

# **Semiparametric Vector Generalized Linear Models**

Estimation and Computation

---

Gabriel Dennis

Honours Defence

Supervisor: Dr Alan Huang

23rd of June 2021

This project had the following goals

- generalise the semiparametric GLM of Huang ([2014](#)) to deal with vector responses
- write code to fit the model
- verify the properties of the model's estimates via simulation

Generalising the semiparametric GLM Huang (2014) involves writing the joint exponentially tilted density in the following form

$$dF_i(\mathbf{y}) = \exp\{b_i + \boldsymbol{\theta}_i^T \mathbf{y}\} dF(\mathbf{y}), \quad i = 1, \dots, n, \mathbf{y} \in \mathbb{R}^K, \quad (1)$$

Generalising the semiparametric GLM Huang (2014) involves writing the joint exponentially tilted density in the following form

$$dF_i(\mathbf{y}) = \exp\{b_i + \boldsymbol{\theta}_i^T \mathbf{y}\} dF(\mathbf{y}), \quad i = 1, \dots, n, \mathbf{y} \in \mathbb{R}^K, \quad (1)$$

with slight changes to the normalising and tilt constraints

$$b(\mathbf{X}_i, \boldsymbol{\beta}, F) = -\log \left\{ \int_{\mathcal{Y}} \exp\{\boldsymbol{\theta}_i^T \mathbf{y}\} dF(\mathbf{y}) \right\}, \quad (2)$$

$$\mu_{(k)}(\mathbf{X}_{(k)}^T \boldsymbol{\beta}_{(k)}) = \int_{\mathcal{Y}} y_{(k)} \exp\{b_i + \boldsymbol{\theta}_i^T \mathbf{y}\} dF(\mathbf{y}), \quad k = 1, \dots, K. \quad (3)$$

To estimate the reference distribution and mean model parameters the semiparametric log-likelihood is

$$\ell(\boldsymbol{\beta}, \boldsymbol{p}) = \sum_{i=1}^n \log p_i + b_i + \boldsymbol{\theta}_i^T \boldsymbol{Y}_i. \quad (4)$$

To estimate the reference distribution and mean model parameters the semiparametric log-likelihood is

$$\ell(\boldsymbol{\beta}, \mathbf{p}) = \sum_{i=1}^n \log p_i + b_i + \boldsymbol{\theta}_i^T \mathbf{Y}_i. \quad (4)$$

This also results in changes to the empirical normalising and tilt constraints

$$1 = \sum_{i=1}^n p_i \exp\{b_j + \boldsymbol{\theta}_j^T \mathbf{Y}_i\}, \quad j = 1, 2, 3, \dots, n \quad (5)$$

$$\mu_{(k)}(\mathbf{X}_{(k)j}^T \boldsymbol{\beta}_{(k)}) = \sum_{i=1}^n p_i Y_{(k)i} \exp\{b_j + \boldsymbol{\theta}_j^T \mathbf{Y}_i\} \quad j = 1, \dots, n, k = 1, \dots, K. \quad (6)$$

The code to fit the model is written in MATLAB using `fmincon`, and uses a formula syntax to specify constraints across different components mean models.

The code to fit the model is written in MATLAB using `fmincon`, and uses a formula syntax to specify constraints across different components mean models.

Separate models

```
model = fit_vspglm(["y_1 ~ x_1", "y_2 ~ x_2"], tbl, links);.
```



The code to fit the model is written in MATLAB using `fmincon`, and uses a formula syntax to specify constraints across different components mean models.

Separate models

```
model = fit_vspglm(["y_1 ~ x_1", "y_2 ~ x_2"], tbl, links);.
```

Different covariates but equal regression coefficients,

```
model = fit_vspglm(["(y_1, y_2) ~ ((x_1&x_2))"], tbl, links);,
```

The code to fit the model is written in MATLAB using `fmincon`, and uses a formula syntax to specify constraints across different components mean models.

Separate models

```
model = fit_vspglm(["y_1 ~ x_1", "y_2 ~ x_2"], tbl, links);.
```

Different covariates but equal regression coefficients,

```
model = fit_vspglm(["(y_1, y_2) ~ ((x_1&x_2))"], tbl, links);,
```

same covariates and regression coefficients,

```
model = fit_vspglm(["(y_1, y_2) ~ x"], tbl, links);
```

Simulations from non-standard generalized linear mixed models and generalized linear mixed effects models showed comparable parameter coverage and type I error rates to those of Huang ([2014](#)) and the multivariate density ratio model of Marchese ([2018](#)), using similar sample sizes.

Tri-variate Poisson simulation.

$$\alpha \sim \mathcal{N}(0, \sigma_0^2), \mathbf{X} \sim \text{U}[-1, 1]^3 \quad \sigma \in \mathbb{R}_+$$

$$Y_1 | \alpha, \mathbf{X}_{(1)} \sim \text{Pois}(\exp(\mathbf{X}_{(1)}^T \boldsymbol{\beta}_{(1)} + \alpha))$$

$$Y_2 | \alpha, \mathbf{X}_{(2)} \sim \text{Pois}(\exp(\mathbf{X}_{(2)}^T \boldsymbol{\beta}_{(2)} + 0.5\alpha))$$

$$Y_3 | \alpha, \mathbf{X}_{(3)} \sim \text{Pois}(\exp(\mathbf{X}_{(3)}^T \boldsymbol{\beta}_{(3)} - 0.3\alpha))$$

**Table 1:** Simulation results using a sample size of  $n = 200$  and  $N = 1000$  simulations

$\beta$	$\hat{\beta}$	$ \beta - \hat{\beta} $	Errors		$p \leq 0.05$	CI		
			$\hat{\sigma}$	$\text{se}(\hat{\beta})$		90%	95%	99%
0.4	0.40	0.004	0.13	0.12	0.88	0.89	0.94	0.99
-0.8	-0.79	0.002	0.13	0.14	0.99	0.93	0.96	0.99
0	0.001	0.001	0.13	0.12	0.05	0.89	0.95	0.99

There are several areas which future work can take place

- reduce the number of parameters optimized over
- increase the numeric stability of the code
- write a R package for the VSPGLM
- verify properties of the VSPGLM theoretically



Huang, Alan (2014). "Joint estimation of the mean and error distribution in generalized linear models". In: *Journal of the American Statistical Association* 109.505, pp. 186–196.



Marchese, Scott (2018). "Semiparametric Regression Methods for Mixed Type Data Analysis". Thesis.

$$\alpha \sim \mathcal{N}(0, \sigma_0^2), \quad \sigma \in \mathbb{R}_+$$

$$\mathbf{X} \sim \text{U}[-1, 1]^3$$

$$Y_1 | \alpha, \mathbf{X}_{(1)} \sim \text{Pois}(\exp(\mathbf{X}_{(1)}^T \boldsymbol{\beta}_{(1)} + 0.4\alpha))$$

$$Y_2 | \alpha, \mathbf{X}_{(2)} \sim \text{Pois}(\exp(\mathbf{X}_{(2)}^T \boldsymbol{\beta}_{(2)} - 0.5\alpha))$$

$$Y_3 | \alpha, \mathbf{X}_{(3)} \sim \text{Pois}(\exp(\mathbf{X}_{(3)}^T \boldsymbol{\beta}_{(3)}))$$

# Trivariate Poisson Simulation Results

**Table 2:** Simulation results for trivariate Poisson model using a sample size of  $n = 200$  and  $N = 1000$  simulations

$\beta$	$\hat{\beta}$	$ \beta - \hat{\beta} $	Errors		$p \leq 0.05$	CI		
			$\hat{\sigma}$	$\text{se}(\hat{\beta})$		90%	95%	99%
0.4	0.40	0.004	0.13	0.12	0.88	0.89	0.94	0.99
-0.8	-0.79	0.002	0.13	0.14	0.99	0.93	0.96	0.99
0	0.001	0.001	0.13	0.12	0.05	0.89	0.95	0.99



## Trivariate Mixed Effects Simulation

$$\alpha \sim \mathcal{N}(0, \sigma_0^2)$$

$$\mathbf{X} \sim \text{U}[-1, 1]^3$$

$$\mathbf{Y}_1 | \mathbf{X}_{(1)}, \alpha \sim \mathcal{N}(\mathbf{X}_{(1)}^T \boldsymbol{\beta}_{(1)} + \alpha, \sigma_1^2)$$

$$\mathbf{Y}_2 | \mathbf{X}_{(2)}, \alpha \sim \text{Pois}(\exp(\mathbf{X}_{(2)}^T \boldsymbol{\beta}_{(2)} - 0.5\alpha))$$

$$\mathbf{Y}_3 | \mathbf{X}_{(3)}, \alpha \sim \text{Gamma}(\lambda, \exp(\mathbf{X}_{(3)}^T \boldsymbol{\beta}_{(3)} + 0.4\alpha))$$

# Trivariate Mixed Effects Simulation Results

**Table 3:** Simulation results for trivariate mixed effects model using a sample size of  $n = 200$  and  $N = 1000$  simulations

Margin	$\beta$	$\hat{\beta}$	$ \beta - \hat{\beta} $	Errors		$p \leq 0.05$	CI		
				$\hat{\sigma}$	$\text{s\!e}(\hat{\beta})$		90%	95%	99%
Normal	1	1.005	0.0058	0.17	0.18	1	0.90	0.95	0.99
Poisson	-0.5	-0.49	0.0005	0.13	0.13	0.96	0.83	0.88	0.93
Gamma	0.4	0.39	0.004	0.12	0.13	0.85	0.89	0.93	0.97

## Multivariate Normal Simulation Results

**Table 4:** Simulation results for bivariate normal model using sample size of  $n = 200$  and  $N = 1000$  simulations

$\beta$	$\hat{\beta}$	$ \beta - \hat{\beta} $	Errors		$p \leq 0.05$	CI		
			$\hat{\sigma}$	$\text{s\bar{e}}(\hat{\beta})$		90%	95%	99%
-1	-0.99	0.0004	0.11	0.11	1	0.91	0.95	0.99
0	-0.0004	0.0004	0.11	0.11	0.05	0.90	0.95	0.99
0.5	0.49	0.009	0.14	0.13	0.95	0.88	0.94	0.98
2.2	2.19	0.003	0.14	0.14	1	0.90	0.94	0.98

**Table 5:** Type 1 errors at significance levels of 0.10, 0.05 and 0.01 using sample sizes of  $n = 75, 150$ , and  $N = 3000$  simulations

$n$	Type 1 Errors		
	0.10	0.05	0.01
75	0.119	0.060	0.012
150	0.097	0.050	0.013