

Risk Parity Portfolios under a Markov Regime-Switching Framework

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MMF Symposium 2018, Blue Mountain, Ontario

January 19, 2018



Mechanical & Industrial Engineering
UNIVERSITY OF TORONTO

What is this presentation about?

- ▶ Markets follow a cyclical behaviour - alternating between bullish and bearish states.
- ▶ If we hold a portfolio, how could we hedge against these cycles?
 - Could we take advantage of this information to improve our models?
- ▶ We will explore a typical asset allocation problem that incorporates the cyclical behaviour of markets via machine learning algorithms.

There are two terms to keep in mind throughout this presentation

- ▶ Hidden Markov Models (HMMs): Will be used to incorporate information about the market state into our portfolio
- ▶ Risk Parity: Will be used to solve our asset allocation problem.

Hidden Markov Models

- ▶ HMMs are popular in Machine Learning as both classification and clustering algorithms.
- ▶ What if we believe that asset returns are governed by hidden market states?
- ▶ Could we use observed market returns to discern between these states and incorporate these dynamics into a portfolio optimization problem?

Hidden Markov Models in Asset Allocation

- ▶ We can use HMMs to improve the estimated parameters used in portfolio optimization.
- ▶ Typical inputs are the asset expected returns and covariance matrix.
- ▶ Estimation errors in expected returns tend to have a significant impact during the optimization process.
- ▶ What effect would a HMM have in a portfolio that does not use expected returns?

Risk Parity Portfolios

- ▶ Risk parity seeks to construct a portfolio based on a risk-weighted basis.
- ▶ Does not require estimated returns as an input, improving stability.
- ▶ Each asset should have the same risk contribution.
- ▶ By design, the resulting portfolio is well-diversified.
- ▶ Risk parity is also known as 'equal risk contribution' (**ERC**).

Here is a summary of the key findings from our research

- ▶ We seek to improve a risk parity portfolio by identifying a regime-dependent risk measure.
- ▶ This would incorporate the market dynamics as an input to the optimization problem.
- ▶ We can also use the Markov model to dictate the portfolio rebalancing policy.
- ▶ Higher risk-adjusted returns can be achieved by reducing exposure to concentrated risk through this improved covariance estimate.

A 2-state model can be explained by this factor model of asset returns

$$\mathbf{r} = l_1(\alpha_1 + \mathbf{V}_1^\top \mathbf{f}_1 + \epsilon_1) + l_2(\alpha_2 + \mathbf{V}_2^\top \mathbf{f}_2 + \epsilon_2).$$

- ▶ $\mathbf{r} \in \mathbb{R}^n$ are returns of the n assets.
- ▶ $l_i = 1$ when the current market state $s_t = i$ and $l_i = 0$ otherwise.
- ▶ $\alpha_i \in \mathbb{R}^n$ is the vector of regression intercepts.
- ▶ $\mathbf{f}_i \sim \mathcal{N}(\phi_i, \mathbf{F}_i) \in \mathbb{R}^m$ is the vector of factor returns.
- ▶ $\mathbf{V}_i \in \mathbb{R}^{m \times n}$ is the matrix of factor loadings.
- ▶ $\epsilon_i \sim \mathcal{N}(\mathbf{0}, \mathbf{D}_i) \in \mathbb{R}^n$ is the vector of residuals.

From this factor model we can derive a **probability-weighted regime-dependent covariance matrix**.

Experimental Setup

- ▶ We use the Fama–French 3-factor model as the basis.
- ▶ We construct portfolios with 48 assets ($n = 48$).
- ▶ 4 different investment strategies were tested:
 - i Nominal ERC
 - ii Robust ERC
 - iii **2-State ERC**
 - iv 3-State ERC
- ▶ We use weekly asset and factor returns from 1985 to 2010.

We conducted two individual experiments

1. Fixed 6-month rebalance policy (22 investment periods)
2. Dynamic rebalance policy
 - ▶ We re-identify the regime every four weeks.
 - ▶ We identify the regime by inspection:
 - ⇒ We check the last entry of the smoothed probabilities from our Markov model.
 - ▶ To mitigate estimation noise, rebalancing occurs only when a sustained change in regime is perceived:
 - ⇒ A change in regime must be consistent for **two** consecutive 4-week periods.

Example of estimated regime changes for the out-of-sample test

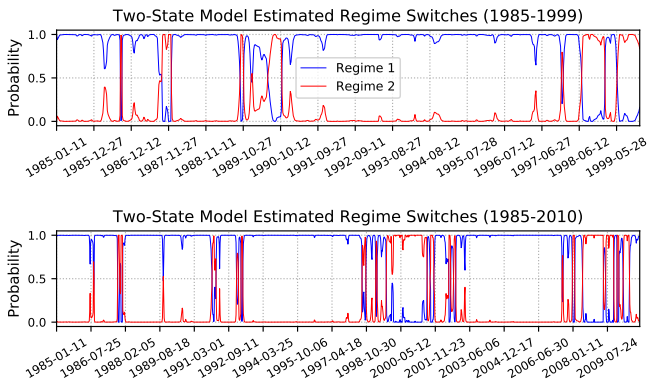


Figure: Estimated regime switches for test period 2000–2010. The probability of having a bullish (bearish) market is shown in blue (red). **Top:** regime switches estimated before the first investment period. **Bottom:** regime switches estimated after the last investment period.

Computational Experiments: Results

Fixed-term rebalance policy

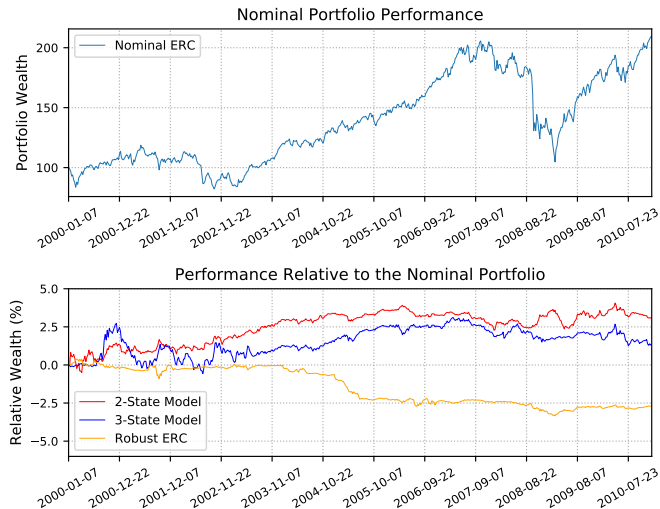


Figure: Evolution of wealth for test period 2000–2010 under a fixed-term rebalance policy. **Top:** evolution of wealth of nominal portfolio. **Bottom:** evolution of wealth relative to the nominal portfolio.

Summary of Results: Fixed-term rebalance policy

	Nominal ERC	Robust	2-State	3-State
Yearly Return	0.0695	0.0668	0.0724	0.0708
Sharpe Ratio	0.0511	0.0488	0.0531	0.0516
Turnover	0.0222	0.0826	0.0722	0.1866

Computational Experiments: Results

Dynamic rebalance policy

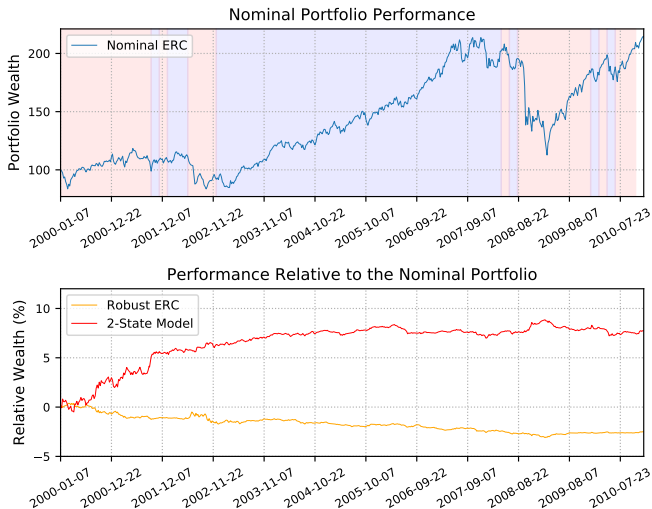


Figure: Evolution of wealth for test period 2000–2010 under a dynamic rebalance policy. **Top:** evolution of absolute wealth of nominal portfolio. **Bottom:** evolution of wealth relative to the nominal portfolio.

Summary of Results: Dynamic rebalance policy

	Nominal	Robust	2-State
Yearly Return	0.0719	0.0694	0.0792
Sharpe Ratio	0.0528	0.0508	0.0579
Turnover	0.0274	0.0516	0.1699

Key takeaways

- ▶ A regime-switching factor model allows for the estimation of regime-dependent covariance matrices.
- ▶ This model is able to capture and incorporate the market dynamics.
- ▶ We can use the Markov model to dictate the portfolio rebalancing policy.
- ▶ Higher risk-adjusted returns can be achieved by improving the quality of our estimated risk measure.
- ▶ The model developed is highly tractable and easily scalable.

Decomposing the risk measure

A decomposable risk measure can be used to derive the risk contribution per asset:

$$\sigma_p = \sqrt{\mathbf{x}^\top \Sigma \mathbf{x}} = \sum_{i=1}^n \sigma_i,$$
$$\sigma_i = x_i \frac{\partial \sigma_p}{\partial x_i} = x_i \frac{(\Sigma \mathbf{x})_i}{\sqrt{\mathbf{x}^\top \Sigma \mathbf{x}}},$$

- ▶ $\mathbf{x} \in \mathbb{R}^n$ is the vector of asset weights.
- ▶ $\Sigma \in \mathbb{R}^{n \times n}$ is the asset covariance matrix.
- ▶ σ_p is the portfolio standard deviation.
- ▶ σ_i is the risk contribution of asset i .

Risk Parity Portfolio Optimization

A typical risk parity formulation seeks to minimize the squared differences between the asset risk contributions:

$$\begin{aligned} \min_{\mathbf{x}} \quad & \sum_{i=1}^n \sum_{j=1}^n (x_i(\boldsymbol{\Sigma}\mathbf{x})_i - x_j(\boldsymbol{\Sigma}\mathbf{x})_j)^2 \\ \text{s.t.} \quad & \mathbf{1}^\top \mathbf{x} = 1, \\ & \mathbf{x} \geq 0. \end{aligned}$$

This yields a unique solution when short-selling is not allowed.

The general factor model of asset returns

$$\mathbf{r} = \boldsymbol{\alpha} + \mathbf{V}^\top \mathbf{f} + \boldsymbol{\epsilon}$$

- ▶ $\mathbf{r} \in \mathbb{R}^n$ are returns of the n assets.
- ▶ $\boldsymbol{\alpha} \in \mathbb{R}^n$ is the vector of regression intercepts.
- ▶ $\mathbf{f} \sim \mathcal{N}(\boldsymbol{\phi}, \mathbf{F}) \in \mathbb{R}^m$ is the vector of factor returns.
- ▶ $\mathbf{V} \in \mathbb{R}^{m \times n}$ is the matrix of factor loadings.
- ▶ $\boldsymbol{\epsilon} \sim \mathcal{N}(\mathbf{0}, \mathbf{D}) \in \mathbb{R}^n$ is the vector of residuals.

Accordingly, the expected returns and covariance matrix are

$$\begin{aligned}\boldsymbol{\mu} &= \boldsymbol{\alpha} + \mathbf{V}^\top \boldsymbol{\phi}, \\ \boldsymbol{\Sigma} &= \mathbf{V}^\top \mathbf{F} \mathbf{V} + \mathbf{D}.\end{aligned}$$

Suppose the factors behave differently under distinct market regimes

- ▶ A 2-state market is governed by the following Markov transition matrix

$$\Gamma = \begin{bmatrix} \gamma_{11} & \gamma_{12} \\ \gamma_{21} & \gamma_{22} \end{bmatrix}$$

- γ_{ij} is the probability of switching from state i to state j .
- ▶ Suppose this matrix dictates whether the asset returns are governed by

$$\mathbf{r} = \alpha_1 + \mathbf{V}_1^\top \mathbf{f}_1 + \epsilon_1,$$

or

$$\mathbf{r} = \alpha_2 + \mathbf{V}_2^\top \mathbf{f}_2 + \epsilon_2.$$

We can capture the market dynamics within a single factor model

- ▶ For a 2-state model, we have the factor model

$$\mathbf{r} = l_1(\boldsymbol{\alpha}_1 + \mathbf{V}_1^\top \mathbf{f}_1 + \boldsymbol{\epsilon}_1) + l_2(\boldsymbol{\alpha}_2 + \mathbf{V}_2^\top \mathbf{f}_2 + \boldsymbol{\epsilon}_2).$$

- $l_i = 1$ when the current market state $s_t = i$ and $l_i = 0$ otherwise.
- ▶ Computing the covariance as $\text{cov}(r_i, r_j) \forall i, j$, we obtain a probability-weighted regime-dependent covariance matrix

$$\begin{aligned} \Sigma_{RS_i} = & \gamma_{i1} \mathbf{V}_1^\top \mathbf{F}_1 \mathbf{V}_1 + \gamma_{i2} \mathbf{V}_2^\top \mathbf{F}_2 \mathbf{V}_2 + \gamma_{i1}(1 - \gamma_{i1}) \mathbf{V}_1^\top \boldsymbol{\phi}_1 \boldsymbol{\phi}_1^\top \mathbf{V}_1 \\ & + \gamma_{i2}(1 - \gamma_{i2}) \mathbf{V}_2^\top \boldsymbol{\phi}_2 \boldsymbol{\phi}_2^\top \mathbf{V}_2 - \gamma_{i1} \gamma_{i2} \mathbf{V}_1^\top \boldsymbol{\phi}_1 \boldsymbol{\phi}_2^\top \mathbf{V}_2 \\ & - \gamma_{i1} \gamma_{i2} \mathbf{V}_2^\top \boldsymbol{\phi}_2 \boldsymbol{\phi}_1^\top \mathbf{V}_1 + \gamma_{i1} \mathbf{D}_1 + \gamma_{i2} \mathbf{D}_2. \end{aligned}$$

- $\gamma_{ij} = \mathbb{E}[l_j \mid s_t = i]$, i.e. the transition probability.

Estimation of regime switches between 1985 – 1999

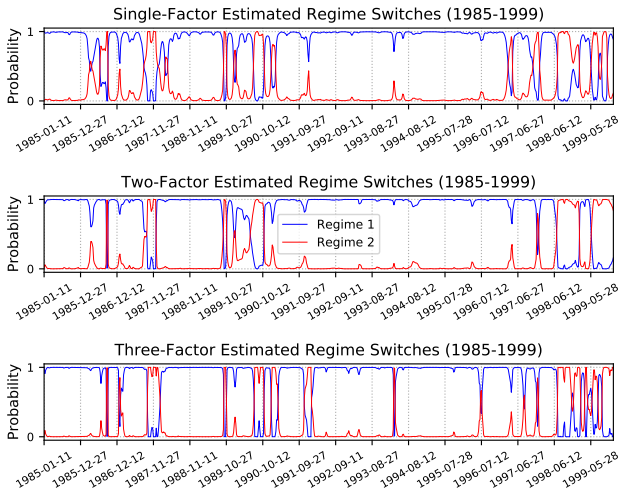


Figure: Estimated regime switches for the period of 1985–1999. The probability of having a bullish (bearish) market is shown in blue (red). **Top:** estimated switches based on the market factor. **Middle:** estimated switches based on the market and the 'high-minus-low' factors. **Bottom:** estimated switches based on all 3 Fama–French factors.