

Presentation at Crash Course on Data Assimilation

Geir Evensen



Available from <https://github.com/geirev/Presentations>

The Variational Inverse Problem

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Simple scalar example

Given the model

$$\begin{aligned}\frac{dx}{dt} &= 1 \\ x(0) &= 0 \\ x(1) &= 2\end{aligned}$$

- Overdetermined problem.
- No solution.

Allowing for errors

Relax model and conditions

$$\frac{dx}{dt} = 1 + q(t)$$

$$x(0) = 0 + a$$

$$x(1) = 2 + b$$

- Underdetermined problem.
- Infinitively many solutions.

Statistical assumption

Statistical null hypothesis, \mathcal{H}_0 :

$$\overline{q(t)} = 0,$$

$$\bar{a} = 0,$$

$$\bar{b} = 0,$$

$$\overline{q(t_1)q(t_2)} = C_0\delta(t_1 - t_2),$$

$$\overline{a^2} = C_0,$$

$$\overline{b^2} = C_0,$$

$$\overline{q(t)a} = 0,$$

$$\overline{ab} = 0,$$

$$\overline{q(t)b} = 0.$$

Seeking a solution that:

- is close to the conditions, and
- almost satisfies the model,

by minimizing error terms.

Penalty function

- Define quadratic penalty function

$$\mathcal{J}[x] = W_0 \int_0^1 \left(\frac{dx}{dt} - 1 \right)^2 dt + W_0 (x(0) - 0)^2 + W_0 (x(1) - 2)^2$$

with $W_0 = C_0^{-1}$.

- Then x is an extremum if

$$\delta \mathcal{J}[x] = \mathcal{J}[x + \delta x] - \mathcal{J}[x] = \mathcal{O}(\delta x^2)$$

when $\delta x \rightarrow 0$.

Variation of penalty function

We have

$$\begin{aligned}\mathcal{J}[x + \delta x] &= W_0 \int_0^1 \left(\frac{dx}{dt} - 1 + \frac{d\delta x}{dt} \right)^2 dt \\ &\quad + W_0 (x(0) - 0 + \delta x(0))^2 + W_0 (x(1) - 2 + \delta x(1))^2\end{aligned}$$

and we must have

$$\int_0^1 \frac{d\delta x}{dt} \left(\frac{dx}{dt} - 1 \right) dt + \delta x(0)(x(0) - 0) + \delta x(1)(x(1) - 2) = 0,$$

From integration by part we get

$$\delta x \left(\frac{dx}{dt} - 1 \right) \Big|_0^1 - \int_0^1 \delta x \frac{d^2 x}{dt^2} dt + \delta x(0)(x(0) - 0) + \delta x(1)(x(1) - 2) = 0.$$

Minimum of penalty function

This gives the following system of equations

$$\begin{aligned}\delta x(0) \left(-\frac{dx}{dt} + 1 + x \right) \Big|_{t=0} &= 0, \\ \delta x(1) \left(\frac{dx}{dt} - 1 + x - 2 \right) \Big|_{t=1} &= 0, \\ \int_0^1 \delta x \left(\frac{d^2x}{dt^2} \right) dt &= 0,\end{aligned}$$

or since δx is arbitrary....

Euler-Lagrange equation

The Euler–Lagrange equation

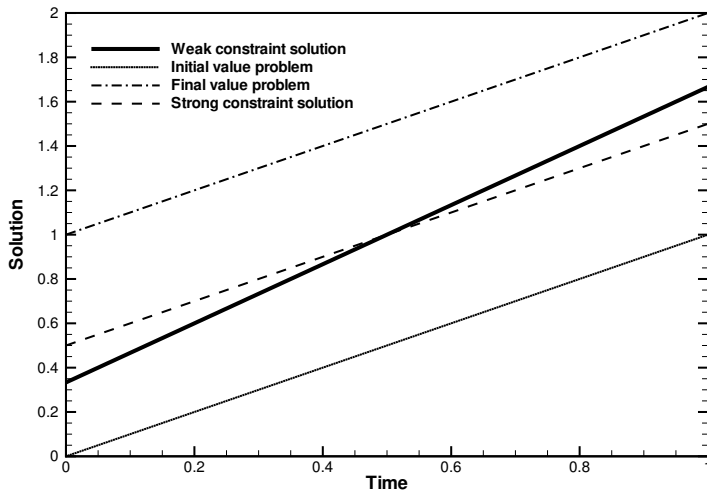
$$\begin{aligned}\frac{dx}{dt} - x &= 1 && \text{for } t = 0, \\ \frac{dx}{dt} + x &= 3 && \text{for } t = 1, \\ \frac{d^2x}{dt^2} &= 0.\end{aligned}$$

- Elliptic boundary value problem in time.
- It has a unique solution.

$$x = c_1 t + c_2,$$

with $c_1 = 4/3$ and $c_2 = 1/3$.

Results



Summary

- Model with conditions has a **unique solution**.
- Additional data makes problem **over determined**.
- Allowing for errors gives **infinitely many solutions**.
- Specify **mean and covariance** for error terms.
- Define **variational inverse problem** for **least-squares solution**.
- Weights are the inverses of the error covariances.
- **Least-squares solution** is defined by **Euler-Lagrange eqs.**
- Boundary value problem in time.
- **Weak-constraint solution**: almost satisfies dynamics and data.
- **Strong-constraint solution**: satisfies dynamics, and close to data.

Bayes' and the data assimilation problem

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Why Bayes Theorem?

- Provides a fundamental *framework* for data assimilation.
- All data-assimilation methods can be derived from Bayes'.

Properties of a probability density function

- The graph of the density function is continuous, since it is defined over a continuous range over a continuous variable.
- The total probability

$$P(x) = \int_{-\infty}^{\infty} f(x)dx = 1$$

- The probability of $x \in [a, b]$ is

$$P(x \in [a, b]) = \int_a^b f(x)dx$$

- And two special cases

$$P(x = c) = \int_c^c f(x)dx = 0 \quad \text{and} \quad \int_{-\infty}^{\infty} \delta(x)dx = 1$$

Also, we have

- The joint probability

$$f(x, y) = f(x)f(y|x) = f(y)f(x|y)$$

- Solving for $f(x|y)$ gives Bayes' theorem

$$f(x|y) = \frac{f(x)f(y|x)}{f(y)}$$

- Bayes states that “the probability of x given y , is equal to the probability of x , times the likelihood of y given x , divided by the probability of y .”
- Here $f(y)$ is a normalization constant so that the integral of $f(x|y)$ becomes one.

Bayes' theorem

Given:

- A state variable \mathbf{x} and its prior pdf: $f(\mathbf{x})$
- A vector of observations \mathbf{d} and their likelihood: $f(\mathbf{d}|\mathbf{x})$
- Bayes' theorem defines the posterior pdf, $f(\mathbf{x}|\mathbf{d})$:

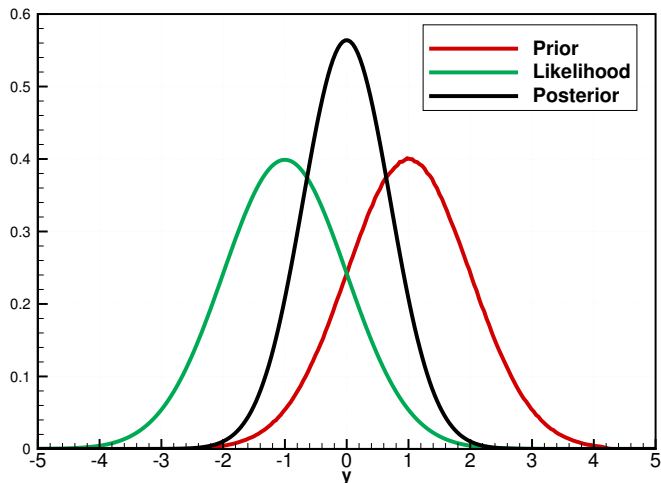
Bayes' theorem

$$f(\mathbf{x}|\mathbf{d}) = \frac{f(\mathbf{x})f(\mathbf{d}|\mathbf{x})}{f(\mathbf{d})}$$

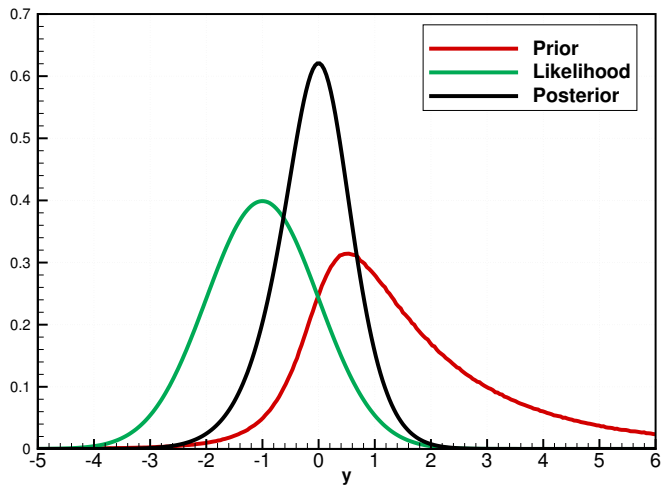
What is the likelihood function: $f(d|x)$

- The likelihood function $f(d|x)$ is the probability of the observed data d for various values of the unknown parameters x .
- The likelihood is used after data are available to describe a plausibility of a parameter value x .
- The likelihood does not have to integrate to one.

Example of using Bayes' theorem



Example of using Bayes' theorem



Revert to inverse problem again

With Gaussian errors

$$\frac{dx}{dt} = 1 + q(t)$$

$$x(0) = 0 + a$$

$$x(1) = 2 + b$$

$$f(q(t)) \propto \exp \left\{ -\frac{1}{2} \frac{q^2(t)}{C_{qq}} \right\}$$

$$f(a) \propto \exp \left\{ -\frac{1}{2} \frac{a^2}{C_{aa}} \right\}$$

$$f(b) \propto \exp \left\{ -\frac{1}{2} \frac{b^2}{C_{bb}} \right\}$$

$$f(a) f \left(\int_0^1 q(t) dt \right) f(b) = \exp \left(-\frac{1}{2} J[x] \right)$$

$$\mathcal{J}[x] = C_{qq}^{-1} \int_0^1 \left(\frac{dx}{dt} - 1 \right)^2 dt + C_{aa}^{-1} (x(0) - 0)^2 + C_{bb}^{-1} (x(1) - 2)^2$$

Thus, Bayes' theorem leads to least-squares variational inverse problem for Gaussian error distributions.

Summary

- Bayes' theorem defines the “ultimate” data-assimilation problem.
- Impossible to solve in high dimensions.
- Gaussian approximation is key and leads to least-squares inverse problem.

Linear estimation theory and update equations

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Best Linear Unbiased Estimator

We have given a prediction x^f and a measurement d :

$$\begin{aligned} x^f &= x^t + p^f, & \overline{p^f} &= 0 & \overline{(p^f)^2} &= C_{xx}^f \\ d &= x^t + \epsilon, & \overline{\epsilon} &= 0 & \overline{(\epsilon)^2} &= C_{dd} & \overline{(\epsilon p^f)} &= 0 \end{aligned}$$

What is the **Best Linear Unbiased Estimator (BLUE)** of x ?

- x^f could be a Forecast or a First-guess solution.
- d is a measurement.

Original BLUE derivation from system control theory

Given

$$\begin{array}{lll}
 x^f = x^t + p^f, & \overline{p^f} = 0 & \overline{(p^f)^2} = C_{xx}^f \\
 x^a = x^t + p^a, & \overline{p^a} = 0 & \overline{(p^a)^2} = C_{xx}^a \\
 d = x^t + \epsilon, & \overline{\epsilon} = 0 & \overline{(\epsilon)^2} = C_{dd}
 \end{array}$$

A linear unbiased estimator is

$$\begin{aligned}
 x^a &= (1 - \alpha)x^f + \alpha d \\
 &= x^f + \alpha(d - x^f)
 \end{aligned}$$

Inserting gives

$$\begin{aligned}
 x^a &= x^t + p^a = x^t + p^f + \alpha(x^t + \epsilon - x^t - p^f) \\
 p^a &= p^f + \alpha(\epsilon - p^f)
 \end{aligned}$$

Derivation

$$\begin{aligned}
 \overline{(p^a)^2} &= C_{xx}^a = \overline{(p^f + \alpha(\epsilon - p^f))^2} \\
 &= \overline{(p^f)^2} + 2\alpha \overline{p^f(\epsilon - p^f)} + \alpha^2 \overline{\epsilon^2 - 2\epsilon p^f + (p^f)^2} \\
 &= C_{xx}^f - 2\alpha C_{xx}^f + \alpha^2 (C_{dd} + C_{xx}^f),
 \end{aligned}$$

Set derivative equal to zero

$$\frac{\partial C_{xx}^a}{\partial \alpha} = -2C_{xx}^f + 2\alpha(C_{dd} + C_{xx}^f) = 0.$$

to get

$$\alpha = \frac{C_{xx}^f}{C_{xx}^f + C_{dd}}$$

Derivation

The optimal (BLUE) solution is

$$x^a = x^f + \frac{C_{xx}^f}{C_{xx}^f + C_{dd}}(d - x^f)$$

Error estimate when inserting for α

$$C_{xx}^a = C_{xx}^f \left(1 - \frac{C_{xx}^f}{C_{xx}^f + C_{dd}} \right)$$

Derivation from Bayes' Theorem

Assuming a Gaussian prior and likelihood

$$f(x) = \exp \left\{ -\frac{1}{2}(x - x^f)C_{xx}^{-1}(x - x^f) \right\}$$
$$f(d|x) = \exp \left\{ -\frac{1}{2}(d - x)C_{dd}^{-1}(d - x) \right\}$$

From Bayes

$$f(x|d) \propto f(x)f(d|x)$$

By taking the logarithm we get the cost function

$$\mathcal{J}(x) = (x - x^f)C_{xx}^{-1}(x - x^f) + (d - x)C_{dd}^{-1}(d - x)$$

Derivation from Bayes' Theorem

Derivative of cost function set to zero

$$\begin{aligned}\frac{1}{2} \frac{\partial \mathcal{J}(x)}{\partial x} &= (x - x^f) C_{xx}^{-1} - (d - x) C_{dd}^{-1} \\ &= x(C_{xx}^{-1} + C_{dd}^{-1}) - x^f C_{xx}^{-1} - d C_{dd}^{-1} = 0\end{aligned}$$

Solve for x

$$\begin{aligned}x &= x^f \frac{C_{xx}^{-1}}{C_{xx}^{-1} + C_{dd}^{-1}} + d \frac{C_{dd}^{-1}}{C_{xx}^{-1} + C_{dd}^{-1}} \\ &= x^f \frac{C_{dd}}{C_{dd} + C_{xx}} + d \frac{C_{xx}}{C_{dd} + C_{xx}} \\ &= x^f + \frac{C_{xx}}{C_{xx} + C_{dd}} (d - x^f)\end{aligned}$$

$$\begin{aligned}&\times \frac{C_{dd} C_{xx}}{C_{dd} C_{xx}} \\ &+ x^f - \frac{C_{dd} + C_{xx}}{C_{dd} + C_{xx}} x^f\end{aligned}$$

Summary

- The BLUE is the optimal way of combining two linear estimates of a parameter.
- We can derive it from Bayes' formula when assuming Gaussian error statistics.

$$f(x|d) \propto f(x)f(d|x)$$

$$\mathcal{J}(x) = (x - x^f)C_{xx}^{-1}(x - x^f) + (d - x)C_{dd}^{-1}(d - x)$$

$$x^a = x^f + \frac{C_{xx}}{C_{xx} + C_{dd}}(d - x^f)$$

$$C_{xx}^a = C_{xx}^f \left(1 - \frac{C_{xx}^f}{C_{xx}^f + C_{dd}} \right)$$

BLUE in vector form

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Vector state and data

Given a state vector $\mathbf{x} \in \mathfrak{R}^n$ and a data vector $\mathbf{d} \in \mathfrak{R}^m$.

$$\begin{array}{lll} \mathbf{x}^f = \mathbf{x}^t + \mathbf{p} & \bar{\mathbf{p}} = 0 & \overline{\mathbf{p}\mathbf{p}^T} = \mathbf{C}_{xx} \\ \mathbf{d} = \mathbf{H}\mathbf{x}^t + \boldsymbol{\epsilon} & \bar{\boldsymbol{\epsilon}} = 0 & \overline{\boldsymbol{\epsilon}\boldsymbol{\epsilon}^T} = \mathbf{C}_{dd} \quad \overline{\mathbf{p}\boldsymbol{\epsilon}^T} = 0 \end{array}$$

where we define the measurement operator $\mathbf{H} \in \mathfrak{R}^{m \times n}$.

As an example consider the case with $m = 2$ and $n = 10$.

$$\begin{aligned} \mathbf{x}^f &= [x_1, x_2, x_3, \dots, x_{10}]^T \\ \mathbf{d} &= [d_1, d_2]^T \end{aligned}$$

with

$$\mathbf{H} = \begin{bmatrix} 0 & 0 & 0 & 1 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 1 & 0 & 0 & 0 \end{bmatrix}$$

implying that $d_1 = x_4^t + \epsilon_1$ and $d_2 = x_7^t + \epsilon_2$.

Start from Bayes' Theorem

Bayes for the state vector \mathbf{x} given the measurement vector \mathbf{d}

$$f(\mathbf{x}|\mathbf{d}) \propto f(\mathbf{x})f(\mathbf{d}|\mathbf{x})$$

Assume Gaussian prior and measurement errors

$$f(\mathbf{x}|\mathbf{d}) \propto \exp -\frac{1}{2} \left\{ (\mathbf{x} - \mathbf{x}^f)^T \mathbf{C}_{xx}^{-1} (\mathbf{x} - \mathbf{x}^f) + (\mathbf{d} - \mathbf{H}\mathbf{x})^T \mathbf{C}_{dd}^{-1} (\mathbf{d} - \mathbf{H}\mathbf{x}) \right\}$$

Maximizing $f(\mathbf{x}|\mathbf{d})$ identical to minimizing

$$\mathcal{J}(\mathbf{x}) = (\mathbf{x} - \mathbf{x}^f)^T \mathbf{C}_{xx}^{-1} (\mathbf{x} - \mathbf{x}^f) + (\mathbf{d} - \mathbf{H}\mathbf{x})^T \mathbf{C}_{dd}^{-1} (\mathbf{d} - \mathbf{H}\mathbf{x})$$

Minimum is defined by zero-gradient of cost function

$$\begin{aligned}\frac{1}{2}\nabla\mathcal{J}(\mathbf{x}) &= \mathbf{C}_{xx}^{-1}(\mathbf{x} - \mathbf{x}^f) - \mathbf{H}^T\mathbf{C}_{dd}^{-1}(\mathbf{d} - \mathbf{H}\mathbf{x}) \\ &= \mathbf{C}_{xx}^{-1}\mathbf{x} - \mathbf{C}_{xx}^{-1}\mathbf{x}^f - \mathbf{H}^T\mathbf{C}_{dd}^{-1}\mathbf{d} + \mathbf{H}^T\mathbf{C}_{dd}^{-1}\mathbf{H}\mathbf{x} = 0\end{aligned}$$

$$\left(\mathbf{C}_{xx}^{-1} + \mathbf{H}^T\mathbf{C}_{dd}^{-1}\mathbf{H}\right)\mathbf{x} = \mathbf{C}_{xx}^{-1}\mathbf{x}^f + \mathbf{H}^T\mathbf{C}_{dd}^{-1}\mathbf{d}$$

$$\mathbf{x} = \left(\mathbf{C}_{xx}^{-1} + \mathbf{H}^T\mathbf{C}_{dd}^{-1}\mathbf{H}\right)^{-1} \left(\mathbf{C}_{xx}^{-1}\mathbf{x}^f + \mathbf{H}^T\mathbf{C}_{dd}^{-1}\mathbf{d}\right)$$

Minimizing solution

State space formulation

$$\mathbf{x} = \left(\mathbf{C}_{xx}^{-1} + \mathbf{H}^T \mathbf{C}_{dd}^{-1} \mathbf{H} \right)^{-1} \mathbf{C}_{xx}^{-1} \mathbf{x}^f + \left(\mathbf{C}_{xx}^{-1} + \mathbf{H}^T \mathbf{C}_{dd}^{-1} \mathbf{H} \right)^{-1} \mathbf{H}^T \mathbf{C}_{dd}^{-1} \mathbf{d}$$

Using the following two lemmas (from the Woodbury matrix identity):

$$\begin{aligned} \left(\mathbf{C}_{xx}^{-1} + \mathbf{H}^T \mathbf{C}_{dd}^{-1} \mathbf{H} \right)^{-1} \mathbf{C}_{xx}^{-1} &= \mathbf{I} - \mathbf{C}_{xx} \mathbf{H}^T (\mathbf{H} \mathbf{C}_{xx} \mathbf{H}^T + \mathbf{C}_{dd})^{-1} \mathbf{H} \\ \left(\mathbf{C}_{xx}^{-1} + \mathbf{H}^T \mathbf{C}_{dd}^{-1} \mathbf{H} \right)^{-1} \mathbf{H}^T \mathbf{C}_{dd}^{-1} &= \mathbf{C}_{xx} \mathbf{H}^T (\mathbf{H} \mathbf{C}_{xx} \mathbf{H}^T + \mathbf{C}_{dd})^{-1} \end{aligned}$$

we obtain the observation space formulation

$$\mathbf{x} = \mathbf{x}^f + \mathbf{C}_{xx} \mathbf{H}^T (\mathbf{H} \mathbf{C}_{xx} \mathbf{H}^T + \mathbf{C}_{dd})^{-1} (\mathbf{d} - \mathbf{H} \mathbf{x}^f)$$

* What is the update error covariance

The error variance for the update is defined as

$$\mathbf{C}_{xx}^a = \overline{(\mathbf{x}^t - \mathbf{x}^a)(\mathbf{x}^t - \mathbf{x}^a)^T}$$

Let us define for simplicity

$$\mathbf{C} = \mathbf{H}\mathbf{C}_{xx}^f\mathbf{H}^T + \mathbf{C}_{dd}$$

$$\mathbf{R} = \mathbf{H}\mathbf{C}_{xx}^f$$

$$\mathbf{h} = \mathbf{d} - \mathbf{H}\mathbf{x}^f$$

$$= \mathbf{H}\mathbf{x}^t + \boldsymbol{\epsilon} - \mathbf{H}\mathbf{x}^f = \mathbf{H}(\mathbf{x}^t - \mathbf{x}^f) + \boldsymbol{\epsilon}$$

And we can write

$$\mathbf{x}^a = \mathbf{x}^f + \mathbf{R}^T\mathbf{C}^{-1}\mathbf{h}$$

* Update error covariance

$$\begin{aligned}
 \mathbf{C}_{xx}^a &= \overline{(\mathbf{x}^t - \mathbf{x}^a)(\mathbf{x}^t - \mathbf{x}^a)^T} \\
 &= \overline{(\mathbf{x}^t - \mathbf{x}^f - \mathbf{R}^T \mathbf{C}^{-1} \mathbf{h})(\mathbf{x}^t - \mathbf{x}^f - \mathbf{R}^T \mathbf{C}^{-1} \mathbf{h})^T} \\
 &= \overline{(\mathbf{x}^t - \mathbf{x}^f)(\mathbf{x}^t - \mathbf{x}^f)^T} & (\rightarrow \mathbf{C}_{xx}^f) \\
 &\quad - 2 \overline{(\mathbf{x}^t - \mathbf{x}^f)(\mathbf{R}^T \mathbf{C}^{-1} \mathbf{h})^T} \\
 &\quad + \overline{(\mathbf{R}^T \mathbf{C}^{-1} \mathbf{h})(\mathbf{R}^T \mathbf{C}^{-1} \mathbf{h})^T}
 \end{aligned}$$

* Second term

$$\begin{aligned}
 & - \overline{2(\mathbf{x}^t - \mathbf{x}^f)(\mathbf{R}^T \mathbf{C}^{-1} \mathbf{h})^T} \\
 & = \overline{-2(\mathbf{x}^t - \mathbf{x}^f) \mathbf{h}^T \mathbf{C}^{-1} \mathbf{R}} \\
 & = \overline{-2(\mathbf{x}^t - \mathbf{x}^f)(\mathbf{H}(\mathbf{x}^t - \mathbf{x}^f) + \boldsymbol{\epsilon})^T \mathbf{C}^{-1} \mathbf{R}} \\
 & = \overline{-2(\mathbf{x}^t - \mathbf{x}^f)(\mathbf{x}^t - \mathbf{x}^f)^T \mathbf{H}^T \mathbf{C}^{-1} \mathbf{R}} \\
 & = -2\mathbf{C}_{xx}^f \mathbf{H}^T \mathbf{C}^{-1} \mathbf{R} \\
 & = -2\mathbf{R}^T \mathbf{C}^{-1} \mathbf{R}
 \end{aligned}$$

* Third term

$$\begin{aligned}
 & \overline{(R^T C^{-1} h)(R^T C^{-1} h)^T} \\
 &= \overline{R^T C^{-1} h h^T C^{-1} R} \\
 &= \overline{R^T C^{-1} (H(x^t - x^f) + \epsilon)(H(x^t - x^f) + \epsilon)^T C^{-1} R} \\
 &= R^T C^{-1} C C^{-1} R \\
 &= R^T C^{-1} R
 \end{aligned}$$

* Update error covariance

$$\begin{aligned}C_{xx}^a &= C_{xx}^f - 2R^T C^{-1} R + R^T C^{-1} R \\&= C_{xx}^f - R^T C^{-1} R \\&= C_{xx}^f - C_{xx}^f H^T \left(H C_{xx}^f H^T + C_{dd} \right)^{-1} H C_{xx}^f\end{aligned}$$

Minimizing solution in the Gaussian case

Kalman filter update equations

$$\mathbf{x}^a = \mathbf{x}^f + \mathbf{C}_{xx}^f \mathbf{H}^T (\mathbf{H} \mathbf{C}_{xx}^f \mathbf{H}^T + \mathbf{C}_{dd})^{-1} (\mathbf{d} - \mathbf{H} \mathbf{x}^f)$$

$$\mathbf{C}_{xx}^a = \mathbf{C}_{xx}^f - \mathbf{C}_{xx}^f \mathbf{H}^T (\mathbf{H} \mathbf{C}_{xx}^f \mathbf{H}^T + \mathbf{C}_{dd})^{-1} \mathbf{H} \mathbf{C}_{xx}^f.$$

Kalman gain matrix

The control theory community defines the Kalman Gain Matrix

$$\mathbf{K} = \mathbf{C}_{xx}^f \mathbf{H}^T (\mathbf{H} \mathbf{C}_{xx}^f \mathbf{H}^T + \mathbf{C}_{dd})^{-1}$$

to obtain a simpler expression of the update:

$$\mathbf{x}^a = \mathbf{x}^f + \mathbf{K}(\mathbf{d} - \mathbf{H}\mathbf{x}^f)$$

$$\mathbf{C}_{xx}^a = \mathbf{C}_{xx}^f - \mathbf{K} \mathbf{H} \mathbf{C}_{xx}^f$$

“Representer” formulation

$$\begin{aligned}\mathbf{x}^a &= \mathbf{x}^f + \mathbf{C}_{xx}^f \mathbf{H}^T (\mathbf{H} \mathbf{C}_{xx}^f \mathbf{H}^T + \mathbf{C}_{dd})^{-1} (\mathbf{d} - \mathbf{H} \mathbf{x}^f) \\ &= \mathbf{x}^f + \mathbf{R}^T \mathbf{C}^{-1} \mathbf{h}\end{aligned}$$

Define $\mathbf{b} = \mathbf{C}^{-1} \mathbf{h}$ as the solution of the linear system

$$(\mathbf{H} \mathbf{C}_{xx}^f \mathbf{H}^T + \mathbf{C}_{dd}) \mathbf{b} = (\mathbf{d} - \mathbf{H} \mathbf{x}^f)$$

So we can write the update as

$$\begin{aligned}\mathbf{x}^a &= \mathbf{x}^f + \mathbf{R}^T \mathbf{b} \\ \mathbf{C}_{xx}^a &= \mathbf{C}_{xx}^f - \mathbf{R}^T \mathbf{C}^{-1} \mathbf{R}\end{aligned}$$

Note that the covariance update is independent of the actual measurement values.

Summary

- Introduced the concept of a measurement operator.
- Derived Kalman filter update equations in the vector case.
- We update both the state estimate and its error covariance.
- We defined the Kalman Gain.
- We heard about the Representer formulation.
- Original minimization problem is of dimension n .
- The KF update reduces the dimension to $m \ll n$.

Sequential and Smoother solutions from Bayes

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Starting from Bayes (again)

$$f(\mathbf{x}|\mathbf{d}) \propto f(\mathbf{x})f(\mathbf{d}|\mathbf{x})$$

- $f(\mathbf{x})$ is the density for the model state in space and time.
- $f(\mathbf{d}|\mathbf{x})$ is the measurement likelihood.

Assume that the model is 1st order Markov process

$$x_i = g(x_{i-1}) + q_i, \quad q_i \leftarrow f(x_i|x_{i-1})$$

- Since the solution x_i only depends on x_{i-1} we can write

$$f(\mathbf{x}) = f(x_0, x_1, \dots, x_k) = f(x_0) \prod_{i=1}^k f(x_i|x_{i-1}).$$

- Valid for most numerical prediction models.

Assume independent data in time

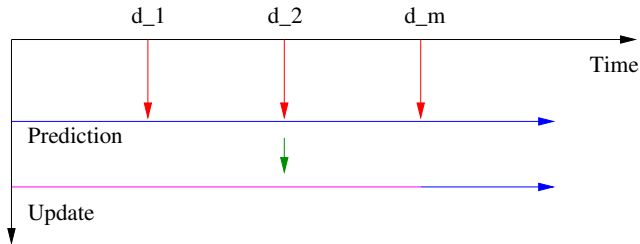
- When measurement errors are uncorrelated in time

$$f(\mathbf{d}|\mathbf{x}) = \prod_{j=1}^m f(d_j|x_j).$$

- Allows for recursive processing of measurements in time

Bayes' becomes

$$f(x_0, x_1, \dots, x_k, |\mathbf{d}) \propto f(x_0) \prod_{i=1}^k f(x_i | x_{i-1}) \prod_{j=1}^m f(d_j | x_j)$$



Rewrite as

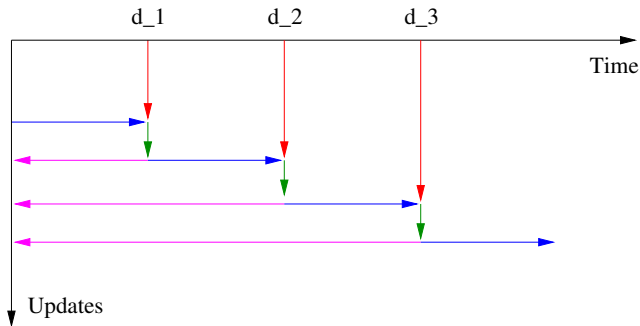
$$\begin{aligned}
 f(\mathbf{x}|\mathbf{d}) &\propto f(\mathbf{x})f(\mathbf{d}|\mathbf{x}) = \\
 &f(x_0) \\
 &f(x_1|x_0)f(d_1|x_1) \\
 &f(x_2|x_1)f(d_2|x_2) \\
 &\vdots \\
 &f(x_k|x_{k-1})f(d_m|x_m) \\
 &f(x_{k+1}|x_k)
 \end{aligned}$$

Recursive “smoother” updates

The recursive idea: "Today's posterior is tomorrow's prior"

$$f(x_0, x_1 | d_1) = f(x_0) f(x_1 | x_0) f(d_1 | x_1)$$

$$f(x_0, x_1, x_2 | d_1, d_2) = f(x_0, x_1 | d_1) f(x_2 | x_1) f(d_2 | x_2)$$

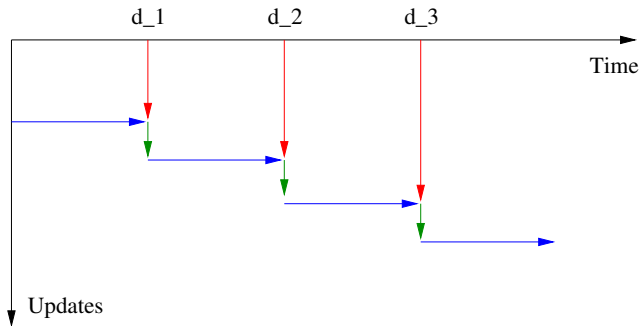


Recursive “filter” updates

Marginal pdfs

$$f(x_1|d_1) = \int_{x_0} f(x_0)f(x_1|x_0)dx_0 f(d_1|x_1) = f(x_1)f(d_1|x_1)$$

$$f(x_2|d_1, d_2) = \int_{x_1} f(x_1|d_1)f(x_2|x_1)dx_1 f(d_2|x_2) = f(x_2|d_1)f(d_2|x_2)$$



Summary

- Assume model is Markov process.
- Assume measurements are independent in time. (Not generally true.)
- We can process independent measurements sequentially in time.
- The solution of one sub-problem is prior for the next one.

Kalman Filter and Extended Kalman Filter

Geir Evensen



Kalman Filter

- Recursively updates model state and uncertainty.
- Variance minimizing update step.
- Estimate improves and uncertainty reduces at each update.

Error propagation

Derivation for linear scalar model

- Evolution of true state

$$x_k^t = Gx_{k-1}^t + q_{k-1}$$

- The model state evolves according to

$$x_k^f = Gx_{k-1}^a$$

- Difference is

$$x_k^t - x_k^f = G(x_{k-1}^t - x_{k-1}^a) + q_{k-1}$$

Predicting the errors

Square difference and take the expectation

$$\overline{(x_k^t - x_k^f)^2} = \overline{G(x_{k-1}^t - x_{k-1}^a)^2 G} + \overline{q_{k-1}^2} + 2\overline{G(x_{k-1}^t - x_{k-1}^a)q_{k-1}}$$

Error covariance evolution equation

$$C_{xx}^f(t_k) = GC_{xx}^a(t_{k-1})G + C_{qq}(t_{k-1}).$$

- Model errors uncorrelated with state error.

The full Kalman Filter (vector form)

Linear model prediction

$$\mathbf{x}_k^f = \mathbf{G}\mathbf{x}_{k-1}^a$$

Error covariance prediction

$$\mathbf{C}_{xx}^f(t_k) = \mathbf{G}\mathbf{C}_{xx}^a(t_{k-1})\mathbf{G}^T + \mathbf{C}_{qq}(t_{k-1}).$$

Analysis update (skipped t_k index)

$$\mathbf{x}_k^a = \mathbf{x}_k^f + \mathbf{C}_{xx}^f \mathbf{H}^T (\mathbf{H}\mathbf{C}_{xx}^f \mathbf{H}^T + \mathbf{C}_{dd})^{-1} (\mathbf{d}_k - \mathbf{H}\mathbf{x}_k^f)$$

Error covariance update (for each t_k)

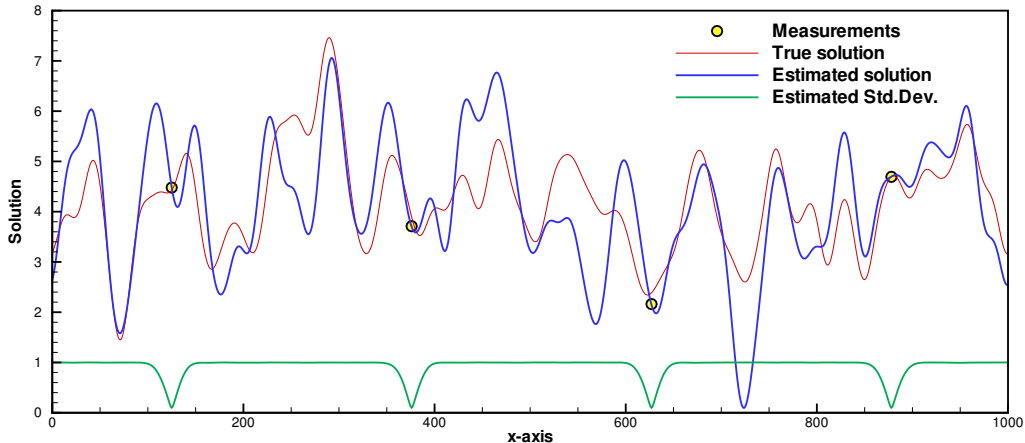
$$\mathbf{C}_{xx}^a = \mathbf{C}_{xx}^f - \mathbf{C}_{xx}^f \mathbf{H}^T \left(\mathbf{H}\mathbf{C}_{xx}^f \mathbf{H}^T + \mathbf{C}_{dd} \right)^{-1} \mathbf{H}\mathbf{C}_{xx}^f.$$

Kalman Filter Example

- Linear advection equation
- Periodic domain
- Random reference solution (truth).
- First guess is reference plus random perturbation.
- Initial variance is 1.0 m^2
- Four measurements every 5 time units.
- Measurement variance is 0.01 m^2 .
- Cases without and including system noise of 0.0004 m^2 .

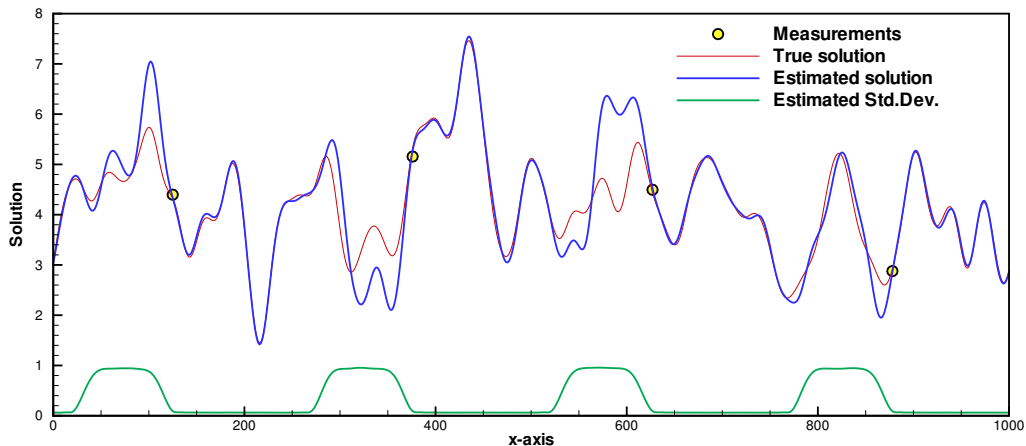
Kalman Filter example: Perfect model

Solution after first update $t = 5.0$



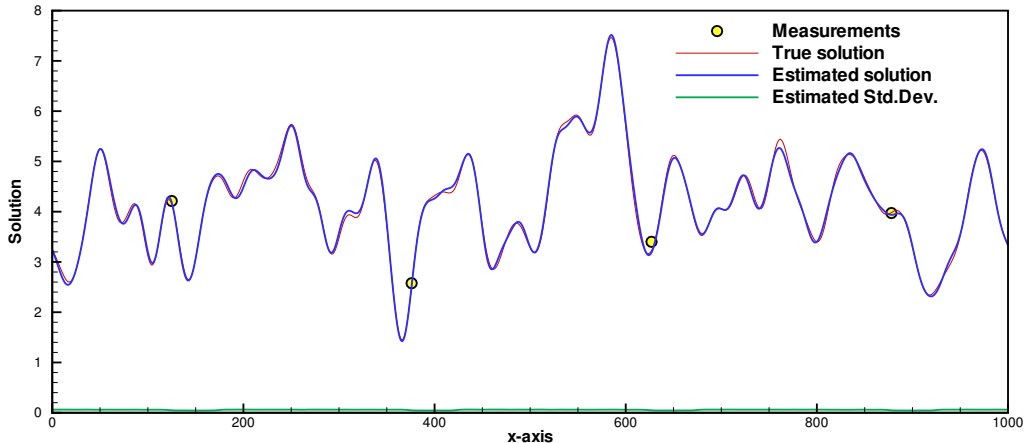
Kalman Filter example: Perfect model

Solution after 30 updates $t = 150.0$



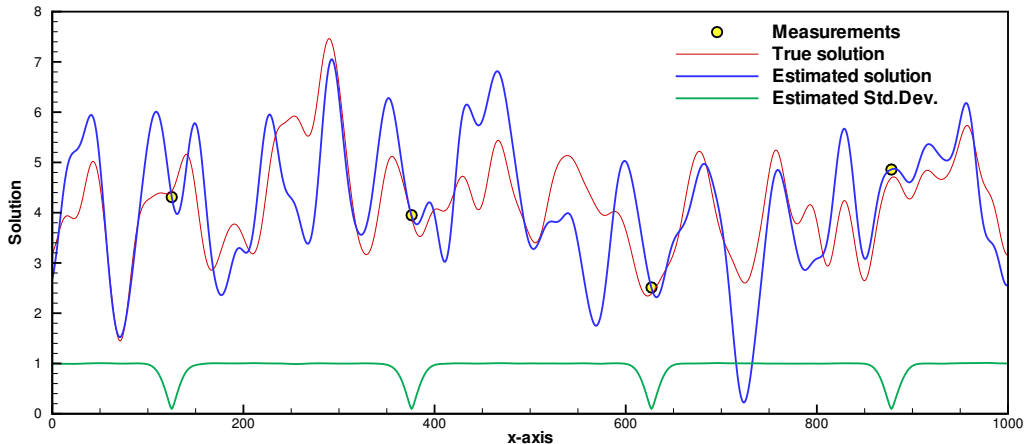
Kalman Filter example: Perfect model

Solution after 60 updates $t = 300.0$



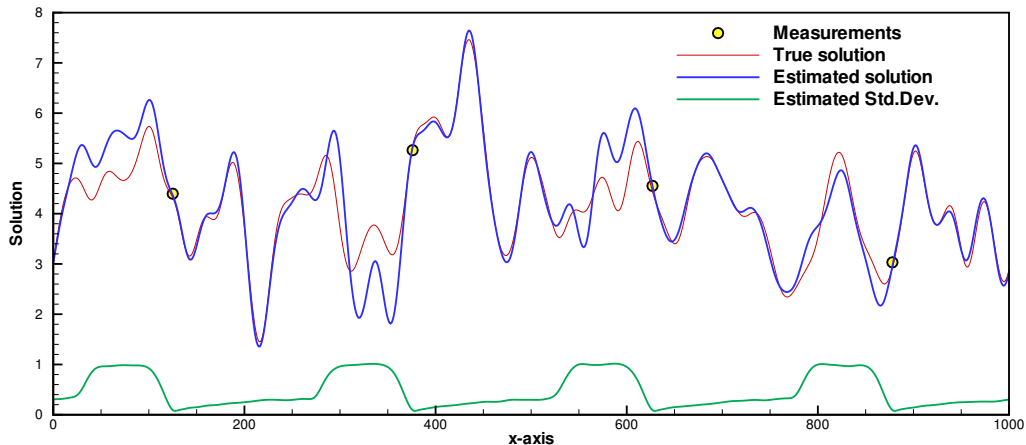
Kalman Filter example: with model error

Solution after first update $t = 5.0$



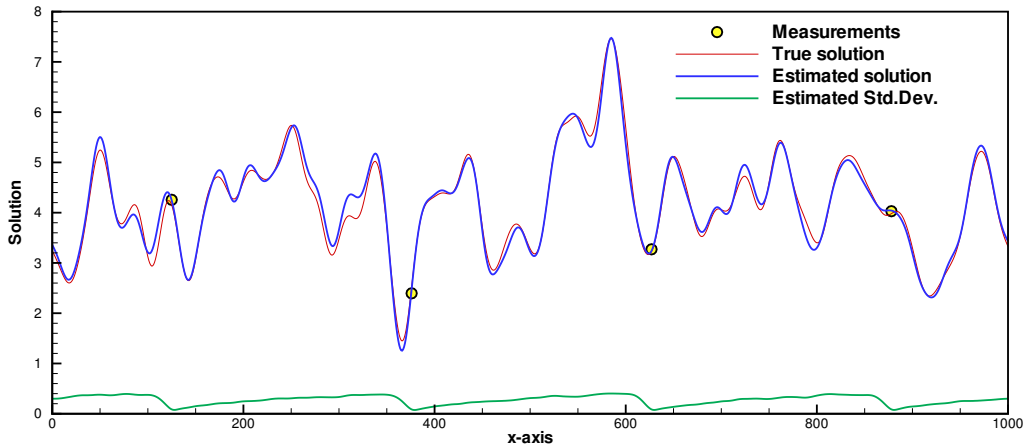
Kalman Filter example: with model error

Solution after 30 updates $t = 150.0$



Kalman Filter example: with model error

Solution after 60 updates $t = 300.0$



Inverse problem revisited

What is the KF solution for the linear inverse problem?

$$\frac{dx}{dt} = 1$$

$$x(0) = 0$$

$$d = x(1) = 2$$

KF solution

Solve initial value problem

$$\frac{dx}{dt} = 1$$

$$x(0) = 0$$

$$\implies x^f(t) = t$$

Predicted error variance

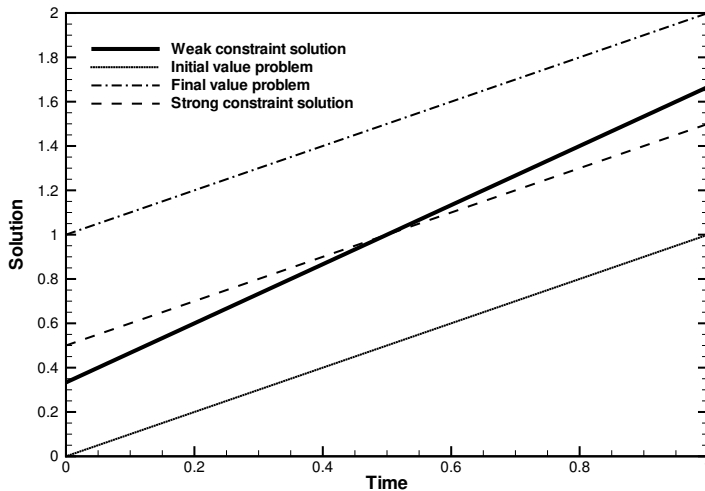
$$C_{xx}^f(1) = C_{xx}^a(0) + C_{qq} = 2C_0$$

Update at $t = 1$ is

$$x^a = x^f + \frac{C_{xx}^f}{C_{dd} + C_{xx}^f}(d - x^f) = 1 + \frac{2C_0}{C_0 + 2C_0}(2 - 1) = 5/3$$

KF solution at final time equals the weak-constraint variational solution

Inverse problem revisited



Nonlinear dynamics

Derivation of Extended Kalman Filter (EKF)

$$\begin{aligned}\mathbf{x}_k^t &= \mathbf{g}(\mathbf{x}_{k-1}^t) + \mathbf{q}_{k-1}, \\ \mathbf{x}_k^f &= \mathbf{g}(\mathbf{x}_{k-1}^a), \\ \mathbf{x}_k^t - \mathbf{x}_k^f &= \mathbf{g}(\mathbf{x}_{k-1}^t) - \mathbf{g}(\mathbf{x}_{k-1}^a) + \mathbf{q}_{k-1}.\end{aligned}$$

Use Taylor expansion

$$\mathbf{g}(\mathbf{x}_{k-1}^t) = \mathbf{g}(\mathbf{x}_{k-1}^a) + \mathbf{G}(\mathbf{x}_{k-1}^a)(\mathbf{x}_{k-1}^t - \mathbf{x}_{k-1}^a) + \frac{1}{2}\mathbf{H}(\mathbf{x}_{k-1}^a)(\mathbf{x}_{k-1}^t - \mathbf{x}_{k-1}^a)^2 + \dots$$

EKF: Derivation

Difference becomes

$$\mathbf{x}_k^t - \mathbf{x}_k^f = \mathbf{G}(\mathbf{x}_{k-1}^a)(\mathbf{x}_{k-1}^t - \mathbf{x}_{k-1}^a) + \frac{1}{2}\mathbf{H}(\mathbf{x}_{k-1}^a)(\mathbf{x}_{k-1}^t - \mathbf{x}_{k-1}^a)^2 + \cdots + \mathbf{q}_{k-1}.$$

By squaring and taking the expectation we get

$$\begin{aligned} \mathbf{C}_{xx}^f(t_k) &= \overline{(\mathbf{x}_k^t - \mathbf{x}_k^f)(\mathbf{x}_k^t - \mathbf{x}_k^f)^T} \\ &= \mathbf{G}(\mathbf{x}_{k-1}^a) \overline{(\mathbf{x}_{k-1}^t - \mathbf{x}_{k-1}^a)(\mathbf{x}_{k-1}^t - \mathbf{x}_{k-1}^a)^T} (\mathbf{G}(\mathbf{x}_{k-1}^a))^T + \cdots + \mathbf{C}_{qq}(t_{k-1}) \\ &= \mathbf{G}(\mathbf{x}_{k-1}^a) \mathbf{C}_{xx}^a(t_{k-1}) (\mathbf{G}(\mathbf{x}_{k-1}^a))^T + \cdots + \mathbf{C}_{qq}(t_{k-1}). \end{aligned}$$

EKF: Error evolution

Close by discarding higher order moments to get

$$\begin{aligned}\mathbf{x}_k^f &= \mathbf{g}(\mathbf{x}_{k-1}^a), \\ \mathbf{C}_{xx}^f(t_k) &\simeq \mathbf{G}_{k-1} \mathbf{C}_{xx}^a(t_{k-1}) \mathbf{G}_{k-1}^T + \mathbf{C}_{qq}(t_{k-1}),\end{aligned}$$

together with standard analysis equations.

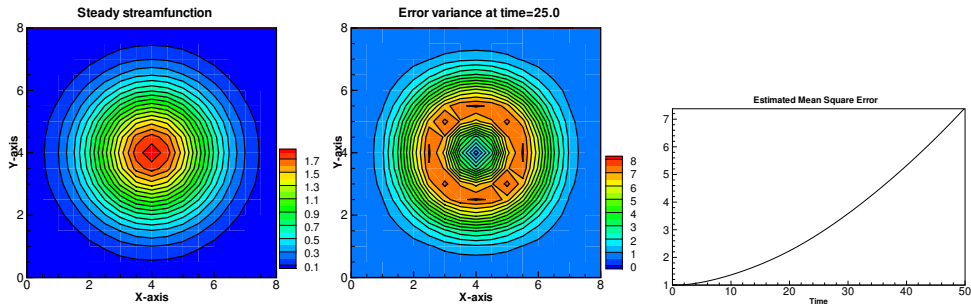
Example of Extended KF

Nonlinear quasi-geostrophic model

- Steady stream function solution.
- Curved and sheared flow.
- Supports instability.
- Initial variance is 1.0.

Example of Extended KF

Results (from Evensen 1992):



- Linear closure approximation not valid!
- Leads to linear instability and exponential error growth.

EKF: Summary

KF is optimal linear filter method!

- Storage of $O(n^2)$ elements.
- Integration of $2n$ models.

EKF applies closure approximation in error covariance equation.

- Requires implementation of tangent linear model.
- Too simple closure may lead to linear instabilities.
- Does not work for strongly nonlinear models.

We need a better alternative!

Ensemble Kalman Filter

A Monte Carlo alternative to KF and EKF

Geir Evensen



The error covariance matrix

Define ensemble covariances around the ensemble mean

$$\begin{aligned} \mathbf{C}_{xx}^f &\simeq \overline{\mathbf{C}}_{xx}^f = \overline{(\mathbf{x}^f - \overline{\mathbf{x}^f})(\mathbf{x}^f - \overline{\mathbf{x}^f})^T} \\ \mathbf{C}_{xx}^a &\simeq \overline{\mathbf{C}}_{xx}^a = \overline{(\mathbf{x}^a - \overline{\mathbf{x}^a})(\mathbf{x}^a - \overline{\mathbf{x}^a})^T} \end{aligned}$$

- The ensemble mean $\overline{\mathbf{x}}$ is the best-guess.
- The ensemble spread defines the error variance.
- The ensemble smoothness defines the error covariance.

Dynamical evolution of error statistics

- Ensemble of models (particles) defines probability $f(\mathbf{x})$.
- Ensemble members evolve according to the model dynamics.

$$d\mathbf{x} = \mathbf{g}(\mathbf{x})dt + d\mathbf{q}.$$

- Probability density evolve according to Kolmogorov's equation.

$$\frac{\partial f}{\partial t} + \sum_i \frac{\partial (g_i f)}{\partial x_i} = \frac{1}{2} \sum_{i,j} C_{ij} \frac{\partial^2 f}{\partial x_i \partial x_j}.$$

- Fundamental equation for evolution of error statistics.
- Integrating a large ensemble of stochastic models is a MC method for solving Kolmogorov's equation

Analysis scheme (1)

Define the model-forecast error-covariance matrix

$$\mathbf{C}_{xx}^f \simeq \overline{\mathbf{C}}_{xx}^f = \overline{(\mathbf{x}^f - \overline{\mathbf{x}^f})(\mathbf{x}^f - \overline{\mathbf{x}^f})^T}.$$

and the measurement error-covariance matrix

$$\mathbf{C}_{dd} \simeq \overline{\mathbf{C}}_{dd} = \overline{\boldsymbol{\epsilon}\boldsymbol{\epsilon}^T}.$$

Create an ensemble of perturbed observations

$$\mathbf{d}_j = \mathbf{d} + \boldsymbol{\epsilon}_j,$$

where \mathbf{d} is the vector of observed values and $\boldsymbol{\epsilon}_j$, is a vector of observation noise.

Analysis scheme (2)

Update each ensemble member according to

$$\begin{aligned}\mathbf{x}_j^a &= \mathbf{x}_j^f + \bar{\mathbf{C}}_{xx}^f \mathbf{H}^T \left(\mathbf{H} \bar{\mathbf{C}}_{xx}^f \mathbf{H}^T + \bar{\mathbf{C}}_{dd} \right)^{-1} \left(d_j - \mathbf{H} \mathbf{x}_j^f \right) \\ &= \mathbf{x}_j^f + \bar{\mathbf{K}} \left(d_j - \mathbf{H} \mathbf{x}_j^f \right)\end{aligned}$$

Thus, the update of the mean becomes

$$\begin{aligned}\bar{\mathbf{x}}^a &= \bar{\mathbf{x}}^f + \bar{\mathbf{C}}_{xx}^f \mathbf{H}^T \left(\mathbf{H} \bar{\mathbf{C}}_{xx}^f \mathbf{H}^T + \bar{\mathbf{C}}_{dd} \right)^{-1} \left(d - \mathbf{H} \bar{\mathbf{x}}^f \right) \\ &= \bar{\mathbf{x}}^f + \bar{\mathbf{K}} \left(d - \mathbf{H} \bar{\mathbf{x}}^f \right)\end{aligned}$$

Analysis scheme (3)

The error covariance update then becomes

$$\begin{aligned}
 \bar{\mathbf{C}}_{xx}^a &= \overline{(\mathbf{x}^a - \bar{\mathbf{x}}^a)(\mathbf{x}^a - \bar{\mathbf{x}}^a)^T} \\
 &= \overline{\left((\mathbf{I} - \bar{\mathbf{K}}\mathbf{H})(\mathbf{x}^f - \bar{\mathbf{x}}^f) + \bar{\mathbf{K}}(\mathbf{d} - \bar{\mathbf{d}}) \right) \left(\dots \right)^T} \\
 &= (\mathbf{I} - \bar{\mathbf{K}}\mathbf{H})\bar{\mathbf{C}}_{xx}^f(\mathbf{I} - \mathbf{H}^T\bar{\mathbf{K}}^T) + \bar{\mathbf{K}}\bar{\mathbf{C}}_{dd}\bar{\mathbf{K}}^T \\
 &= \bar{\mathbf{C}}_{xx}^f - \bar{\mathbf{K}}\mathbf{H}\bar{\mathbf{C}}_{xx}^f - \bar{\mathbf{C}}_{xx}^f\mathbf{H}^T\bar{\mathbf{K}}^T + \bar{\mathbf{K}}(\mathbf{H}\bar{\mathbf{C}}_{xx}^f\mathbf{H}^T + \bar{\mathbf{C}}_{dd})\bar{\mathbf{K}}^T \\
 &= (\mathbf{I} - \bar{\mathbf{K}}\mathbf{H})\bar{\mathbf{C}}_{xx}^f \\
 &= \bar{\mathbf{C}}_{xx}^f - \bar{\mathbf{C}}_{xx}^f\mathbf{H}^T \left(\mathbf{H}\bar{\mathbf{C}}_{xx}^f\mathbf{H}^T + \bar{\mathbf{C}}_{dd} \right)^{-1} \mathbf{H}\bar{\mathbf{C}}_{xx}^f
 \end{aligned}$$

Note that we need to perturb observations to have $\bar{\mathbf{C}}_{dd} = \overline{(\mathbf{d} - \bar{\mathbf{d}})(\mathbf{d} - \bar{\mathbf{d}})^T}$ (Burgers et al., 1998)

Ensemble Kalman Filter (EnKF)

- Represents error statistics using an ensemble of model states.
- Evolves error statistics by ensemble integrations.
- “Variance minimizing” analysis scheme operating on the ensemble.

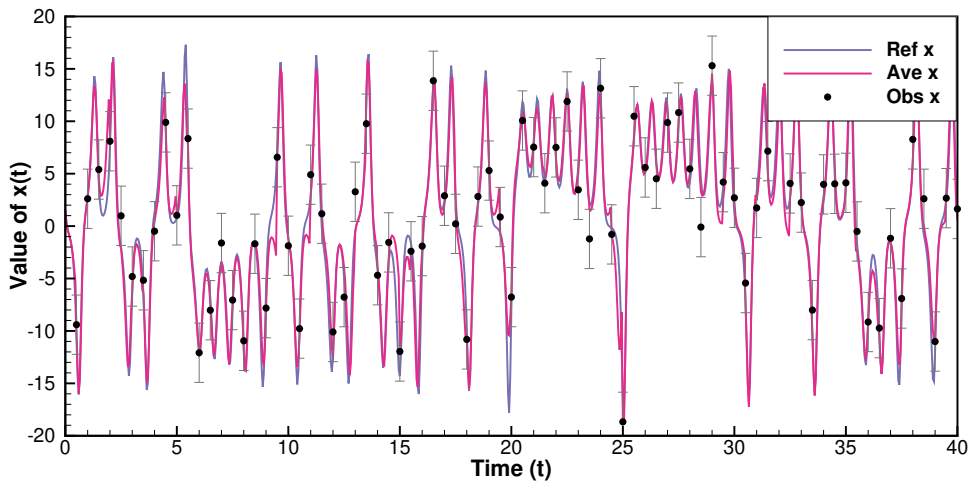


- Monte Carlo, low rank, ensemble subspace method.
- Linear model: EnKF converges to the KF with large ensemble size.
- Fully nonlinear error evolution, contrary to EKF.
- Assumption of Gaussian statistics in analysis scheme.

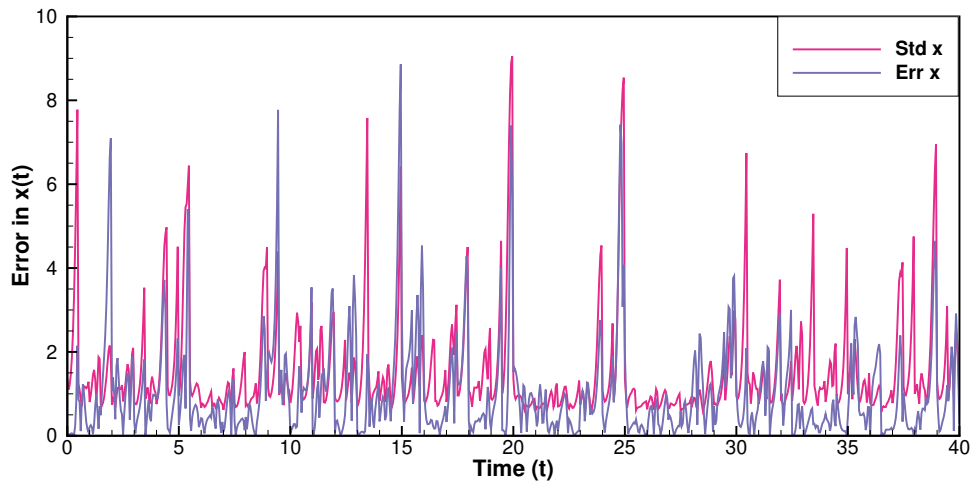
Example: Lorenz model

- Application with the chaotic Lorenz model.
- Illustrates properties with highly nonlinear dynamical models.
- From Evensen (1997), MWR.

EnKF solution



EnKF error variance



Analysis equation (1)

- Define the ensemble matrix

$$\mathbf{A} = (\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_N) \in \mathbb{R}^{n \times N}.$$

- The ensemble mean is (defining $\mathbf{1}_N \in \mathbb{R}^{N \times N} \equiv 1/N$)

$$\bar{\mathbf{A}} = \mathbf{A} \mathbf{1}_N.$$

- The ensemble perturbations become

$$\mathbf{A}' = \mathbf{A} - \bar{\mathbf{A}} = \mathbf{A}(\mathbf{I} - \mathbf{1}_N).$$

- The ensemble covariance matrix $\bar{\mathbf{C}}_{xx} \in \mathbb{R}^{n \times n}$ becomes

$$\bar{\mathbf{C}}_{xx} = \frac{\mathbf{A}'(\mathbf{A}')^T}{N-1}.$$

Analysis equation (2)

- Given a vector of measurements $\mathbf{d} \in \mathbb{R}^m$, define

$$\mathbf{d}_j = \mathbf{d} + \boldsymbol{\epsilon}_j, \quad j = 1, \dots, N,$$

stored in

$$\mathbf{D} = (\mathbf{d}_1, \mathbf{d}_2, \dots, \mathbf{d}_N) \in \mathbb{R}^{m \times N}.$$

- The ensemble perturbations are stored in

$$\mathbf{E} = (\boldsymbol{\epsilon}_1, \boldsymbol{\epsilon}_2, \dots, \boldsymbol{\epsilon}_N) \in \mathbb{R}^{m \times N},$$

thus, the measurement error covariance matrix becomes

$$\overline{\mathbf{C}}_{dd} = \frac{\mathbf{E}\mathbf{E}^T}{N-1}.$$

Analysis equation (3)

The analysis equation

$$\mathbf{A}^a = \mathbf{A} + \overline{\mathbf{C}}_{xx} \mathbf{H}^T \left(\mathbf{H} \overline{\mathbf{C}}_{xx} \mathbf{H}^T + \overline{\mathbf{C}}_{dd} \right)^{-1} (\mathbf{D} - \mathbf{H}\mathbf{A}).$$

can now be written

$$\mathbf{A}^a = \mathbf{A} + \mathbf{A}' (\mathbf{H}\mathbf{A}')^T \left((\mathbf{H}\mathbf{A}') (\mathbf{H}\mathbf{A}')^T + \mathbf{E}\mathbf{E}^T \right)^{-1} (\mathbf{D} - \mathbf{H}\mathbf{A}).$$

The update is expressed entirely in terms of the ensemble

Define $\mathbf{S} = \mathbf{H}\mathbf{A}'$

$$\mathbf{A}^a = \mathbf{A} + \mathbf{A}' \mathbf{S}^T \left(\mathbf{S}\mathbf{S}^T + \mathbf{E}\mathbf{E}^T \right)^{-1} (\mathbf{D} - \mathbf{H}\mathbf{A}).$$

Analysis equation (4)

Define $\mathbf{C} = \mathbf{S}\mathbf{S}^T + \mathbf{E}\mathbf{E}^T$ and the innovations $\mathbf{D}' = \mathbf{D} - \mathbf{H}\mathbf{A}$.

$$\begin{aligned}\mathbf{A}^a &= \mathbf{A} + \mathbf{A}'\mathbf{S}^T \left(\mathbf{S}\mathbf{S}^T + \mathbf{E}\mathbf{E}^T \right)^{-1} (\mathbf{D} - \mathbf{H}\mathbf{A}). \\ &= \mathbf{A} + \mathbf{A}'\mathbf{S}^T \mathbf{C}^{-1} \mathbf{D}' \\ &= \mathbf{A} + \mathbf{A}(\mathbf{I} - \mathbf{1}_N)\mathbf{S}^T \mathbf{C}^{-1} \mathbf{D}' \\ &= \mathbf{A} \left(\mathbf{I} + \mathbf{S}^T \mathbf{C}^{-1} \mathbf{D}' \right) \\ &= \mathbf{A}\mathbf{X}\end{aligned}$$

where we have used

- $\mathbf{A}' = \mathbf{A}(\mathbf{I} - \mathbf{1}_N)$.
- $\mathbf{1}_N \mathbf{S}^T \equiv \mathbf{0}$.

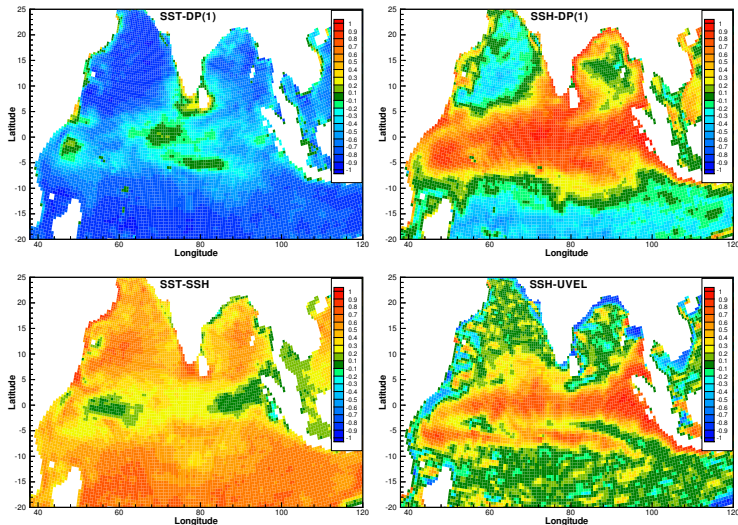
Remarks

- \bar{C}_{xx} is never computed.
- Even $H\bar{C}_{xx}H^T = SS^T$ need not be computed.
- Analysis may be interpreted as:
 - ▶ combination of forecast ensemble members, or,
 - ▶ forecast plus combination of covariance functions.
- Accuracy of analysis is determined by:
 - ▶ the accuracy of X ,
 - ▶ the properties of the ensemble space.
- For a linear model, any choice of X will result in an analysis which is also a solution of the model.

Examples of ensemble statistics

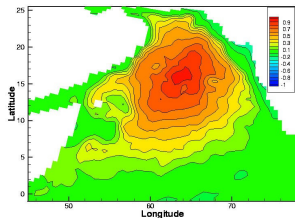
- Taken from Haugen et al. (2002), Ocean Dynamics.
- OGCM (MICOM) for the Indian Ocean.
- Assimilation of SST and SLA data.

Spatial correlations

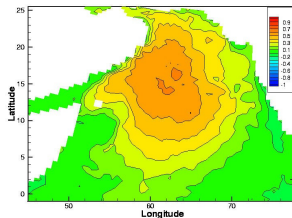


Correlation functions

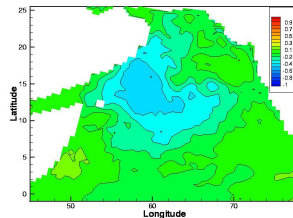
SSH-SSH



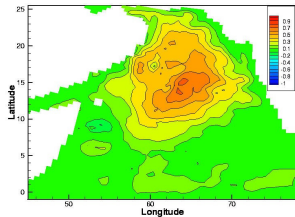
SSH-SST



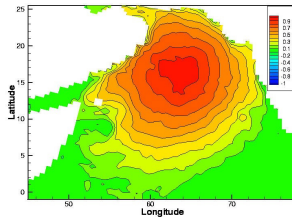
SSH-DP



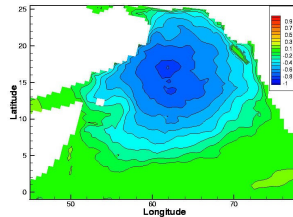
SST-SSH



SST-SST

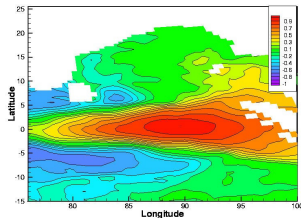


SST-DP

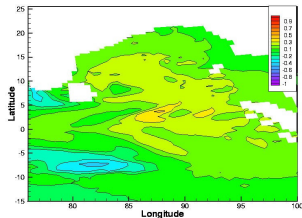


Correlation functions

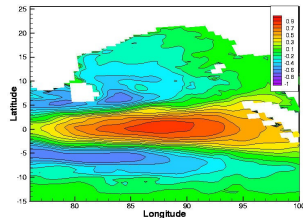
SSH-SSH



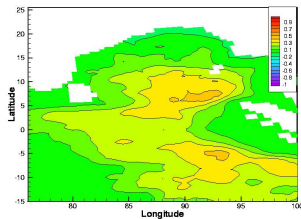
SSH-SST



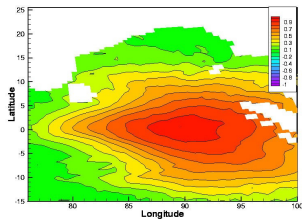
SSH-DP



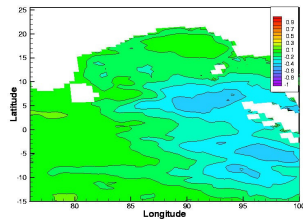
SST-SSH



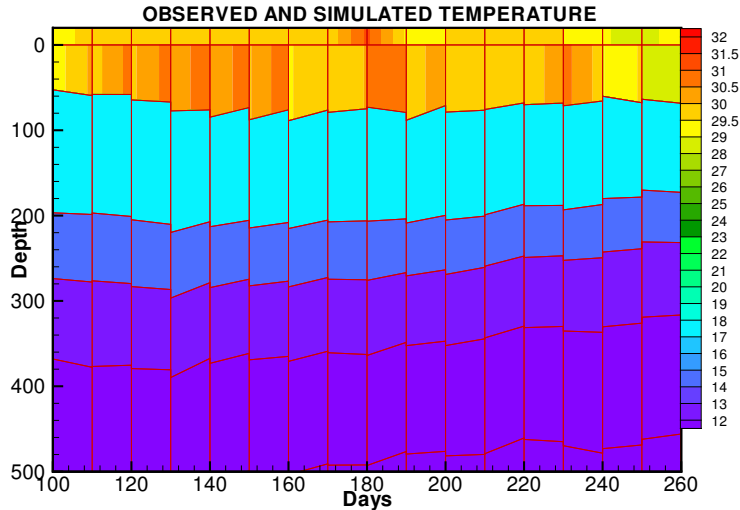
SST-SST



SST-DP



Time-Depth: Temperature



Computational aspects (1)

Analysis scheme

$$\mathbf{A}^a = \mathbf{A}\mathbf{X} = \mathbf{A} \left(\mathbf{I} + \mathbf{S}^T \mathbf{C}^{-1} \mathbf{D}' \right)$$

How to compute the inverse

$$\mathbf{C}^{-1} = \left(\mathbf{S}\mathbf{S}^T + \mathbf{E}\mathbf{E}^T \right)^{-1} \quad (\approx \mathbf{Z}\mathbf{\Lambda}^+\mathbf{Z}^T)$$

- Low rank ($N - 1$).
- Dimension is number of measurements m .
- Direct inversion requires $O(m^3)$ computations.

Subspace inversion

- Why invert m -dimensional matrix when solving for N coefficients?

$$\begin{aligned}
 & (SS^T + EE^T) \\
 & \approx SS^T + (SS^+)EE^T(SS^+)^T \\
 & = U\Sigma(I_N + \Sigma^+U^TEE^TU(\Sigma^+)^T)\Sigma^TU^T \\
 & = U\Sigma(I_N + Z\Lambda Z^T)\Sigma^TU^T \\
 & = U\Sigma Z(I_N + \Lambda)Z^T\Sigma^TU^T.
 \end{aligned}$$

$$(SS^T + EE^T)^{-1} \approx U(\Sigma^+)^TZ(I_N + \Lambda)^{-1}(U(\Sigma^+)^TZ)^T$$

- Define singular value decomposition $S = U\Sigma V^T$.
- Pseudo inverse $S^+ = V\Sigma^+U^T$.
- $(SS^+)E$ is orthogonal projection of E onto the S space.
- Cost is $O(mN^2)$.

Square-Root Schemes

Mean updated using

$$\overline{\mathbf{x}}^a = \overline{\mathbf{x}}^f + \overline{\mathbf{C}}_{xx}^f \mathbf{H}^T \left(\mathbf{H} \overline{\mathbf{C}}_{xx}^f \mathbf{H}^T + \overline{\mathbf{C}}_{dd} \right)^{-1} \left(\mathbf{d} - \mathbf{H} \overline{\mathbf{x}}^f \right).$$

Perturbations updated using factorization of posterior variance

$$\overline{\mathbf{C}}_{xx}^a = \overline{\mathbf{C}}_{xx}^f - \overline{\mathbf{C}}_{xx}^f \mathbf{H}^T \left(\mathbf{H} \overline{\mathbf{C}}_{xx}^f \mathbf{H}^T + \overline{\mathbf{C}}_{dd} \right)^{-1} \mathbf{H} \overline{\mathbf{C}}_{xx}^f.$$

Ensemble notation and simple illustration

$$\mathbf{A}^a \mathbf{A}^{a'T} = \mathbf{A}' \left(\mathbf{I} - \mathbf{S}^T \mathbf{C}^{-1} \mathbf{S} \right) \mathbf{A}'^T$$

Square-Root Schemes

Ensemble notation and simple illustration

$$\mathbf{A}^{\text{a}'} \mathbf{A}^{\text{a}'\text{T}} = \mathbf{A}' \left(\mathbf{I} - \mathbf{S}^{\text{T}} \mathbf{C}^{-1} \mathbf{S} \right) \mathbf{A}'^{\text{T}}$$

$$= \mathbf{A}' \left(\mathbf{Z} \mathbf{\Lambda} \mathbf{Z}^{\text{T}} \right) \mathbf{A}'^{\text{T}}$$

$$= \mathbf{A}' \mathbf{Z} \mathbf{\Lambda}^{\frac{1}{2}} \left(\mathbf{A}' \mathbf{Z} \mathbf{\Lambda}^{\frac{1}{2}} \right)^{\text{T}}$$

Non-symmetrical square root

$$= \mathbf{A}' \mathbf{Z} \mathbf{\Lambda}^{\frac{1}{2}} \mathbf{Z} \left(\mathbf{A}' \mathbf{Z} \mathbf{\Lambda}^{\frac{1}{2}} \mathbf{Z} \right)^{\text{T}}$$

Symmetrical square root

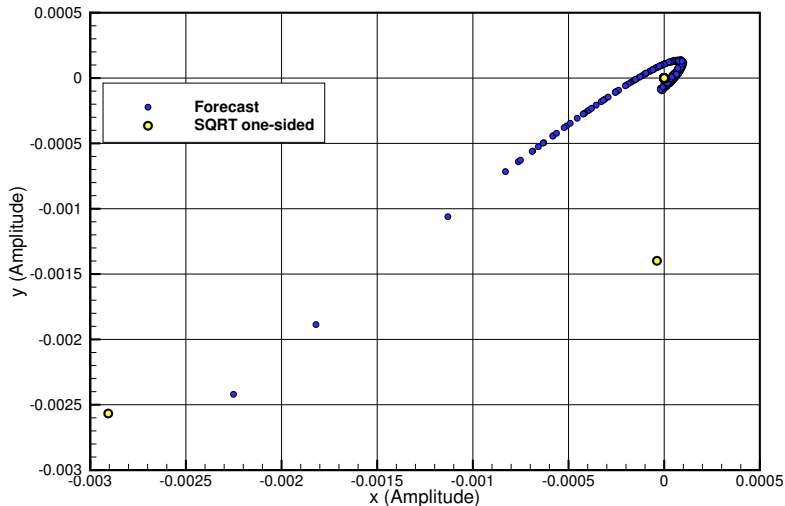
Update becomes

$$\mathbf{A}^{\text{a}'} = \mathbf{A}' \mathbf{Z} \mathbf{\Lambda}^{\frac{1}{2}} \mathbf{Z}$$

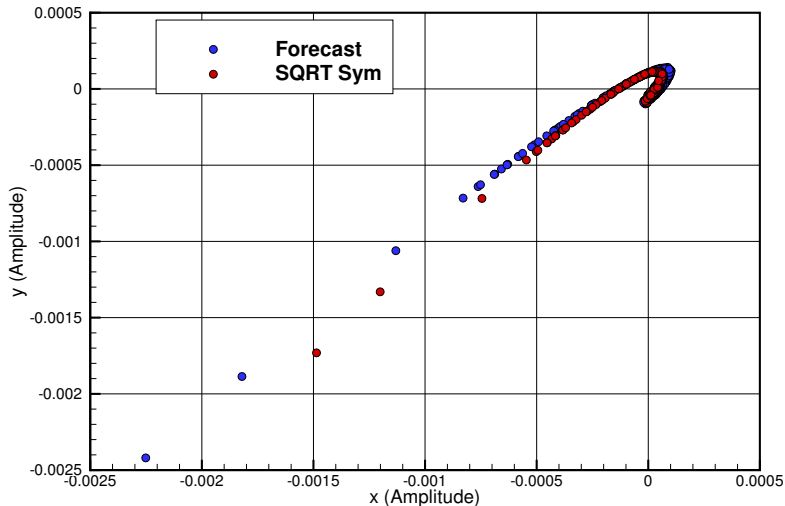
or when including mean preserving random rotation $\mathbf{\Phi} \mathbf{\Phi}^{\text{T}} = \mathbf{I}$

$$\mathbf{A}^{\text{a}'} = \mathbf{A}' \mathbf{Z} \mathbf{\Lambda}^{\frac{1}{2}} \mathbf{Z} \mathbf{\Phi}$$

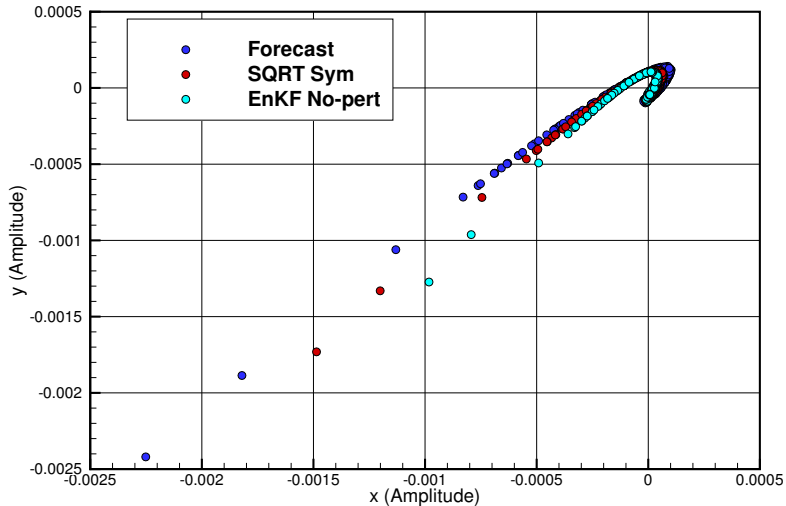
Square-root schemes: Non symmetrical



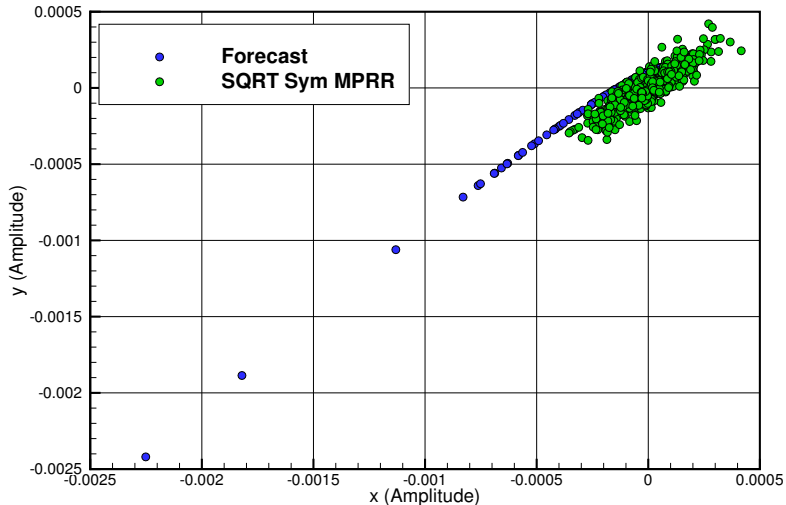
Square-root schemes: Symmetrical



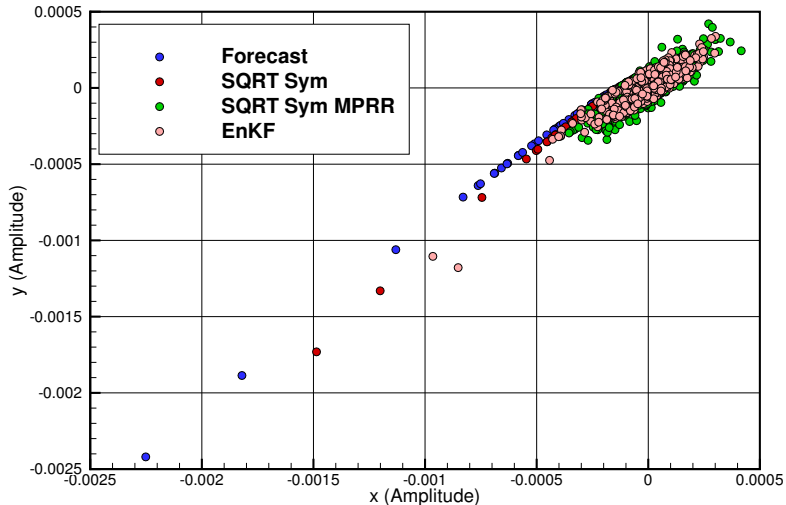
Square-root schemes vs no-pert EnKF



Square root schemes: Symmetrical with MPRR



Square root schemes vs EnKF



Summary

- The EnKF has worked well with highly nonlinear dynamics.
- The EnKF has worked well with high-dimensional models.
- There is no linearization in the evolution of error statistics.
- Major approximation is Gaussian assumption in update step.
- Another approximation is limited ensemble size.

More on ensemble methods

Geir Evensen



Ensemble methods

EnKF: Ensemble Kalman Filter

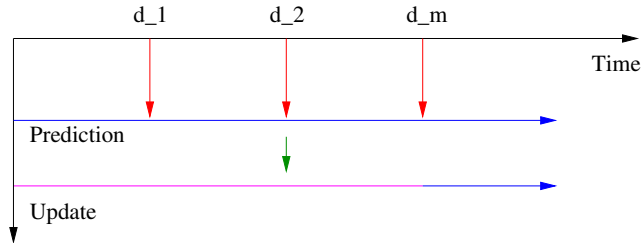
ES: Ensemble Smoother

EnKS: Ensemble Kalman Smoother

- Ensemble representation for pdfs.
- Ensemble prediction for time evolution of pdfs.
- Linear ensemble analysis scheme.

ES: The Ensemble Smoother

Smoother solution processing all data in one go.

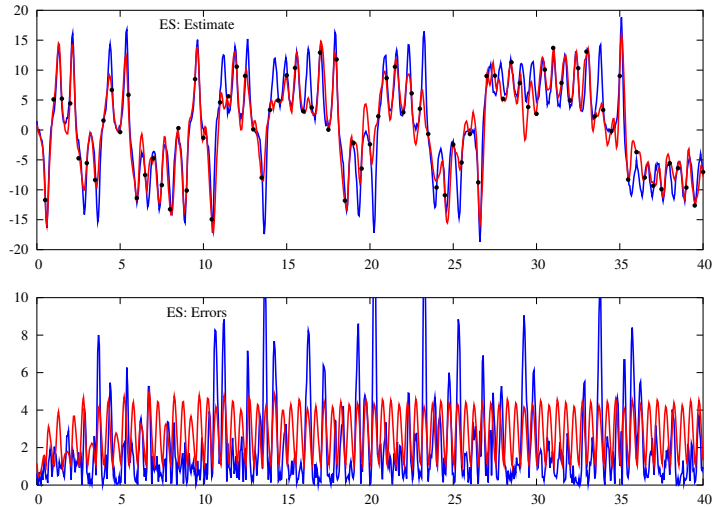


ES: summary

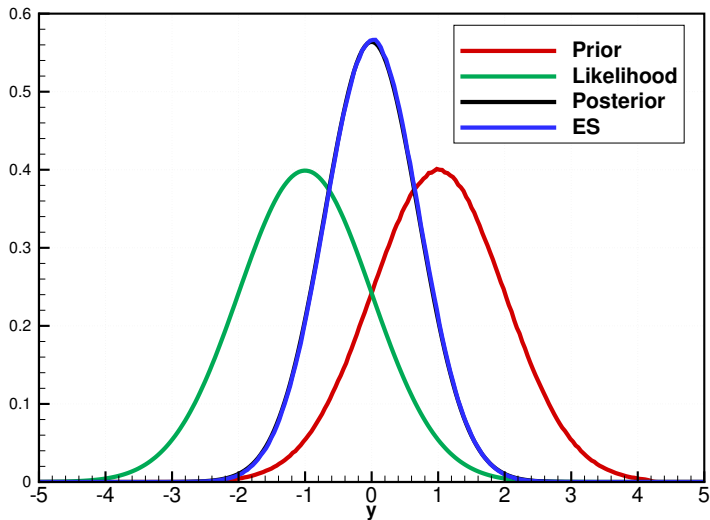
Gauss–Markov interpolation in space and time.

- Creates an ensemble for the model prediction.
- Assumes Gaussian pdf for model prediction.
- Computes variance minimizing ensemble analysis.
- Exact solution for linear problems.

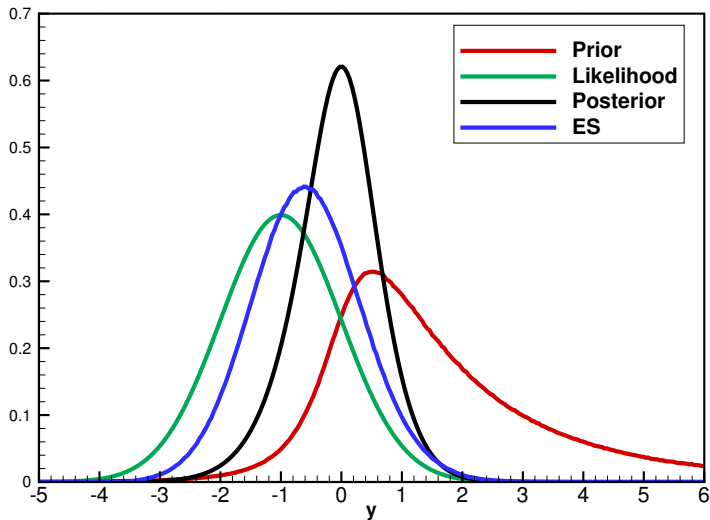
ES: Example with Lorenz equations



ES vs Bayes' (Gaussian prior)

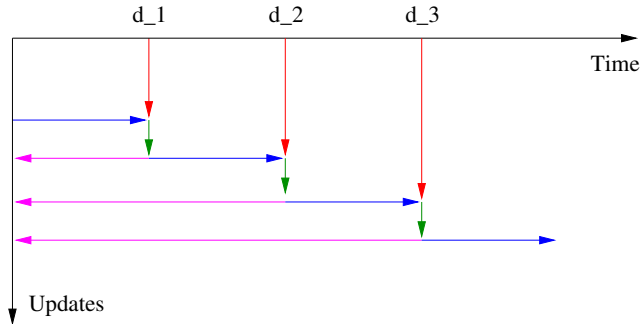


ES vs Bayes' (non-Gaussian prior)

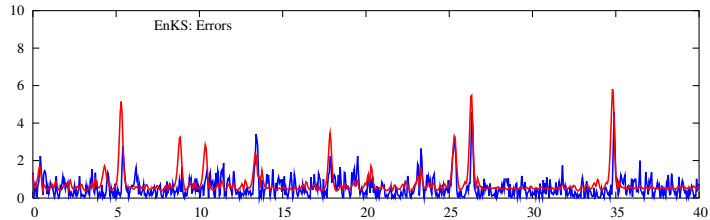
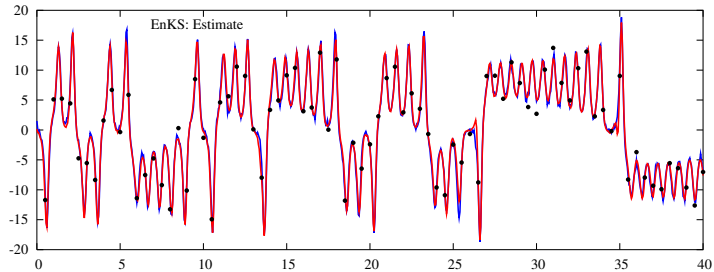


EnKS: The ensemble Kalman smoother

- Smoother solution with sequential processing of data



EnKS solution

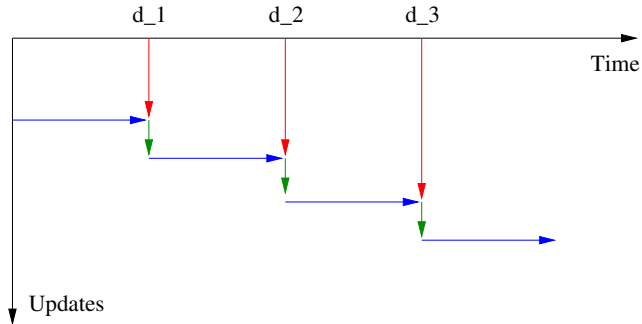


EnKS summary

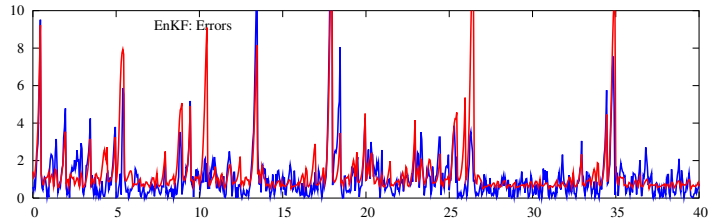
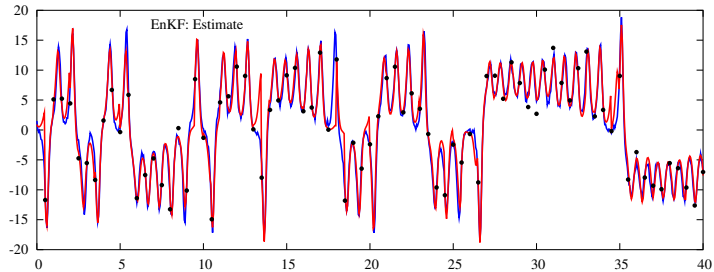
- ES and EnKS give identical results for linear models.
- EnKS is superior to the ES with nonlinear models.
 - ▶ Sequential processing of measurements introduces “Gaussianity”.
 - ▶ Ensemble is kept close to the true state.

EnKF: Ensemble Kalman Filter

- Filtering solution



EnKF solution



State and parameter estimation

- Scalar model for x with parameter α .

$$\frac{\partial x}{\partial t} = 1 - \alpha + q,$$

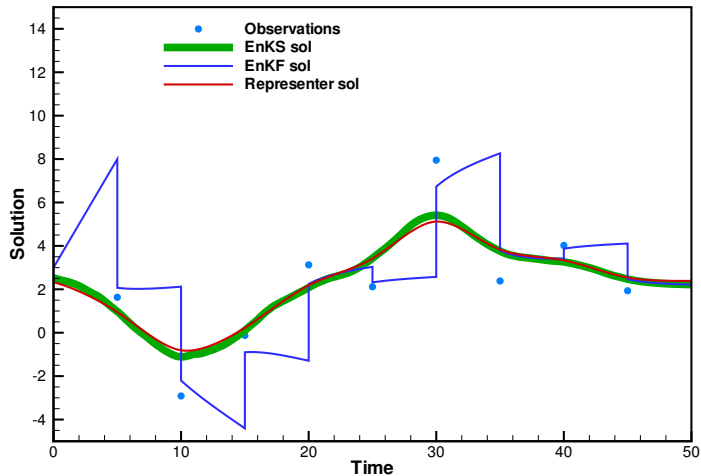
$$x(t = 0) = 3 + a,$$

$$\alpha = 0 + \alpha',$$

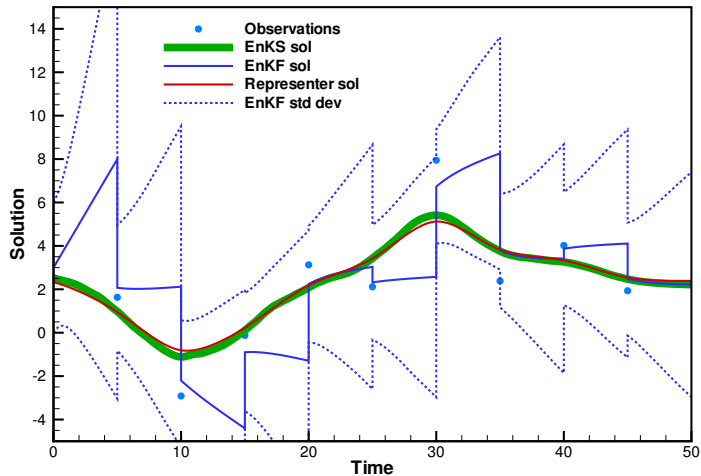
$$\mathcal{M}(x) = d + \epsilon.$$

- True parameter value is $\alpha = 1$.
- Truly linear model.
- Solved using EnKF, EnKS and Representer methods.
- Exponential time correlation for model errors.

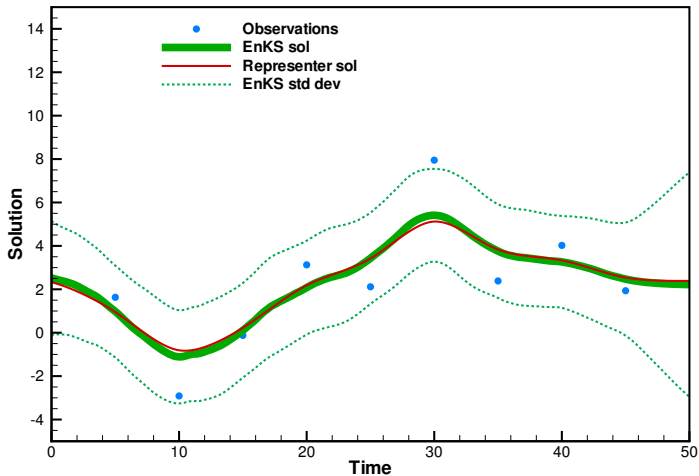
State and parameter estimation



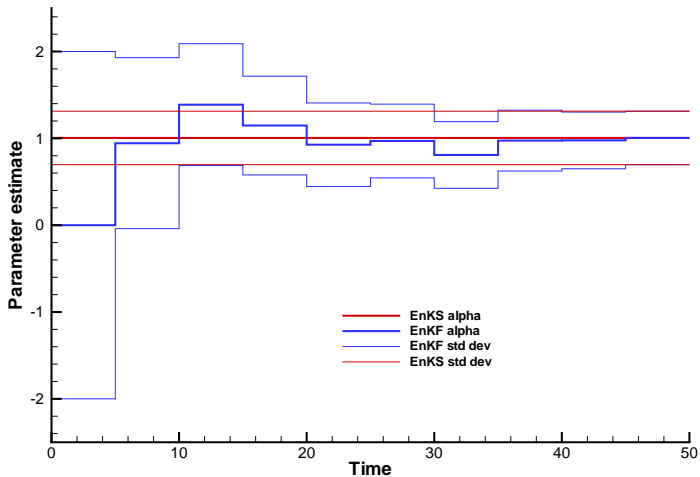
State and parameter estimation



State and parameter estimation



Estimate of parameter



Burgers, G., P. J. van Leeuwen, and G. Evensen. Analysis scheme in the ensemble Kalman filter. *Mon. Weather Rev.*, 126:1719–1724, 1998.

Haugen, V. E., O. M. Johannessen, and G. Evensen. Indian Ocean: Validation of the Miami Isopycnal Coordinate Ocean Model and ENSO events during 1958–1998. *J. Geophys. Res.*, 107 (C5):11–1–11–23, 2002.