

Homework 4

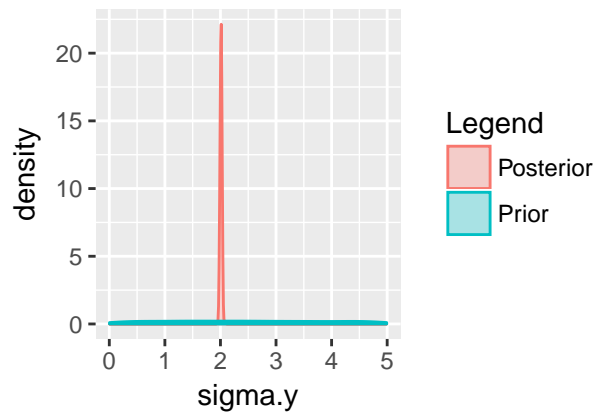
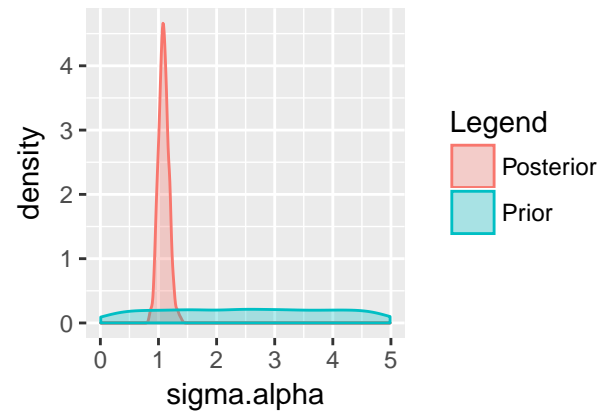
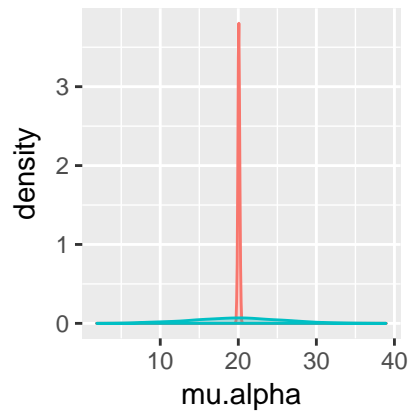
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Problem 1

$$m_{\sigma_y} = 5, m_{\sigma_{\alpha}} = 5, \mu_0 = 20, \sigma_{\mu_0}^2 = 36$$

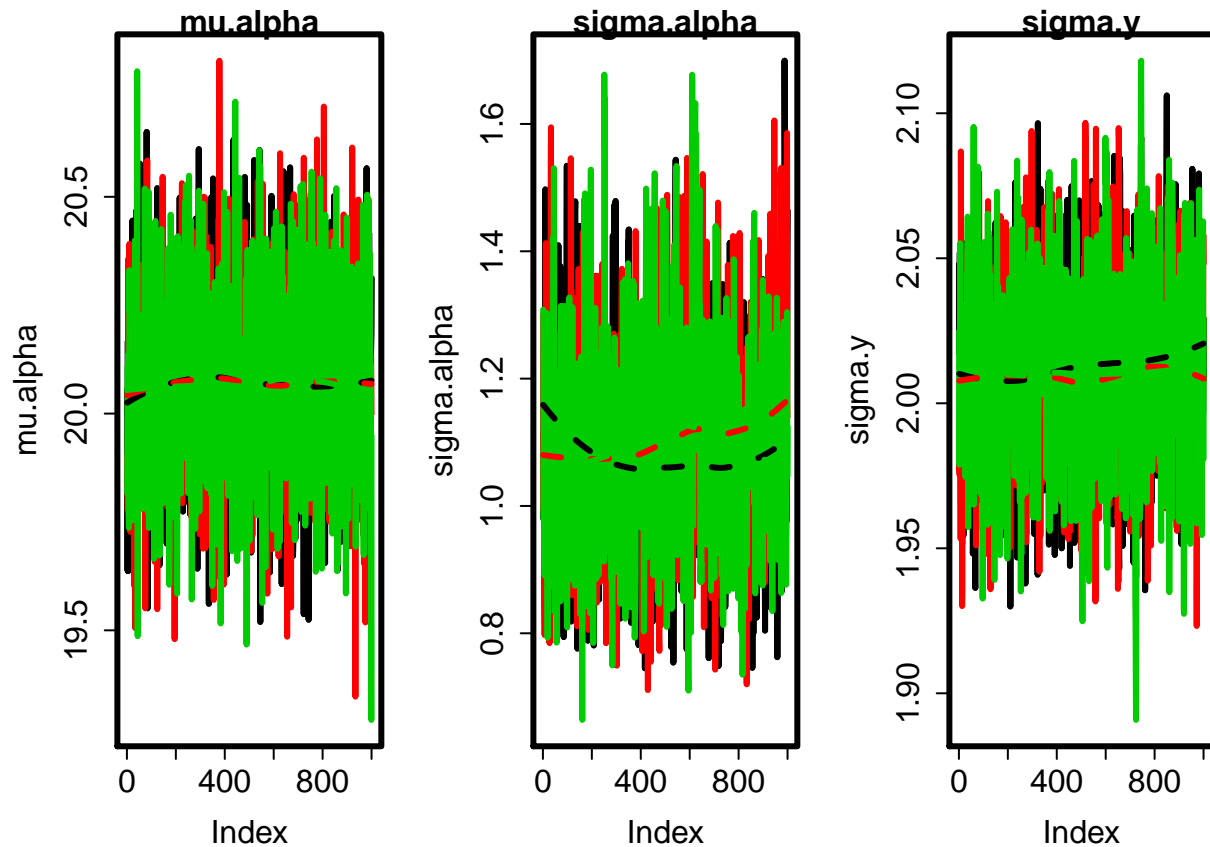
```
model <-  
"model {  
  for (i in 1:n){  
    y.i[i] ~ dnorm(alpha.j[getj.i[i]],tau.y)  
  }  
  
  for (j in 1:J){  
    alpha.j[j] ~ dnorm(mu.alpha,tau.alpha)  
  }  
  
  tau.y <- pow(sigma.y, -2)  
  tau.alpha <- pow(sigma.alpha, -2)  
  
  mu.alpha ~ dnorm(20,1/6^2)  
  sigma.y ~ dunif(0,5)  
  sigma.alpha ~ dunif(0,5)  
}"
```

```
## Compiling model graph  
##   Resolving undeclared variables  
##   Allocating nodes  
## Graph information:  
##   Observed stochastic nodes: 2400  
##   Unobserved stochastic nodes: 43  
##   Total graph size: 4856  
##  
## Initializing model
```



Problem 2

```
##          n.eff  Rhat
## mu.alpha    3000 1.001
## sigma.alpha   560 1.004
## sigma.y     2200 1.001
```



The high N_{eff} values show that the chains have low auto-covariance, which shows that the chains are able to explore the full parameter space. The \hat{R} value of 1 for all params show that all the chains have mixed well, and they did not get stuck in local modes.

Problem 3

a)

```
print (c(mean(mcmc.array[,, "mu.alpha"]), quantile(mcmc.array[,, "mu.alpha"], c(.025, .975))) )

##          2.5%    97.5%
## 20.06783 19.68833 20.48143

print (c(mean(mcmc.array[,, "sigma.alpha"]), quantile(mcmc.array[,, "sigma.alpha"], c(.025, .975))) )

##          2.5%    97.5%
## 1.0880475 0.8260481 1.4396968

print (c(mean(mcmc.array[,, "sigma.y"]), quantile(mcmc.array[,, "sigma.y"], c(.025, .975))) )

##          2.5%    97.5%
## 2.010845 1.953183 2.069122
```

b)

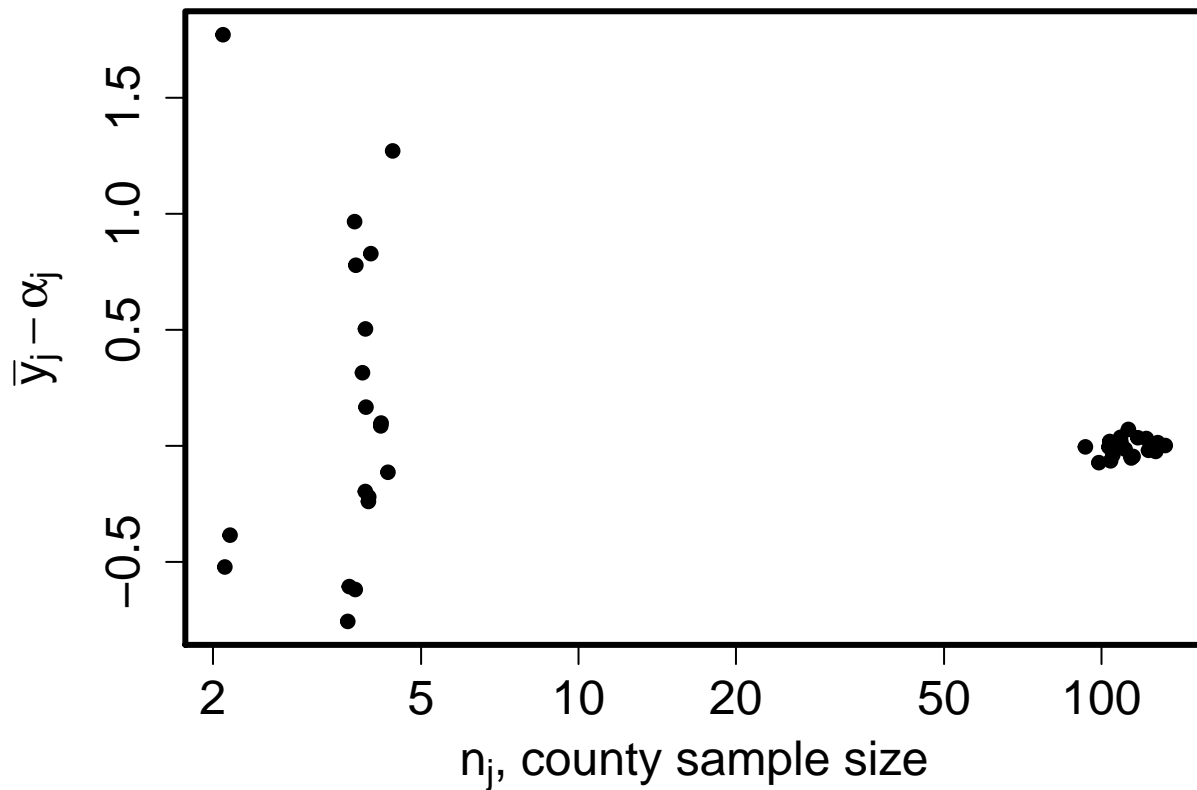
```
print (round(mod0$BUGSoutput$summary[c("alpha.j[1]"), c("mean", "sd", "2.5%", "97.5%")], 1))

## mean    sd  2.5% 97.5%
## 21.1   0.9  19.4  22.8
```

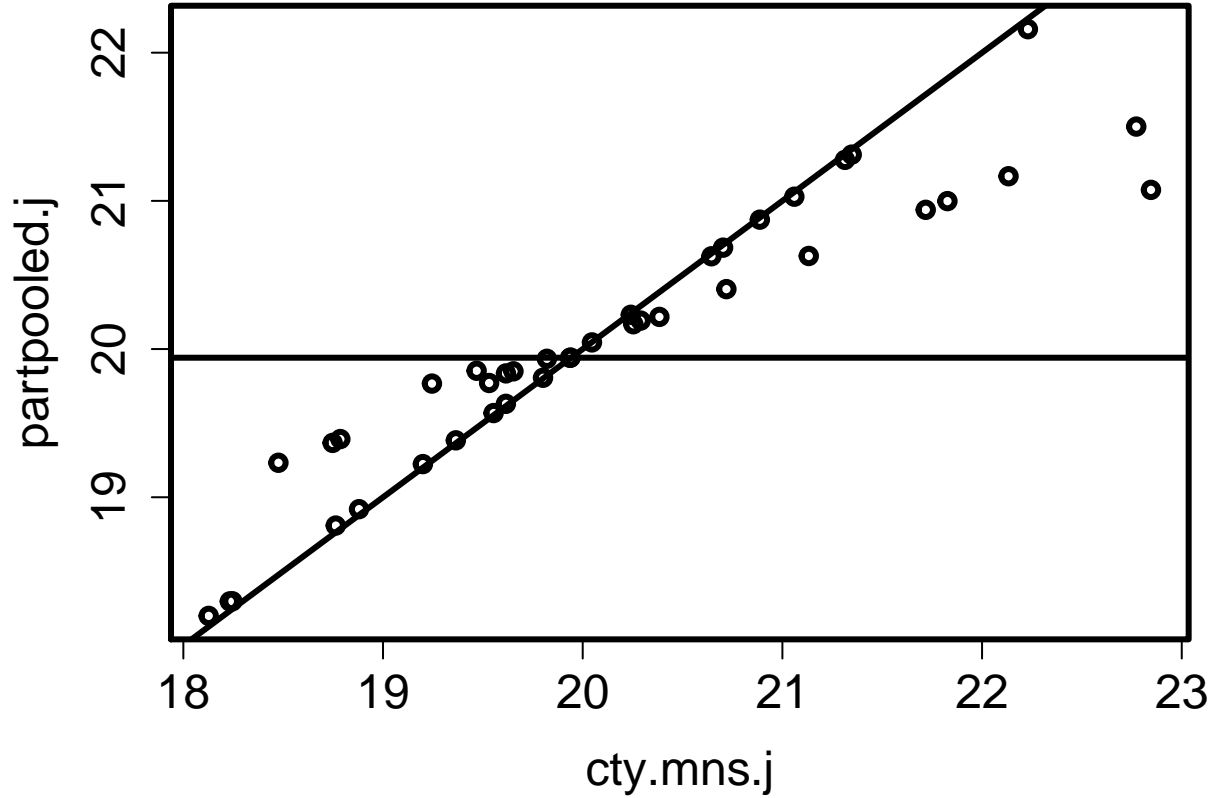
Problem 4

```
par(lwd = 3, cex.axis = 1.5, cex.lab = 1.5, cex.main = 1.5, mar = c(5,5,1,1))
lims <- range(cty.mns.j-partpooled.j)
plot(c(cty.mns.j-partpooled.j) ~ sample.size.jittered.j, type = "p",
      ylab = expression(bar(y)[j] - alpha[j]),
      xlab = expression(paste(n[j], ", county sample size")), pch = 20,
      log = "x",
      ylim = lims)

abline(h=ybarbar)
```



```
plot(cty.mns.j,partpooled.j,abline(h=ybarbar))
abline(0,1)
```



We can see that there is more variability of \bar{y}_j around α_j when sample size is small, which makes sense because the data does not pull α_j towards \bar{y}_j . We can from the second plot that as \bar{y}_j gets larger in magnitude relative to the overall mean, the α_j is pulled down towards the overall mean. Similarly, as \bar{y} gets small in magnitude relative to the overall mean, α_j gets pulled upwards towards the overall-mean. This is the “shrinkage” effect.

Problem 5

Consider

$$p(\alpha_j|y, \mu_\alpha, \sigma_y, \sigma_\alpha) \propto p(y|\alpha_j, \sigma_y)p(\alpha_j|\mu_\alpha, \sigma_\alpha)$$

$$p(\alpha_j|y, \mu_\alpha, \sigma_y, \sigma_\alpha) \propto \left(\frac{1}{\sqrt{2\pi}\sigma_y}\right)^n e^{-\frac{1}{2\sigma_y^2} \sum_i (y_i - \alpha_j)^2} \frac{1}{\sqrt{2\pi}\sigma_\alpha} e^{-\frac{1}{2\sigma_\alpha^2} (\alpha_j - \mu_\alpha)^2}$$

For now, let's just focus in on the terms involving e^x

we see the exponent can be written as

$$-\frac{1}{2\sigma_y^2} \sum_i (y_i - \alpha_j)^2 - \frac{1}{2\sigma_\alpha^2} (\alpha_j - \mu_\alpha)^2$$

Expanding out the square we get

$$-\frac{1}{2\sigma_y^2} \sum_i (y^2 - 2y\alpha_j + \alpha_j^2) - \frac{1}{2\sigma_\alpha^2} (\alpha_j^2 - 2\alpha_j\mu_\alpha + \mu_\alpha^2)$$

$$-\frac{1}{2\sigma_y^2} \sum_i y^2 + \frac{n_j \bar{y}_j \alpha_j}{\sigma_y^2} - \frac{n_j \alpha_j^2}{2\sigma_y^2} - \frac{1}{2\sigma_\alpha^2} \alpha_j^2 + \frac{1}{\sigma_\alpha^2} \alpha_j \mu_\alpha - \frac{1}{2\sigma_\alpha^2} \mu_\alpha^2$$

Collecting terms on α_j we can re-write this as

$$\begin{aligned} & -\frac{n_j \alpha_j^2}{2\sigma_y^2} - \frac{1}{2\sigma_\alpha^2} \alpha_j^2 + \frac{1}{\sigma_\alpha^2} \alpha_j \mu_\alpha + \frac{n_j \bar{y}_j \alpha_j}{\sigma_y^2} - \frac{1}{2\sigma_\alpha^2} \mu_\alpha^2 - \frac{1}{2\sigma_y^2} \sum_i y^2 \\ & -\alpha_j^2 \left(\frac{n_j}{2\sigma_y^2} + \frac{1}{2\sigma_\alpha^2} \right) + \alpha_j \left(\frac{1}{\sigma_\alpha^2} \mu_\alpha + \frac{n_j \bar{y}_j}{\sigma_y^2} \right) - \frac{1}{2\sigma_\alpha^2} \mu_\alpha^2 - \frac{1}{2\sigma_y^2} \sum_i y^2 \end{aligned}$$

We see this is almost a quadratic in α_j , that is we almost have a form of

$$a^2 \alpha_j^2 - b \alpha_j + c^2 - c^2 = (a \alpha_j - c)^2$$

If we take

$$b = 2ac \rightarrow c = \frac{b}{2a}$$

We know from above that $b = \left(\frac{1}{\sigma_\alpha^2} \mu_\alpha + \frac{n_j \bar{y}_j}{\sigma_y^2} \right)$

and $a = \sqrt{\frac{n_j}{2\sigma_y^2} + \frac{1}{2\sigma_\alpha^2}}$

So we can choose

$$\begin{aligned} c &= \frac{\left(\frac{1}{\sigma_\alpha^2} \mu_\alpha + \frac{n_j \bar{y}_j}{\sigma_y^2} \right)}{\sqrt{2 \left(\frac{n_j}{2\sigma_y^2} + \frac{1}{2\sigma_\alpha^2} \right)}} \\ &= \frac{\left(\frac{1}{\sigma_\alpha^2} \mu_\alpha + \frac{n_j \bar{y}_j}{\sigma_y^2} \right)}{\sqrt{\frac{n_j}{\sigma_y^2} + \frac{1}{\sigma_\alpha^2}}} \end{aligned}$$

plugging this back into the exp we get

$$p(\alpha_j | y, \mu_\alpha, \sigma_y, \sigma_\alpha) \propto \frac{1}{C} e^{-\left(\sqrt{\frac{n_j}{2\sigma_y^2} + \frac{1}{2\sigma_\alpha^2}} \alpha_j - \frac{\left(\frac{1}{\sigma_\alpha^2} \mu_\alpha + \frac{n_j \bar{y}_j}{\sigma_y^2} \right)}{\sqrt{\frac{n_j}{\sigma_y^2} + \frac{1}{\sigma_\alpha^2}}} \right)^2}$$

$$p(\alpha_j | y, \mu_\alpha, \sigma_y, \sigma_\alpha) \propto \frac{1}{C} e^{-\frac{1}{2 \left(\frac{n_j}{\sigma_y^2} + \frac{1}{\sigma_\alpha^2} \right)} \left(\alpha_j - \frac{\left(\frac{1}{\sigma_\alpha^2} \mu_\alpha + \frac{n_j \bar{y}_j}{\sigma_y^2} \right)}{\frac{n_j}{\sigma_y^2} + \frac{1}{\sigma_\alpha^2}} \right)^2}$$

$$p(\alpha_j | y, \mu_\alpha, \sigma_y, \sigma_\alpha) \propto \frac{1}{C} e^{-\frac{1}{2} \frac{1}{\frac{n_j}{\sigma_y^2} + \frac{1}{\sigma_\alpha^2}} \left(\alpha_j - \frac{\left(\frac{1}{\sigma_\alpha^2} \mu_\alpha + \frac{n_j \bar{y}_j}{\sigma_y^2} \right)}{\frac{n_j}{\sigma_y^2} + \frac{1}{\sigma_\alpha^2}} \right)^2}$$

which we can see is the kernel of a normal distribution with mean

$$m = \frac{\left(\frac{1}{\sigma_\alpha^2} \mu_\alpha + \frac{n_j \bar{y}_j}{\sigma_y^2} \right)}{\frac{n_j}{\sigma_y^2} + \frac{1}{\sigma_\alpha^2}}$$

and variance

$$v = \frac{1}{\frac{n_j}{\sigma_y^2} + \frac{1}{\sigma_\alpha^2}}$$