Financial Time Series

Lecture 2

Linear time series

• A time series $\{r_t\}$ is linear if it can be written as

$$r_t = \mu + \sum_{i=0}^{\infty} \psi_i a_{t-i}$$

Here $\mu=E[r_t]$, $\psi_0=1$ and $\{a_t\}$ is white noise, i.e. independent with mean zero and constant variance $\sigma^2{}_a$

• Note that $\{r_t\}$ is (strictly) stationary by definition

Linear time series

• So for the series $\{r_t\}$ we have

•
$$Var(r_t) = \sigma^2 \sum_{i=0}^{\infty} (\psi_i)^2$$

•
$$\gamma_l = Cov(r_t, r_{t-l}) = \sigma_a^2 \sum_{i=0}^{\infty} \psi_i \psi_{i+l}$$

•
$$\rho_l = \frac{\gamma_l}{\gamma_0} = \frac{\sum_{i=0}^{\infty} \psi_i \psi_{i+l}}{\sum_{i=0}^{\infty} (\psi_i)^2}$$

Autoregressive models

 If we for an observed series find significant one lag autocorrelation it may be useful to model the series as

$$r_t = \varphi_0 + \varphi_1 r_{t-1} + a_t$$

Here $\{a_t\}$ is white noise, i.e. independent with mean zero and constant variance σ^2_a

AR(1)

• So for the AR(1) series $\{r_t\}$ as on the previous slide

•
$$Var(r_t) = \gamma_0 = \frac{\sigma^2_a}{1 - (\varphi_1)^2}$$

•
$$\gamma_l = Cov(r_t, r_{t-l}) = \varphi_1 \gamma_{l-1}, l > 0$$

•
$$\rho_l = \varphi_1 \rho_{l-1} = (\varphi_1)^l, l > 0$$

AR-processes

A zero mean AR(p)-process is defined by

$$r_t = \varphi_1 r_{t-1} + \dots + \varphi_p r_{t-p} + a_t$$

May be useful in modeling e.g., stock returns

• To be able to do this we need to estimate $\varphi_1, \dots, \varphi_p$...

MA models

The series defined by

$$r_t = \mu + a_t - \theta_1 a_{t-1} - \dots - \theta_q a_{t-q}$$

is called an MA(q) series/process.

Stationary by definition

MA(q)

It holds that

•
$$E[r_t] = \mu$$

•
$$Var(r_t) = \left(1 + \sum_{i=1}^{q} (\theta_i)^2\right) \sigma_a^2$$

•
$$\gamma_l = \{\mu = 0\} = \sigma_a^2 \sum_{k=|l|}^q (-\theta_{k-|l|}) (-\theta_k)$$

where $-\theta_0 = 1$ and $\theta_k = 0$, k > q

Reflection

• Note that the autocorrelation cuts off after lag q

• Note also that for an MA(1) we may use $a_t = r_t + \theta_1 a_{t-1}$ and write

$$a_t = r_t + \theta_1 r_{t-1} + (\theta_1)^2 r_{t-2} + \cdots$$

• In order for this to make sense it must hold that $|\theta_1| < 1$ in which case the MA(1) is invertible

Choosing AR or MA or both?

 The choice of model may be made using autocorrelation functions (ACF) and partial autocorrelation functions (PACF)

 Typically, a slowly decaying ACF and an ACF with "q peaks and then zero" indicate AR and MA(q) models, respectively

Below we will explain PACF

Estimation of AR-parameters

- Estimation may be done in several different ways, but we need to decide on what number, p, of parameters to use
- This can be done using partial autocorrelations (PACF)

$$\alpha_h = \varphi_{hh}$$

where φ_{hh} is the parameter φ_h in the AR(h) model

Intuition

• For an AR(p) process we may think of the the PACF at lag 5, say, as the correlation between X_t and X_{t+5} not explained by their common correlations with X_{t+1}, \ldots, X_{t+4}

 If the PACF at a certain lag is statistically significant the AR model of choice will have order no smaller than that lag

• For an AR(1) model, $r_t = \varphi r_{t-1} + a_t$, we know that the ACF is

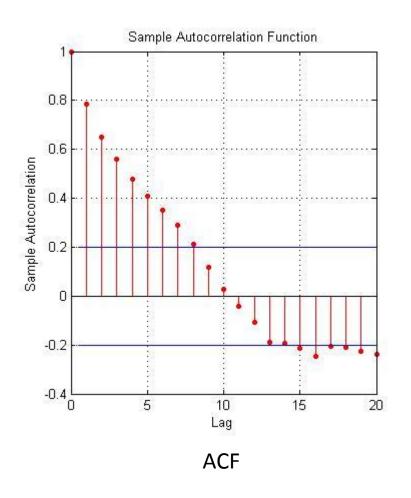
$$\rho_h = \varphi^{|h|}$$

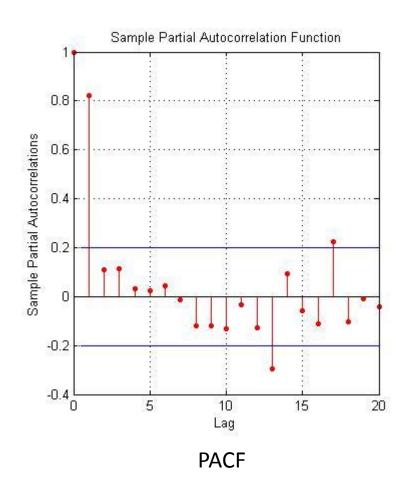
One can show that the PACF is

$$\alpha_1 = \varphi$$
 and $\alpha_h = 0$ for $h > 0$

 So the correlation for larger lags is explained by the lag one correlation

Example AR(1) with $\varphi = 0.8$





To find the φ -parameters

• For the AR(h) model we may use the Yule-Walker equations:

$$\gamma = \Gamma \varphi$$

where
$$\gamma = [\gamma_1 \cdots \gamma_h]^T$$
 and $\Gamma_{ij} = \gamma_{i-j}$

• Numerically this is done by replacing γ_h with

$$\hat{\gamma}_h = \frac{1}{n} \sum_{j=1}^{n-n} (x_{j+h} - \bar{x})(x_j - \bar{x})$$

Note the

• We need to invert the "gamma matrix" to find $\hat{\varphi}$

• Choosing 1/n instead of the "intuitive" 1/(n-h) in the sample autocovariance will guarantee invertibility

To determine the order p

We need to decide up to which order the PACF is statistically significant

• At 95% confidence we want our estimate of φ_{hh} within $\pm 1.96/\sqrt{T}$ where T is the no. of observations

In matlab you can use the "parcorr" function

Checking the model

- We want residuals to behave like WN
- So, for our fitted model

$$\hat{r}_t = \hat{\varphi}_1 r_{t-1} + \dots + \hat{\varphi}_p r_{t-p}$$

we create the residuals

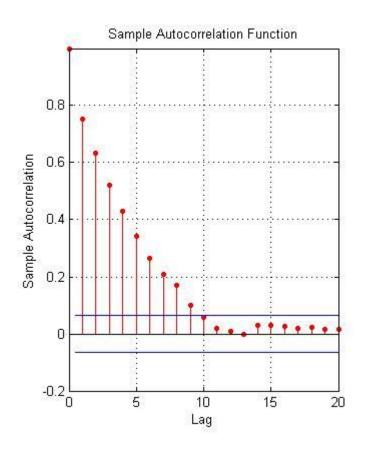
$$e_t = \hat{r}_t - r_t$$

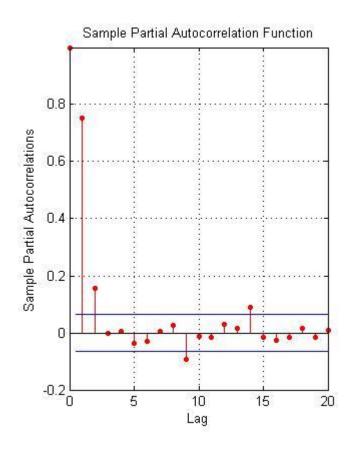
Checking the model

Apply graphical, Ljung-Box or Portmeanteau procedures

• From the residuals we also get an estimate of the WN variance: $\hat{\sigma}^2{}_a = \frac{1}{T-p} \sum_{i=p+1}^T e_i{}^2$

• Given some time series, we find





- Based on the plots it seems reasonable to fit an AR(2)-process
- To do this, using the Yule-Walker equations, we need the sample covariances at lags 0,1,2
- Using the cov command in matlab we find the elements of the estimated covariance matrix, $\hat{\Gamma}$ and the estimated covariance vector $\hat{\gamma}$

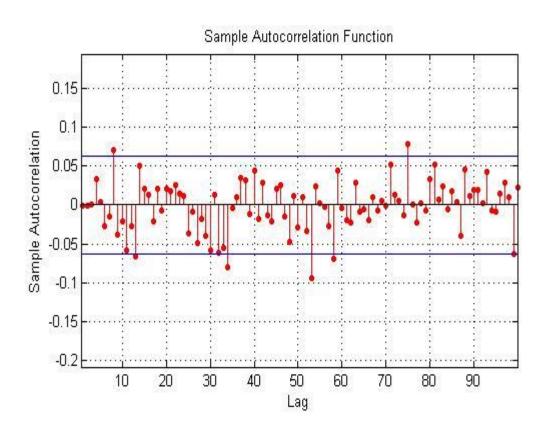
- We have $\hat{\Gamma}_{11} = \hat{\Gamma}_{22} = \hat{\gamma}_0 = 2.3456$, $\hat{\Gamma}_{12} = \hat{\Gamma}_{21} = \hat{\gamma}_1 = 1.7667$ and $\hat{\gamma}_2 = 1.4882$
- We may invert $\hat{\Gamma}$ using the inv command

• Finally, we find $\hat{\varphi} = \hat{\Gamma}^{-1} \hat{\gamma} = [0.6362 \ 0.1553]^T$

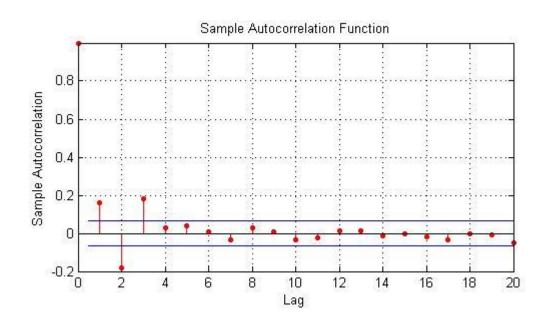
So our estimated series, i.e., predictions are

$$\hat{r}_t = 0.6362r_{t-1} + 0.1553r_{t-2}$$

• If the fit is good, we will see uncorrelated residuals $e_t = r_t - \hat{r}_t$



Identifying MA



Identifying MA

Above, we see the ACF for an MA(3)

 Generally, we may use the ACF and decide on using MA if the ACF is significant up to a certain lag

The last lag at which the ACF is significant is q

Reflection

 To identify the number of parameters to use in AR estimation we use the PACF

 To identify the number of parameters to use in MA estimation we use the ACF

Estimating MA parameters

 In the example above, we use the ACF to decide that an MA(3) seems reasonable

• To use this for predictions, we need to estimate θ_1 , θ_2 and θ_3

 This can be done using the auto covariance function, ML or "The innovations algorithm" or...

Estimation using the autocovariance

 We know that for an MA(q) process the autocovarince function is

$$\gamma_h = \sigma^2_a \sum_{k=|h|}^q \left(-\theta_{k-|h|}\right) \left(-\theta_k\right)$$

Estimation using the autocovariance

So for the MA(3), we have

$$\gamma_{0} = \sigma_{a}^{2} (1 + \theta_{1}^{2} + \theta_{2}^{2} + \theta_{3}^{2})$$

$$\gamma_{1} = \sigma_{a}^{2} (-\theta_{1} + \theta_{1}\theta_{2} + \theta_{2}\theta_{3})$$

$$\gamma_{2} = \sigma_{a}^{2} (-\theta_{2} + \theta_{1}\theta_{3})$$

$$\gamma_{3} = -\sigma_{a}^{2} \theta_{3}$$

To solve these equations

• We may use "fsolve" in matlab replacing γ by $\widehat{\gamma}$

• Using 1000 observations of the MA(3) process above (which was in fact $r_t=a_t-(-0.6)a_{t-1}-0.4a_{t-2}-(-0.2)a_{t-3}$ and $a_t \sim N(0,1)$) we find

$$\hat{\theta}_1 = -1.0603, \hat{\theta}_2 = 0.3520, \hat{\theta}_3 = -0.0982,$$

$$\hat{\sigma}^2 = 0.6620$$

But, using 10000 or 100000 obs...

• For 10000 obs, we get

$$\hat{\theta}_1 = -0.6287, \hat{\theta}_2 = 0.4133, \hat{\theta}_3 = -0.2174,$$

$$\hat{\sigma}^2 = 0.9655$$

And for 100000

$$\hat{\theta}_1 = -0.6048, \hat{\theta}_2 = 0.4058, \hat{\theta}_3 = -0.1951,$$

$$\hat{\sigma}^2 = 0.9971$$

Innovations algorithm

• Recursive procedure which also gives us an estimate of the WN variance σ^2_a

 Uses the sample covariances at lags up to the same order as the process we want to fit

 Gives "rough estimates" that may be used as starting values for refined methods, e.g., ML or least squares

Innovations algoritm (in general)

- Start with $\hat{v}_0 = \hat{\gamma}(0)$
- For $m=1,\ldots,n-1$ and $k=0,\ldots,m-1$, proceed using

$$\hat{\theta}_{m,m-k} = \frac{1}{\hat{v}_k} \left[\hat{\gamma}(m-k) - \sum_{j=0}^{k-1} \hat{\theta}_{m,m-j} \hat{\theta}_{k,k-j} \hat{v}_j \right]$$

$$\widehat{v}_m = \widehat{\gamma}(0) - \sum_{j=0}^{m-1} \widehat{\theta}^2_{m,m-j} \, \widehat{v}_j$$

Innovations algoritm

• The estimate of the WN variance is \widehat{v}_m

Innovations algorithm

• So, to fit an AR(3), we need $\hat{\gamma}(0)$, $\hat{\gamma}(1)$, ...

 Using the cov command in matlab we use the covariance matrix with covariances for 99 lags

Innovations algorithm

• So
$$\hat{v}_0 = \hat{\gamma}(0) = 1.4337$$

•
$$\hat{\theta}_{11} = \frac{\hat{\gamma}(1)}{\hat{\gamma}(0)} = 0.1588$$

•
$$\hat{v}_1 = \hat{\gamma}(0) - \hat{\theta}^2_{11} \hat{v}_0 = 1.3975$$

•
$$\hat{\theta}_{22} = \frac{\hat{\gamma}(2)}{\hat{\gamma}(0)} = -0.1811$$

•
$$\hat{\theta}_{21} = \frac{1}{\hat{v}_1} [\hat{\gamma}(1) - \hat{\theta}_{22} \hat{\theta}_{11} \hat{v}_0] = 0.1924$$

•
$$\hat{v}_2 = \hat{\gamma}(0) - \hat{\theta}^2_{22}\hat{v}_0 - \hat{\theta}^2_{21}\hat{v}_1 = 1.3349$$

Innovations algorithm

•
$$\hat{\theta}_{33} = \frac{\hat{\gamma}(3)}{\hat{\gamma}(0)} = 0.1802$$

•
$$\hat{\theta}_{32} = \frac{1}{\hat{v}_1} [\hat{\gamma}(2) - \hat{\theta}_{33} \hat{\theta}_{11} \hat{v}_0] = -0.2151$$

•
$$\hat{\theta}_{31} = \frac{1}{\hat{v}_2} [\hat{\gamma}(1) - \hat{\theta}_{33} \hat{\theta}_{22} \hat{v}_0 - \hat{\theta}_{32} \hat{\theta}_{21} \hat{v}_1] = 0.2490$$

•
$$\hat{v}_3 = \hat{\gamma}(0) - \hat{\theta}^2_{33}\hat{v}_0 - \hat{\theta}^2_{32}\hat{v}_1 - \hat{\theta}^2_{31}\hat{v}_2 = 1.2397$$

Innovations algorithm

 Proceeding in 99 steps (until estimates "do not fluctuate too much") we find

$$\hat{ heta}_{99,1}=0.4216,\,\hat{ heta}_{99,2}=-0.2902,\,\hat{ heta}_{99,3}=0.1885$$
 and $\hat{v}_{99}=1.2290$

Predictions

 For the AR(3) example, with the estimates from the autocovariance procedure based on 10000 obs

$$\hat{r}_t = 0.6287a_{t-1} - 0.4133a_{t-2} + 0.2174a_{t-3}$$

How do we find (estimate) the WN components?

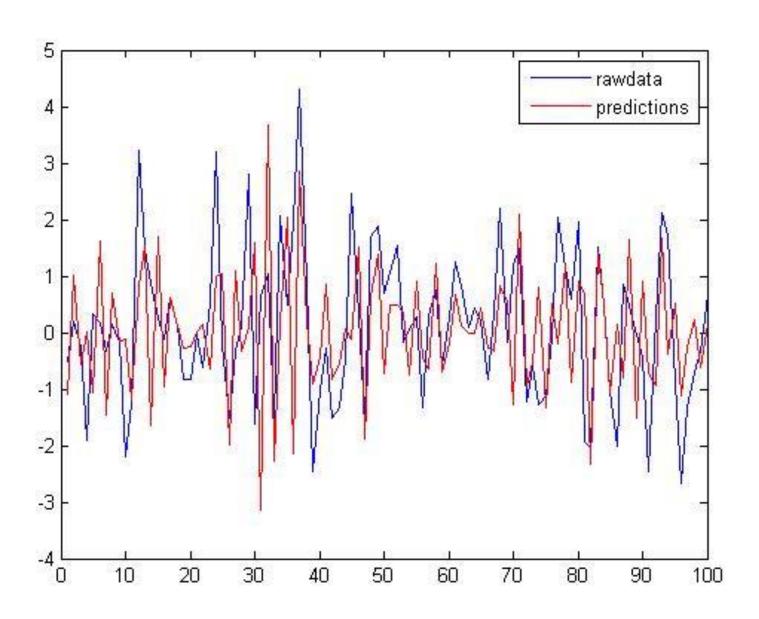
Predictions

We may use

$$a_1 = r_1 a_t = r_t - 0.6287 a_{t-1}$$

Below we find a plot of predictions and observed values

Predictions



Alternative prediction method

Generally, we are looking for linear predictors

$$\hat{r}_t = \alpha_0 + \sum_{i=1}^n \alpha_i r_{t-i}$$

that minimize

$$E[(\hat{r}_t - r_t)^2]$$

Alternative prediction method

 The optimal "alphas" may be found by numerically minimizing (fminsearch in matlab)

$$\sum_{k=0}^{q+n} \left[\sum_{l=\max\{0,k-q\}}^{\min\{k,n\}} \alpha_l \hat{\theta}_{k-l} \right]^2$$

Characteristic roots and business cycles

 The roots of the characteristic polynomial of AR(p) model:

$$\varphi(z) = 1 - \varphi_1 z - \dots - \varphi_p z^p$$

can be used to describe long-term behvaiour of the series

Example

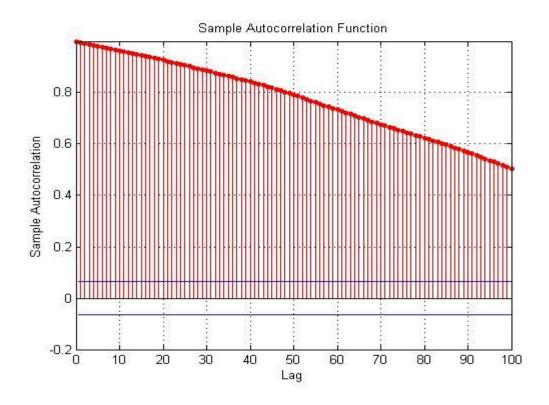
• For an AR(2), we may solve the equation

$$1 - \varphi_1 x - \varphi_2 x^2 = 0$$

- If the roots are real the ACF will behave like a mixture of two exponential decays
- If the roots are complex numbers the ACF will behave like damping sine and cosine waves

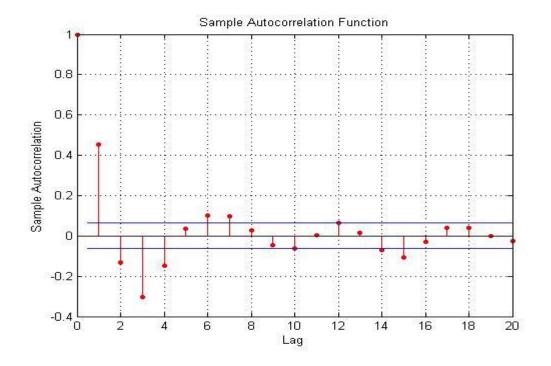
Example

• For $r_t = 0.6r_{t-1} + 0.4r_{t-2} + a_t$, "the roots" are 1 and 2.5



Example

• For $r_t=0.6r_{t-1}-0.4r_{t-2}+a_t$ "the roots" are $\frac{3}{4}\pm\frac{\sqrt{31}}{4}i$



Business cycles

 The complex roots indicate business cycles for which the average lengths, k, are given by cycles

$$k = \frac{2\pi}{\arccos(\varphi_1/(2\sqrt{-\varphi_2}))}$$