

Metamodeling for Bias Estimation of Biological Reference Points Under Two-Parameter SRRs

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■ Low Contrast

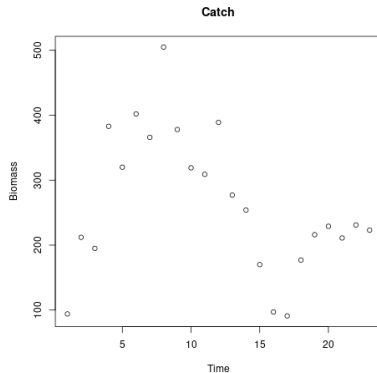
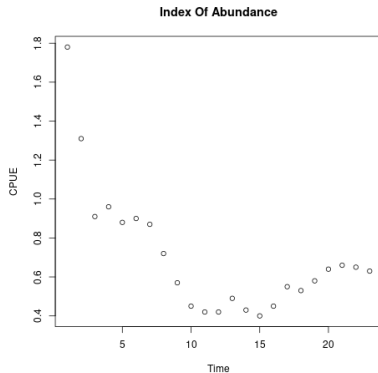
■ High Contrast

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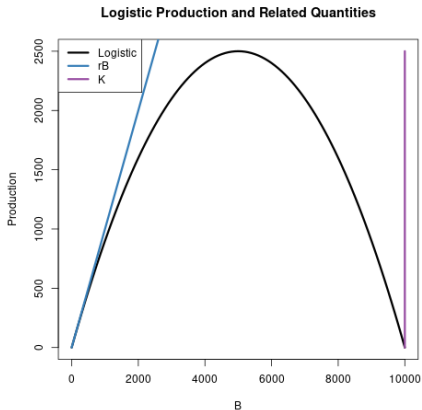
$$I_t = qB_t e^{\epsilon} \quad \epsilon \sim N(0, \sigma^2)$$

$$\frac{dB(t)}{dt} = P(B(t); \theta) - C(t)$$

Schaefer Model

$$P_{\theta}(B) = rB \left(1 - \frac{B}{K}\right)$$

$$\theta = (r, K)$$

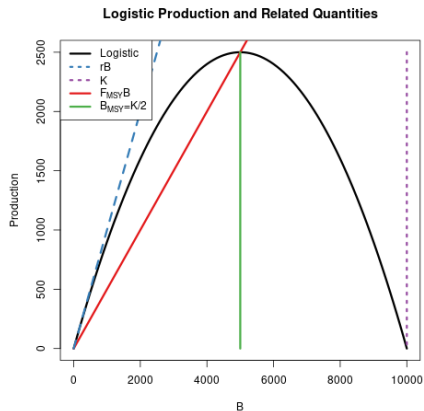


Schaefer Reference Points

$$F^* = \frac{r}{2}$$

$$\frac{B^*}{B_0} = \frac{1}{2}$$

$$MSY = \frac{rK}{4}$$



Conceptually:

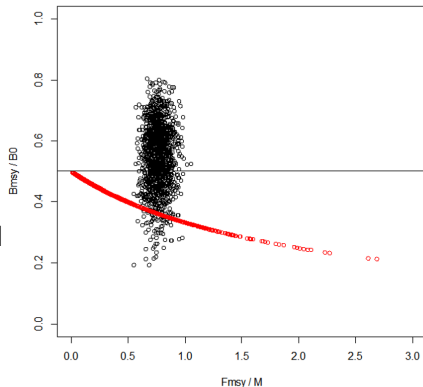
$$F^* \in \mathbb{R}^+ \quad \frac{B^*}{\bar{B}(0)} \in (0, 1)$$

Mangel et al. 2013, CJFAS:

- BH Model:

$$F^* \in \mathbb{R}^+ \quad \frac{B^*}{\bar{B}(0)} = \frac{1}{F^*/M+2}$$

- Similar Constraint for Ricker and other 2 Parameter Curves



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Schaefer Model:

$$F^* \in \mathbb{R}^+ \quad \frac{B^*}{\bar{B}(0)} = \frac{1}{2}$$

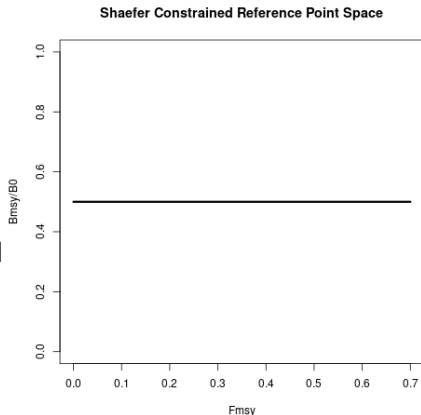


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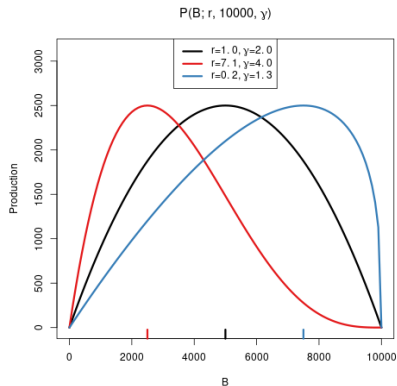
5 End

Pella-Tomlinson Production Model

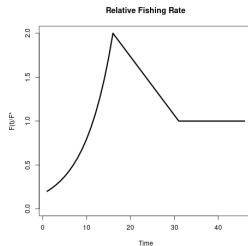
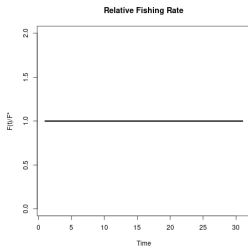
$$I(t) \sim LN(qB(t), \sigma^2)$$
$$\frac{dB(t)}{dt} = P_{\theta}(B(t)) - F(t)B(t)$$

$$P_{\theta}(B) = \frac{rB}{\gamma - 1} \left(1 - \frac{B}{K}\right)^{\gamma-1}$$
$$\theta = (r, K, \gamma)$$

$\gamma = 2 \Rightarrow$ Schaefer Model

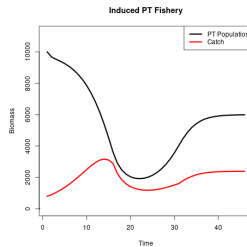
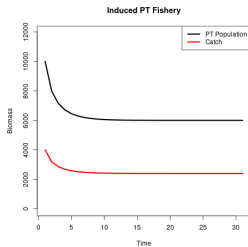
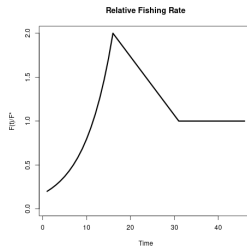
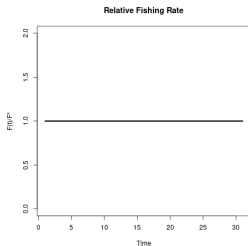


Catch

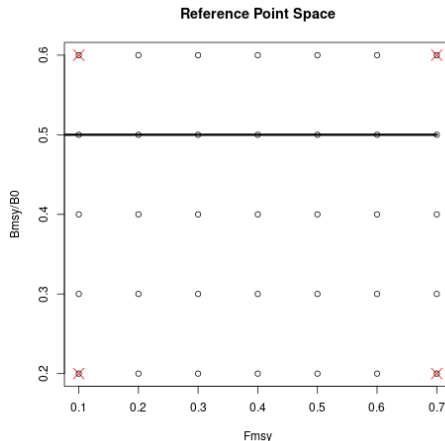


$$\begin{aligned} C(t) &= F(t)B(t) \\ &= F^* \left(\frac{F(t)}{F^*} \right) B(t) \end{aligned}$$

Catch

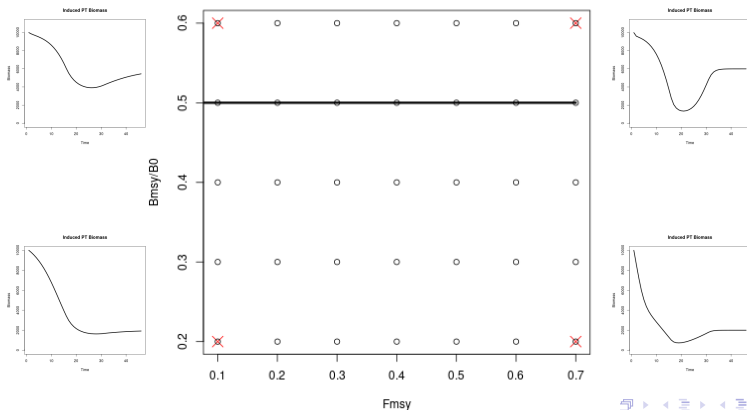


$$\theta = \left[r = F^* \left(\frac{1 - \frac{B^*}{\bar{B}(0)}}{\frac{B^*}{\bar{B}(0)}} \right) \left(1 - \frac{B^*}{\bar{B}(0)} \right)^{\left(\frac{\frac{B^*}{\bar{B}(0)} - 1}{\frac{B^*}{\bar{B}(0)}} \right)}, K = 10000, \gamma = \frac{1}{\frac{B^*}{\bar{B}(0)}} \right]$$



$$\theta = \left[r = F^* \left(\frac{1 - \frac{B^*}{\bar{B}(0)}}{\frac{B^*}{\bar{B}(0)}} \right) \left(1 - \frac{B^*}{\bar{B}(0)} \right)^{\left(\frac{\frac{B^*}{\bar{B}(0)} - 1}{\frac{B^*}{\bar{B}(0)}} \right)}, K = 10000, \gamma = \frac{1}{\frac{B^*}{\bar{B}(0)}} \right]$$

Reference Point Space



Metamodel

$$\underbrace{\left(F^*, \frac{B^*}{\bar{B}(0)} \right)}_{\text{PT Truth}} \xrightarrow{\text{GP}} \underbrace{\left(\hat{F}^*, \frac{\hat{B}^*}{\bar{B}(0)} \right)}_{\text{Shaefer Estimate}}$$

- GP interpolates over degrees of RP model misspecification.
- Propagation of estimate uncertainty smooths bias estimation.
- Explicitly highlights trade-offs induced in RPs.

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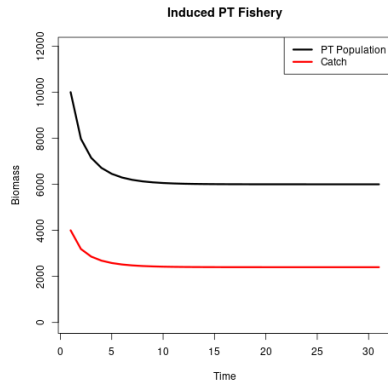
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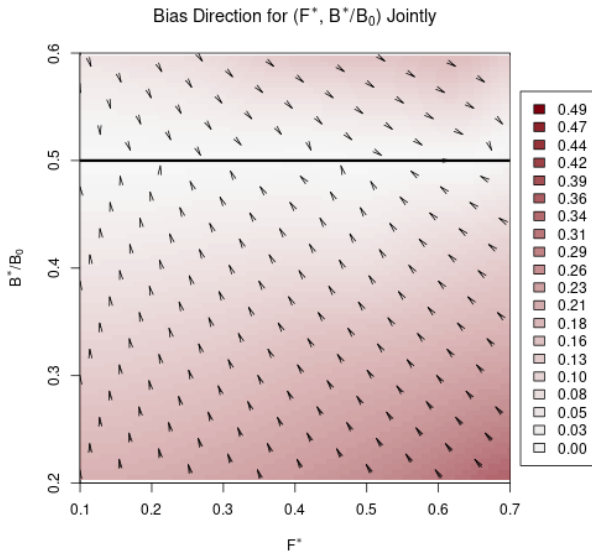
4 Proposals

■ Growth & Productivity

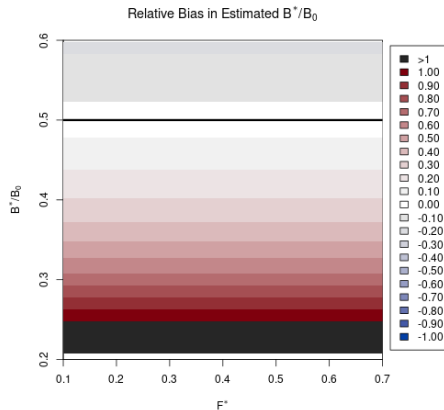
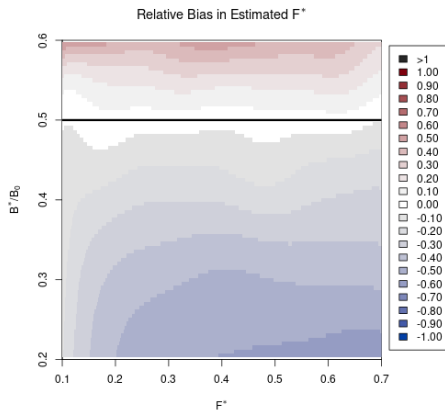
■ Catch Interpolation

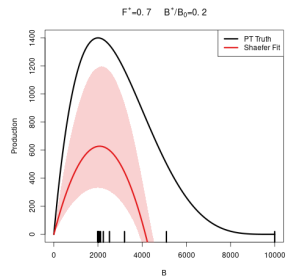
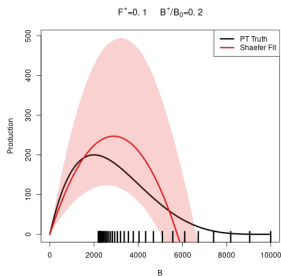
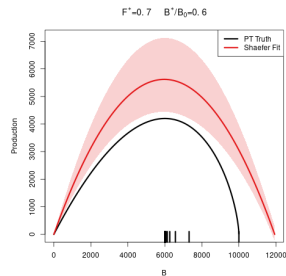
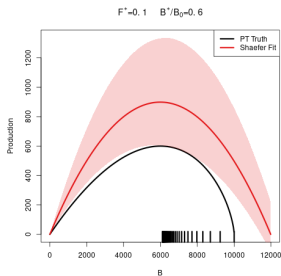
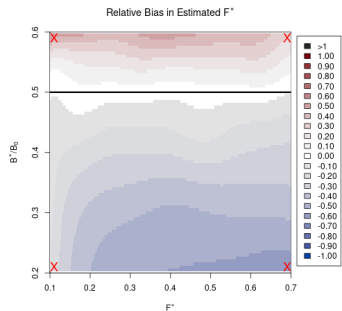
5 End

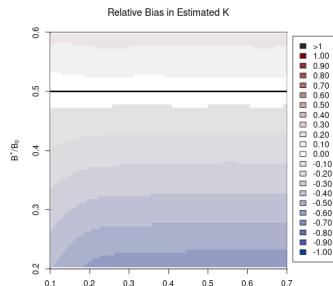
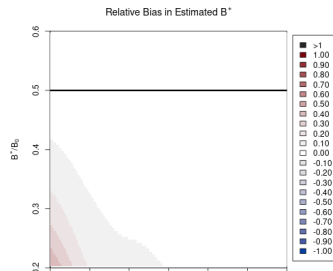
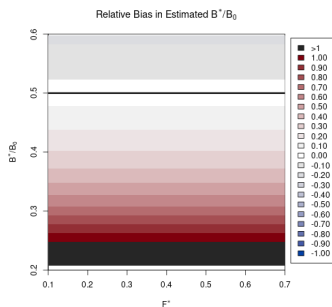




Components of Bias







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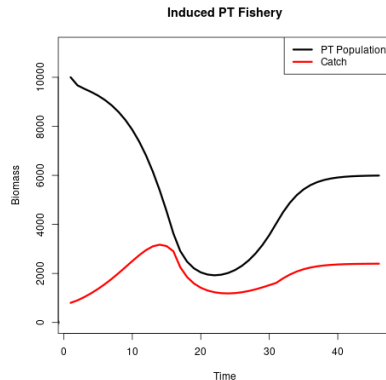
■ High Contrast

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■ Growth & Productivity

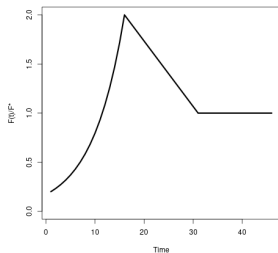
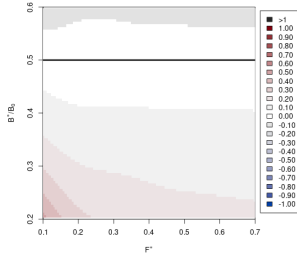
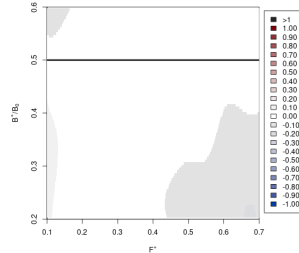
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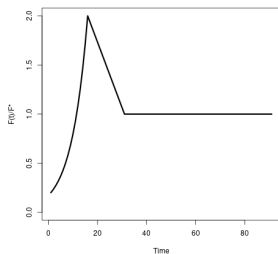
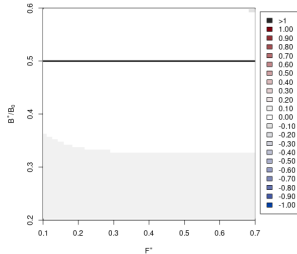
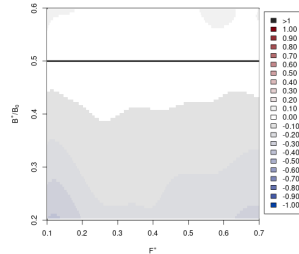


High Contrast

Relative Fishing Rate

Relative Bias in Estimated B^* Relative Bias in Estimated F^* 

Relative Fishing Rate

Relative Bias in Estimated B^* Relative Bias in Estimated F^* 

Summary

- Given unbiased estimation (i.e. MLE), model misspecification is a necessary but not sufficient condition for inducing bias.
- Different data informs different parts of the production function differently
- In the overconstrained setting we pay for our modeling mistakes in bias. In practice, when the true model is not known and the Schaefer model is unlikely to be correctly specified, one should at best expect to only estimate either B^* or F^* correctly depending on the particular degree of model misspecification. The observed contrast then serves to distribute the available information among B^* and F^* .
- F^* bias is strongly contrast dependent
 - ⇒ Bias depends on how similar the modeled and true production functions can be at the observed biomasses
- A rich simulation-based method for describing global RP bias

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Productivity Extension

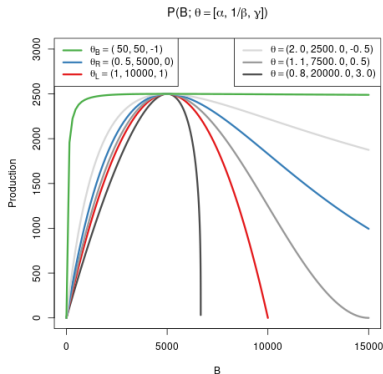
$$\frac{dB}{dt} = P(B; \theta) - (M + F)B$$

$$P(B; [\alpha, \beta, \gamma]) = \alpha B(1 - \beta\gamma B)^{\frac{1}{\gamma}}$$

$\gamma = -1 \Rightarrow$ Beverton-Holt

$\gamma = 1 \Rightarrow$ Logistic

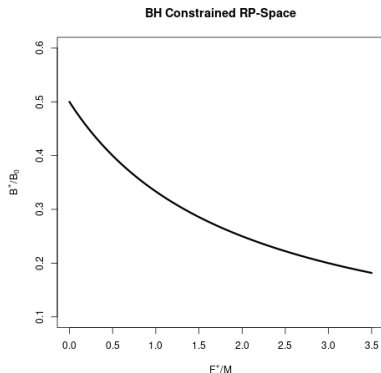
$\gamma \rightarrow 0 \Rightarrow$ Ricker



Productivity Extension

$$P_{\text{BH}}(B; [\alpha, \beta, -1]) = \frac{\alpha B}{(1 + \beta B)}$$

$$\frac{B^*}{\bar{B}(0)} = \frac{1}{\frac{F^*}{M} + 2}$$



Growth Extension

$$\begin{aligned}\frac{dB}{dt} &= \overbrace{w(a_0)R(B; \theta)}^{\text{Recruitment Biomass}} + \overbrace{\kappa [w_\infty N - B]}^{\text{Net Growth}} - \overbrace{(M + F)B}^{\text{Mortality}} \\ \frac{dN}{dt} &= R(B; \theta) - (M + F)N\end{aligned}$$

$$\begin{aligned}R(B; [\alpha, \beta, \gamma]) &= \alpha B(t - a_0)(1 - \beta\gamma B(t - a_0))^{\frac{1}{\gamma}} \\ w(a) &= w_\infty(1 - e^{-\kappa a})\end{aligned}$$

bullets of primary points of individual growth and maturity

Outline

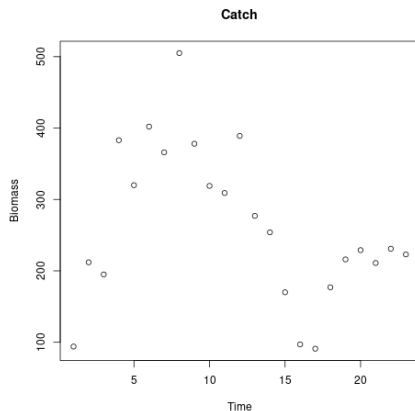
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Catch Interpolation

$$t \in \mathbb{R}^+ \quad \tau = \lceil t \rceil - 1$$

$$\mathbb{E}[y(t)] = \int_{\tau}^t x(t^*) dt^*$$

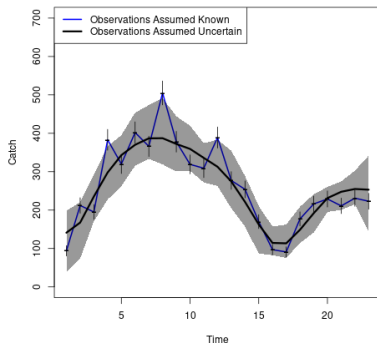
$$x(t) = \beta_0 + \sum_{j=1}^{T-1} \beta_j (t - \tau_j) \mathbf{1}_{t > \tau_j}$$



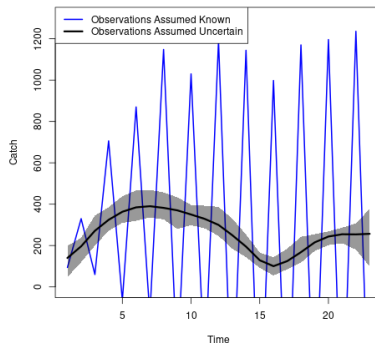
$$y(\tau_i) = \beta_0 + \sum_{j=1}^{i-1} \beta_j \left[\left(\frac{\tau_i^2}{2} - \tau_j \tau_i \right) \mathbb{1}_{\tau_i > \tau_j} - \left(\frac{\tau_{i-1}^2}{2} - \tau_j \tau_{i-1} \right) \mathbb{1}_{\tau_{i-1} > \tau_j} \right] + \epsilon_i$$

$$\beta_j \sim N(0, \phi) \quad \phi \sim \text{Half-Cauchy}(0, 1) \quad \epsilon_i \sim N(0, \sigma_i^2)$$

Observed Catch with Predictive Interpolations



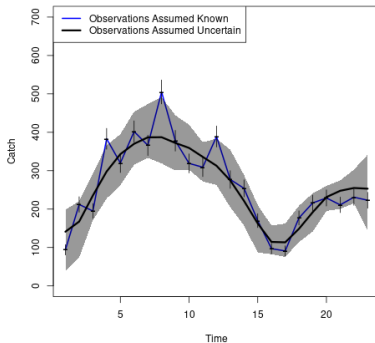
Interpolated Instantaneous Catch



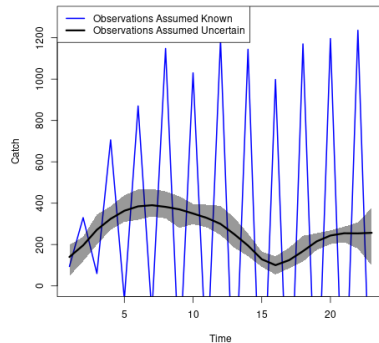
$$\frac{dB}{dt} = P_{\theta}(B(t)) - C(t)$$

$$B(\tau + 1) \approx B(\tau) + P_{\theta}(B(\tau)) - C(\tau)$$

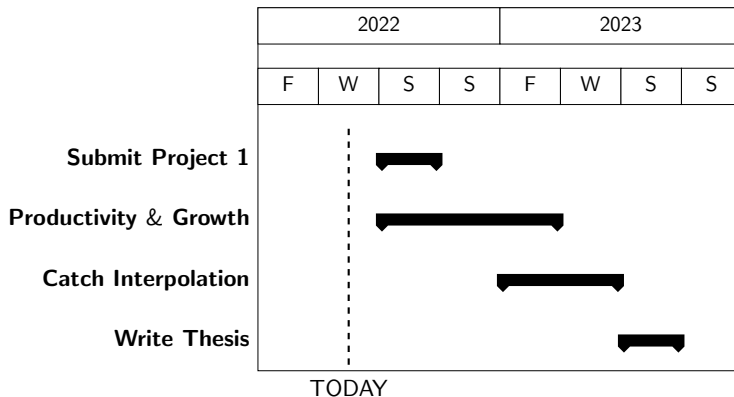
Observed Catch with Predictive Interpolations



Interpolated Instantaneous Catch



Timeline



Thanks and Acknowledgements NOAA, Sea Grant Ecetra

$$\hat{\mu} = \widehat{\log(r)} \quad - \text{ or } - \quad \hat{\mu} = \widehat{\log(K)}$$

$$\mathbf{x} = \left(F^*, \frac{B^*}{\bar{B}(0)} \right)$$

$$\hat{\mu} = \beta_0 + \beta' \mathbf{x} + f(\mathbf{x}) + \epsilon$$

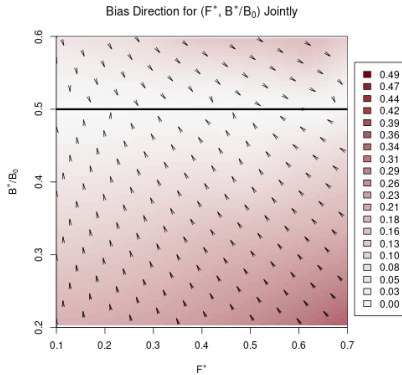
$$f(\mathbf{x}) \sim \text{GP}(0, \tau^2 R(\mathbf{x}, \mathbf{x}'))$$

$$\epsilon_i \sim \text{N}(0, \hat{\omega}_i).$$

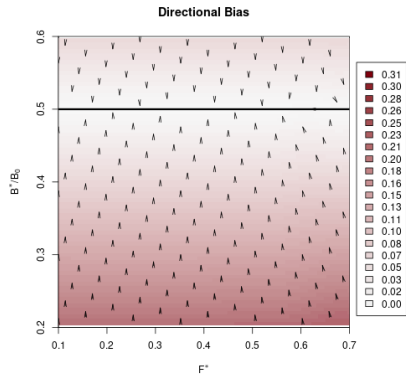
$$R(\mathbf{x}, \mathbf{x}') = \exp \left(\sum_{j=1}^2 \frac{-(x_j - x'_j)^2}{2\ell_j^2} \right)$$

Cross-Covariogram

Low Contrast



High Contrast



$$\frac{B^*}{\bar{B}(0)} = \frac{\left(\frac{\alpha}{M+F^*}\right)^{\frac{1}{\gamma}} - 1}{\left(\frac{\alpha}{M}\right)^{\frac{1}{\gamma}} - 1}$$
$$\alpha = (M + F^*) \left[1 - \frac{1}{\gamma} \left(\frac{F^*}{M + F^*} \right) \right]^{-\gamma}$$
$$\beta = \frac{1}{\gamma \bar{B}(0)} \left(1 - \left(\frac{\alpha}{M} \right)^{\frac{1}{\gamma}} \right)$$

$$\frac{dB}{dt} = P_{\theta}(B(t)) - C(t)$$

$$B(\tau + 1) \approx B(\tau) + P_{\theta}(B(\tau)) - C(\tau)$$

