# 함수추정의 응용 및 실습

Assignment #1

서울대학교 통계학과 2017-11362 박건도

2022년 04월 05일

# 1. Reflection boundary condition

(a) R function for smoothing by a moving average

```
move1 <- function(yy, mm){
  nd <- length(yy)
  yyr <- yy[nd:(nd-mm+1)]
  yyl <- yy[mm:1]
  y <- c(yyl, yy, yyr)
  ey <- c()
  for (ind in 1:nd){
    ey[ind] <- mean(y[ind:(ind+2*mm)])
  }
  ey
}</pre>
```

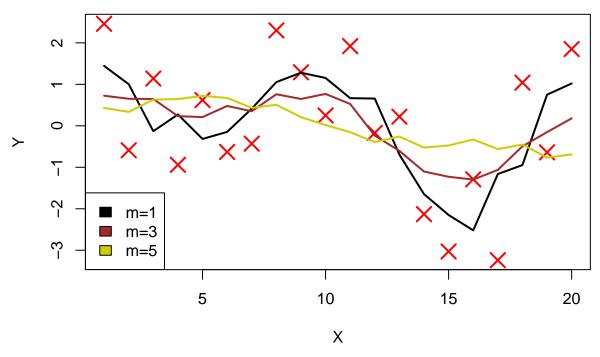
### (b) R function for smoothing by a binomial filter

```
bf1 <- function(yy, mm){
  nd <- length(yy)
  mm2 <- mm * 0.5
  yyw1 <- yy
  yyw2 <- yy
  rlim <- mm2
  yyr <- NULL
  count <- 0
  while(rlim > nd) {
    yyw1 <- rev(yyw1)</pre>
```

```
yyr <- c(yyr, yyw1)</pre>
    rlim <- rlim - nd
    count <- count + 1</pre>
  switch(count %% 2 + 1,
          yyr \leftarrow c(yyr, yy[nd:(nd - rlim + 1)]),
          yyr <- c(yyr, yy[1:rlim]))</pre>
  llim <- mm2
  yyl <- NULL
  while (llim > nd) {
    yyw2 <- rev(yyw2)</pre>
   yy1 \leftarrow c(yyw2, yy1)
   llim <- llim - nd
  }
  switch(count %% 2 + 1,
          yyl \leftarrow c(yy[llim:1], yyl),
          yyl \leftarrow c(yy[(nd - llim + 1):nd], yyl))
  y2 \leftarrow matrix(c(yyl, yy, yyr), ncol = 1)
  ww <- matrix(0, ncol = nd + mm, nrow = nd + mm)</pre>
  imat <- row(ww)</pre>
  jmat <- col(ww)</pre>
  check <- 0 <= (mm2 + imat - jmat) & (mm2 + imat - jmat) <= mm
  ww[check] <- exp(lgamma(mm +1) -</pre>
                        lgamma(mm2 + imat[check] - jmat[check] + 1) -
                        lgamma(mm2 - imat[check] + jmat[check] + 1) -
                        mm * logb(2))
  ey <- ww %*% y2
  ey \leftarrow as.vector(ey[(mm2 + 1):(nd + mm2)])
  еу
}
```

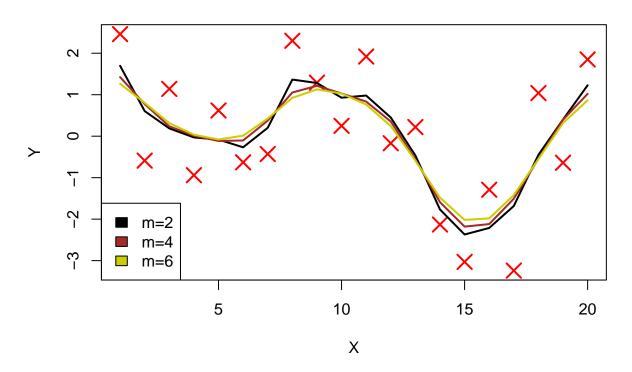
#### (c) smooth 20 data.

## mooving average



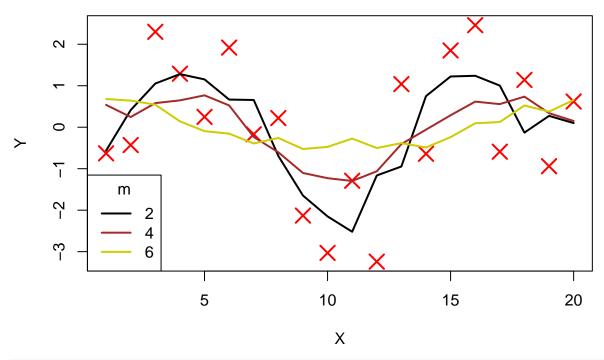
```
plot(xx, yy, xlab="X", ylab="Y", main="binomial filter", type="n")
points(xx, yy, pch =4, lwd = 2, col = "red", cex = 2)
lines(xx, bf1(yy, 2), lwd=2)
lines(xx, bf1(yy, 4), lwd=2, col="brown")
lines(xx, bf1(yy, 6), lwd=2, col="yellow3")
```

## binomial filter

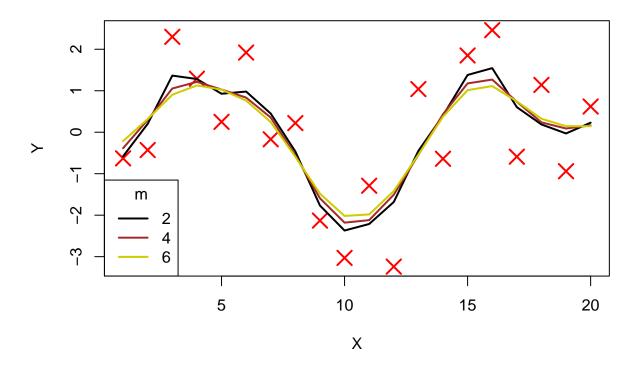


### (d) smooth shifted 20 data.

## mooving average



## binomial filter



# 2. Fitting a spline by the least squares

### (a) knots

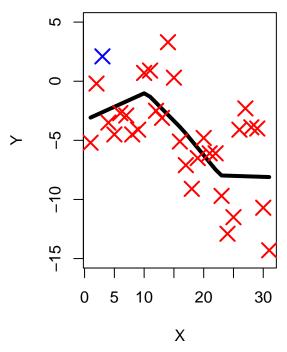
```
spline1 <- function(kk, deg, yy){
  data1 <- data.frame(x=1:length(yy), y=yy)
  fit.lm <- lm(y ~ bs(x, knots=kk, degree=deg), data=data1)
  fitted.values(fit.lm)
}

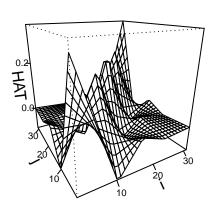
nd <- 31
xx <- 1:nd
yy <- scan("WAK2.CSV")
knots <- c(10.3, 16.5, 22.8)

ey <- spline1(knots, 1, yy)

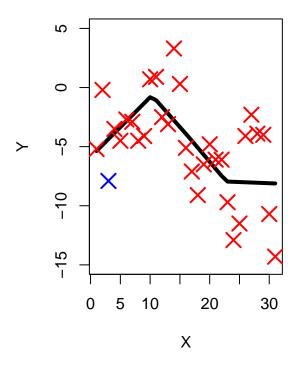
par(mfrow = c(1,2))
# fitted graph
plot(xx, ey, type = "n", ylim = c(-15, 5),</pre>
```

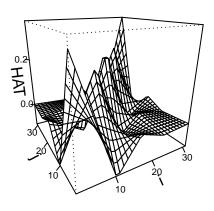
## **Spline**





### Spline with movement of data





## 3. Local linear regression

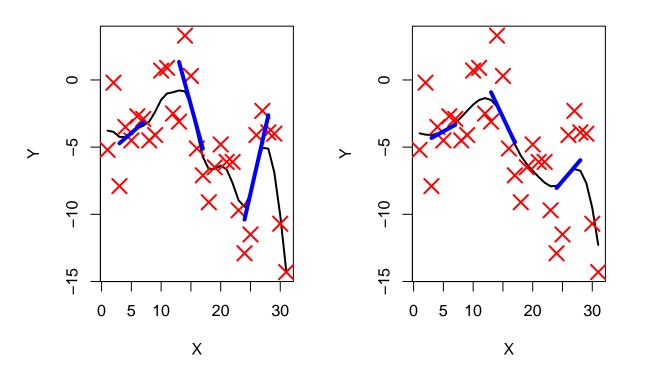
### (a) R function

```
lline<-function(yy, hh) {
    llin <- function(ex1, xdata, ydata, band) {
        wts <- exp((-0.5 * (ex1 - xdata)^2)/band^2)
        data1 <- data.frame(x = xdata, y = ydata, www = wts)
        fit.lm <- lm(y ~ x, data = data1, weights = www)
        est <- fit.lm$coef[1] + fit.lm$coef[2] * ex1
        list(est=est, coef=fit.lm$coef)
    }
    nd <- length(yy)
    xx <- seq(from = 1, by = 1, length = nd)
    xxmat <- matrix(xx, ncol = 1)</pre>
```

```
par(mfrow=c(1,length(hh)))
  for (h in hh){
    eys <- apply(xxmat, 1, llin, xdata = xx, ydata = yy,
               band = h)
    ey <- sapply(eys, function(yy) yy$est)</pre>
    ey <- as.vector(ey)</pre>
    plot(xx, yy, type = "n", xlab = "X", ylab = "Y",
         main=paste0("local linear regression(h=",h,")"))
    lines(xx, ey, lwd = 2)
    points(xx, yy, pch =4, lwd = 2, col = "red", cex = 2)
    for (ind in c(5, 15, 26)){
      xx_sub <- (ind-2):(ind+2)</pre>
      yy_sub <- eys[[ind]]$coef[1] + eys[[ind]]$coef[2] * xx_sub</pre>
      lines(xx_sub, yy_sub, lwd=4, col="blue")
    }
  }
  еу
}
hh \leftarrow c(1.5, 2.5)
ey <- lline(yy, hh)</pre>
```

# local linear regression(h=1.5)

# local linear regression(h=2.5)

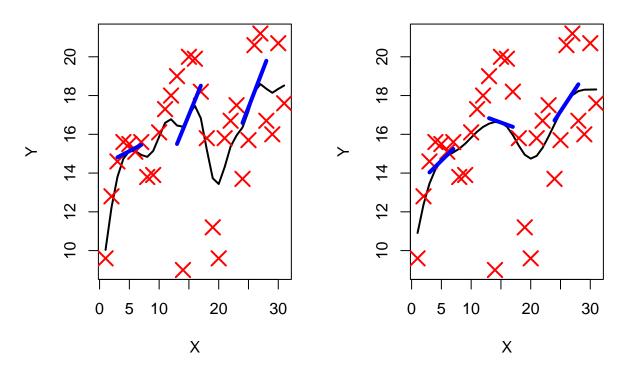


### (b) various bandwidth

```
yy <- c(9.6, 12.8, 14.6, 15.6, 15.5, 15.1, 15.6, 13.8, 13.9, 16.1, 17.3, 18, 19, 9, 20, 19.9, 18.2, 15.8, 11.2, 9.6, 15.8, 16.7, 17.5, 13.7, 15.7, 20.6, 21.2, 16.7, 16, 20.7, 17.6) ey <- lline(yy, hh)
```

## local linear regression(h=1.5)

# local linear regression(h=2.5)



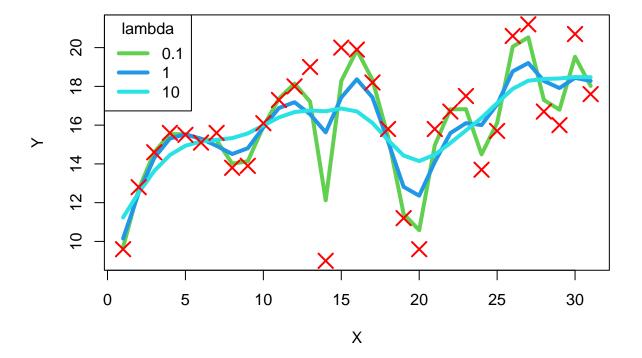
# 4. Smoothing spline

### (a) various smoothing parameter

```
smspe<-function(yy, lambda)
{
#(1)
   nd <- length(yy)
#(2)
   ss <- c(1, -2, 1, rep(0, nd - 3))
   ss <- rbind(ss, c(-2, 5, -4, 1, rep(0, length = nd - 4)))
   for(ii in 1:(nd - 4)) {
      ss <- rbind(ss, c(rep(0, ii - 1), 1, -4, 6, -4, 1, rep(0, nd - ii - 4)))
   }
}</pre>
```

```
ss <- rbind(ss, c(rep(0, length = nd - 4), 1, -4, 5, -2))
  ss <- rbind(ss, c(rep(0, length = nd - 3), 1, -2, 1))
#(3)
  ssi <- diag(nd) + lambda * ss
#(4)
  ey <- solve(ssi, yy)</pre>
  ey <- as.vector(ey)</pre>
#(5)
  return(ey)
plot(xx, yy, type = "n", xlab = "X", ylab = "Y", main="smoothing spline")
for (lambda in 10^(-1:1)){
  ey <- smspe(yy, lambda)</pre>
 lines(xx, ey, lwd = 4, col=log10(lambda)+4)
}
points(xx, yy, pch =4, lwd = 2, col = "red", cex = 2)
legend("topleft", legend=c(0.1, 1, 10), col=3:5, lty=1, lwd=4, title="lambda")
```

## smoothing spline



#### (b) movement of data

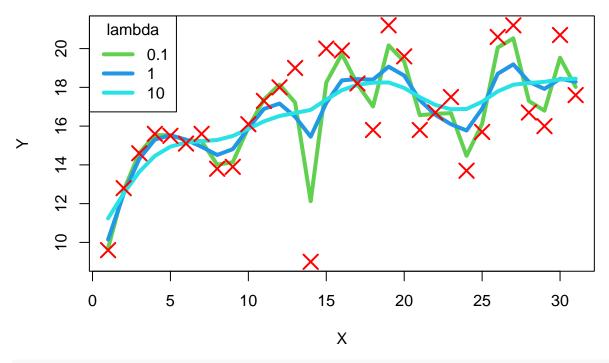
```
# movement of data
yy[19] <- yy[19] + 10

yy[20] <- yy[20] + 10

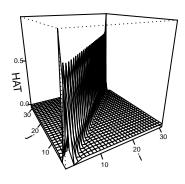
plot(xx, yy, type = "n",xlab = "X", ylab = "Y", main="smoothing spline")
for (lambda in 10^(-1:1)){
    ey <- smspe(yy, lambda)
    lines(xx, ey, lwd = 4, col=log10(lambda)+4)
}

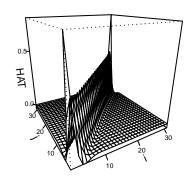
points(xx, yy, pch =4, lwd = 2, col = "red", cex = 2)
legend("topleft", legend=c(0.1, 1, 10), col=3:5, lty=1, lwd=4, title="lambda")</pre>
```

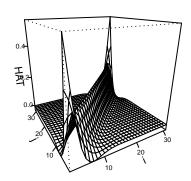
## smoothing spline



```
# hat matrix
par(mfrow=c(1,3),mai = c(0,0.2,1,0.2), oma = c(0,0,0,0))
for (lambda in 10^(-1:1)){
    ww <- apply(diag(nd), 2, smspe, lambda=lambda)
    persp(1:31, 1:31, ww, xlab = "i", ylab = "j", zlab = "HAT",
        lab = c(3, 3, 3), theta = -30, phi = 20,
        ticktype = "detailed", nticks=3,cex.lab=1,cex.axis=0.6,
        main=paste0('lambda = ',lambda))
}</pre>
```

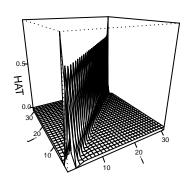


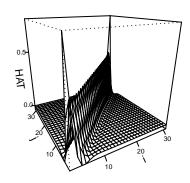


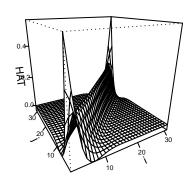


### (c) hat matrix

lambda = 0.1 lambda = 1 lambda = 10







# 5. Hat matrix of a simple regression

(a)

$$E = \sum_{i=1}^{n} (\alpha_{o} + \alpha_{i} \chi_{i} - Y_{i})^{2}$$

$$\begin{cases} \frac{\partial E}{\partial \alpha_0} = \sum_{i=1}^n 2 (\alpha_0 + \alpha_1 X_i - Y_i) = 0 \\ \frac{\partial E}{\partial \alpha_1} = \sum_{i=1}^n 2 X_i (\alpha_0 + \alpha_1 X_i - Y_i) = 0 \end{cases}$$

$$n\hat{a_o} + \hat{a_i} \sum x_i - \sum Y_i = 0 \implies \hat{a_o} = \overline{Y} - \hat{a_i} \overline{x}$$

$$\hat{\alpha_0} \sum X_i + \hat{\alpha_1} \sum X_i^2 - \sum X_i Y_i = 0 \quad \text{Sax} = \sum (X_i - \widehat{X})^2 = \sum X_i^2 - n \, \overline{X}^2, \quad \text{Say} = \sum (X_i - \widehat{X}) (Y_i - \widehat{Y}) = \sum X_i Y_i - n \, \overline{X} \, \overline{Y}$$

$$\Rightarrow (\overline{Y} - \hat{\alpha}, \overline{X}) \cdot n \overline{X} + \hat{\alpha}_1 (S_{XX} + n \overline{X}^2) - (S_{XY} + n \overline{X} \overline{Y}) = 0$$

$$\therefore \hat{\Omega}_i = \frac{Sxy}{Sxx}$$

$$cf) H = \begin{bmatrix} \frac{\partial^2 E}{\partial \alpha_0^2} & \frac{\partial^2 E}{\partial \alpha_0 \partial \alpha_1} \\ \frac{\partial^2 E}{\partial \alpha_0 \partial \alpha_0} & \frac{\partial^2 E}{\partial \alpha_0} \end{bmatrix} = 2 \begin{bmatrix} n & \sum X_1 \\ \sum X_1 & \sum X_1^2 \end{bmatrix} \qquad n \sum X_1^2 = 0 \implies \text{convex}$$

(b)

From (a), 
$$\hat{Y}_k = \hat{a_0} + \hat{a_1} X_k = \overline{Y} + \frac{Sxy}{Sxx} (X_k - \overline{X})$$

$$\overline{Y}^{*} = \overline{Y} + \frac{1}{n} \Delta Y_{K}$$
  $S_{XY}^{*} = \sum_{i \neq k} (X_{i} - \overline{X})(Y_{i} - \overline{Y}^{*}) + (X_{K} - \overline{X})(Y_{K} + \Delta Y_{K} - \overline{Y}^{*})$ 

$$=\sum_{i \neq k} X_i Y_i - \widehat{Y}^{*} \sum_{i \neq k} X_i - \widehat{X} \sum_{i \neq k} Y_i + (n-1) \widehat{X} \widehat{Y}^{*} + X_k Y_k + X_k \Delta Y_k - \widehat{Y}^{*} X_k - \widehat{X} Y_k - \widehat{X} \Delta Y_k + \widehat{X} \widehat{Y}^{*}$$

$$= \sum_{i=1}^{r-1} X_i Y_i - \overline{Y}_{*} u \underline{X} - \underline{X} u \underline{A}_{*} + u \underline{X} \underline{A}_{*} + \chi \kappa \nabla_{k}$$

$$\dot{Y}_{k} + \Delta \hat{Y}_{k} = \bar{Y}^{k} + \frac{S_{xy}^{k}}{S_{xx}} (X_{k} - \bar{X})$$

$$= \widehat{\gamma} + \frac{1}{n} \Delta \gamma_k + \frac{(\chi_{k-\overline{\chi}}) S_{xy} + (\chi_{k-\overline{\chi}})^2 \Delta \gamma_k}{S_{xx}}$$

(c)

$$\Delta \hat{Y}_{E} = (\hat{Y}_{E} + \Delta \hat{Y}_{E}) - \hat{Y}_{E} = \left[ \hat{Y} + \frac{1}{n} \Delta Y_{E} + \frac{(X_{E} - \bar{X})S_{AY} + (X_{E} - \bar{X})^{2} \Delta Y_{E}}{S_{AX}} \right] - \left[ \hat{Y} + \frac{S_{AY}}{S_{AX}} (X_{E} - \bar{X}) \right]$$

$$= \frac{1}{n} \Delta Y_{E} + \frac{\Delta Y_{E}}{S_{AX}} (X_{E} - \bar{X})^{2}$$

(d)

With equispaced predictor,

$$\hat{Y_{k}} = \sum_{i=1}^{j=1} [H]^{\kappa_{i}} Y_{i} \Rightarrow Y_{k} \text{ on } \Delta Y_{k} \stackrel{?}{\leq} CHOHE, \quad \hat{Y_{k}} + \Delta \hat{Y_{k}} = \sum_{i\neq k} [H]^{\kappa_{i}} Y_{i} + [H]^{\kappa_{k}} (Y_{k} + \Delta \hat{Y_{k}})$$

۵/km 1을 대입하면 [H] للاها إلى ضور كيل.

$$: [H]^{kk} = \frac{1}{i} + \frac{2^{xx}}{(x^{k-x})^{x}}$$

(e)

$$\textstyle \sum_{i=1}^{n} [\mathsf{H}]_{ii} = \sum_{i=1}^{n} \frac{1}{n} + \frac{(X_k - \bar{X})^2}{S_{xx}} = 2.$$