

Review

$$\mathbf{P}(A \mid B) = \frac{\mathbf{P}(A \cap B)}{\mathbf{P}(B)}, \quad \text{assuming } \mathbf{P}(B) > 0$$

- Multiplication rule:

$$\mathbf{P}(A \cap B) = \mathbf{P}(B) \cdot \mathbf{P}(A \mid B) = \mathbf{P}(A) \cdot \mathbf{P}(B \mid A)$$

- Total probability theorem:

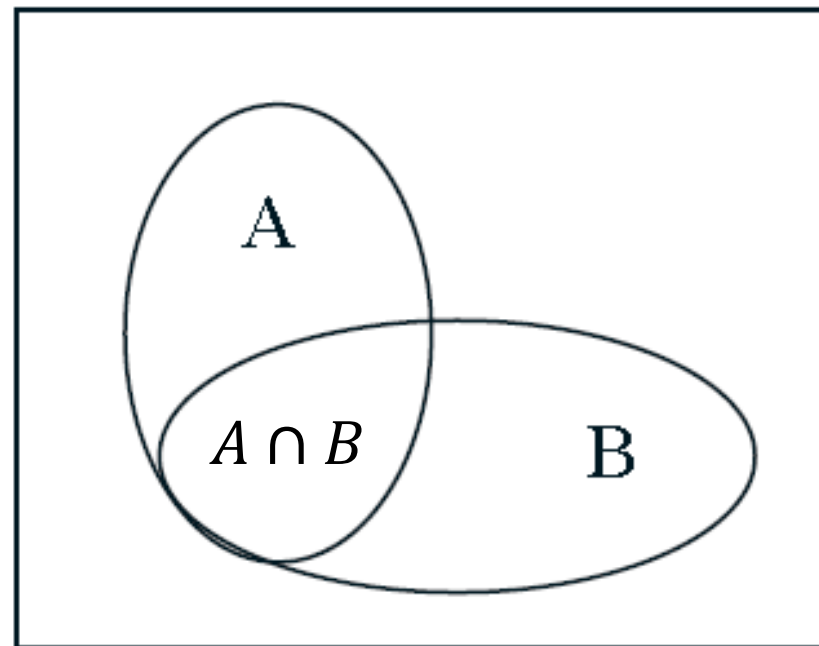
$$\mathbf{P}(B) = \mathbf{P}(A)\mathbf{P}(B \mid A) + \mathbf{P}(A^c)\mathbf{P}(B \mid A^c)$$

- Bayes rule:

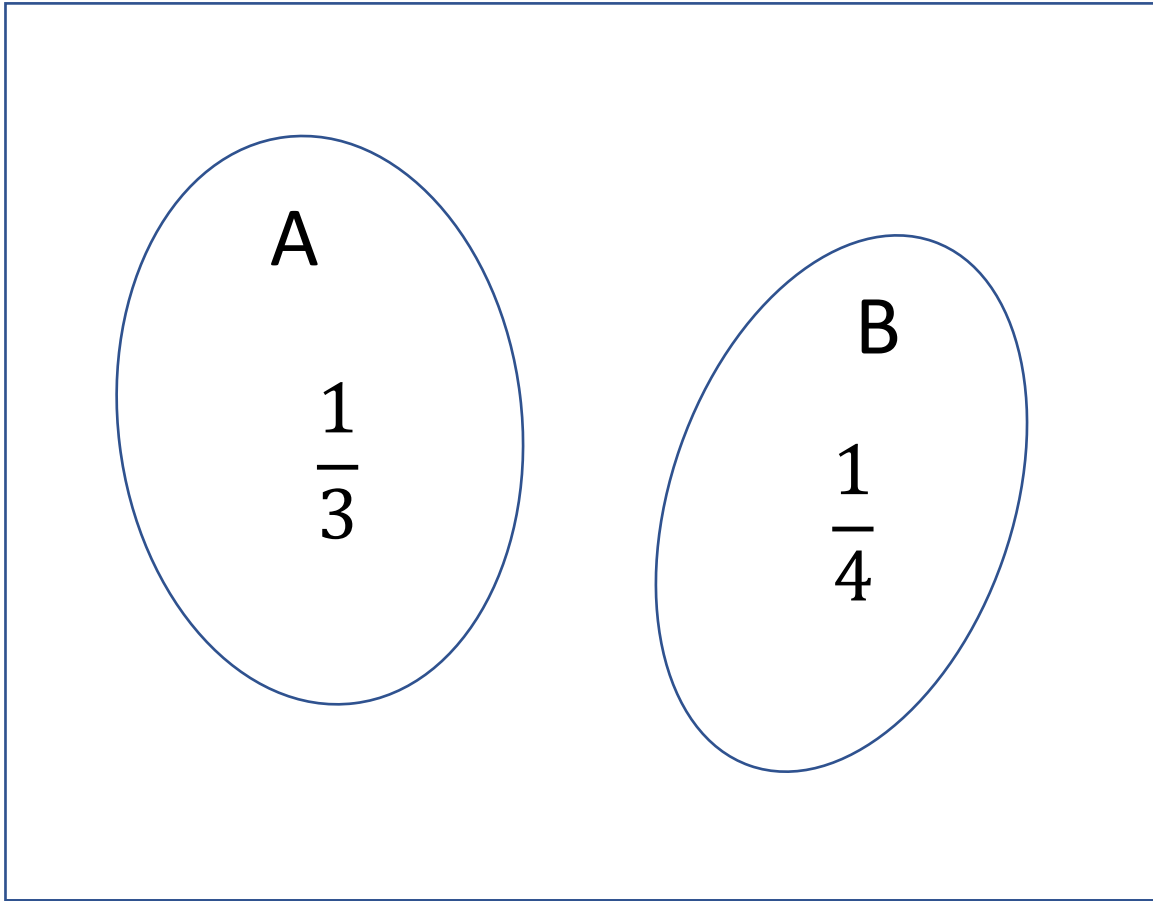
$$\mathbf{P}(A_i \mid B) = \frac{\mathbf{P}(A_i)\mathbf{P}(B \mid A_i)}{\mathbf{P}(B)}$$

Independence of two events

- **“Defn:”** $P(B | A) = P(B)$
 - “occurrence of A provides no information about B 's occurrence”
- Recall that $P(A \cap B) = P(A) \cdot P(B | A)$
- **Defn:** $P(A \cap B) = P(A) \cdot P(B)$
- Symmetric with respect to A and B
 - applies even if $P(A) = 0$
 - implies $P(A | B) = P(A)$



INDEPENDENCE is not to be confused with DISJOINTNESS



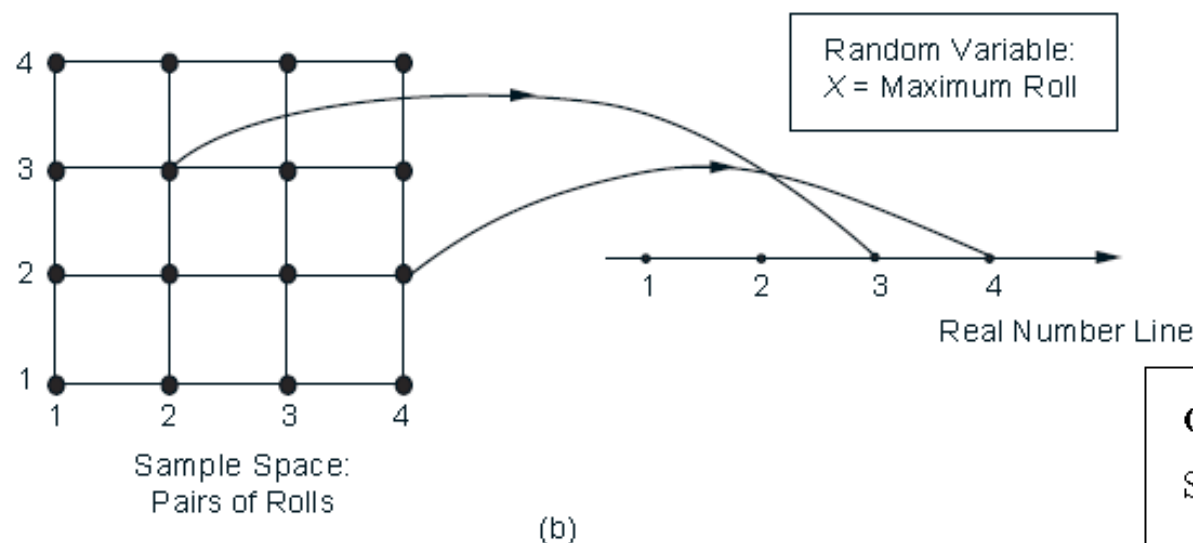
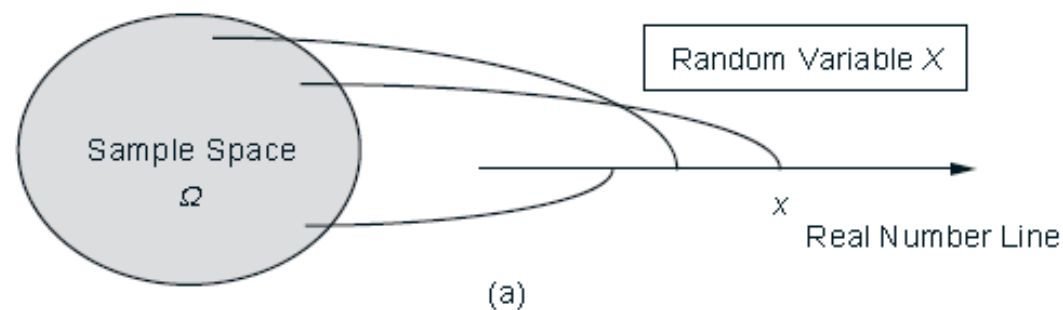
$$P(A \cap B) = 0$$
$$P(A)P(B) = \frac{1}{12}$$

NOT EQUAL

$$P(A) = \frac{1}{3}$$
$$P(A|B) = 0 \neq P(A)$$

NOT EQUAL

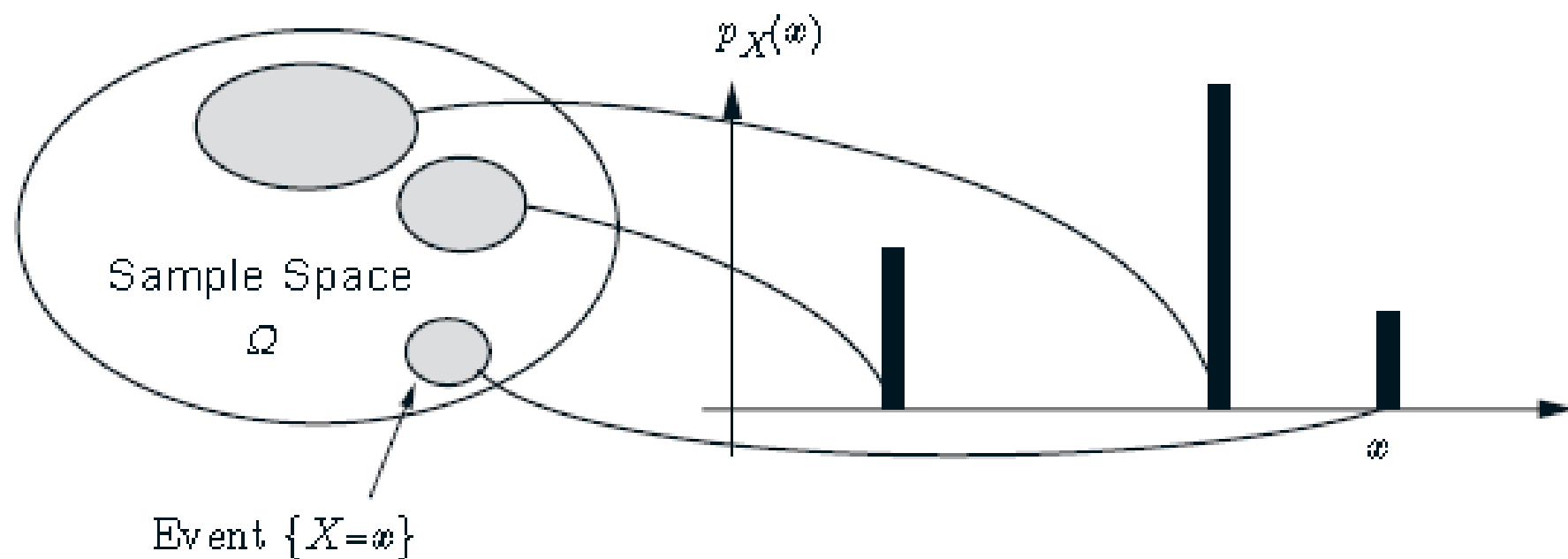
Discrete Random Variables



Concepts Related to Discrete Random Variables

Starting with a probabilistic model of an experiment:

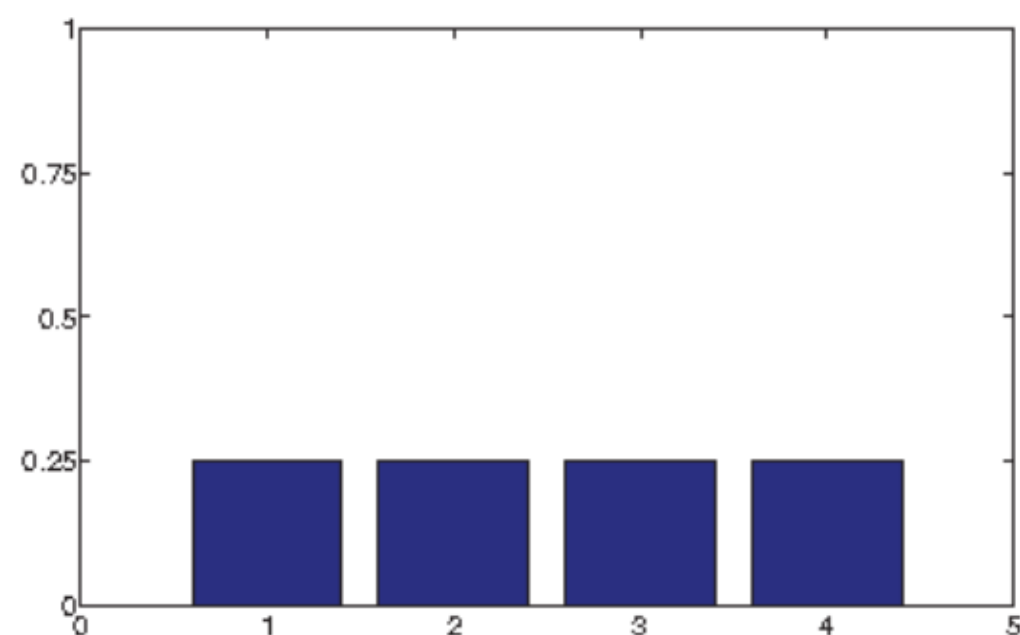
- A **discrete random variable** is a real-valued function of the outcome of the experiment that can take a finite or countably infinite number of values.
- A (discrete) random variable has an associated **probability mass function** (PMF), which gives the probability of each numerical value that the random variable can take.
- A **function of a random variable** defines another random variable, whose PMF can be obtained from the PMF of the original random variable.



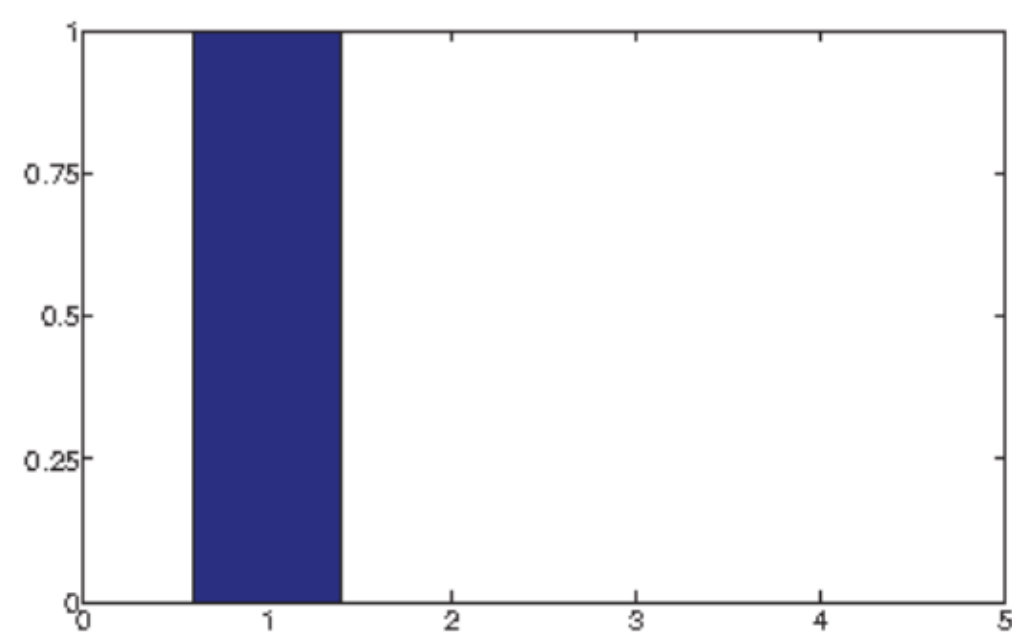
- Notation:

$$\begin{aligned}
 p_X(x) &= \mathbf{P}(X = x) \\
 &= \mathbf{P}(\{\omega \in \Omega \text{ s.t. } X(\omega) = x\})
 \end{aligned}$$

- $p_X(x) \geq 0 \quad \sum_{\omega} p_X(x) = 1$

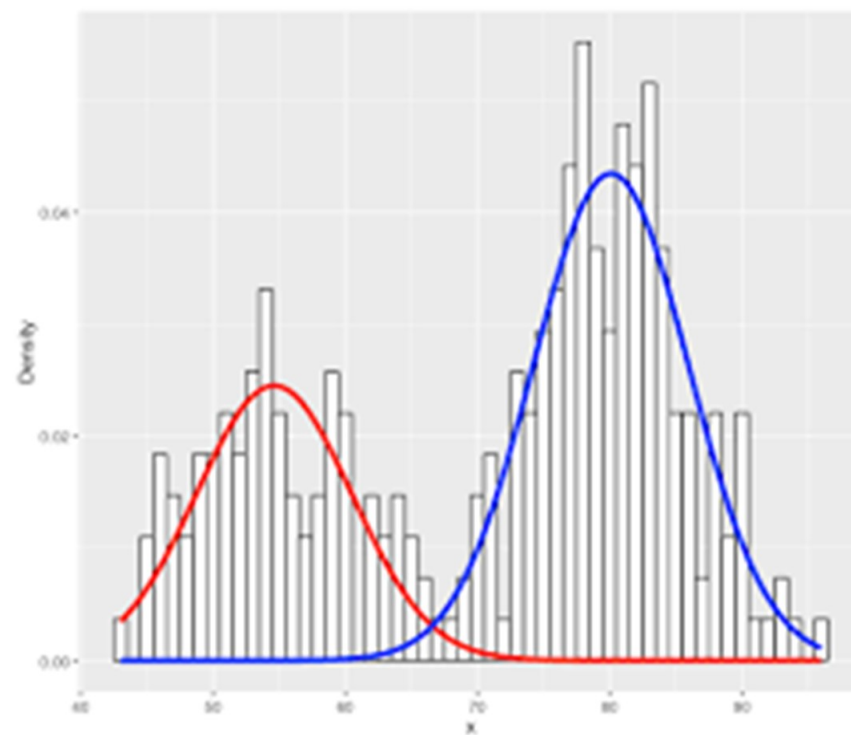
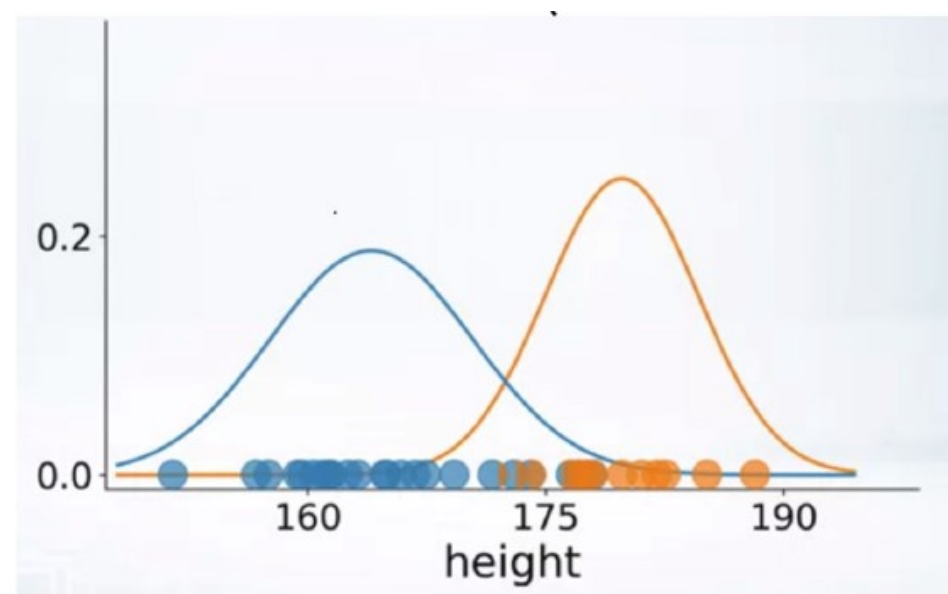
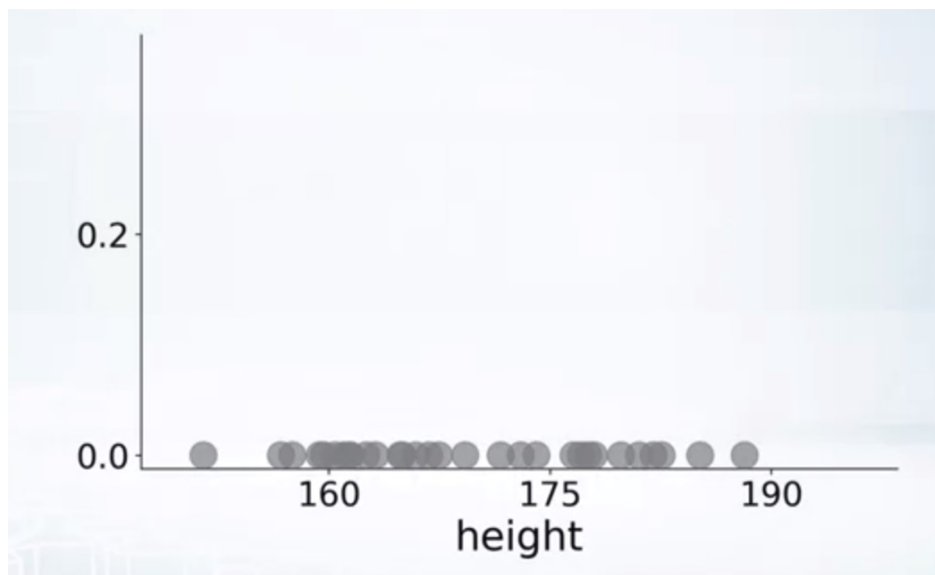


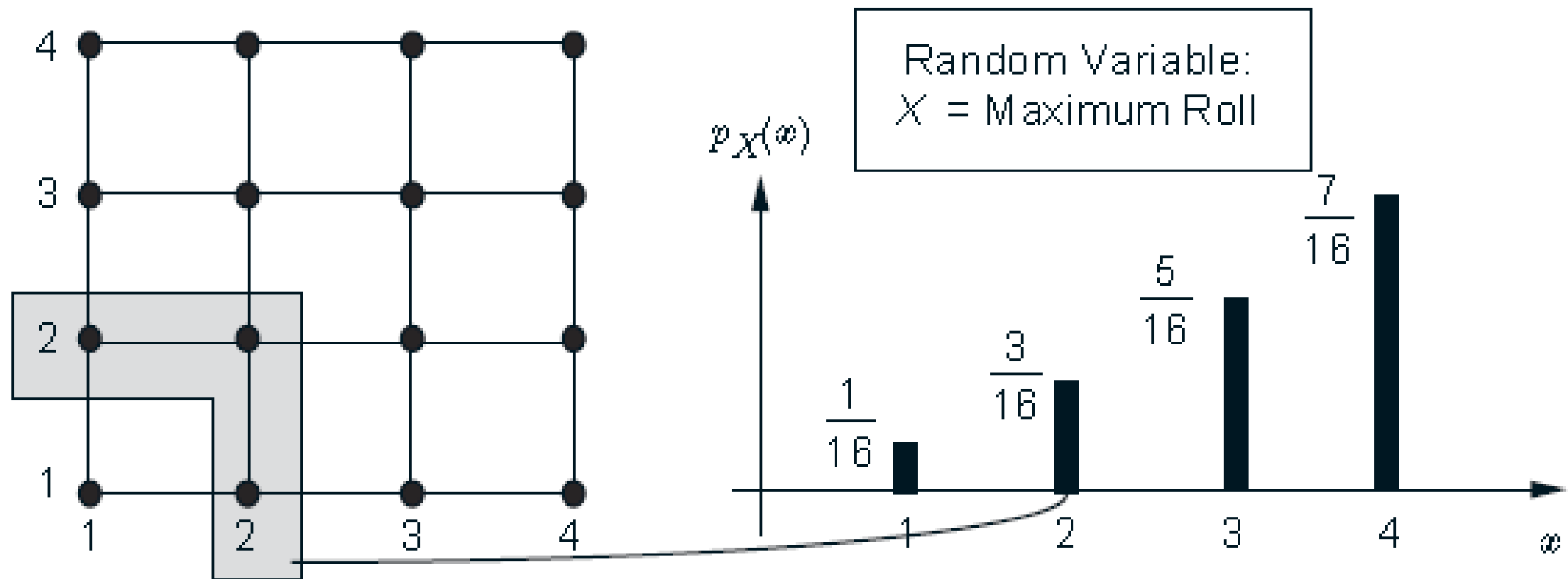
(a)



(b)

Figure 2.1 (A) a uniform distribution on $\{1, 2, 3, 4\}$, with $p(x = k) = 1/4$. (b) a degenerate distribution $p(x) = 1$ if $x = 1$ and $p(x) = 0$ if $x \in \{2, 3, 4\}$. Figure generated by `discreteProbDistFig`.





(b)

The Bernoulli Random Variable

$$X = \begin{cases} 1 & \text{if a head,} \\ 0 & \text{if a tail.} \end{cases}$$

Its PMF is

$$p_X(x) = \begin{cases} p & \text{if } x = 1, \\ 1 - p & \text{if } x = 0. \end{cases}$$

INDEPENDENT AND IDENTICALLY DISTRIBUTED IID

Each attempt is independent and has the same probability of success: the attempts are INDEPENDENT AND IDENTICALLY DISTRIBUTED or IID.

GEOMETRIC DISTRIBUTION

Example 5.2.1. Suppose we flip a fair coin several times in a row until we get a head. Let X be the total number of flips. What is the distribution over X ?

$$P(X = n) = (1/2)^n \text{ for positive integer } n.$$

Example 5.2.2. Suppose we perform the same type of experiment as in the previous example, but our coin isn't fair: the probability of getting a head is some value p . Now what is the distribution over X , the number of flips required to get a head?

$$P(TT \cdots TH)$$

$$P(X = n) = (1 - p)^{n-1} p \quad \text{for positive integer } n.$$

GEOMETRIC DISTRIBUTION

$$P(X = n) = (1 - p)^{n-1} p$$

for positive integer n .

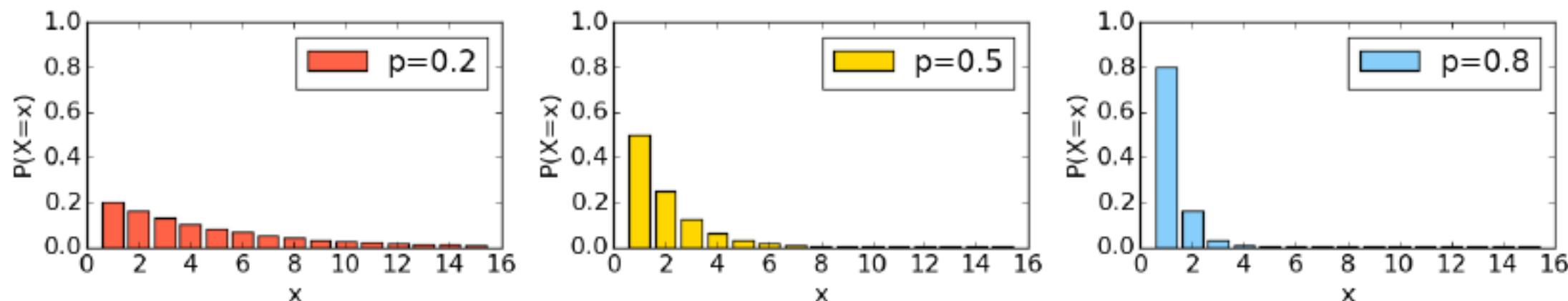


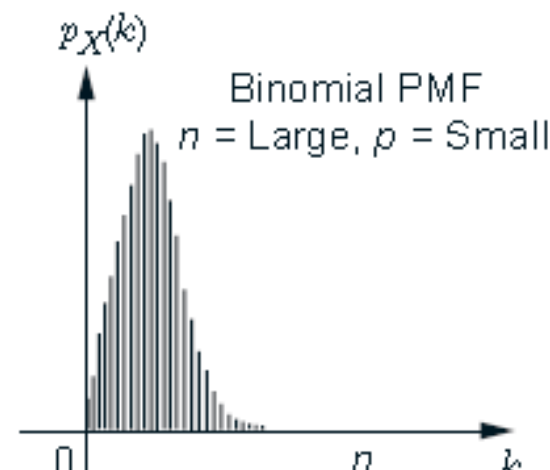
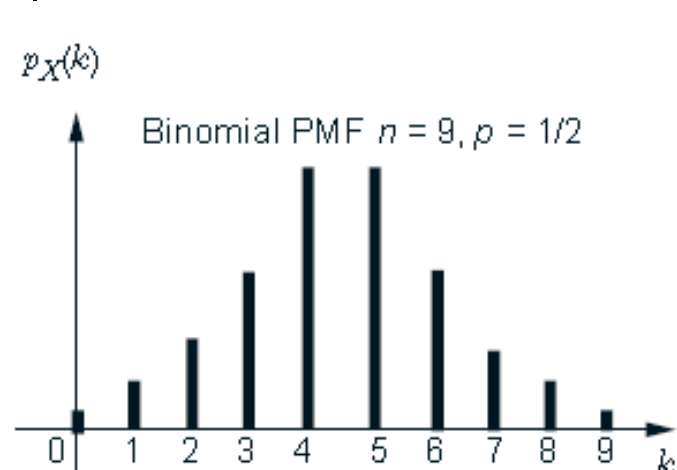
Figure 4: Three examples of geometric distributions, with parameter values of 0.2, 0.5, and 0.8. Possible values of X extend to infinity, but we only show values up to 15.

The Binomial Random Variable

A biased coin is tossed n times. At each toss, the coin comes up a head with probability p , and a tail with probability $1 - p$, independently of prior tosses. Let X be the number of heads in the n -toss sequence. We refer to X as a **binomial** random variable **with parameters n and p** . The PMF of X consists of the binomial probabilities that were calculated in Section 1.4:

$$p_X(k) = \mathbf{P}(X = k) = \binom{n}{k} p^k (1 - p)^{n-k}, \quad k = 0, 1, \dots, n.$$

$$\sum_{k=0}^n \binom{n}{k} p^k (1 - p)^{n-k} = 1$$



Marginal Distribution Calculation → Integrating “out” the unwanted variable

verbose	concise	
$P(X = x) = 1 - P(X \neq x)$	$P(x) = 1 - P(\neg x)$	complement (18)
$P(X = x) = \sum_y P(X = x, Y = y)$	$P(x) = \sum_y P(x, y)$	law of total prob (19)
$P(X = x Y = y) = \frac{P(X = x, Y = y)}{P(Y = y)}$	$P(x y) = \frac{P(x, y)}{P(y)}$	defn of cond prob (20)
$P(X = x, Y = y) = P(X = x Y = y)P(y)$	$P(x, y) = P(x y)P(y)$	product rule (21)
$P(X = x Y = y) = \frac{P(Y = y X = x)P(X = x)}{P(Y = y)}$	$P(x y) = \frac{P(y x)P(x)}{P(y)}$	Bayes' Rule (22)

JOINT PMFS OF MULTIPLE RANDOM VARIABLES

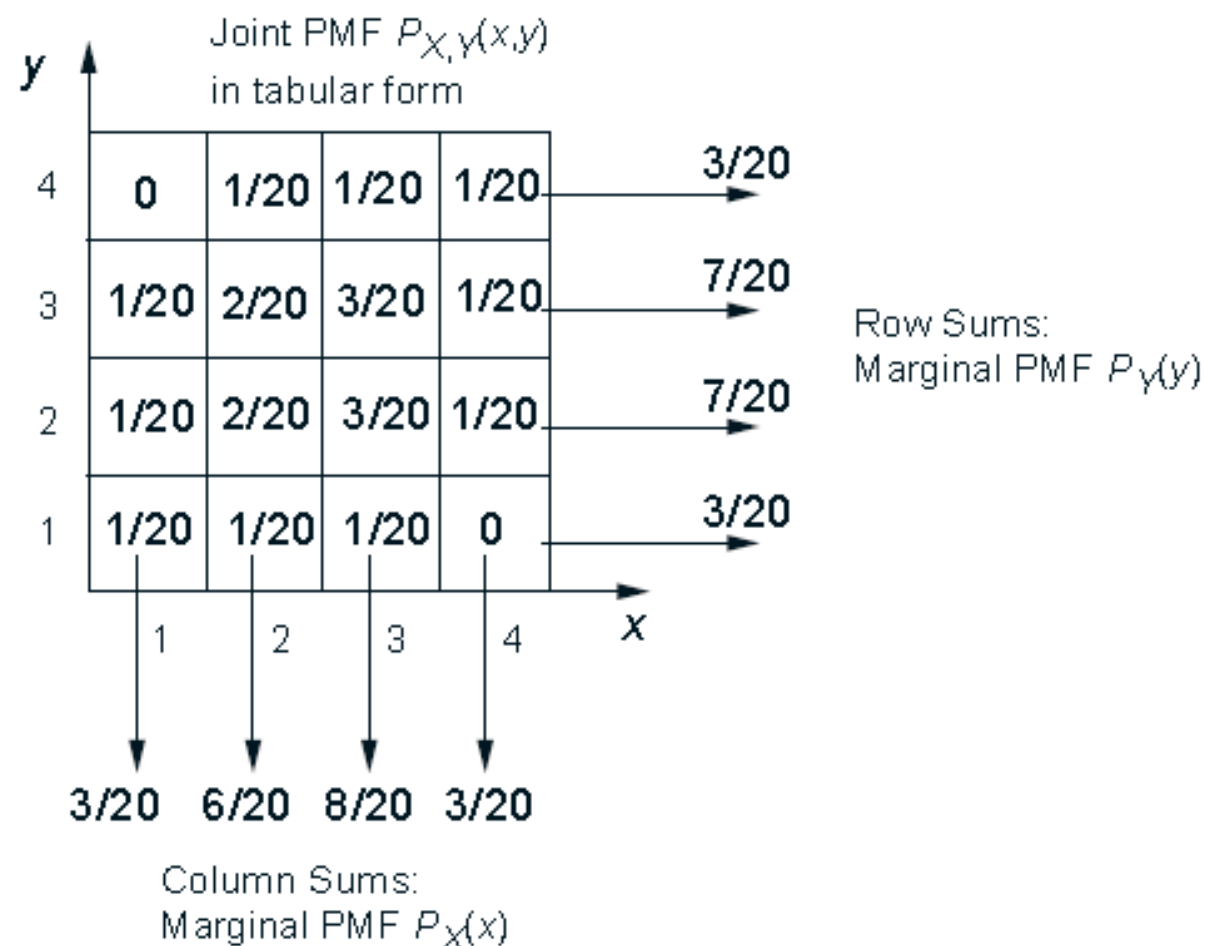
$$p_{X,Y}(x, y) = \mathbf{P}(X = x, Y = y)$$

$$\mathbf{P}((X, Y) \in A) = \sum_{(x,y) \in A} p_{X,Y}(x, y).$$

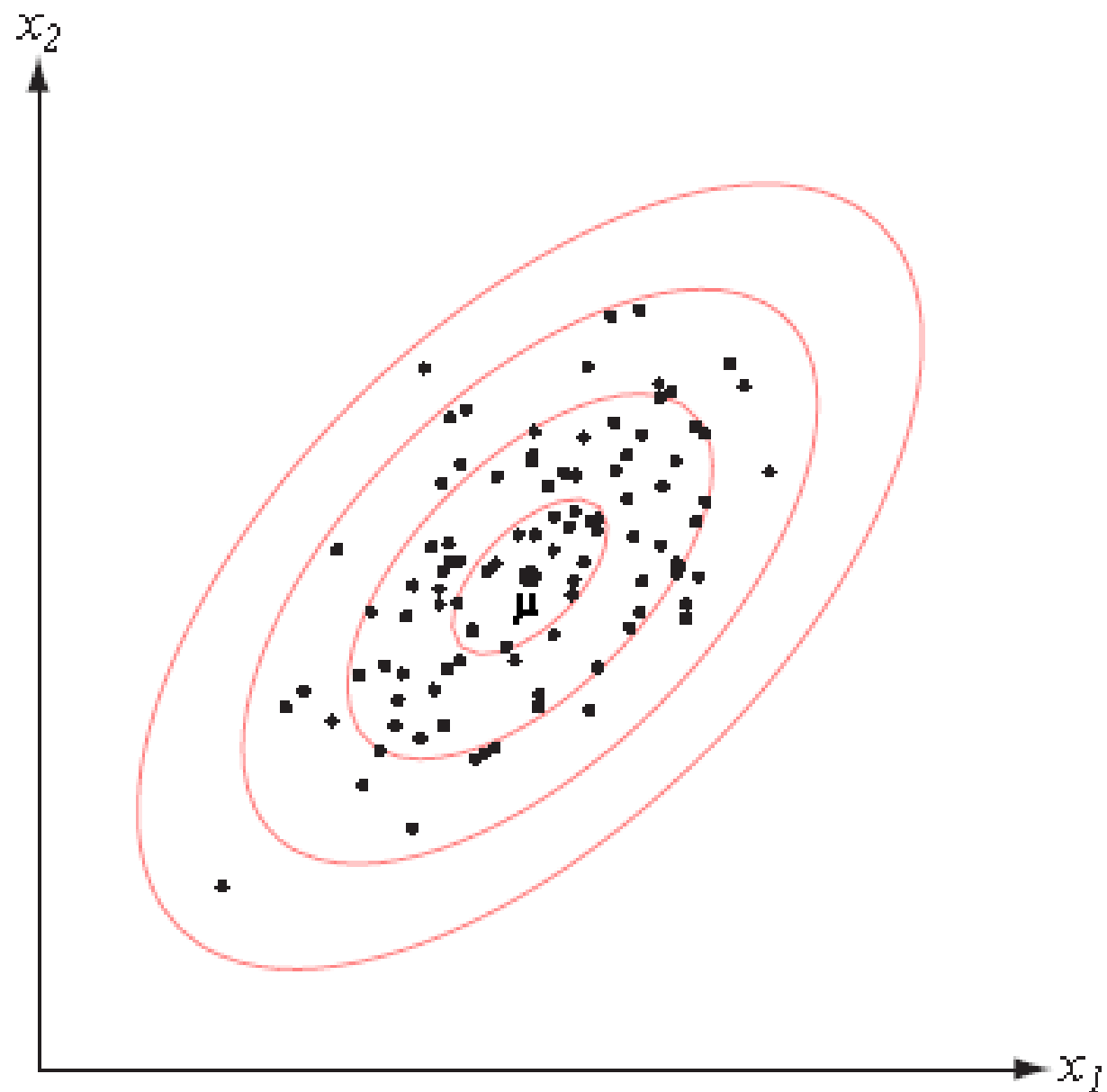
marginal PMFs:

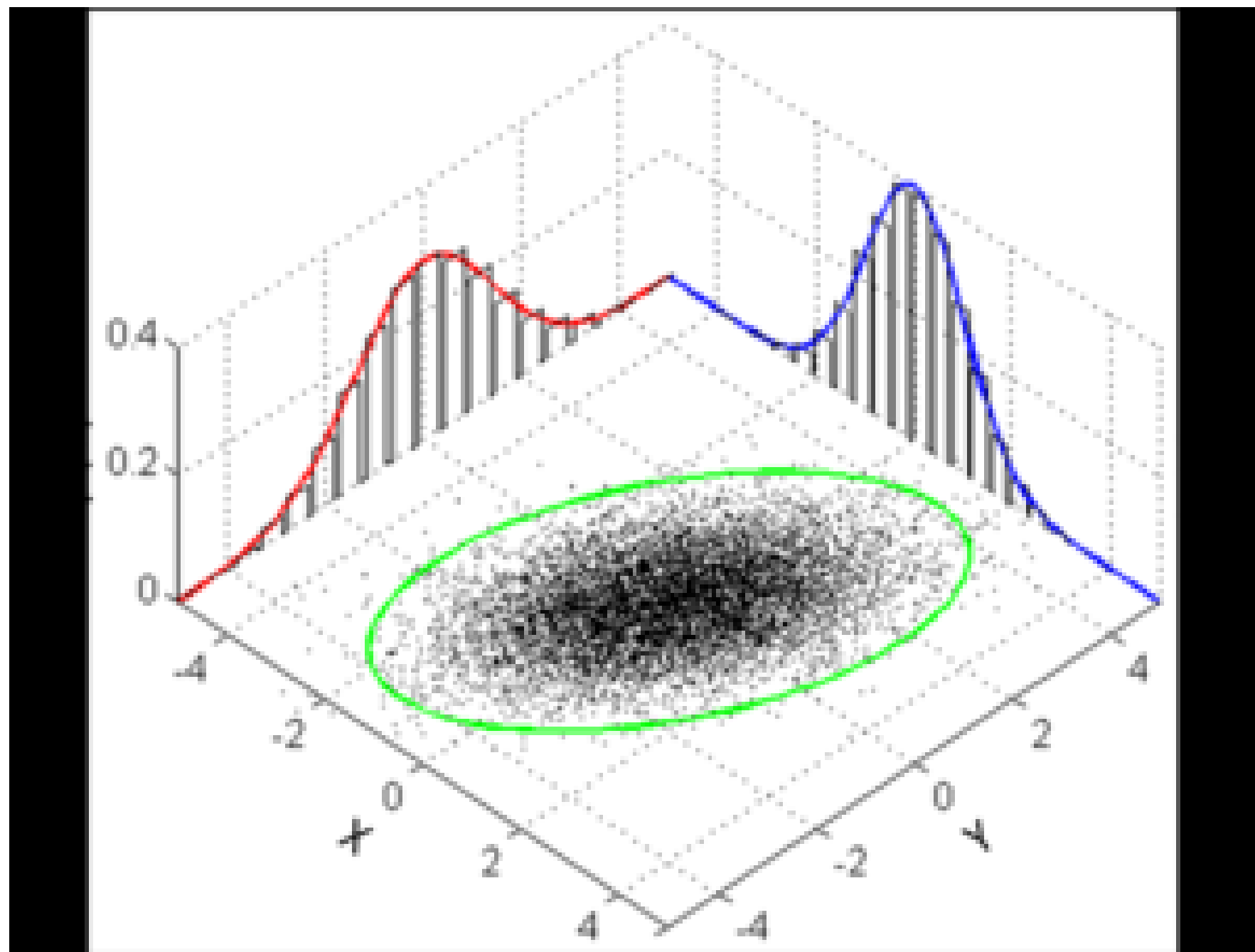
$$p_X(x) = \sum_y p_{X,Y}(x, y),$$

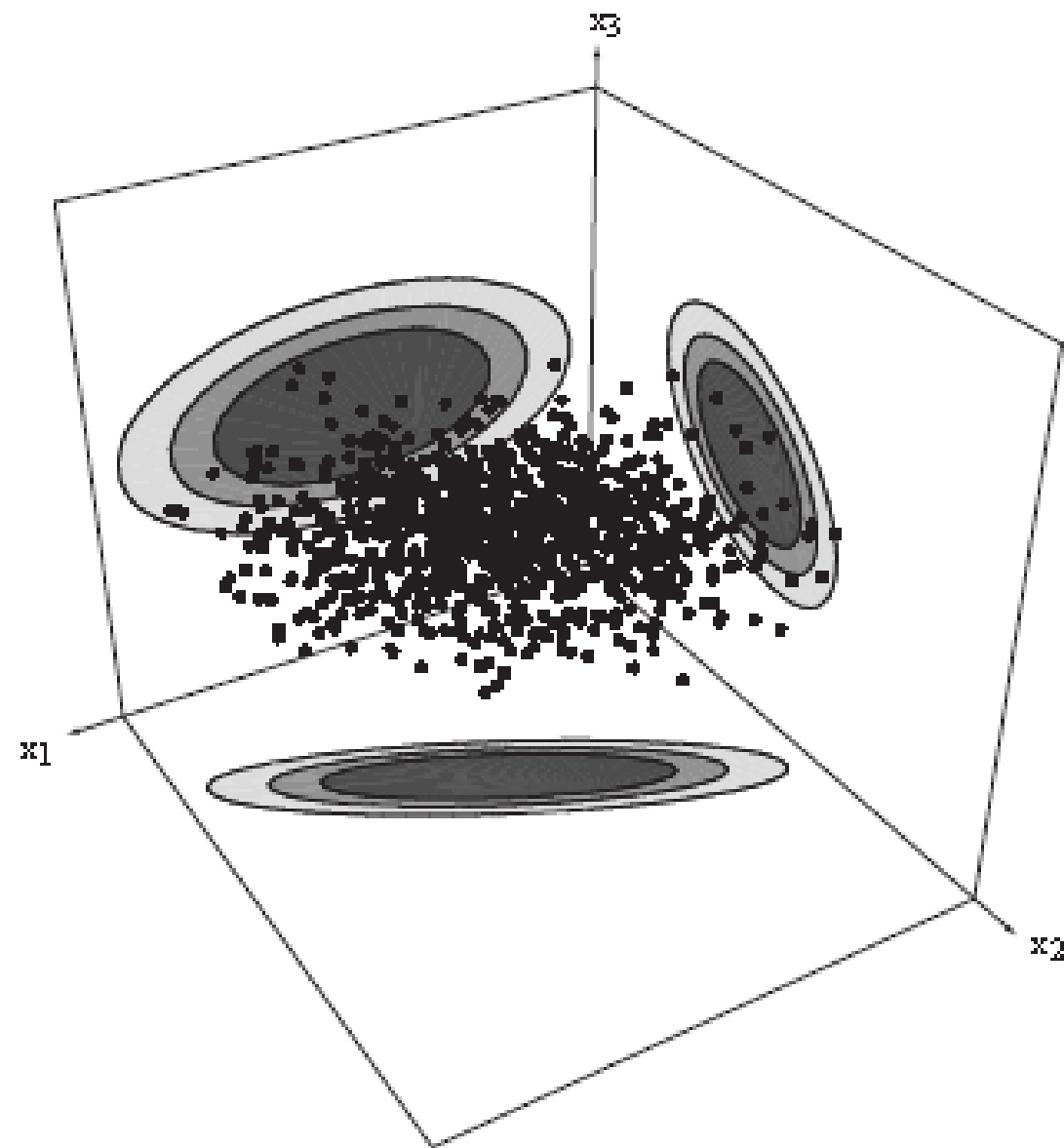
$$p_Y(y) = \sum_x p_{X,Y}(x, y).$$



$$\mathbf{P}((X, Y) \in A) = \sum_{(x, y) \in A} p_{X, Y}(x, y).$$

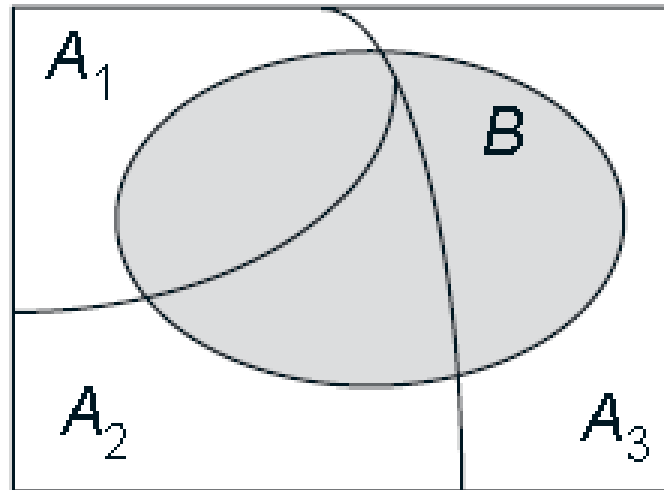






$$p(x) = \int_{-\infty}^{\infty} p(x, y) dy$$

law of tot prob, continuous version



$$B = (A_1 \cap B) \cup \dots \cup (A_n \cap B).$$

$$\mathbf{P}(B) = \mathbf{P}(A_1 \cap B) + \dots + \mathbf{P}(A_n \cap B).$$

$$P(B) = \sum_{i=1}^n P(A_i, B)$$

if X_1, X_2, \dots, X_n are **continuous random variables**, then the marginal **probability density function** should be

$$f_{X_i}(x_i) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} f(x_1, x_2, \dots, x_n) dx_1 dx_2 \dots dx_{i-1} dx_{i+1} \dots dx_n.$$

EXPECTED VALUE EXPECTATION

MEAN
 μ

Expectation

We define the **expected value** (also called the **expectation** or the **mean**) of a random variable X , with PMF $p_X(x)$, by[†]

$$\mathbb{E}[X] = \sum_x xp_X(x).$$

Expected Value Rule for Functions of Random Variables

Let X be a random variable with PMF $p_X(x)$, and let $g(X)$ be a real-valued function of X . Then, the expected value of the random variable $g(X)$ is given by

$$\mathbb{E}[g(X)] = \sum_x g(x)p_X(x).$$

$$M = \frac{m_1 k_1 + m_2 k_2 + \cdots + m_r k_r}{k}.$$

$$p_i \approx \frac{k_i}{k}, \quad i = 1, \dots, r.$$

$$M = \frac{m_1 k_1 + m_2 k_2 + \cdots + m_r k_r}{k} \approx m_1 p_1 + m_2 p_2 + \cdots + m_r p_r.$$

$$M = \sum_{i=1}^n m_i p_i \quad \Rightarrow \quad \mathbb{E}[X] = \sum_x x p_X(x).$$

Expected Value Rule for Functions of Random Variables

Let X be a random variable with PMF $p_X(x)$, and let $g(X)$ be a real-valued function of X . Then, the expected value of the random variable $g(X)$ is given by

$$\mathbf{E}[g(X)] = \sum_x g(x)p_X(x).$$

$$\begin{aligned}\mathbf{E}[g(X)] &= \mathbf{E}[Y] \\ &= \sum_y yp_Y(y) \\ &= \sum_y y \sum_{\{x \mid g(x)=y\}} p_X(x) \\ &= \sum_y \sum_{\{x \mid g(x)=y\}} yp_X(x) \\ &= \sum_y \sum_{\{x \mid g(x)=y\}} g(x)p_X(x) \\ &= \sum_x g(x)p_X(x).\end{aligned}$$

VARIANCE

$$\text{var}(X) = \mathbf{E}[(X - \mathbf{E}[X])^2].$$

STANDARD DEVIATION

σ

$$\sigma_X = \sqrt{\text{var}(X)}.$$

$$\text{var}(X) = \mathbf{E}[(X - \mathbf{E}[X])^2] = \sum_x (x - \mathbf{E}[X])^2 p_X(x).$$

Variance in Terms of Moments Expression

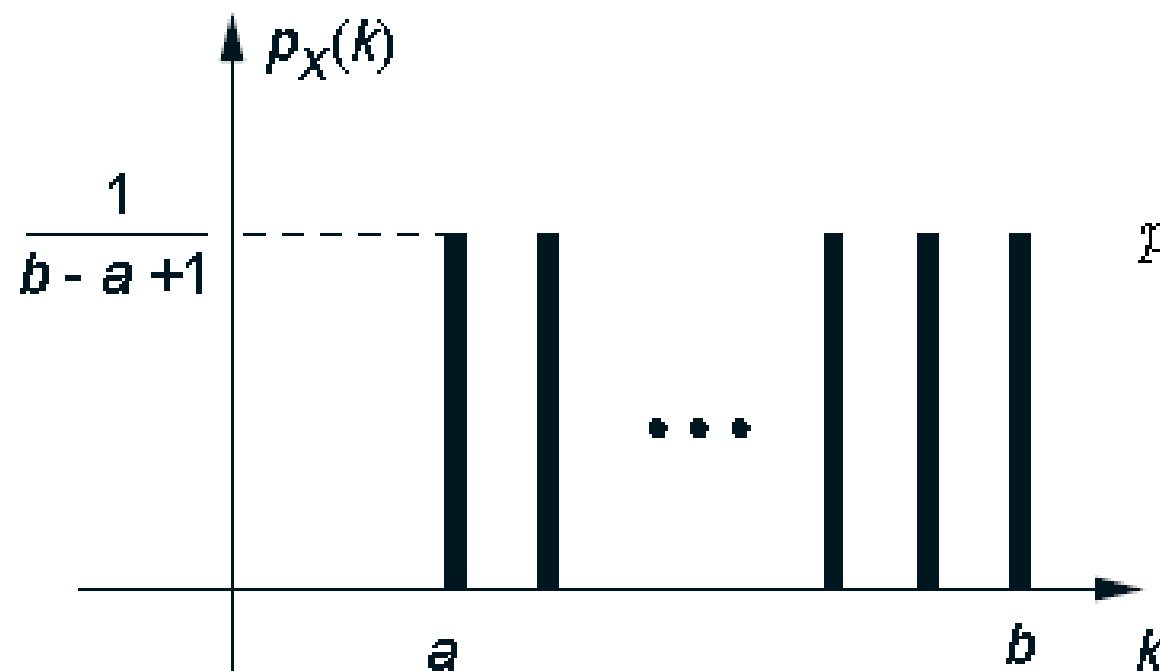
$$\text{var}(X) = \mathbf{E}[X^2] - (\mathbf{E}[X])^2.$$

n th moment

$$\mathbf{E}[X^n] = \sum_x x^n p_X(x),$$

Discrete Uniform Random Variable.

roll of a fair six-sided die $p_X(k) = \begin{cases} 1/6 & \text{if } k = 1, 2, 3, 4, 5, 6, \\ 0 & \text{otherwise.} \end{cases}$



$$p_X(k) = \begin{cases} \frac{1}{b-a+1} & \text{if } k = a, a+1, \dots, b, \\ 0 & \text{otherwise,} \end{cases}$$

$$E[X] = \frac{a+b}{2},$$

$$\text{var}(X) = \frac{(b-a)(b-a+1)}{12}.$$