

## Interval Estimation

$$P(\hat{\theta} - \epsilon < \theta < \hat{\theta} + \epsilon)$$

Point Estimation: Estimating an unknown population parameter  $\theta$  by a statistic  $\hat{\theta}(x_1, x_2, \dots, x_n)$   $P(\theta - \epsilon < \hat{\theta} < \theta + \epsilon) \approx 1, \forall \theta \in \Theta$

Disadvantage: Gives no information regarding the actual error committed in an estimation.

Interval Estimation: Choose two estimators A and B simultaneously for the unknown parameter  $\theta$  so that interval  $(a, b)$  contains  $\theta$  with certain probability, where  $a, b$  are the corresponding estimates.  $P(A < \theta < B) = 1 - \epsilon$  OR  $1 - \alpha$

Confidence Interval :  $\theta$  : unknown population parameter  
 $\theta \in \Theta$  (set of admissible values)

We know the functional form of the distribution function of the population r.v.  $X$ , i.e.  $F_X(x)$  but it contains the unknown population parameter  $\theta$ .

Let,  $(x_1, x_2, \dots, x_n)$  be a random sample of size  $n$  and  $0 < \epsilon < 1$  be a given no. Then if there exist two statistics  $a = a(x_1, x_2, \dots, x_n)$  and  $b = b(x_1, x_2, \dots, x_n)$  s.t.  $P(A < \theta < B) = 1 - \epsilon$

where  $A, B$  are the r.v.'s corr. to  $a$  and  $b$

the interval  $(a, b)$  is called a confidence interval  
of confidence coefficient  $(1-\epsilon)$ .

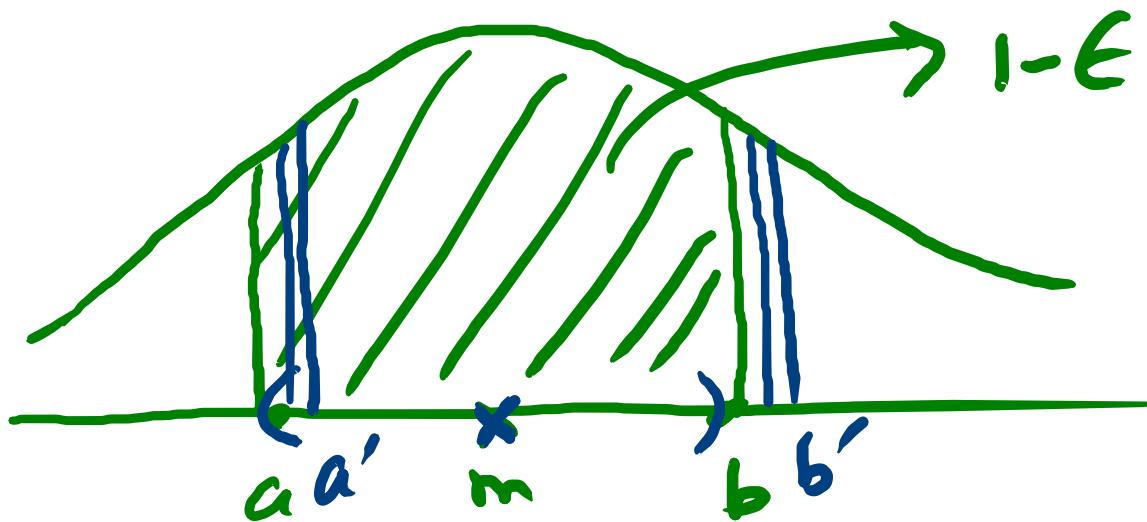
Note: 1.  $(a, b)$  is also called  $(1-\epsilon) \times 100\%$   
confidence interval.

2.  $a, b$  are called respectively the lower  
and upper limit of C.I. of  $\theta$ .

$$P(A < \theta < B) = 1 - \epsilon$$

$\downarrow$  fixed

If  $\epsilon = 0.05$   
 $1 - \epsilon = 0.95$   
 $100 \times 0.95\% = 95\%$   
C.I.



$$P(A < m < B) = 1 - \epsilon$$

Note: ③ There may be infinite number of confidence intervals with confidence coefficient  $1 - \epsilon$ .

④ If  $(a_1, b_1)$  and  $(a_2, b_2)$  be two C.I.s for  $\theta$  with confidence coefficient  $1 - \epsilon$  then we choose the one with minimum length, i.e. we choose  $(a_1, b_1)$  if  $b_1 - a_1 < b_2 - a_2$ .

## Frequency Interpretation of Confidence Intervals :

$$P(A < \theta < B) = 1 - \epsilon$$

Let  $(x_1^{(i)}, x_2^{(i)}, \dots, x_n^{(i)})$  be a random sample  
of size n

$$i=1, 2, \dots, N$$

Let,  $(a_1, b_1), (a_2, b_2), \dots, (a_N, b_N)$  be the corresponding  
confidence intervals.

If N be large then  $\theta$  will belong to approximately  
 $(1-\epsilon)N$  of the computed intervals.

i.e.  $(a_i, b_i)$  will include  $\theta$  approximately  $(1-\epsilon)N$  times out of N times.

## Method of Computing C.I. for an Unknown Population Parameter

To find a confidence interval of an unknown population parameter  $\theta$  with confidence coeff.  $1-\epsilon$ .

Let,  $(x_1, x_2, \dots, x_n)$  be a random sample of size  $n$ .

I. Choose  $t = t(x_1, x_2, \dots, x_n)$  such that it contains  $\theta$  explicitly (in its expression) and it does not contain any other unknown population parameters  $\theta_1, \theta_2, \dots, \theta_k$

II. Sampling distribution of  $t$  is known and independent of all parameters  $\theta, \theta_1, \theta_2, \dots, \theta_k$ .

$$U = \frac{\bar{X} - \theta}{\sigma/\sqrt{n}} \sim N(0, 1)$$

Let,  $f_T$  be the p.d.f. of  $T$  (corr. to  $t$ ).

Now we choose  $\check{u} = u(\epsilon)$  and  $\check{v} = v(\epsilon)$  s.t.

(from Statistical Tables for the distribution of  $T$ )

$$P(u < T < v) = 1 - \epsilon \Leftrightarrow \int_u^v f_T(t) dt = 1 - \epsilon$$

If possible

$$P(A < \theta < B) = 1 - \epsilon$$

thus  $(a, b) \rightarrow$  C.I. with confidence coeff.  $(1 - \epsilon)$ .

① Find  $100 \times (1-\epsilon)\%$  confidence interval for the parameters  $m$  of a normal  $(m, \sigma)$  population.

$$\epsilon = 0.05$$

Sol: Case I:  $\sigma$  is known

$$1-\epsilon = 0.95$$

95%  $\rightarrow$  C.I.

Choose statistic  $U = \frac{\bar{X} - m}{\sigma/\sqrt{n}}$  whose sampling

distribution is  $N(0, 1)$ . Choose points  $\pm u_\epsilon$  (from  
symmetrically about the origin s.t.

the table of

Standard  
normal  
distribution

$$P(-u_\epsilon < U < u_\epsilon) = 1-\epsilon$$

$$\Rightarrow P\left(-\frac{u_\epsilon}{\sigma/\sqrt{n}} < \frac{\bar{X} - m}{\sigma/\sqrt{n}} < u_\epsilon\right) = 1-\epsilon$$

$$\Rightarrow P\left(\bar{X} - \frac{\sigma u_\epsilon}{\sqrt{n}} < m < \bar{X} + \frac{\sigma u_\epsilon}{\sqrt{n}}\right) = 1-\epsilon$$

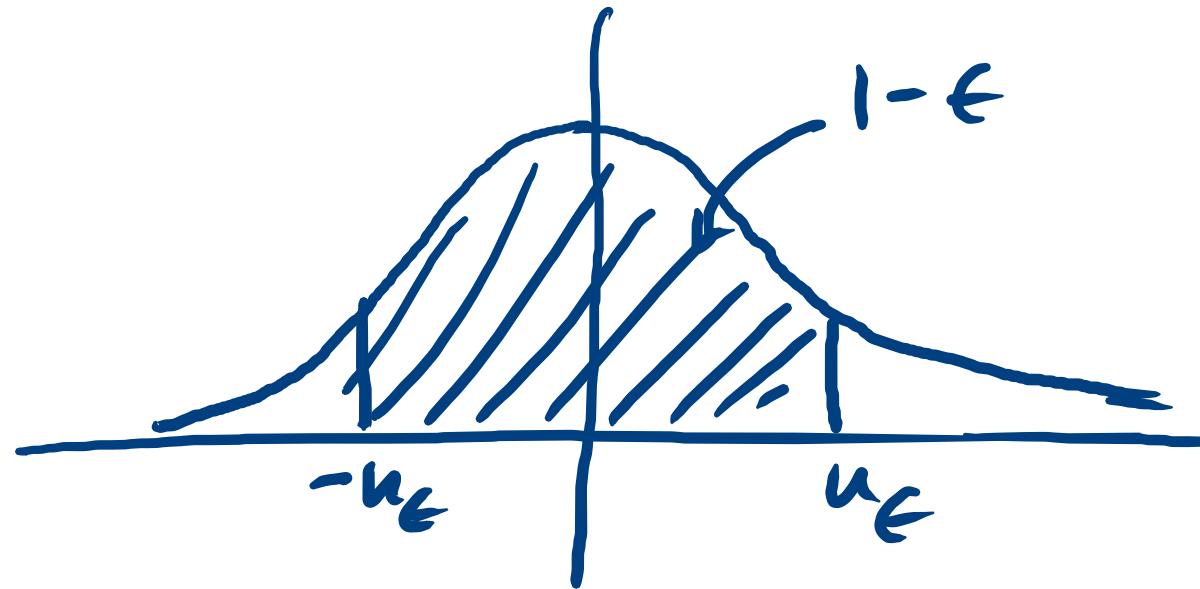
The C.I. for  $m$  is

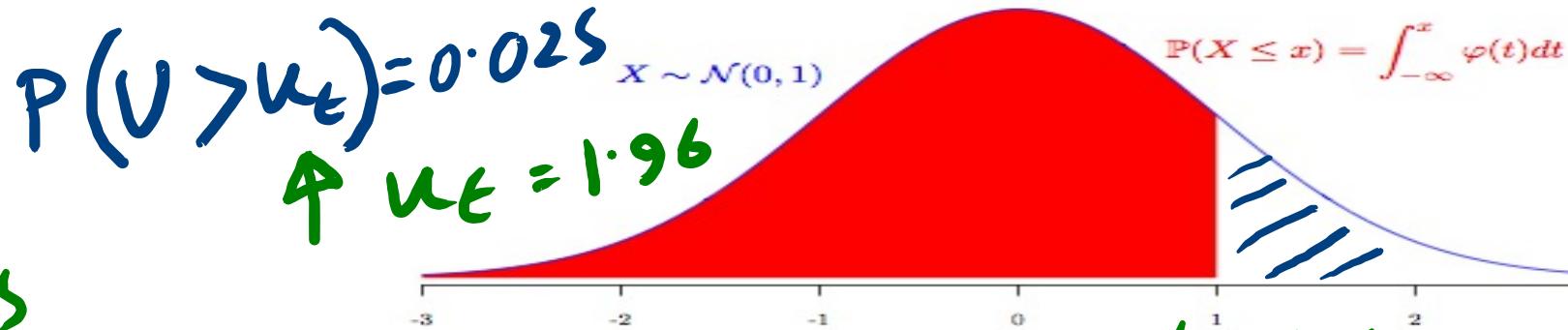
$$\left(\bar{x} - \frac{\sigma u_\epsilon}{\sqrt{n}}, \bar{x} + \frac{\sigma u_\epsilon}{\sqrt{n}}\right)$$

when  $u_\epsilon$  is given

by  $P(|U| < u_\epsilon) = 1 - \epsilon$

$$\Rightarrow P(U > u_\epsilon) = \frac{\epsilon}{2}$$





$$1 - 0.025 \\ = 0.975$$

→

|     | 0.00   | 0.01   | 0.02   | 0.03   | 0.04   | 0.05   | 0.06   | 0.07   | 0.08   | 0.09   |
|-----|--------|--------|--------|--------|--------|--------|--------|--------|--------|--------|
| 0.0 | 0.5000 | 0.5040 | 0.5080 | 0.5120 | 0.5160 | 0.5199 | 0.5239 | 0.5279 | 0.5319 | 0.5359 |
| 0.1 | 0.5398 | 0.5438 | 0.5478 | 0.5517 | 0.5557 | 0.5596 | 0.5636 | 0.5675 | 0.5714 | 0.5753 |
| 0.2 | 0.5793 | 0.5832 | 0.5871 | 0.5910 | 0.5948 | 0.5987 | 0.6026 | 0.6064 | 0.6103 | 0.6141 |
| 0.3 | 0.6179 | 0.6217 | 0.6255 | 0.6293 | 0.6331 | 0.6368 | 0.6406 | 0.6443 | 0.6480 | 0.6517 |
| 0.4 | 0.6554 | 0.6591 | 0.6628 | 0.6664 | 0.6700 | 0.6736 | 0.6772 | 0.6808 | 0.6844 | 0.6879 |
| 0.5 | 0.6915 | 0.6950 | 0.6985 | 0.7019 | 0.7054 | 0.7088 | 0.7123 | 0.7157 | 0.7190 | 0.7224 |
| 0.6 | 0.7257 | 0.7291 | 0.7324 | 0.7357 | 0.7389 | 0.7422 | 0.7454 | 0.7486 | 0.7517 | 0.7549 |
| 0.7 | 0.7580 | 0.7611 | 0.7642 | 0.7673 | 0.7704 | 0.7734 | 0.7764 | 0.7794 | 0.7823 | 0.7852 |
| 0.8 | 0.7881 | 0.7910 | 0.7939 | 0.7967 | 0.7995 | 0.8023 | 0.8051 | 0.8078 | 0.8106 | 0.8133 |
| 0.9 | 0.8159 | 0.8186 | 0.8212 | 0.8238 | 0.8264 | 0.8289 | 0.8315 | 0.8340 | 0.8365 | 0.8389 |
| 1.0 | 0.8413 | 0.8438 | 0.8461 | 0.8485 | 0.8508 | 0.8531 | 0.8554 | 0.8577 | 0.8599 | 0.8621 |
| 1.1 | 0.8643 | 0.8665 | 0.8686 | 0.8708 | 0.8729 | 0.8749 | 0.8770 | 0.8790 | 0.8810 | 0.8830 |
| 1.2 | 0.8849 | 0.8869 | 0.8888 | 0.8907 | 0.8925 | 0.8944 | 0.8962 | 0.8980 | 0.8997 | 0.9015 |
| 1.3 | 0.9032 | 0.9049 | 0.9066 | 0.9082 | 0.9099 | 0.9115 | 0.9131 | 0.9147 | 0.9162 | 0.9177 |
| 1.4 | 0.9192 | 0.9207 | 0.9222 | 0.9236 | 0.9251 | 0.9265 | 0.9279 | 0.9292 | 0.9306 | 0.9319 |
| 1.5 | 0.9332 | 0.9345 | 0.9357 | 0.9370 | 0.9382 | 0.9394 | 0.9406 | 0.9418 | 0.9429 | 0.9441 |
| 1.6 | 0.9452 | 0.9463 | 0.9474 | 0.9484 | 0.9495 | 0.9505 | 0.9515 | 0.9525 | 0.9535 | 0.9545 |
| 1.7 | 0.9554 | 0.9564 | 0.9573 | 0.9582 | 0.9591 | 0.9599 | 0.9608 | 0.9616 | 0.9625 | 0.9633 |
| 1.8 | 0.9641 | 0.9649 | 0.9656 | 0.9664 | 0.9671 | 0.9678 | 0.9686 | 0.9693 | 0.9699 | 0.9706 |
| 1.9 | 0.9713 | 0.9719 | 0.9726 | 0.9732 | 0.9738 | 0.9744 | 0.9750 | 0.9756 | 0.9761 | 0.9767 |
| 2.0 | 0.9772 | 0.9778 | 0.9783 | 0.9788 | 0.9793 | 0.9798 | 0.9803 | 0.9808 | 0.9812 | 0.9817 |
| 2.1 | 0.9821 | 0.9826 | 0.9830 | 0.9834 | 0.9838 | 0.9842 | 0.9846 | 0.9850 | 0.9854 | 0.9857 |
| 2.2 | 0.9861 | 0.9864 | 0.9868 | 0.9871 | 0.9875 | 0.9878 | 0.9881 | 0.9884 | 0.9887 | 0.9890 |
| 2.3 | 0.9893 | 0.9896 | 0.9898 | 0.9901 | 0.9904 | 0.9906 | 0.9909 | 0.9911 | 0.9913 | 0.9916 |
| 2.4 | 0.9918 | 0.9920 | 0.9922 | 0.9925 | 0.9927 | 0.9929 | 0.9931 | 0.9932 | 0.9934 | 0.9936 |
| 2.5 | 0.9938 | 0.9940 | 0.9941 | 0.9943 | 0.9945 | 0.9946 | 0.9948 | 0.9949 | 0.9951 | 0.9952 |
| 2.6 | 0.9953 | 0.9955 | 0.9956 | 0.9957 | 0.9959 | 0.9960 | 0.9961 | 0.9962 | 0.9963 | 0.9964 |
| 2.7 | 0.9965 | 0.9966 | 0.9967 | 0.9968 | 0.9969 | 0.9970 | 0.9971 | 0.9972 | 0.9973 | 0.9974 |
| 2.8 | 0.9974 | 0.9975 | 0.9976 | 0.9977 | 0.9977 | 0.9978 | 0.9979 | 0.9979 | 0.9980 | 0.9981 |
| 2.9 | 0.9981 | 0.9982 | 0.9982 | 0.9983 | 0.9984 | 0.9984 | 0.9985 | 0.9985 | 0.9986 | 0.9986 |
| 3.0 | 0.9987 | 0.9987 | 0.9987 | 0.9988 | 0.9988 | 0.9989 | 0.9989 | 0.9989 | 0.9990 | 0.9990 |