

$E[\hat{\theta}]^2 = M\left[\left(\frac{2}{3\theta}\right)\left(-\frac{x}{\theta}\right) - \frac{1}{\theta^2}\right]$   
 $\frac{1}{\theta^2} = \frac{2}{\theta^2} - \frac{2}{\theta} + \frac{1}{\theta^2} = \theta^2$   
 ab]  $\Rightarrow$  пуском. no gar-ye

$\int_0^{+\infty} \frac{e^{-x/\theta}}{\theta} dx = 1$

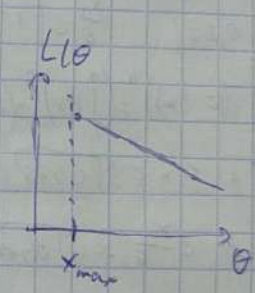
$\hat{\theta}: \frac{\theta^2}{3} \geq \frac{\theta^2}{3}$

no gar-ye  $\hat{\theta}_1$  app,  $\theta$   $D[\hat{g}] = \frac{g'(\theta)}{h I(\theta)}$   
 no t. f!  $\rightarrow \hat{\theta}$  app  $\Rightarrow \hat{\theta}_1$  re app.

$(TS) \mathcal{F} \sim R[\theta, 2\theta]$

$p(x) = \frac{1}{\theta} \{0 \leq x \leq 2\theta\} + 0 \{x \in \mathbb{R} \setminus [0, 2\theta]\}$

DMN:  
 $I_1 = \int_0^{2\theta} \frac{x}{\theta} dx = \frac{x^2}{2\theta} \Big|_0^{2\theta} = 2\theta - \frac{1}{2}\theta = \frac{3}{2}\theta$   
 $I_2 = \int_0^{2\theta} \frac{x^2}{\theta} dx = \frac{x^3}{3\theta} \Big|_0^{2\theta} = \frac{8}{3}\theta^2 - \frac{1}{3}\theta^2 = \frac{7}{3}\theta^2$   
 $D[\mathcal{F}] = \frac{7}{3}\theta^2 - \frac{9}{4}\theta^2 = \frac{1}{12}\theta^2$



$\tilde{x}_1 = \bar{x}$   
 $\tilde{\theta}_1 = \frac{2}{3} \bar{x}$   
 $\frac{2}{3} \tilde{\theta}_1 = \bar{x}$

$\hat{\theta}_1: M[\tilde{\theta}_1] = \frac{2}{3} M[\bar{x}] = \frac{2}{3n} \sum_{i=1}^n M[x_i] = \theta$

$D[\tilde{\theta}_1] = D\left[\frac{2}{3} \bar{x}\right] = \frac{4 \cdot \eta}{9n^2} \cdot \frac{1}{12} \theta^2 = \frac{\theta^2}{12n} \rightarrow 0, n \rightarrow \infty \Rightarrow \text{est. no gar-ye.}$

DMN:  $p(x, \theta) = \begin{cases} \frac{1}{\theta} & x \in [0, 2\theta] \\ 0 & \text{elsewhere} \end{cases}$



$$X_{\max}: p(t) = \frac{h}{\theta} \left( \frac{t}{\theta} - \frac{\theta}{\theta} \right)^{n+1} \cdot \mathbb{I}_{t \in [0, 2\theta]} + 0 \cdot \mathbb{I}_{t \in \mathbb{R} \setminus [0, 2\theta]}$$

$$\begin{aligned} \tilde{\theta}_2: M[\tilde{\theta}_2] &= \frac{1}{2} M[X_{\max}] + \frac{1}{2\theta^n} \int_0^{2\theta} (t-\theta)^{n+1} dt = \left\{ \begin{array}{l} k=t-\theta \\ t=k+\theta \end{array} \right\} = \\ &= \frac{n}{2\theta^n} \int_0^\theta k^{n+1} (k+\theta) dk = \frac{n}{2\theta^n(n+1)} k^{n+1} \Big|_0^\theta + \frac{h}{2(h+1)\theta^{n+1}} k^n \Big|_0^\theta = \frac{\theta n}{2(n+1)} + \frac{\theta}{2} = \frac{\theta(2n+1)}{2(n+1)} \\ \tilde{\theta}_2' &= \frac{2(n+1)}{2n+1} \tilde{\theta}_2 - \text{несмысл} \end{aligned}$$

$$D[\tilde{\theta}_2'] = \frac{n(n+1)}{(2n+1)^2} D[\tilde{\theta}_2] = \frac{(n+1)^2}{(2n+1)^2} D[X_{\max}]$$

$$M[X_{\max}] = \frac{\theta(2n+1)}{n+1}$$

$$\begin{aligned} M[X_{\max}^2] &= \frac{n}{\theta^n} \int_0^\theta k^{n+1} (k^2 + 2k\theta + \theta^2) dk = \frac{n}{\theta^n} \left[ \frac{\theta^{n+2}}{n+2} + \frac{2\theta^{n+2}}{n+1} + \frac{\theta^{n+2}}{n} \right] = \\ &= n \left[ \frac{\theta^2(n^2+1)}{n^2+3n+2} + 2\theta^2 \frac{(n^2+2n)}{n^2+3n+2} + \theta^2 \frac{(n^2+3n+2)}{n^2+3n+2} \right] = \theta^2 \left( \frac{4n^2+8n+2}{n^2+3n+2} \right) \end{aligned}$$

$$D[X_{\max}] = \theta^2 \left( \frac{4n^2+8n+2}{n^2+3n+2} - \frac{4n^2+4n+1}{n^2+2n+1} \right) = \theta^2 \left( \frac{n}{n^2+n^2(n+1)} \right)$$

$$D[\tilde{\theta}_2'] = \frac{(n+1)^2}{(2n+1)^2} \frac{n}{(n+1)^2(n+2)} \theta^2 \rightarrow 0, \quad n \rightarrow \infty \quad \text{convergence}$$

экспериментально:

$$\forall \theta \in \mathbb{R}^+ \hookrightarrow D[\tilde{\theta}_1] > D[\tilde{\theta}_2]$$

$$\frac{\theta^2}{27n} > \frac{n}{(2n+1)^2(n+2)} \text{ при } n \geq 4$$

$$x: \mathbb{R}^+ [0, 2\theta] \quad \frac{x_1}{\theta} \in [1, 2]$$

$$p(x_{\max}) = (F(x))' = \left( \int_1^x dx \right)' = (x-1)'$$

$$x = \theta$$

$$\frac{\sqrt{n} \cdot \frac{2}{3} \sqrt{\hat{\alpha}_2 - \hat{\alpha}_1^2} - \frac{2}{3} \sqrt{\hat{\alpha}_2 - \hat{\alpha}_1^2}}{\frac{2}{3} \sqrt{\hat{\alpha}_2 - \hat{\alpha}_1^2}} =$$

$$\sqrt{0.025} + 1 < x < \sqrt{0.975} + 1$$

$$\sqrt{0.025} + 1 \leq \frac{x_{\max}}{\theta} \leq \sqrt{0.975} + 1$$

$$\frac{x_{\max}}{\sqrt{0.975} + 1} < \theta < \frac{x_{\max}}{\sqrt{0.025} + 1}$$

$$\text{по ОММ: } \sqrt{n} \frac{g(2) - g(1)}{\alpha(x)} \sim N(0, 1)$$

$$-1.96 < \sqrt{n} \frac{\tilde{\theta} - \theta}{\frac{2}{3} \sqrt{\hat{\alpha}_2 - \hat{\alpha}_1^2}} < 1.96$$

$$\sigma'(x) = \sqrt{\sigma^2 g(x) g'(x)} = \frac{2}{3} \sqrt{\hat{\alpha}_2 - \hat{\alpha}_1^2}$$

$$- \frac{1.96}{\sqrt{n}} \cdot \frac{2}{3} \sqrt{\hat{\alpha}_2 - \hat{\alpha}_1^2} + \tilde{\theta} < \tilde{\theta} < \frac{1.96}{\sqrt{n}} \cdot \frac{2}{3} \sqrt{\hat{\alpha}_2 - \hat{\alpha}_1^2} + \tilde{\theta}$$



2, 20}

(14)  $g \in (0, N)$

$$f \sim \frac{\theta}{2} \{(-1, 1) \setminus \{0\}\} + \frac{1-\theta}{2} \{0\} + \frac{1-\theta}{2} \{2\}$$

$$\frac{\theta}{2} = \frac{\theta(2n+1)}{2(n+1)}$$

$$\mu_1 = \mu[f] = \int_{-\infty}^{\infty} x p(x, \theta) dx = \int_{-1}^1 \frac{\theta}{2} x dx + \frac{1-\theta}{2} \cdot 0 + \frac{1-\theta}{2} \cdot 2 = 1-\theta$$

$$\mu_2 = \mu[f^2] = \frac{\theta}{2} \int_{-1}^1 x^2 dx + \frac{1-\theta}{2} \cdot 2^2 = \frac{\theta}{3} + 2(1-\theta) = 2 - \frac{5}{3}\theta$$

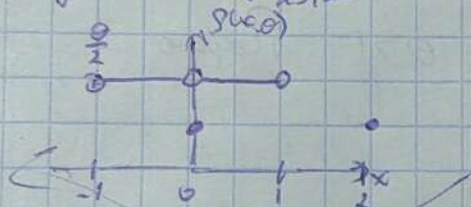
$$\mu_2 = \mu_2 - \mu^2 = D[f] = 2 - \frac{5}{3}\theta - \theta^2 + 2\theta - 1 = 1 + \frac{\theta}{3} - \theta^2$$

$$OMM: \mu_1(\theta) = \bar{\mu}_1 = \bar{x} \Rightarrow 1-\theta = \bar{x} \Rightarrow \tilde{\theta} = 1-\bar{x}$$

$$\mu[\tilde{\theta}_1] = \mu[1-\bar{x}] = 1 - \mu[f] = 1 - (1-\theta) = \theta \Rightarrow \text{recovery}$$

$$D[\tilde{\theta}_1] = D[1-\bar{x}] = \frac{1}{n} D[f] \rightarrow 0 \quad \text{con no var. yet.}$$

as OMM



$$L = \left(\frac{\theta}{2}\right)^{n-m_1-m_2} \left(\frac{1-\theta}{2}\right)^{m_1+m_2}$$

$$\ln L = (n-m_1-m_2) \ln \frac{\theta}{2} + (m_1+m_2) \ln \frac{1-\theta}{2}$$

suppose

$$(\ln L)' = \frac{n-m_1-m_2}{\theta} + \frac{m_1+m_2}{\theta-1} = \frac{n\theta - 1 + m_1 + m_2}{\theta(\theta-1)} = 0 \Rightarrow \theta = 1 - \frac{1}{n+1}$$

$$(\ln L)'' = \frac{m_1+m_2-n}{\theta^2} - \frac{m_1+m_2}{(\theta-1)^2} = \frac{(m_1+m_2-n)(\theta-1)^2 - (m_1+m_2)\theta^2}{\theta^2(\theta-1)^2} = \frac{1}{2} \frac{(n_1+n_2)(n_1+n_2-1)}{n\theta^2(\theta-1)^2} < 0$$

$$\theta = 1 - \frac{1}{n+1} \Rightarrow \tilde{\theta}_2 = 1 - \frac{1}{n+1}$$

$$\mu[\tilde{\theta}_2] = 1 - \mu[\frac{1}{n+1}] = 1 - \frac{1-\theta}{2} = \frac{1+\theta}{2} = \theta \Rightarrow \text{recovery}$$

$$D[\tilde{\theta}_2] = D[\frac{1}{n+1}] = \frac{(1-\theta)(\theta)}{n} \rightarrow 0, \text{ con}$$

c) Fisher k-Var

$$1) f(x, \theta) \in C^{\infty}([0, 1])$$

$$2) \frac{\partial}{\partial \theta} \int_{\mathbb{R}} f(x, \theta) dx = \int_{\mathbb{R}} \frac{\partial}{\partial \theta} f(x, \theta) dx \quad \text{a (0)}$$

$$\int_{-\infty}^{\infty} \frac{\partial}{\partial \theta} f(x, \theta) dx = \frac{1}{2} \int_{-1}^1 dx \left( \frac{1-\theta}{2} \right)' \theta + \left( \frac{1-\theta}{2} \right)' \theta = 1 - \frac{1}{2} - \frac{1}{2} = 0$$



3)  $I(\theta) \in C(0,1)$ ,  $I(\theta) > 0$  на  $(0,1)$ .

$$\frac{\partial \ln p(x, \theta)}{\partial \theta} = \frac{\partial}{\partial \theta} \left( \ln \frac{\theta}{2} \right) = \frac{1}{\theta} \quad \frac{\partial \ln p_{12}}{\partial \theta} = \frac{\partial}{\partial \theta} \left( \ln \left( \frac{1-\theta}{2} \right) \right) = -\frac{1}{1-\theta}$$

$$I(\theta) = \int_0^1 \frac{1}{2\theta} dx = \frac{1}{\theta(1-\theta)} \left( \frac{1-\theta}{2} \right) \cdot 2 = \frac{1}{\theta} + \frac{1}{1-\theta} = \frac{1}{\theta(1-\theta)} \geq 0 \text{ на } (0,1)$$

$\tilde{\theta}_1 = 1 - \bar{x}$  - несмещ,  $D[\tilde{\theta}_1]$  - оптимальная дисперсия  $\in (0,1)$   $\tilde{\theta}_1$  - пер

$\tilde{\theta}_2 = 1 - \tilde{\theta}_1 = \bar{x}$  - несмещ,  $D[\tilde{\theta}_2]$  - оптимальная дисперсия  $\in (0,1)$   $\tilde{\theta}_2$  - пер

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$$\theta_2: D[\tilde{\theta}_2] \geq \frac{1}{n I(\theta)} = \frac{\theta(1-\theta)}{n} \quad \text{ногай. уел. справедлива}$$

$\Rightarrow \tilde{\theta}_1$  неэф. т.к.  $\tilde{\theta}_1 \neq \tilde{\theta}_2$

(Т6)

$$p(x) = \begin{cases} \frac{\theta-1}{x^\theta}, & x \geq 1 \\ 0, & x < 1 \end{cases}, \quad \theta > 1 \quad \text{Парето}$$

$$L(x, \theta) = \frac{(\theta-1)^n}{\prod_{i=1}^n x_i^\theta}$$

$$\ln L = n \ln(\theta-1) - \theta \sum_{i=1}^n \ln x_i$$

$$(\ln L)' = \frac{n}{\theta-1} - \sum_{i=1}^n \ln x_i = 0$$

$$\tilde{\theta} = \ln \frac{1}{\ln x_i} \quad \text{т.к. } \frac{1}{\ln x_i}$$

$$(\ln L)'' = \frac{-n}{(\theta-1)^2} < 0 \Rightarrow \theta = \frac{1}{\ln x_i} - \text{м. макс.}$$

$$\left( \int \frac{\theta-1}{x^\theta} dx \right)' = x^{1-\theta} \ln x \Rightarrow \text{ногай. уел. неэф.}$$

$$\int \frac{\partial}{\partial \theta} \left( \frac{\theta-1}{x^\theta} \right) dx = x^{1-\theta} \ln x$$

$$\int_1^\infty \frac{\theta-1}{x^\theta} dx = -\frac{1}{x^{\theta-1}} + 1 = \frac{1}{2} \quad \tilde{x} = 2^{\frac{1}{\theta-1}}$$

$$g(\tilde{\theta}) = 2^{\frac{1}{\theta-1}}$$

$$\sqrt{n} \cdot \frac{g(\tilde{\theta}) - g(\theta)}{\sigma(\theta)} \sim N(0,1) \quad \sigma(\theta) \sim \sigma(\tilde{\theta})$$



$$\sigma'(\theta) = \sqrt{\nabla' g(\theta) I^{-1}(\theta) \nabla g(\theta)}$$

$$I(\theta) = E[(\ln p(\theta))^2] = E\left[\left(\frac{1}{\theta-1} - \ln x\right)^2\right] = \int_1^{\infty} \left(\frac{1}{\theta-1} - \ln x\right)^2 p(x, \theta) dx =$$

$$= \int_1^{\infty} \left(\frac{1}{\theta-1} - \ln x\right)^2 \frac{x^{\theta-1}}{x^2} dx = \frac{1}{(\theta-1)^2} \text{ since } \theta > 1$$

$$\nabla g(\theta) = \frac{-\ln 2 \cdot 2^{\frac{1}{\theta-1}}}{\theta-1} \quad \sigma'(\theta) = \frac{-\ln 2 \cdot 2^{\frac{1}{\theta-1}}}{\theta-1}$$

$$\sqrt{n} \frac{g(\hat{\theta}) - g(\theta)}{\sigma'(\hat{\theta})} \sim N(0, 1)$$

$$\frac{1.96 \sigma'(\hat{\theta})}{\sqrt{n}} + g(\hat{\theta}) < g(\theta) < \frac{-1.96 \sigma'(\hat{\theta})}{\sqrt{n}} + g(\hat{\theta})$$

$$\sqrt{n} \frac{\hat{\theta} - \theta}{\sigma(\hat{\theta})} \sim N(0, 1)$$

$$\sigma(\theta) \stackrel{P}{\rightarrow} \sigma'(\theta)$$

$$\sigma'(\theta) \approx -1 \quad \sqrt{n} \frac{\hat{\theta} - \theta}{\sigma'(\hat{\theta})} \sim N(0, 1)$$

$$\frac{-1.96(-1)}{\sqrt{n}} + 1 + \frac{1}{\ln x_i} < \theta < \frac{1.96(-1)}{\sqrt{n}} + 1 + \frac{1}{\ln x_i}$$