

Improving Catch Estimation Methods in Sparsely Sampled Mixed-Stock Fisheries.

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Introduction

- I am Nick
- describe the california spp comps. port sampling data for modeling
- describe our modeling efforts for estimating spp comps.

MCATs in Time

- **Top Panel:** Number of samples in rockfish market categories (1978-2015)
 - Colors represent different market categories
 - Thickness shows the number of samples
- **Bottom Panel:** Number of rockfish market categories
 - Count the colors
 - ~20 mcats in the late 70s
 - ~50 mcats in the recent times
- **Middle Panel:** Average number of samples per stratum
 - Find samples for each stratum (mcat, gear, port, year, qtr)
 - Average them
- 1978-1982
- 1983-1990

78-82 Bars

- **Top Panel:**
 - For each market category accounting for 99% of landings
 - * (blue) Proportion landings by weight
 - * (red) Proportion samples by #
- **Bottom Panel:** Aggregated Species Compositions
 - Colors represent 13 select species (others grey)
 - Number above is the # species present
 - Hatching is MCAT nominal species
- MCATs not pure
 - often nominal species is not even the major species
 - * BCAC
 - * BRWN
- Sampling Opportunistic
 - Sampling co-occurs with landings
 - Often as more species are present there are more samples
 - This is lucky for modeling
 - * More samples than parameters (largely driven by spp)
 - * Most landings are modeled (78-82: 96.8%)
- No sampling south of Conception

83-90 Bars

- Same picture but 83-90
 - (blue) Proportion landings by weight
 - (red) Proportion samples by #
 - Aggregated Species Compositions
- Top 99% of landings in more market categories
 - MCATs still largely impure
- 83-90: 98.3% of landings modeled

Likelihood Forms

- First modeling choice: Pick a Likelihood
- Shelton et al. 2012 Fit Multinomial via the Multinomial-Poisson trans.
 - Piece together independent Poissons
- We are not limited to Multinomial distribution
 - quantify uncertainty (residual variability)
 - consider modeling overdispersion
 - additional parameter (ϕ) to disentangle mean from variance
- y_{ij} : i^{th} sample of the j^{th} species' integer weight
- Remove all other modeling decisions by modeling a single stratum
 - MCAT 250
 - Monterey
 - Trawl
 - 1982/Q2

Likelihood Graphs

- Fit models and look at how they predict
- **Left Panel:** 95% HDI from each model along side observed sppComp data
 - black horizontal lines are observed species comps
 - blue is Poisson (i.e. Multinomial) Model
 - red is Binomial
 - green is the Negative Binomial Model
 - yellow is the Beta-binomial Model
- **Right Panel:** Entire Beta-binomial predictive distribution
- Overdispersion is present (spp comps from $[0,1]$)
- ~50 observations => 2.5 missing in 95% interval
 - Maybe NB missing a few to many, and BB missing a few to few
 - BB certainly finding the most variance
 - split intervals but... very appropriate density

Likelihood Table

- Consider MSE, DIC, WAIC, and Marginal Likelihood Bayesian Model Prob.
- Varied model selection criterion (Nothing is perfect!)
- Consistent and large support for the Overdispersion Models
 - Most support for BB
- Moving forward I develop the BB model

Beta-Binomial Model

- A Full Operationalized Model!
- $y_{ijklm\eta}$: i^{th} sample of the j^{th} species' integer weight, in the k^{th} port, caught with the l^{th} gear, in the η^{th} quarter, of year m , for a particular market category.
- Stratum μ linked to θ and observed cluster size (n)
- Stratum σ^2 is largely a function of μ but with overdispersion ρ
 - $\rho \rightarrow 0$: Binomial variance
 - $\rho \rightarrow 1$: n times Binomial variance
- Modeling of θ (all predictors are categorical):
 - Intercept
 - Additive offsets for: Species, Port, Gear
 - Consider multiple time models

Time Models

- Bayesian Modeling
 - Heirarchical v. Random Effect Disclaimer
- (M1) Fixed main effect time model
 - No pooling
- (M2) Random main effect time model
 - years/quarter pool separately
- (M3) Random main effects + random interaction
- (M4) Random interactions jointly pooled
- (M5) Random interactions quarterly variances pooling across years
- (M6) Random interactions yearly variances pooling across quarters

Priors

- Very diffuse priors
- Main effects diffuse Normals
- ρ transformed to be a real number
 - $\text{logit}(\rho) \rightarrow (-3.91, 3.91)$
 - $\rho \rightarrow (0.02, 0.98)$
- Any heirarchical variance gets the same IG prior
 - Considered others:
 - * $\sqrt{v} \sim \text{Half-Cauchy}(10^{-2})$
 - * $\sqrt{v} \sim \text{Unif}(0, 10^5)$

Beta-Binomial Fits

- Fit model separately in 78-82 and 83-90 and compare model selection criterion
- 78-82:
 - Consistent support for more pooling
 - All measures point to (M4)
- 83-90:
 - Consistent support for interaction models
 - Uncertainty between (M3), (M4), and (M5)
 - Lesser support for (M6)
- We fit model (M4) everywhere
 - Stable and relatively fast model to fit
 - Given its support in 78-82, I am drawn to (M4)
 - * Each time period seems to have a mind of its own

$$\begin{aligned}\beta_{m\eta}^{(t)} &= \beta_m^{(y)} + \beta_\eta^{(q)} + \beta_{m\eta}^{(y:q)} \\ \beta_m^{(y)} &\sim N(0, 32) \\ \beta_\eta^{(q)} &\sim N(0, 32) \\ \beta_{m\eta}^{(y:q)} &\sim N(0, v)\end{aligned}$$

?? LUNCH ??

Posterior Predictive Species Comps.

- Having settled on (M4) in both time periods, how do we build species comps?
- Inference results in samples from posterior distribution $P(\mu_{jklm\eta}, \sigma_{jklm\eta}^2 | y)$
- Run samples back through BB likelihood to compute Monte Carlo integral and get posterior predictive distribution of sampled weight.
- Use draws from model posterior predictive weight to compute species comp. distribution
 - Plot shows average species compositions
 - Full distribution for y^* as well as π^*
 - Each sample sums to 1 and $\sum_j \mathbb{E}[\pi_j^*] = 1$
- By adding an unobserved latent time period we can make out-of-sample predictions
 - (M4): unobserved $\beta^{(y)*}$ and $\beta^{(q)*}$

Single Quarter Hindcast

- Recall for 1978-1982 there was no sampling south of point conception.
- Adding an unobserved year and quarter
 - make predictions for each species in each combo of:
 - * three observed gear groups
 - * three southern port complexes

78-82 Prediction

- Modeled MCATs
- MCATs in the rows (**ordered by landings**) w/ 3 nominal HDI prediction levels
 - For each stratum of each MCAT compare data to prediction intervals
 - Observed level should match Nomial
 - Prediction higher than nominal => Overfitting
 - Prediction lower => Underfitting (not enough residual variance)
- Most do well
 - Average performance is reasonable
 - Note this is a unweighted, simple, average
 - More accurate would weight average by samples at each stratum
- Particularly well in heavily landed stratum
 - correlation of sampling effort w/ landings
- Widow is a wild child
 - only example that is off by more than 5% points at any level

83-90 Prediction

- Same Table
 - Modeled MCATs (**ordered by landings**) w/ 3 nominal pred. levels
- Again most do well
- Recall landings were spread across more MCATs in 83-90
 - Enough samples to also model more MCATs
- Blackgill, Yellowtail, Cowcod: off by 5% points at some level
 - Negligable Landings

Speciating Landings

- $\lambda_{.klm\eta}$ is reported on landing receipts
- $\lambda_{jklm\eta}^*$ stored in DB
- Aggregate to any level
 - across quarter, port complex, gear group
 - Also MCAT (I ran out of index variables :/)
- E.J. will show the speciated time series with predictive intervals
 - summed across MCAT
 - as it might be used in assessment

BMA Story

- Mentioned partial pooling thru time via heirarchical modeling
- But present system also pools in space
 - Given sparsity, it's entirely possible that we also need spatial pooling
- I show MSE to demonstrate the biase/variance trade off
 - Pooling directly exchanges sample size (postior variance) with bias
 - **Far Right** Least Bias
 - **Far Left** Most Bias, but most data for small posterior variance
 - A practicle solution is somewhere in between
- [Bell number] Idea: Try all partions of port complexes
- $B_{10} = 115975$
 - Too many
 - add Spatial Modeling Constraints
 - Biogeography viewed through the lens of human behavior
 - * sampling behavior
 - * fisherman behavior
- $\bar{B}_{10} = 61136$
 - Require partitions to be “small”
 - No super-grouping greater than 3 port complexes
 - points close in space behave similarly (smoothness)
- $\hat{B}_{10} = 512$
 - Require continuous partitions (no leap frogging)
 - Like a GP continuity constraint
- $\hat{\hat{B}}_{10} = 274$
 - Together we have “small” and “continuous” partitions
 - smoothness and continuity
 - a computationally manageable set of models to compute

BMA Math

- Defines a candidate model set
- We could just pick the single “best” model
 - defining “best” is hard
 - model selection criterion are imperfect
- Prediction results are averaged results

78-82 BMA Results

- Describe plot
- Recall no sampling in the south
 - All latent structure filled in by predictive distribution in the south
- 250:
 - Marginal model probability
 - * N1: $32+14+13+12=71\%$
 - * N2: $2+2+2+2=8\%$
- 253:
 - Central Block
 - BRG/BDG
 - Lump/Split CRS and ERK (among top 5 models; 58% model weight)
 - * Split: 0.3448276
 - * Lump: 0.6551724
- 269:
 - Sold on the BRG-BDG break

83-90 BMA Results

- Recall BRG-OSF Missing data
- 250:
 - Missing data story
 - * Lump or
 - * Quarantine
- 956:
 - Lump/Split CRS and ERK
 - A break at Cape Mendicino
 - BRG/BDG/OSF Quarantine os missing data
- 269:
 - Piont Conception Break
 - Cape Mendicino Break

Conclusions

- Using Bayesian models we have:
 - Account for overdispersion
 - Estimate uncertainty (full distribution)
 - Formal Mechanisms for pooling
 - provide structure for making out-of-sample prediction
- Future Modeling
 - Explore additional predictore in θ
 - * Landing weighting
 - * Vessel Effects
 - * Speceies:Gear interactions
 - Overdispersion Multivariate models
 - * Dirichelette-Multinomial Model
 - Maybe Time Series Models
 - Cluster and integrate out spatial parameters via DP?