• If the spectrum of a random variable *X* is finite or countably infinite then the distribution of *X* is called a *discrete distribution*.

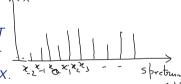
- If the spectrum of a random variable X is finite or countably infinite then the distribution of X is called a discrete distribution.
- Let spectrum of X: $T = \{x_i : i = 0, \pm 1, \pm 2, ...\}$ with

$$\dots < x_{-2} < x_{-1} < x_0 < x_1 < x_2 < \dots$$

The function $f_X : \mathbb{R} \to [0,1]$, defined as

$$f_X(x) = \begin{cases} P(X = x_i), & x = x_i \in T \\ 0, & \text{elsewhere.} \end{cases}$$

is called the *probability mass function* (p.m.f.) of X.



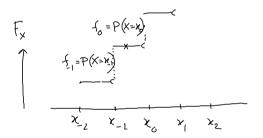
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•
$$F_X(x) = \sum_{\substack{x_j \le x_i \\ \hat{i}}} P(X = x_j) \text{ if } x_i \le x \angle x_{i+1}$$

$$= \sum_{\substack{j = -\infty \\ \hat{j} = -\infty}} f_X(x_j)$$

•
$$F_X(x) = \sum_{x_j \le x_i} P(X = x_j)$$
 if $x_i \le x \angle x_{i+1}$

• F_X is a step function with steps $f_i = P(X = x_i)$ for $i = 0, \pm 1, \pm 2, ...$



Discrete Distribution: Properties

1.
$$\sum_{j=-\infty}^{\infty} f_j = 1$$
 We have, $F_X(\infty) = 1$

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Discrete Distribution: Properties

$$1. \sum_{j=-\infty}^{\infty} f_j = 1$$

2. At each non-spectrum point a, P(X = a) = 0

$$P(X=a) = F_X(a) - F_X(a-0)$$

let, Ixk, xk+1 & spectrum of X s.t.

$$F_{x}(a) = \sum_{j=-\infty}^{k} f_{x}(x_{j}), \quad F(a-o) = \sum_{j=-\infty}^{k} f_{x}(x_{j})$$

$$\Rightarrow P(x-a) = \emptyset.$$

Discrete Distribution: Properties

$$1. \sum_{j=-\infty}^{\infty} f_j = 1$$

2. At each non-spectrum point a, P(X = a) = 0

3.
$$P(a < X \le b) = \sum_{a < x_i \le b} f_X(x_i)$$

$$P(a < X \le b) = F_X(b) - F_X(a)$$

Discrete Distribution: Examples

1. Binomial (n,p) Distribution

$$\begin{array}{l} \text{Spectrum of } \times = \left\{0, 1, 2, \dots, n\right\} \\ \text{Probabilt mano fn.} \left(p.m.f \right) \\ \text{f}_{\chi}(x) = \binom{n}{\chi} p^{\chi} \binom{1-p}{n-\chi}, \quad \chi=0, 1, 2, \dots \\ = 0, \quad \text{elsewhen} \\ \sum_{\chi=0}^{n} \binom{n}{\lambda} p^{\chi} \binom{1-p}{n-\chi} = \binom{p+(1-p)}{n-\chi} = 1. \end{array}$$

Discrete Distribution: Examples

2. Poisson (μ) Distribution

Continuous Distribution

Definition

The distribution of a random variable X is said to be continuous if

- \checkmark 1. the distribution function F_X is continuous
 - 2. $\frac{d}{dx}F_X(x) = F_X'(x)$ is piecewise continuous in $(-\infty,\infty)$

<u>Define</u>: $f_X : \mathbb{R} \to [0,1]$ as $f_X(x) = \frac{d}{dx} F_X(x)$ which is called the <u>probability</u> <u>density function</u> (p.d.f.) of X.

1.
$$f_X \ge 0$$

 $f_X(x) \ge 0 \quad \forall x \in \mathbb{R}$

$$f_{x}(x) = \frac{d}{dx} F_{x}(x) \ge 0$$
 Since $F_{x}(x)$ is mon. inc. fr.

1. $f_X \ge 0$

2.
$$P(a < X \le b) = \int_{a}^{b} f_X(x) dx$$

$$P(a < X \leq b) = F_{X}(b) - F_{X}(a)$$
.

(i)
$$f_{X}(x) = \frac{d}{dx} F_{X}(x)$$
 (primitive exists)

(1)
$$f_{x}(x)$$
 has finite no. of jmp discontinuities (1st kind)
in $[a,b] \Rightarrow f_{x}(x)$ is $[a,b] = f_{x}(x)$

$$\Rightarrow$$
 Applying Fundamental the of Integral Calculus
$$\int_{a}^{b} f_{x}(x) dx = F_{x}(b) - F_{x}(a) = P(a < x \le b).$$

$$a$$
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1.
$$f_X \ge 0$$

2.
$$P(a < X \le b) = \int_{a}^{b} f_X(x) dx$$

3.
$$F_X(x) = \int_{-\infty}^x f_X(x) dx$$

$$F(-\infty) = \bigcirc \Rightarrow F_X(x) = \int_{-\infty}^x f_X(x) dx.$$

1.
$$f_X \ge 0$$

2.
$$P(a < X \le b) = \int_{a}^{b} f_X(x) dx$$

3.
$$F_X(x) = \int_{-\infty}^{x} f_X(x) dx$$

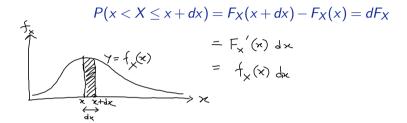
4.
$$\int_{-\infty}^{\infty} f_X(x) dx = 1 \qquad \text{Since} \qquad F_{\chi} (\omega) = 1.$$

- 1. $f_X \ge 0$
- 2. $P(a < X \le b) = \int_{a}^{b} f_X(x) dx$
- 3. $F_X(x) = \int_{-\infty}^{x} f_X(x) dx$
- $4. \int_{-\infty}^{\infty} f_X(x) dx = 1$
- 5. P(X = a) = 0 for a given constant a

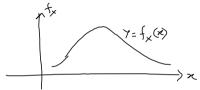
$$P(X=a) = F_x(a) - F_x(a-o) = 0$$
 Since F_x is continuon.

- 1. $f_X \ge 0$
- 2. $P(a < X \le b) = \int_a^b f_X(x) dx$
- $\vee 3. F_X(x) = \int_{-\infty}^{x} f_X(x) dx$
 - $4. \int_{-\infty}^{\infty} f_X(x) dx = 1$
 - 5. P(X = a) = 0 for a given constant a
 - 6. **Converse statement:** Every non-negative, real-valued, piecewise-continuous function f that is integrable in $(-\infty,\infty)$ and satisfies $\int_{-\infty}^{\infty} f(x) dx = 1$, is the probability density function of a continuous distribution.

7. **Probability Differential:** Let X has continuous distribution. In differential notation we write:



8. **Density Curve:** The curve $y = f_X(x)$ is called the probability density curve of the corresponding continuous distribution.



1. Uniform (a, b)

$$f_{x}(x) = \frac{1}{b-a}, \quad a < x < b$$

$$= 0, \quad e | seuhene$$

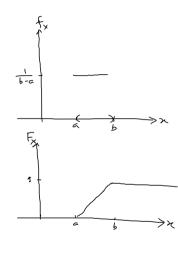
$$\frac{Nolz!}{b-a} \quad \frac{1}{b-a}$$

$$\int_{-\infty}^{\infty} f_{x}(x) dx = 1.$$
Fig.

$$\frac{x \leqslant \alpha}{\alpha \leqslant x \leqslant b} F_{x}(x) = 0$$

$$\frac{x \leqslant \alpha}{-x} F_{x}(x) = \int_{-\infty}^{x} f_{x}(x) dx = \int_{-\infty}^{x} \frac{1}{b-\alpha} dx$$

$$= \frac{x - \alpha}{b-\alpha}$$



2. Normal
$$(m, \sigma)$$
, $\times \sim N_{\sigma rmd}(m, \sigma)$

$$\int_{X} (x) = \frac{1}{\sqrt{2\pi}\sigma} e^{\frac{(x-m)^{2}}{2\sigma^{2}}}, \quad -\infty < x < \infty = \frac{1}{\sqrt{\pi}} \sqrt{\pi}$$
where $\sigma = \sqrt{\frac{1}{2}} \sqrt{\frac{1}{2}}$

$$\frac{N_{\sigma} k^{2}}{\sqrt{1}} = \sqrt{\frac{1}{2}} \sqrt{\frac{1}{2}}$$

In partial of m=0, $\sigma=1$ normal distribution will be called Standard Normal Bistribution. denit for: $\Rightarrow (x) = \frac{1}{\sqrt{2\pi}} = \frac{x^2}{2}$, $\frac{1}{\sqrt{2\pi}} \left(\frac{x}{\sqrt{2\pi}} \right) = \frac{1}{\sqrt{2\pi}} \left(\frac{x}{\sqrt{2\pi}} \right)^{2} = \frac{x^{2}}{\sqrt{2\pi}} dx$

3. Cauchy (λ, μ) Distribution

$$f_{\chi}(x) = \frac{1}{\pi} \cdot \frac{\lambda}{\lambda^2 + (x - \mu)^2} \quad , \quad -\alpha < x < \infty$$

$$\lambda > 0, \quad \mu \quad \text{as } t = parents$$

$$0 \quad f_{\chi}(x) \geq 0 \quad \forall x$$

$$0 f_{x}(x) \geq 0 \forall x$$

4. Gamma (1) distribution

or,
$$\times \sim V(l)$$
, $l > 0$

$$f_{\chi}(x) = \frac{1}{\Gamma(l)} e^{-\chi} x^{l-1}, \quad 0 < x < \infty$$

$$= 0, \quad \text{elsewhen.}$$

$$\underline{Note:} \quad () f_{\chi}(x) \ge 0 \quad \& \quad (i) \quad (f_{\chi}(x) \le x - 1).$$

5. Beta (I, m) distribution of 1st kind

$$\begin{array}{ll}
\times \sim \beta_{1}(l,m) & (l,m) \\
f_{\chi}(x) = \frac{1}{B(l,m)} \times^{l-1} (1-x)^{m-1}, & 0 < x < 1.
\end{array}$$

$$= 0, \quad \text{elsewh}.$$

6. Beta (I, m) distribution of 2nd kind

$$\times \sim \beta_{2}(\ell, m), \qquad (\ell, m > 0)$$

$$f_{x}(x) = \frac{1}{B(\ell, m)} \frac{x^{\ell-1}}{(1+x)^{\ell+m}}, \quad 0 < x < \infty$$

 $P(X=1) = \frac{1}{6}$, $P(X=2) = \frac{1}{6}$, $P(X=3) = \frac{1}{2}$

Prop @

Let X: con. to score Spection of X: {0, 1, 2, 3}

Let R be to r.v. corr to distance of hit from O $\oint_{\mathcal{D}} (r) = \frac{2}{\pi} \frac{1}{1+r^2} , \quad 0 \le r < \infty$

$$f_{R}(r) = \frac{2}{\pi} \frac{1}{1+r^{2}}, \quad 0 \le r < \infty$$

$$= 0, \quad \text{elseul}$$

$$P(X=0) = P(R > \sqrt{3}) = 1 - P(R \le \sqrt{3}) = 1 - \frac{2}{n} \frac{1}{1+r^{2}}$$

 $P(X=0) = P(R > \sqrt{3}) = 1 - P(R \le \sqrt{3}) = 1 - \frac{2}{n} \int_{-1+r^2}^{1} dr$ $=1-\frac{2}{\pi}\left[\tan^{-1}x\right]^{\frac{1}{3}}$

$$f_{R}(r) = \frac{2}{\pi} \frac{1}{1+r^{2}}, \quad 0 \le r < \infty$$

$$= 0, \quad \text{elseul}$$

$$P(X=0) = P(R) \sqrt{3} = 1 - P(R \le \sqrt{3}) = 1 - \frac{2}{\pi} \sqrt{\frac{1+r^{2}}{1+r^{2}}}$$

To find
$$P(x=0)$$
, $P(x=1)$,...

to distance of hit from O
, $O \le Y < \infty$

 $= 1 - \frac{2}{71} \frac{\pi}{3} = \frac{1}{3}$