entropy flow networks. ICLR 2017. (*=equal contribution)

Thank you!

Table of Contents

End of the talk.
Questions?

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Backup slides

GCLDS: Supervised case, GCGLM

- Given count data $x_i \in \mathbb{N}$, and associated covariates $z_i \in \mathbb{R}^p$, build Generalized Count Generalized Linear Model (GCGLM).

$$x_i \sim \mathcal{GC}(\theta(z_i), g(\cdot)), \quad \text{where} \quad \theta(z_i) = z_i\beta.$$

- Reminder: generalized count distribution

$$p_{\mathcal{GC}}(x; \theta, g(\cdot)) \propto \frac{\exp(\theta \cdot x + g(x))}{x!}, \quad x \in \mathbb{N}$$

GCLDS: special cases of GCGLM

Model Name	Typical Parameterization	GCGLM Parametrization
Logistic regression	$P(x = k) = \dfrac{\exp\left(k(\alpha + z\beta)\right)}{1 + \exp(\alpha + x\beta)}$	$g(k) = \alpha k; k = 0, 1$
Poisson regression	$P(x = k) = \dfrac{\lambda^k}{k!}\exp(-\lambda);$ $\lambda = \exp(\alpha + z\beta)$	$g(k) = \alpha k$
Adjacent category regression	$\dfrac{P(x = k + 1)}{P(x = k)} = \exp(\alpha_k + z\beta)$	$g(k) = \displaystyle\sum_{i=1}^{k}(\alpha_{i-1} + \log i);$ $k = 0, 1, ..., K$
Negative binomial regression	$P(x = k) = \dfrac{(k + r - 1)!}{k!(r - 1)!}(1 - p)^r p^k$ $p = \exp(\alpha + z\beta)$	$g(k) = \alpha k + \log\left(k + r - 1\right)!$
COM-Poisson regression	$P(x = k) = \dfrac{\lambda^k}{(k!)^\nu} / \displaystyle\sum_{j=1}^{+\infty} \dfrac{\lambda^j}{(j!)^\nu}$ $\lambda = \exp(\alpha + z\beta)$	$g(k) = \alpha k + (1 - \nu)\log k!$

MEFN: Normalizing flow structures

- Rezende and Mohamed (2015) Proposes two specific families of transformations for variational inference

- Planar flow

$$f_i(\mathbf{z}) = \mathbf{z} + \mathbf{u}_i h(\mathbf{w}_i^T \mathbf{z} + b_i),$$

where $b_i \in \mathbb{R}$, $\mathbf{u}_i, \mathbf{w}_i \in \mathbb{R}^d$ and h is an activation function.

- Circular flow

$$f_i(\mathbf{z}) = \mathbf{z} + \beta_i h(\alpha_i, r_i)(\mathbf{z} - \mathbf{z}_i'),$$

$\beta_i \in \mathbb{R}$, $\alpha_i > 0$, $\mathbf{z}_i' \in \mathbb{R}^d$, $h(\alpha, r) = 1/(\alpha + r)$ and
$r_i = ||\mathbf{z} - \mathbf{z}_i'||$.

MEFN: real-nvp normalizing flow structure

- Used in Dinh, Sohl-Dickstein, and Bengio (2016) for image denstiy estimation.
- Affine coupling layer: split variable $\mathbf{z} \in \mathbb{R}^D$ into $\mathbf{z}_1 \in \mathbb{R}^d$ and $\mathbf{z}_2 \in \mathbb{R}^{D-d}$. linearly transform $\mathbf{z}_2$ given $\mathbf{z}_1$.

$$f \begin{pmatrix} \mathbf{z}_1 \\ \mathbf{z}_2 \end{pmatrix} = \begin{pmatrix} \mathbf{z}_1 \\ \mathbf{z}_2 \odot \exp\left(s(\mathbf{z}_1)\right) + t(\mathbf{z}_1) \end{pmatrix}$$

  where $\odot$ is element-wise product.
 $s(\mathbf{z}_1), t(\mathbf{z}_1) : \mathbb{R}^d \to \mathbb{R}^{D-d}$.
- The transformation family is flexible because
    - s and t can be complex while maintaining tractability (inversion and Jacobian computation).
    - Partitioning $\mathbf{z}$ into $(\mathbf{z}_1, \mathbf{z}_2)$ can be chosen arbitrarily.
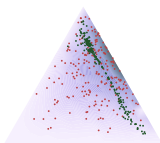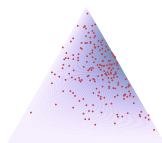
Experiment: Dirichlet

- Dirichlet distribution is the ME distribution on a simplex
  $\mathcal{S} = \{\mathbf{z} = (z_1, \ldots, z_{d-1}) : z_i \geq 0 \text{ and } \sum_{k=1}^{d-1} z_k \leq 1\}$ with
  expetation on the log of each coordinate
  $E[\log Z_k] = \kappa_k (k = 1, ..., d)$, where $Z_d = 1 - \sum_{k=1}^{d-1} Z_k$.
- 10 layers of planar flow (Rezende and Mohamed 2015).
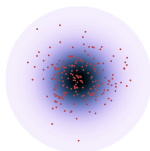
Initial distribution p_0

MEFN result $p_{\phi*}$ (Control case: moment matching)

Ground truth $p*$