

Department of Computer Science ETH Zürich

Evolutionary Dynamics

Assignment #02

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1.1 Simpson and Shannon index

1.1.1 a

Since all draws are independent the probability to observe the same type in k draws is simply a k – time multiplication (Product Rule) of individual probabilities pi (which is in essence a relative frequency of type i), so it it is p^k

Then we have to sum up such probability for each strain i (additivity axiom, each draw is mutually exclusive)

So, finally

$$D_k = \sum_{i=1}^n p_i^k$$

We remember from our final ODE model, that after immune response reaches its equilibrium at levels $x_i^* = \frac{cv_i}{b+uv}$ and $z^* = \frac{kv}{b+uv}$ and with this values we could have a single ODE for virus load:

$$v' = \frac{v}{b + uv}[rb - v(cpD + kq - ru)]$$

Solution of this ODE (i.e. behavior of v(t)) depends on parameters and, in general, has three regimes: immediate disease, chronic infection and disease after long asymptomatic period. A certain combination of these parameters called "antigenic diversity threshold"

$$D<\frac{ru-kq}{cp}$$

and it puts general equation above out of equilibrium, so causes uncontrolled grows of v(t). The only part of this equation which is changing during the course of the disease is D (genetic similarity of virus). It is actually close to 1 right after the infection and then gradually goes to smaller values (more different virus strains) till reaches its threshold value, when immune system cannot control virus anymore. From mathematical point of view increasing of dv/dt is smaller with large D.

1.1.2 b

For an uniform distribution of n strains:

$$H = -\sum_{i=1}^{n} p_i \log(p_i) \text{ if } p_1 = p_2 = \dots = p_n$$

$$H = -(p_i \log(p_i) + p_{i+1} \log(p_{i+1}) + \dots + p_n \log(p_n)) = -np \log(p) =$$

$$H = -p_i (\log p_i^n) = -np_i \log p_i \text{ and as } p_i = \frac{1}{n} \text{ for each } i$$

$$H = -n * \frac{1}{n} * \log \frac{1}{n} = -\log n^{-1} = \log n$$

For n = 2:

$$H = -p_1 \log p_1 - p_2 \log p_2$$

For $p_2 = 1 - p_1$:

$$H = -p_1 \log p_1 - (1 - p_1) \log(1 - p_1)$$

$$H' = \log(1 - p_1) - \log p_1$$

To find a maximum/minimum (stable point of H): H' = 0

$$0 = -\log p_1 + \log(1 - p_1)$$

$$\log p_1 = \log(1 - p_1)$$

$$p_1 = 1 - p_1$$

$$p_1 = \frac{1}{2}$$

$$P_2 = 1 - \frac{1}{2} = \frac{1}{2}$$

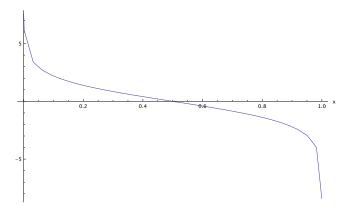


Figure 1.1: H' function representation

To describe wheter the point $p_1 = \frac{1}{2}$ is a maximum for H, we take the second derivative: H'' < 0:

$$H'' = -\frac{1}{p_1 \ln 10} - \frac{1}{(1 - p_1) \ln 10} = -\frac{1}{\ln 10} \left(\frac{1}{p_1} + \frac{1}{1 - p_1} \right)$$

at
$$p_1 = \frac{1}{2}$$
: $H'' = -\frac{1}{\ln 10} \left(\frac{1}{\frac{1}{2}} + \frac{1}{\frac{1}{2}} \right) = -\frac{1}{\ln 10} * 4 = -4 \ln 10 < 0$

Hence, for n=2, the uniform distribution maximizes the Shannon index.

2.1 Epidemiological dynamics and basic reproductive ratio

2.1.1 a

$$x(0) = x_0$$

$$y(0) = y_0$$

$$z(0) = 0$$

$$x' = bN - \beta xy - bx$$

$$y' = \beta xy - (b+r)y$$

$$z' = ry - bZ$$

$$\begin{split} \frac{dZ(t)}{dt} &= ry(t) - bZ(t) \stackrel{t=0}{\Longrightarrow} \frac{dZ(0)}{dt} = ry(0) - bZ(0) \\ \frac{d}{dt}(0) &= ry_0 - b(0) \\ 0 &= ry_0 \to \boxed{y_0 = 0} \end{split}$$

$$\frac{dy(t)}{dt} = \beta x(t) - (b+r)y(t) \stackrel{t=0}{\Longrightarrow} \frac{dy(0)}{dt} = \beta x(0)y(0) - (b+r)y(0)$$

$$\frac{d}{dt}(0) = \beta x_0 y_0 - (b+r)y_0$$

$$x_0 = \frac{b+r}{\beta}$$

$$\frac{dx(t)}{dt} = bN - \beta x(t)y - bx(t) \stackrel{t=0}{\Longrightarrow} \frac{dX(0)}{dt} = bN - \beta x(0)y(0) - bx(0)$$
$$0 = bN - bx_0$$
$$N = x_0 = \frac{b+r}{\beta}$$

If $N > N_C = \frac{b+r}{\beta}$, y can grow and sequentially Z can grow as well. If y = 0, then x converges to the uninfected equilibrium N, while y can only grow if x exceeds the threshold $\frac{b+r}{\beta}$. Hence, the pathogen carrying the disease can only invade its host if this threshold is smaller than N.

$$\frac{b+r}{\beta} < N \to \frac{N}{\frac{b+r}{\beta}} > 1 \to N \frac{\beta}{b+r} > 1$$

$$R_0 = N \frac{\beta}{b+r} \to \begin{cases} R_0 < 1 \to \text{ the parasite cannot spread} \\ R_0 > 1 \to \text{ the parasite can spread and invade the host} \end{cases}$$

2.1.2 b

$$f(x) = \frac{dx}{dt} = bN - \beta xy - bx$$

$$\Rightarrow bN - \beta x^* y^* - bx^* = 0$$

$$g(y) = \frac{dy}{dt} = \beta xy - (b+r)y$$

$$\Rightarrow \beta x^* y^* - (b+r)y^* \to x^* = \frac{b+r}{\beta}$$

$$h(Z) = \frac{dZ}{dt} = ry - bZ$$

$$\Rightarrow ry^* - bZ^* = 0 \to Z^* = \frac{ry^*}{b}$$

$$x^* = \frac{b+r}{\beta} = \frac{2+4}{3} = 2 \to \boxed{x^* = 2}$$

$$bN - \beta x^* y^* - bx^* = 0 \to 2000 - 6y^* - 4 = 0 \to y^* = \frac{1996}{6} \to \boxed{y^* = \frac{998}{3} \approx 332.66}$$

$$Z^* = \frac{r}{b}y^* = 2y^* = \frac{1996}{3} \to \boxed{Z^* = \frac{1996}{3} \approx 665.33333333}$$

To find stability of equilibrium |f'(x)| < 1:

$$\frac{df(x)}{dx} = -b - \beta y \qquad \rightarrow 998 - 2 = -1000 \qquad \rightarrow |f'(x)| = 1000 > 1 \qquad \rightarrow x^* \text{ is not stable}$$

$$\frac{dg(y)}{dy} = \beta x - (b+r) \qquad \rightarrow g'(y) = 6 - 6 = 0 \qquad \rightarrow |g'(y)| < 1 \qquad \rightarrow y^* \text{ is stable}$$

$$\frac{dh(Z)}{dZ} = -b \qquad \rightarrow h'(Z) = -2 \qquad \rightarrow |h'(Z)| > 1 \qquad \rightarrow Z^* \text{ is not stable}$$

2.1.3 c

x-null clines: 0 = 2000 - 3xy - 2xy-null clines: 0 = 3xy - 6y

Solutions:

$$x_1 = 2, y_1 = \frac{997}{3}$$

 $x_2 = \frac{2000}{3}, y_2 = 0$

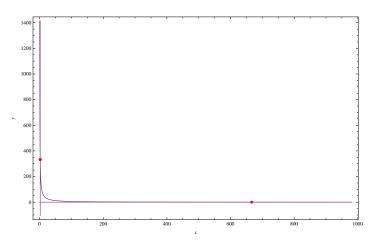


Figure 2.2: Null clines graph

3.1 Random walk

3.1.1 a

$$\begin{split} E[X(t)|X(0) &= i] = E[X(0) + \sum_{s=1}^{t} \Delta(t)] \\ &= E[i] + E\left[\sum_{s=1}^{t} \Delta(s)\right] \\ &= i + E\left[\sum_{s=1}^{t} \left[1 * \frac{a}{2} + 0 * (1 - a) - 1 * \frac{a}{2}\right]\right] \\ E[X(t)|X(0) &= i] = i + sE(\Delta(1)) = i + s * 0 = i \end{split}$$

3.1.2 b

$$\begin{split} Var[X(t)|X(0) = i] &= atVar(X(t)) = Var(x(0) + \Delta(1) + \dots + \Delta(t)) \\ &= Var(x(0) + Var(\Delta(1)) + \dots + Var(\Delta(t))) \\ &= Var(X(0)) + t * Var(\Delta(1)) = \\ &\text{as } Var(\Delta(1)) = (1 - 0)^2 * \frac{a}{2} + (1 - 0)^2 * \frac{a}{2} + 0^2(1 - a) = a \\ &Var(X(t)|X(0) = i) = Var(i) + t * Var(\Delta(1)) = 0 + t * a = at \end{split}$$

4.1 Neutral Moran process

4.1.1 a

Transition matrix of Neural Moran process is (assume p = i/N):

$$\begin{split} P_{0,0} &= 1 \\ P_{i,i-1} &= (\frac{N-1}{i})(i/N) = (1-p)(p) \\ P_{i,i+1} &= (i/N)(\frac{N-1}{i}) = (p)(1-p) \\ P_{i,i} &= 1 - P_{i,i-1} - P_{i,i+1} = 1 - (1-p)p - p(1-p) = 2p^2 - 2p + 1 = p^2 + (1-p)^2 \end{split}$$

From *law of total expectation:*

$$\begin{split} E[X(t)|X(t-1) &= i] = (i-1)P_{i,i-1} + iP_{i,i} + (i+1)P_{i,i+1} \\ &= (i-1)(1-p)(p) + i(p^2 + (1-p)^2) + (i+1)(p)(1-p) \\ &= (i-1+i+1)p(1-p) + i(p^2 + (1-p)^2) \\ &= 2ip(1-p) + i(p^2 + (1-p)^2) \\ &= 2ip - 2ip^2 + ip^2 - 2ip + i = i \end{split}$$

Then, by doing repeatedly

$$E_Y[Y] = E_Z[E_Y[Y|Z]] = E[Z] \Rightarrow E[X(t)] = E[X(0)]$$
 (4.1)

So, E[X(t)|X(0) = i] = i

4.1.2 b

$$V_t = Var(X(t)|X(0) = i)$$
(4.2)

Therefore,

$$V_1 = Var(X(1)|X(0) = i) = E[X(1)^2|X(0) = i] - E[X(1)|X(0) = i]^2$$

= $(i-1)^2 p(1-p) + i^2 (p^2 + (1-p)^2) + (i+1)^2 p(1-p) - i^2$
= $2p(1-p) = 2(i/N)(1-i/N)$

Var(X(t)|X(t-1)=i)=Var(X(1)|X(0)=i), since (X(t)|X(t-1)=i) and (X(1)|X(0)=i) are identically distributed.

From law of total variance:

$$Var[Y] = E[Y^2] - E[Y]^2 = E[E[Y^2|X]] - E[E[Y|X]]^2$$

$$= E[Var[Y|X] + E[Y|X]^2] - E[E[Y|X]]^2$$

$$= E[Var[Y|X]] + (E[E[Y|X]^2] - E[E[Y|X]]^2)$$

$$= E[Var[Y|X]] + Var[E[Y|X]]$$

Hence,

$$Var[Y] = E[Var[Y|Z]] + Var[E[Y|Z]]$$

$$= E[(2Z/N)(1 - Z/N)] + Var(Z)$$

$$= (2E[Z]/N)(1 - E[Z]/N) + (1 - 2N^2)Var(Z)$$

So, we have Y = (X(t)|X(0) = i) and Z = (X(t-1)|X(0) = i)

$$\begin{split} Var[Y] &= (2E[Z]/N)(1-E[Z]/N) + (1-2N^2)Var(Z) \Rightarrow \\ Var[X(t)|X(0) &= i] &= (2E[X(t-1)]/N)(1-E[X(t-1)]/N) + (1-2N^2)Var(X(t-1)|X(0) &= i) \\ &= (2i/N)(1-i/N) + (1-2N^2)Var(X(t-1)|X(0) &= i) \Rightarrow \\ Var[X(t)|X(0) &= i] &= V_1 + (1-2N^2)Var(X(t-1)|X(0) &= i) \Rightarrow \\ V_t &= V_1 + (1-2N^2)V_{t-1} \end{split}$$

Finally (X(0) = i),

$$V_t - \frac{V_1}{2/N^2} = (1 - 2/N^2)(V_{t-1} - \frac{V_1}{2/N^2}) = (1 - 2/N^2)^{t-1}(V_1 - \frac{V_1}{2/N^2}) \Rightarrow$$

$$V_t = V_1 \frac{1 - (1 - 2/N^2)^t}{2/N^2}$$

4.1.3 c

By definition $V_1 = 2P_{i,i-1} = 2P_{i,i+1}$, in terms of problem 3 $V_1 = 2 * \frac{q}{2} = q$

$$\lim_{N \to \inf} \frac{1 - 1(1 - 2/N^2)^t}{2/N^2}$$

Using l'Hopital's rule to calculate the limit:

$$\lim_{N \to \inf} \frac{1 - 1(1 - 2/N^2)^t}{2/N^2} = \lim_{N \to \inf} \frac{-(t(1 - 2/N^2)^{t-1} * 4/N^3)}{-4/N^3} = t(1 - 2/N^2)^{t-1}$$

This gives us:

$$\lim_{N \to \inf} \frac{1 - 1(1 - 2/N^2)^t}{2/N^2} = t$$

Therefore, for large N , we have that Var(X(t)|X(0)=1)=at, which is our result in problem 3 part b.

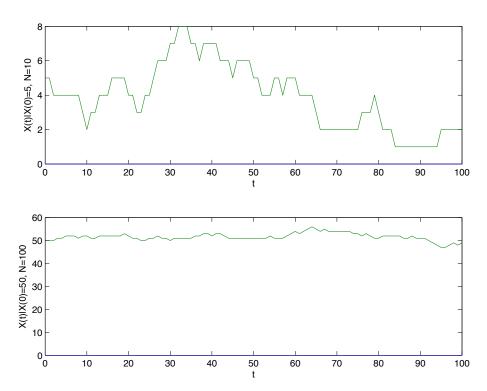


Figure 4.3: Neutral Moran processes, with N=10 and N=100

4.1.4 d

```
1 % Moran for N=10

2 m = 100;

3 n = 10;

4 i = 5;

5 t = 100;

6

7 % Initialization

8 Y_10 = zeros(m,t+

9 % Repetition of m
    % Initialization of the trajectories vector
Y_10 = zeros(m,t+1);
% Repetition of moran process for m times
10
     for o=1:m
          Y_10(m,:) = moran(n,t,i);
     end
12
13
14
15
    % Moran for N=100
    m = 100;
n = 100;
i = 50;
t = 100;
16
17
18
19
20
21
22
    % Initialization of the trajectories vector
Y_100 = zeros(m,t+1);
23
     \mbox{\ensuremath{\mbox{\%}}} Repetition of moran process for m times
24
25
26
27
    Y_100(m,:) = moran(n,t,i);
28
29
     % Plotting
30
     subplot(2,1,1);
31
32
    plot(0:100, Y_10)
xlabel('t');
ylabel('X(t)|X(0)=5, \( \)N=10');
33
34
35
     subplot(2,1,2);
36 plot(0:100, Y_100)
```

```
37 | xlabel('t');

38 | ylabel('X(t)|X(0)=50, N=100');

39 | print('-depsc2', 'plot_moran');
```

Listing 4.1: ./script.m

```
function [ X ] = moran( n, t, i )
   \mbox{\ensuremath{\mbox{MORAN}}} Generation of the Moran vector containing values X(t) from t=0 to
   t=t. Possible states 0,1,\ldots,n and initial i.
   % Initialization of the vector X
   X = zeros(1, t+1);
   % Initial state i
   X(1) = i;
10
   \% Generation of the choices per each moment t. Rand() istantiates a mxn
11
   % matrix, each cell with a random number following the uniform distribution
   % in the interval 0-1
14
15
   cl = round(1+(n-1)*rand(1, t+1));
   cd = round(1+(n-1)*rand(1, t+1));
16
17
   % Computation of the cells in the vector X according to the cl and cd
18
19
20
   for t=2:t+1
       21
22
23
           % Assing the previous position to the current position
24
           X(t) = X(t-1);
25
26
       \begin{array}{ll} \textbf{elseif} & (\texttt{cl(t)} \iff \texttt{X(t-1)} \&\& \ \textbf{cd(t)} \ > \ \texttt{X(t-1)}) \end{array}
27
28
           % Assing to the current position the previous position incremented
29
           X(t) = X(t-1)+1;
30
31
       elseif (cl(t) > X(t-1) && cd(t) <= X(t-1))
32
33
34
35
36
           \mbox{\ensuremath{\%}} Assing to the current position the previous position decremented
           X(t) = X(t-1)-1;
       end
37
   end
38
   end
```

Listing 4.2: ./moran.m

5.1 Absorption in a birth-death process

5.1.1 a

i) \vec{x} is the probability of ending up in state N for $x_i \forall i = 1 ... N$ $t \to \inf$, $P_{ij} \to x_i \to x_i$ is unique stationary distribution.

The absorption probabilities are given by the right-hand eigenvector associated with the largest eigen value, which is one, because p is stochastich matrix.

 $\vec{x}P = \vec{x}$

$$x = (x_0, \dots, x_n)^t$$
$$y = x_i - x_{i-1}$$
$$\gamma = \frac{\beta_i}{\alpha_i}$$

$$p = \begin{bmatrix} 1 & 0 & 0 & \cdots & 0 & 0 & 0 \\ \beta_1 & 1 - \alpha_1 - \beta_1 & \alpha_1 & \cdots & 0 & 0 & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots & \vdots & \vdots \\ 0 & 0 & 0 & \cdots & \beta_{N-1} & 1 - \alpha_{N-1} - \beta_{N-1} & \alpha_{N-1} \\ 0 & 0 & 0 & \cdots & 0 & 0 & 1 \end{bmatrix}$$

$$x_0 = 0$$

$$x_i = p_{i,i-1}x_{i-1} + p_{i,i}x_i + p_{i,i+1}x_{i+1} = \beta_i x_{i-1} (1 - \alpha_i - \beta_i) x_i + \alpha_i x_{i+1}$$

$$x_N = 1$$

$$x_{i} = \beta_{i}x_{i-1}(1 - \alpha_{i} - \beta_{i})x_{i} + \alpha_{i}x_{i+1}$$

$$x_{i-1} = \frac{(\alpha_{i} + \beta_{i})x_{i} - \alpha_{i}x_{i+1}}{\beta_{i}}$$

$$\to y_{i} = x_{i} - x_{i-1} = \beta_{i}x_{i-1} + (1 - \alpha_{i} - \beta_{i})x_{i} + \alpha_{i}x_{i+1} - \frac{(\alpha_{i} + \beta_{i})x_{i} - \alpha_{i}x_{i+1}}{\beta_{i}}$$

$$\to y_{i} = \frac{(\beta_{i} - \alpha_{i} - \beta_{i})x_{i} + \alpha_{i}x_{i+1}}{\beta_{i}}$$

$$\to y_{i} = \frac{\alpha_{i}x_{i+1} - \alpha_{i}x_{i}}{\beta_{i}} = \frac{\alpha_{i}}{\beta_{i}}(x_{i+1} - x_{i}) = \frac{\alpha_{i}}{\beta_{i}}y_{i+1} = \frac{1}{\gamma_{i}}y_{i+1}$$

$$\to y_{i+1} = \gamma_{i}y_{i}$$

ii) Show that $\sum_{i=1}^{0} y_i = x_l$

$$\sum_{i=1}^{l} y_i = x_1 - x_0 + x_2 - x_1 + \dots + x_l - x_{l-1} = x_l - x_0 = x_l \to \boxed{\sum_{i=1}^{l} y_i = x_l}$$

We can conclude $\sum_{i=1}^{N} y_i = x_n = 1$

iii) Show that $x_l = (1 + \sum_{j=1}^{l-1} \prod_{k=1}^{j} \gamma_k) x_1$

form point i) of this section: $y_1 = x_1$, $y_2 = \gamma_1 y_1 = \gamma_1 x_1$, $y_3 = \gamma_2 y_2 = \gamma_1 \gamma_2 x_1$

$$\sum_{i=1}^{l} y_i = x_1 + \gamma_1 x_1 + \gamma_1 \gamma_2 x_1 + \dots + (\gamma_1 \dots \gamma_{l-1}) x_1 = (1 + \sum_{j=1}^{l-1} \prod_{k=1}^{j} \gamma_k) x_1) = x_l$$

$$\to x_{l} = (1 + \sum_{j=1}^{l-1} \prod_{k=1}^{j} \gamma_{k}) x_{1}$$

We can conclude
$$x_N = (1 + \sum\limits_{j=1}^{l-1} \prod\limits_{k=1}^{j} \gamma_k) x_1 = 1$$

$$\rightarrow \boxed{x_1 = \frac{1}{1 + \sum\limits_{j=1}^{l-1} \prod\limits_{k=1}^{j} \gamma_K}}$$

$$x_i = x_1(1 + \sum_{j=1}^{i-1} \prod_{k=1}^{j} \gamma_K))$$

$$\Rightarrow x_{i} = \frac{1 + \sum_{j=1}^{i-1} \prod_{k=1}^{j} \gamma_{K}}{1 + \sum_{j=1}^{N-1} \prod_{k=1}^{j} \gamma_{K}}$$

5.1.2 b

$$x_{i} = \frac{1 + \sum_{j=1}^{i-1} \prod_{k=1}^{j} \gamma_{K}}{1 + \sum_{j=1}^{N-1} \prod_{k=1}^{j} \gamma_{K}}$$
$$\gamma_{i} = \frac{P_{i,i-1}}{P_{i,i+1}} = \frac{1}{r}$$

$$x_{i} = \frac{1 + \frac{1}{r} + \frac{1}{r^{2}} + \frac{1}{r^{3}} + \dots + \frac{1}{r^{i-1}}}{1 + \frac{1}{r} + \frac{1}{r^{2}} + \frac{1}{r^{3}} + \dots + \frac{1}{r^{N-1}}} = \frac{1 - \frac{1}{r^{i}}}{1 - \frac{1}{r^{N}}} \to \boxed{x_{i} = \frac{1 - \frac{1}{r^{i}}}{1 - \frac{1}{r^{N}}}}$$
$$\to \boxed{x_{1} = \frac{1 - \frac{1}{r}}{1 - \frac{1}{r}}}$$

$$\lim_{r \to 1} \frac{f(r)}{g(r)} = \lim_{r \to 1} \frac{f'(r)}{g'(r)} = \lim_{r \to 1} \frac{\frac{1}{r^2}}{\frac{Nr^{N-1}}{r^{2N}}} = \frac{1}{N} \to \boxed{\lim_{r \to 1} \frac{1 - \frac{1}{r}}{1 - \frac{1}{r^N}} = \frac{1}{N}}$$