

ECONhw5

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3/31/2022

```
library(dplyr)

##
## Attaching package: 'dplyr'

## The following objects are masked from 'package:stats':
##
##   filter, lag

## The following objects are masked from 'package:base':
##
##   intersect, setdiff, setequal, union

data <- read.csv("C:/Users/Nick/Downloads/pntsprd.csv")

#2.1

data$sprdcvr

## [1] 1 1 1 0 0 1 0 1 0 1 1 1 1 0 0 0 0 1 0 0 1 0 0 1 1 0 1 1 1 1 1 1 1 0
## [38] 0 0 1
## [75] 1 0 1 0 1 0 0 1 0 1 1 0 1 1 0 1 0 1 1 0 1 1 1 1 0 0 1 1 1 1 1 1 1 1
## [112] 1 1 0
## [149] 0 1 0 1 1 1 0 0 0 0 0 0 1 0 0 1 0 0 0 0 0 0 0 0 0 0 1 0 1 1 1 1 1 0 1
## [186] 0 1 1
## [223] 0 0 0 0 0 0 0 1 0 0 1 1 0 0 0 0 0 1 1 0 1 1 1 1 0 1 1 1 1 0 1 0 1 0
## [260] 1 0 1
## [297] 1 0 0 0 0 0 0 1 0 1 1 1 1 1 0 1 0 0 0 1 0 0 0 1 1 1 1 0 0 1 0 0 0 0
## [334] 1 0 1
## [371] 0 0 1 0 1 1 0 0 1 0 1 0 1 1 0 1 0 1 0 0 0 1 1 1 0 0 1 1 1 1 0 0 1
## [408] 0 1 0
```

```
## [445] 0 0 1 1 0 1 0 0 0 1 0 1 1 1 1 0 0 0 0 0 1 1 0 0 0 0 1 1 0 1 1 0 0
0 1 0
## [482] 0 0 0 1 1 1 0 0 0 0 1 0 1 0 0 0 1 1 0 1 0 0 1 1 0 1 0 1 0 1 0 1
1 1 1
## [519] 0 1 0 0 0 1 1 0 1 0 0 1 1 1 1 1 0 1 1 0 1 1 1 0 1 0 0 0 1 0 0 1 1 0
0
```

```
t.test(data$sprdcvr,mu = .5, paired = F, conf.level=.9)
```

```
##
## One Sample t-test
##
## data: data$sprdcvr
## t = 0.7226, df = 552, p-value = 0.4702
## alternative hypothesis: true mean is not equal to 0.5
## 90 percent confidence interval:
## 0.4803236 0.5504178
## sample estimates:
## mean of x
## 0.5153707
```

Fail to reject the null hypothesis as the p-value is greater than .05.

#2.2

```
gamesonneutralcourt = sum(data$neutral)
gamesonneutralcourt
```

```
## [1] 35
```

#2.3

```
lm.fit <- lm(data$sprdcvr ~ data$favhome + data$neutral + data$fav25 +
data$und25)
summary(lm.fit)
```

```
##
## Call:
## lm(formula = data$sprdcvr ~ data$favhome + data$neutral + data$fav25 +
##      data$und25)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -0.6072 -0.5242  0.3928  0.4758  0.5339
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   0.48957    0.04476  10.938  <2e-16 ***
## data$favhome   0.03459    0.04972   0.696   0.487
## data$neutral   0.11762    0.09466   1.242   0.215
## data$fav25    -0.02347    0.05019  -0.468   0.640
```

```
## data$und25    0.01787    0.09188    0.195    0.846
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.5012 on 548 degrees of freedom
## Multiple R-squared:  0.0034, Adjusted R-squared:  -0.003874
## F-statistic: 0.4674 on 4 and 548 DF,  p-value: 0.7597

# Load libraries
library("lmtest")

## Loading required package: zoo

##
## Attaching package: 'zoo'

## The following objects are masked from 'package:base':
##
##      as.Date, as.Date.numeric

library("sandwich")

# Robust t test
hetero.fit <- coeftest(lm.fit, vcov = vcovHC(lm.fit, type = "HC0"))
hetero.fit

##
## t test of coefficients:
##
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  0.489567   0.044648 10.9651  <2e-16 ***
## data$favhome  0.034591   0.049597  0.6974   0.4858
## data$neutral  0.117618   0.092701  1.2688   0.2051
## data$fav25   -0.023467   0.050154 -0.4679   0.6400
## data$und25    0.017873   0.089655  0.1994   0.8421
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

summary(hetero.fit)

##      Estimate      Std. Error      t value      Pr(>|t|)
## Min.      : -0.02347  Min.      : 0.04465  Min.      : -0.4679  Min.      : 0.0000
## 1st Qu.:  0.01787  1st Qu.: 0.04960  1st Qu.:  0.1993  1st Qu.: 0.2051
## Median :  0.03459  Median : 0.05015  Median :  0.6974  Median : 0.4858
## Mean   :  0.12724  Mean   : 0.06535  Mean   :  2.5325  Mean   : 0.4346
## 3rd Qu.:  0.11762  3rd Qu.: 0.08966  3rd Qu.:  1.2688  3rd Qu.: 0.6400
## Max.    :  0.48957  Max.    : 0.09270  Max.    : 10.9651  Max.    : 0.8421

# Standard errors do not significantly change for any estimators when
# assuming heteroskedasticity robust model. heteroskedasticity robust model is
# more significant (lower p-values).
```

#2.4

If all of the estimators equal zero ($B_1 = B_2 = B_3 = B_4 = 0$), the OLS fit will just be a horizontal line across the data with a y value at the mean of the regressand (sprdcvr). Since sprdcvr is binary and all estimators equal zero, variance of error will not be dependent on any regressor, suggesting homoskedasticity.

#2.5

None of the statistics in any of the tests performed on the model above have shown any significance of relationship between the regressors and if the game will cover the spread. This suggests that the market of sports betting is efficient and that one cannot use available data to get advantage in calling sports bets.

#2.6

lm.fit\$fitted.values

##	1	2	3	4	5	6	7
8							
##	0.4660991	0.5241576	0.5241576	0.5241576	0.4895665	0.5837171	0.5241576
0.5006902							
##	9	10	11	12	13	14	15
16							
##	0.4895665	0.4895665	0.5241576	0.5241576	0.4660991	0.5241576	0.4895665
0.5006902							
##	17	18	19	20	21	22	23
24							
##	0.5241576	0.5241576	0.4895665	0.5241576	0.5241576	0.4895665	0.5241576
0.5006902							
##	25	26	27	28	29	30	31
32							
##	0.6071845	0.4895665	0.5241576	0.5241576	0.5006902	0.5241576	0.5006902
0.5837171							
##	33	34	35	36	37	38	39
40							
##	0.6071845	0.5241576	0.5241576	0.6071845	0.5241576	0.6071845	0.5241576
0.6071845							
##	41	42	43	44	45	46	47
48							
##	0.5241576	0.6071845	0.4895665	0.6071845	0.6071845	0.5241576	0.5837171
0.5241576							
##	49	50	51	52	53	54	55
56							
##	0.5006902	0.5241576	0.5241576	0.5006902	0.4660991	0.5241576	0.5241576
0.5241576							
##	57	58	59	60	61	62	63
64							

0.4895665 0.5241576 0.5006902 0.6071845 0.6071845 0.4895665 0.5241576
0.5241576
65 66 67 68 69 70 71
72
0.6015899 0.5241576 0.5241576 0.5241576 0.5241576 0.5241576 0.6071845
0.4895665
73 74 75 76 77 78 79
80
0.4660991 0.5241576 0.4895665 0.4895665 0.5241576 0.4895665 0.5241576
0.4895665
81 82 83 84 85 86 87
88
0.4895665 0.5185630 0.4660991 0.5241576 0.5006902 0.5006902 0.5241576
0.5241576
89 90 91 92 93 94 95
96
0.5006902 0.6071845 0.5241576 0.6071845 0.5241576 0.5241576 0.6071845
0.5241576
97 98 99 100 101 102 103
104
0.6071845 0.5006902 0.5241576 0.5241576 0.5241576 0.5241576 0.6015899
0.5420304
105 106 107 108 109 110 111
112
0.4660991 0.5241576 0.5241576 0.5006902 0.5241576 0.5241576 0.5241576
0.4660991
113 114 115 116 117 118 119
120
0.5241576 0.4660991 0.5241576 0.5241576 0.5837171 0.5241576 0.5185630
0.5241576
121 122 123 124 125 126 127
128
0.5241576 0.5241576 0.5241576 0.5241576 0.5241576 0.5241576 0.5006902
0.5241576
129 130 131 132 133 134 135
136
0.4895665 0.5241576 0.5241576 0.4895665 0.5241576 0.5241576 0.5241576
0.5241576
137 138 139 140 141 142 143
144
0.5241576 0.5241576 0.5241576 0.4895665 0.4895665 0.6071845 0.6071845
0.5241576
145 146 147 148 149 150 151
152
0.5241576 0.6071845 0.4895665 0.5241576 0.5420304 0.5241576 0.4660991
0.5241576
153 154 155 156 157 158 159
160
0.5837171 0.5241576 0.5006902 0.5837171 0.5006902 0.4895665 0.4895665
0.5241576

##	161	162	163	164	165	166	167
168							
##	0.5006902	0.4660991	0.5420304	0.4660991	0.4660991	0.4660991	0.4895665
0.5006902							
##	169	170	171	172	173	174	175
176							
##	0.5241576	0.4895665	0.5241576	0.5241576	0.5006902	0.5241576	0.4660991
0.4895665							
##	177	178	179	180	181	182	183
184							
##	0.5241576	0.5241576	0.4895665	0.4895665	0.5241576	0.5241576	0.5241576
0.4895665							
##	185	186	187	188	189	190	191
192							
##	0.5241576	0.5241576	0.5241576	0.4895665	0.5241576	0.5241576	0.5241576
0.5006902							
##	193	194	195	196	197	198	199
200							
##	0.5241576	0.5241576	0.5241576	0.4895665	0.5006902	0.5241576	0.5241576
0.5006902							
##	201	202	203	204	205	206	207
208							
##	0.5241576	0.5006902	0.4895665	0.5241576	0.5241576	0.4895665	0.5241576
0.5241576							
##	209	210	211	212	213	214	215
216							
##	0.5241576	0.5185630	0.5006902	0.5241576	0.5241576	0.5241576	0.5241576
0.4895665							
##	217	218	219	220	221	222	223
224							
##	0.5241576	0.5241576	0.5241576	0.4895665	0.5241576	0.5241576	0.5241576
0.5241576							
##	225	226	227	228	229	230	231
232							
##	0.5241576	0.5241576	0.4839719	0.5241576	0.5006902	0.4895665	0.5006902
0.5241576							
##	233	234	235	236	237	238	239
240							
##	0.4895665	0.4895665	0.5006902	0.4895665	0.4895665	0.5006902	0.4660991
0.5241576							
##	241	242	243	244	245	246	247
248							
##	0.5006902	0.5006902	0.5241576	0.5241576	0.4895665	0.5241576	0.5241576
0.5837171							
##	249	250	251	252	253	254	255
256							
##	0.4895665	0.5241576	0.5241576	0.5006902	0.5420304	0.5241576	0.5006902
0.5006902							
##	257	258	259	260	261	262	263
264							

0.6071845 0.5241576 0.5837171 0.6071845 0.4895665 0.4895665 0.5241576
0.4895665
265 266 267 268 269 270 271
272
0.5241576 0.4895665 0.5241576 0.4895665 0.5241576 0.4895665 0.5241576
0.5241576
273 274 275 276 277 278 279
280
0.5241576 0.5241576 0.5241576 0.5241576 0.5006902 0.5241576 0.5006902
0.6015899
281 282 283 284 285 286 287
288
0.5837171 0.5006902 0.4895665 0.4660991 0.4895665 0.5241576 0.5241576
0.5241576
289 290 291 292 293 294 295
296
0.4839719 0.5241576 0.5241576 0.5006902 0.5420304 0.4895665 0.5241576
0.5241576
297 298 299 300 301 302 303
304
0.5185630 0.5006902 0.4895665 0.5006902 0.4895665 0.4660991 0.5241576
0.5006902
305 306 307 308 309 310 311
312
0.4660991 0.4660991 0.5241576 0.5006902 0.5241576 0.5241576 0.5241576
0.5241576
313 314 315 316 317 318 319
320
0.5185630 0.5006902 0.5420304 0.5241576 0.5006902 0.5006902 0.5241576
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321 322 323 324 325 326 327
328
0.4895665 0.5185630 0.5185630 0.5006902 0.5185630 0.4895665 0.5241576
0.5241576
329 330 331 332 333 334 335
336
0.4660991 0.4895665 0.5006902 0.5241576 0.5241576 0.4660991 0.5241576
0.5006902
337 338 339 340 341 342 343
344
0.4660991 0.5241576 0.4895665 0.5241576 0.4895665 0.5420304 0.4839719
0.4895665
345 346 347 348 349 350 351
352
0.5241576 0.5420304 0.5006902 0.5006902 0.5006902 0.4895665 0.4895665
0.5241576
353 354 355 356 357 358 359
360
0.5241576 0.5420304 0.4660991 0.5241576 0.5241576 0.5006902 0.5006902
0.5006902

##	361	362	363	364	365	366	367
368							
##	0.5241576	0.5241576	0.4895665	0.4660991	0.4839719	0.5241576	0.5420304
0.5241576							
##	369	370	371	372	373	374	375
376							
##	0.5241576	0.4660991	0.5006902	0.4895665	0.5241576	0.4895665	0.4660991
0.5241576							
##	377	378	379	380	381	382	383
384							
##	0.5241576	0.5241576	0.5241576	0.4660991	0.4895665	0.5241576	0.5241576
0.5241576							
##	385	386	387	388	389	390	391
392							
##	0.5241576	0.5241576	0.5006902	0.5241576	0.5241576	0.5241576	0.4895665
0.4895665							
##	393	394	395	396	397	398	399
400							
##	0.5241576	0.5185630	0.5241576	0.5185630	0.5006902	0.5006902	0.5241576
0.4895665							
##	401	402	403	404	405	406	407
408							
##	0.5006902	0.5241576	0.5241576	0.4895665	0.5241576	0.5241576	0.5241576
0.5185630							
##	409	410	411	412	413	414	415
416							
##	0.5241576	0.5241576	0.5241576	0.5241576	0.4895665	0.4895665	0.5241576
0.5006902							
##	417	418	419	420	421	422	423
424							
##	0.5006902	0.5241576	0.5241576	0.5241576	0.4895665	0.5241576	0.4895665
0.5241576							
##	425	426	427	428	429	430	431
432							
##	0.4660991	0.4895665	0.5006902	0.5241576	0.5241576	0.5241576	0.4895665
0.4660991							
##	433	434	435	436	437	438	439
440							
##	0.5241576	0.5241576	0.5241576	0.5006902	0.5006902	0.5241576	0.5241576
0.5241576							
##	441	442	443	444	445	446	447
448							
##	0.5241576	0.5241576	0.5241576	0.5241576	0.5006902	0.5241576	0.4895665
0.5241576							
##	449	450	451	452	453	454	455
456							
##	0.5241576	0.5241576	0.4895665	0.5420304	0.5241576	0.5241576	0.4895665
0.5185630							
##	457	458	459	460	461	462	463
464							

0.4660991 0.5241576 0.5241576 0.5241576 0.5241576 0.5241576 0.4895665
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465 466 467 468 469 470 471
472
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473 474 475 476 477 478 479
480
0.5185630 0.6071845 0.4895665 0.5241576 0.4895665 0.5241576 0.4895665
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488
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489 490 491 492 493 494 495
496
0.5241576 0.5241576 0.5006902 0.5241576 0.5241576 0.5241576 0.4895665
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497 498 499 500 501 502 503
504
0.5241576 0.5241576 0.4895665 0.4660991 0.4895665 0.4895665 0.4660991
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529 530 531 532 533 534 535
536
0.5837171 0.4660991 0.5241576 0.4660991 0.5241576 0.5241576 0.5241576
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537 538 539 540 541 542 543
544
0.5241576 0.5006902 0.4895665 0.4895665 0.4660991 0.5185630 0.5241576
0.5241576
545 546 547 548 549 550 551
552
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0.4660991
553
0.4660991

#2.7

```
library(foreign)
```

```
logit <- glm(data$sprdcvr ~ data$favhome + data$neutral + data$fav25 +  
data$und25, family=binomial(link="logit"))
```

```
logit
```

```
##
```

```
## Call: glm(formula = data$sprdcvr ~ data$favhome + data$neutral +  
data$fav25 +
```

```
## data$und25, family = binomial(link = "logit"))
```

```
##
```

```
## Coefficients:
```

```
## (Intercept) data$favhome data$neutral data$fav25 data$und25
```

```
## -0.04169 0.13845 0.47618 -0.09424 0.07177
```

```
##
```

```
## Degrees of Freedom: 552 Total (i.e. Null); 548 Residual
```

```
## Null Deviance: 766.1
```

```
## Residual Deviance: 764.2 AIC: 774.2
```

```
summary(logit)
```

```
##
```

```
## Call:
```

```
## glm(formula = data$sprdcvr ~ data$favhome + data$neutral + data$fav25 +  
## data$und25, family = binomial(link = "logit"))
```

```
##
```

```
## Deviance Residuals:
```

```
## Min 1Q Median 3Q Max
```

```
## -1.3666 -1.2188 0.9993 1.1366 1.2356
```

```
##
```

```
## Coefficients:
```

```
## Estimate Std. Error z value Pr(>|z|)
```

```
## (Intercept) -0.04169 0.17873 -0.233 0.816
```

```
## data$favhome 0.13845 0.19855 0.697 0.486
```

```
## data$neutral 0.47618 0.38413 1.240 0.215
```

```
## data$fav25 -0.09424 0.20070 -0.470 0.639
```

```
## data$und25 0.07177 0.36783 0.195 0.845
```

```
##
```

```
## (Dispersion parameter for binomial family taken to be 1)
```

```
##
```

```
## Null deviance: 766.10 on 552 degrees of freedom
```

```
## Residual deviance: 764.21 on 548 degrees of freedom
```

```
## AIC: 774.21
```

```
##
```

```
## Number of Fisher Scoring iterations: 3
```

```
probit <- glm(data$sprdcvr ~ data$favhome + data$neutral + data$fav25 +  
data$und25, family=binomial(link="probit"))
```

```
probit
```

```
##
## Call: glm(formula = data$sprdcvr ~ data$favhome + data$neutral +
data$fav25 +
## data$und25, family = binomial(link = "probit"))
##
## Coefficients:
## (Intercept) data$favhome data$neutral data$fav25 data$und25
## -0.02606 0.08665 0.29798 -0.05916 0.04591
##
## Degrees of Freedom: 552 Total (i.e. Null); 548 Residual
## Null Deviance: 766.1
## Residual Deviance: 764.2 AIC: 774.2

summary(probit)

##
## Call:
## glm(formula = data$sprdcvr ~ data$favhome + data$neutral + data$fav25 +
## data$und25, family = binomial(link = "probit"))
##
## Deviance Residuals:
## Min 1Q Median 3Q Max
## -1.367 -1.219 0.999 1.137 1.236
##
## Coefficients:
## Estimate Std. Error z value Pr(>|z|)
## (Intercept) -0.02606 0.11197 -0.233 0.816
## data$favhome 0.08665 0.12439 0.697 0.486
## data$neutral 0.29798 0.23901 1.247 0.213
## data$fav25 -0.05916 0.12566 -0.471 0.638
## data$und25 0.04591 0.23020 0.199 0.842
##
## (Dispersion parameter for binomial family taken to be 1)
##
## Null deviance: 766.10 on 552 degrees of freedom
## Residual deviance: 764.21 on 548 degrees of freedom
## AIC: 774.21
##
## Number of Fisher Scoring iterations: 3

# Probit and Logit models still return similar results without statistical
significance.
```