HW2 Stat-comp (due Sun, Oct 15th in D2L)

1) Maximum likelihood estimation and inference with the exponential distribution

Recall that the density function of an exponential random variable is

$$f(x_i|\lambda) = \lambda e^{-\lambda x_i}$$

where $x_i \geq 0$ is the random variable, and $\lambda > 0$ is a rate parameter.

The expected value and variance of the random variables are $E[X] = \frac{1}{\lambda}$ and $Var[X] = \frac{1}{\lambda^2}$.

The following code simulates 50 IID draws from an exponential distribution

```
set.seed(195021)
x=rexp(n=50,rate=2)
```

The maximum likelihood estimate of λ has a closed form. Indeed, for a random sample of IID exponentially distributed random variables

$$P(X_1 = x_1, X_2 = x_2, ..., X_n = x_n | \lambda) = \prod_{i=1}^n \lambda e^{-\lambda x_i} = \lambda^n e^{-\lambda \sum_{i=1}^n x_i} = \lambda^n e^{-\lambda n\bar{x}}$$

Therefore, the log-likelihood is, $l(\lambda|x_1,...,x_n) = nlog(\lambda) - \lambda n\bar{x}$.

The derivative of the log-likelihood with respect to λ is

 $\frac{dl}{d\lambda} = \frac{n}{\lambda} - n\bar{x}$. Setting this derivative equal to zero, and solving for $\hat{\lambda}$ gives $\hat{\lambda} = \frac{1}{\bar{x}}$

- 1.1) Use optimize() to estimate λ , compare your estimate with $\frac{1}{\bar{x}}$.
- 1.2) Use numerical methods to provide an approximate 95% CI for your estimate.

Recall that we can approximate the sampling variance of the maximum likelihood estimate using inverse of the second derivative of the negative log-likelihood evaluated at the ML estimate, also known as the Hessian.

The function optimize() does not provide a Hessian. However, you can use the hessian() function of the numDeriv R-package to obtain a numerical approximation to the second order derivative of the log-Likelihood at the ML estimate. To install this package you can use

```
#install.packages(pkg='numDeriv',repos='https://cran.r-project.org/')
library(numDeriv)
```

To evaluate the Hessian you can use

```
H=hessian(fn=negLogLik,y=y,x=your_ml_estimate)
```

Above, negLogLik is a function to evaluate the -logLik for the model, y is the data and your_ml_estimate is the estimate you obtained in 1.1.

To get (an approximation to) the variance of the ML estimate you can use

Finally, once you have an estimate of the variance you can use estimate +/-1.96*SE to get an approximate 95% CI.

2) CIs for Predictions from Logistic Regression

Recall that in a logistic regression model, the log-odds are parameterized as

$$log\left[\frac{\theta_i}{(1-\theta_i)}\right] = \mathbf{x}_i'\beta = \eta_i \tag{1}$$

The sampling variance of $\mathbf{x}_i'\hat{\boldsymbol{\beta}} = \hat{\eta}_i$ is $Var(\hat{\eta}_i) = \mathbf{x}_i'\mathbf{V}\mathbf{x}_i$, where \mathbf{V} is the (co)variance matrix of the estimated effects; therefore, a SE and an approximate 95%CI for η_i can be obtained using

$$SE(\hat{\eta_i}) = \sqrt{\mathbf{x}_i' \mathbf{V} \mathbf{x}_i} \text{ and } CI : \mathbf{x}_i' \hat{\beta} + / -1.96 \times SE(\hat{\eta_i}).$$

Because the inverse-logit is a monotonic map, we can then obtain a 95% CI for the predicted probabilities by applying the inverse logit, $\theta_i = \frac{e^{\eta_i}}{1+e^{\eta_i}}$, to the bounds of the CI for the linear predictor.

- Using the gout data set, fit a logistic regression for gout using sex, age, and race as predictors (for this you can use glm(), don't forget the link!).
- From the fitted model, derive predictions and SEs for the linear predictor as well as predictions and SEs in the probability scale.

D	C	Λ	Predicted	OF OZ CLID	Pred. Prob	95%CI Prob.
Race	Sex	Age	LP	95%CI LP	Scale	Scale
White	Male	55				
White	Female	55				
Black	Male	55				
Black	Female	55				

Hint:

- predict(fm,se.fit=TRUE,newdata=tmp) returns predictions for the linear predictor, and
- predict(fm,se.fit=TRUE,newdata=tmp,type="response") returns predictions and SEs in the probability scale.

Above, fm is the fitted logistic regression model, and newdata=tmp is a data.frame containing the values of the predictors of the model at which you want to predict.