

COMMUNITY DETECTION FROM LOW-RANK EXCITATIONS OF A GRAPH FILTER

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ABSTRACT

This paper considers the problem of inferring the topology of a graph from noisy outputs of an unknown graph filter excited by low-rank signals. Limited by this low-rank structure, we focus on solving the *community detection* problem, whose aim is to partition the node set of the unknown graph into subsets with high edge densities. We propose to detect the communities by applying spectral clustering on the low-rank output covariance matrix. To analyze the performance, we show that the low-rank covariance yields a sketch of the eigenvectors of the unknown graph. Importantly, we provide theoretical bounds on the error introduced by this sketching procedure based on spectral features of the graph filter involved. Finally, our theoretical findings are validated via numerical experiments.

Index Terms— graph signal processing, graph filter, topology identification, low rank excitation, community detection

1. INTRODUCTION

The emerging fields of Network Science and Big Data have broadened the scope of signal processing to also include the analysis of discrete signals defined on graphs [1–3]. Under the assumption that the signal properties are related to the topology of the graph where they are supported, the goal of graph signal processing (GSP) is to develop algorithms that leverage this relational structure in various ways, including inferring these relationships when they are only partially observed [3]. A suitable way to accomplish these objectives is to rely on the so-called graph-shift operator (GSO), which is a matrix that reflects the local connectivity of the graph and generalizes the shift operator used in time series analysis [2]. Most GSP works assume that the GSO (hence the graph) is known, and then analyze how the algebraic and spectral characteristics of the GSO impact the properties of the signals and filters defined on such a graph. This approach led to the extension of classical problems to the realm of graphs signals such as sampling [4, 5] and reconstruction/interpolation [6, 7]. Our focus here is to investigate how to use graph signals information to infer aspects of the underlying topology. In particular, in this paper we focus on inferring from a graph signal the community structure [8] of the underlying graph.

Network topology inference from a set of (graph-signal) observations is a prominent problem in Network Science [3] with classical approaches for its solution including partial correlations [9], Gaussian graphical models [10], and structural equation models [11, 12], among others. More recently, GSP-based network inference methods emerged that take a system identification approach to solving the network inference problem. Specifically, they postulate the existence of a network as a latent underlying structure and of a model that ties the observations to a network dynamical process defined

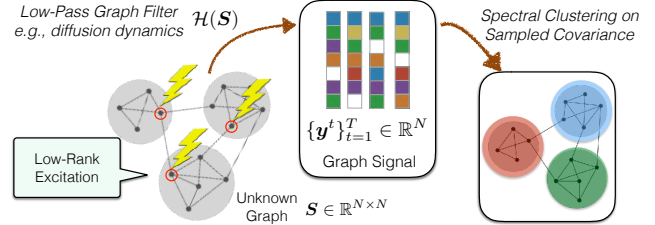


Fig. 1. Overview of the approach in this paper. We observe graph signals that are outputs of a low-pass graph filter subject to low-rank excitations, e.g., when the excitations affect a small number of nodes as illustrated above. From these limited observations, we apply spectral clustering to recover communities of the *unknown* graph.

on such a graph [13–21]. In order to infer the graph topology from the observed graph signals, the approaches we just mentioned rely on different assumptions such as smoothness of the observed signals [13–15], richness of the input signals [16–18], or partial knowledge of the involved filter or dynamic on the graph [19–21].

In this paper, we model the graph signals observed as the outputs of an unknown graph filter subject to low-rank excitations *i.e.*, inputs that belong to a low-dimensional space or that are sparse [see Fig. 1]. For example, we may observe the outcome of a rumor that originates from a few sources (sparse excitation) and evolves via an opinion dynamics in the network (modeled by the unknown graph filter). Under this weaker set of assumptions, in general, the exact inference of the underlying graph is infeasible. Nonetheless, we show that the *community structure* of this unknown graph can be recovered in this laxer setting. To show this result, we rely on (i) the concept of *sketching* and the theory behind its recent application to the acceleration of spectral clustering algorithms [22, 23]; (ii) the concept of *low-pass graph filters* which generalizes a number of graph processes including diffusion. Our contributions are twofold. First, we show that the covariance matrix of the observed graph signals retains a sketch of the principal eigenvectors of the underlying GSO. Second and more importantly, we provide a theoretical characterization of the performance attainable when spectral clustering is applied to the output covariance matrix of the GSO. Using as a benchmark the communities that spectral clustering would output when perfect knowledge of the underlying graph is available, we bound the error in the detection of communities in terms of the spectral properties of the graph filter involved. Consequently, this bound can be used to identify linear network dynamics – modeled as graph filters – that are amenable to the problem of community detection.

Notation. We use boldfaced lower-case (*resp.* upper-cased) letters to denote vectors (*resp.* matrices). For a vector \mathbf{x} , the notation x_i denotes its i th element. For a matrix \mathbf{X} , the notation X_{ij} denotes its (i, j) th element whereas $[\mathbf{X}]_{i,:}$ denotes its i th row vector. Also, $\mathcal{R}(\mathbf{X}) \subseteq \mathbb{R}^N$ denotes the range space of $\mathbf{X} \in \mathbb{R}^{N \times M}$. For a sym-

metric matrix \mathbf{E} , $\beta_i(\mathbf{E})$ denotes its i th largest eigenvalue. For a matrix $\mathbf{M} \in \mathbb{R}^{P \times N}$, it admits the partition $\mathbf{M} = [\mathbf{M}_K \ \mathbf{M}_{N-K}]$ with \mathbf{M}_K (resp. \mathbf{M}_{N-K}) denotes the matrix consisting of the *left-most* K (resp. *right-most* $N - K$) columns of \mathbf{M} . Similarly, $\mathbf{m} \in \mathbb{R}^N$ is partitioned into $\mathbf{m} = [\mathbf{m}_K; \mathbf{m}_{N-K}]$, where \mathbf{m}_K (resp. \mathbf{m}_{N-K}) consists of its *top* K (resp. *bottom* $N - K$) elements.

2. GRAPH SIGNALS AND COMMUNITY DETECTION

Next, we formally introduce the GSP terminology that will be used throughout the paper and we briefly present the classical approach of spectral clustering for community detection.

Graph signals and graph filters. Let $G = (V, E)$ denote an undirected graph with a set of nodes $V := \{1, \dots, N\}$ and a set of links $E \subseteq V \times V$, such that if node i is connected to j , then both (i, j) and (j, i) belong to E . We define the adjacency matrix $\mathbf{A} \in \mathbb{R}^{N \times N}$ as a matrix with non-zero elements $A_{ij} = A_{ji}$ if and only if $(i, j) \in E$. The adjacency matrix can be used to define the degree matrix $\mathbf{D} := \text{diag}(\mathbf{A}\mathbf{1})$ and the (combinatorial) Laplacian matrix as $\mathbf{L} := \mathbf{D} - \mathbf{A}$. Graph signals defined on the nodes of G are functions $f : V \rightarrow \mathbb{R}$, equivalently represented as vectors $\mathbf{x} = [x_1, \dots, x_N]^\top \in \mathbb{R}^N$, where x_i denotes the signal value at node i . The graph G is endowed with the so-called graph shift operator $\mathbf{S} \in \mathbb{R}^{N \times N}$, a matrix whose entry S_{ij} can be non-zero only if $i = j$ or if $(i, j) \in E$. In this paper, we consider the case in which the GSO is given by the graph Laplacian, i.e. $\mathbf{S} = \mathbf{L}$, and has an eigenvalue decomposition given by $\mathbf{S} = \mathbf{V}\mathbf{\Lambda}\mathbf{V}^\top$ where $\mathbf{\Lambda} = \text{diag}(\lambda)$ and λ_i is sorted in an *ascending order* $0 = \lambda_1 \leq \lambda_2 \leq \dots \leq \lambda_N$.

The shift \mathbf{S} can be used to define linear graph-signal operators denominated *graph filters* that have the form

$$\mathcal{H}(\mathbf{S}) := \sum_{\ell=0}^{L-1} h_\ell \mathbf{S}^\ell = \mathbf{V} \left(\sum_{\ell=0}^{L-1} h_\ell \mathbf{\Lambda}^\ell \right) \mathbf{V}^\top. \quad (1)$$

For a given input \mathbf{z} , the output of the filter is simply $\mathbf{x} = \mathcal{H}(\mathbf{S})\mathbf{z}$. The coefficients of the filter are collected into $\mathbf{h} := [h_0, \dots, h_{L-1}]^\top$. Graph filters are of particular interest because they represent linear transformations that can be implemented locally and, thus, can be used to model linear network dynamics. Leveraging the spectral decomposition of \mathbf{S} , graph filters and signals can be represented in the frequency domain. To be precise, let us use the eigenvalues of \mathbf{S} to define the $N \times L$ Vandermonde matrix $\mathbf{\Psi}$, where $\Psi_{ij} := (\lambda_i)^{j-1}$. Then, the frequency representations of a *signal* \mathbf{z} and of a *filter* \mathbf{h} are defined as $\tilde{\mathbf{z}} := \mathbf{V}^\top \mathbf{z}$ and $\tilde{\mathbf{h}} := \mathbf{\Psi} \mathbf{h}$. Exploiting such representations and with \odot denoting the elementwise product, the filter's output $\mathbf{x} = \mathcal{H}(\mathbf{S})\mathbf{z}$ in the frequency domain is given by

$$\tilde{\mathbf{x}} = \text{diag}(\mathbf{\Psi} \mathbf{h}) \mathbf{V}^\top \mathbf{z} = \tilde{\mathbf{h}} \odot \tilde{\mathbf{z}}. \quad (2)$$

This is analogous to the convolution theorem for temporal signals.

Low-pass graph filters. Out of all possible filters, low-pass graph filters play a fundamental role for community detection. Formally, we say that a graph filter is *low-pass* and (K, η) -separable if

$$\tilde{h}_{K+1}/\tilde{h}_K \leq \eta < 1, \quad \tilde{h}_N \geq 0 \quad \text{and} \quad \tilde{h}_i \geq \tilde{h}_{i+1}, \quad (3)$$

for all $i = 1, \dots, N - 1$. Notice that since $\tilde{h}_i \geq \tilde{h}_{i+1}$, the eigenvectors of $\mathcal{H}(\mathbf{S})$ corresponding to the largest K eigenvalues are the left-most K vectors in \mathbf{V} , denoted by \mathbf{V}_K . Common examples of low-pass and (K, η) -separable graph filters (for all K) [15, 16] include $\mathcal{H}_1(\mathbf{S}) = (\mathbf{I} + \mathbf{S})^{-1}$, $\mathcal{H}_2(\mathbf{S}) = \exp(-\mathbf{S})$, and $\mathcal{H}_3(\mathbf{S}) = (\mathbf{I} - \alpha \mathbf{S})^{L-1}$ for $0 < \alpha < 1/\lambda_N$. Filter $\mathcal{H}_1(\mathbf{S})$ can be regarded as the natural extension of a single pole IIR filter to the graph domain,

and $\mathcal{H}_3(\mathbf{S})$ is an FIR filter modeling the effect of $L - 1$ steps of a diffusion process.

If $\eta \approx 0$, i.e., the frequency response declines sharply from the K th to the $(K+1)$ th frequency, $\mathcal{H}(\mathbf{S})$ is approximately rank- K . In fact, Section 3 reveals that a small value of η is essential to obtain a tight theoretical guarantee on detecting K communities from the unknown graph. We remark that a small value of η can typically be achieved by increasing the graph filter order L .

Community detection via spectral clustering. For an undirected graph $G = (V, E)$, a *community* is intuitively a set of nodes with higher edge density among themselves than to nodes outside the set. Thus, the problem of K -community detection amounts to finding K such disjoint communities $\mathcal{C}_1, \dots, \mathcal{C}_K$ such that $V = \mathcal{C}_1 \cup \dots \cup \mathcal{C}_K$. A well-known way to establish a sense of optimal partition of the node sets into communities is that of minimizing the *ratio cut* [24]. Formally, defining the cut weight between two node sets as the sum of edge weights between them, i.e., $\mathbf{A}_{\mathcal{C}_i, \mathcal{C}_j} := \sum_{k \in \mathcal{C}_i, k' \in \mathcal{C}_j} A_{kk'}$. The ratio cut of a given partition is defined as

$$\text{RatioCut}(\mathcal{C}_1, \dots, \mathcal{C}_K) := (1/2) \sum_{i=1}^K \mathbf{A}_{\mathcal{C}_i, \bar{\mathcal{C}}_i} / |\mathcal{C}_i|, \quad (4)$$

where $\bar{\mathcal{C}}_i$ is the complement of \mathcal{C}_i . Unfortunately, the ratio-cut minimization problem is hard due to its combinatorial nature. As a remedy, a widely used heuristic is the *spectral clustering* method [25] which can be seen as a convex relaxation of the ratio-cut minimization problem. Spectral clustering can be described as a two-step process¹. First, every node is projected onto \mathbb{R}^K , where the coordinates of node i are given by $[V_{i1}, \dots, V_{iK}]$, i.e., the i th row vector of the left-most K columns of \mathbf{V} . Secondly, the K -means clustering method [26] is applied to the resulting projected space to obtain the sought K communities.

3. COMMUNITY DETECTION FROM LOW-RANK DATA

Our goal is to detect communities in a graph when we do not have access to the graph itself – nor, equivalently, the GSO \mathbf{S} – but rather to a set of graph signals defined on it. Formally, we observe a set of graph signals $\mathbf{y}^t \in \mathbb{R}^N$, $t = 1, \dots, T$ that are noisy observations of the output of a low-pass and (K, η) -separable graph filter (1):

$$\mathbf{y}^t = \mathbf{x}^t + \mathbf{w}^t \quad \text{and} \quad \mathbf{x}^t = \mathcal{H}(\mathbf{S})\mathbf{z}^t, \quad (5)$$

where the observation noise is $\mathbf{w}^t \sim \mathcal{N}(0, \sigma^2 \mathbf{I})$ and the excitation signal $\mathbf{z}^t \in \mathbb{R}^N$ is zero mean.

Unlike previous work [16] that tackles the problem of identifying \mathbf{S} from a *full-rank excitation* model where $\mathbb{E}[\mathbf{z}^t(\mathbf{z}^t)^\top] = \mathbf{I}$, this paper focuses on low-rank excitation in which

$$\mathbf{C}_z = \mathbb{E}[\mathbf{z}^t(\mathbf{z}^t)^\top] = \mathbf{B}\mathbf{B}^\top, \quad (6)$$

where $\mathbf{B} \in \mathbb{R}^{N \times R}$ with $R < N$. Practical scenarios that lead to (6) include the cases where the input signal \mathbf{z}^t is applied on only a subset of R nodes and where the number of variations in the excitation signal is limited to R modes. Under this low-rank excitation model, the covariance matrix of a general output signal \mathbf{x}^t is given by

$$\mathbf{C}_x = \mathbb{E}[\mathbf{x}^t(\mathbf{x}^t)^\top] = \mathcal{H}(\mathbf{S})\mathbf{B}\mathbf{B}^\top\mathcal{H}(\mathbf{S})^\top, \quad (7)$$

whose rank is at most R . Inferring \mathbf{S} from \mathbf{C}_x is challenging for two reasons: i) we do not assume perfect knowledge of the specific filter coefficients \mathbf{h} defining \mathcal{H} , and ii) information is in general lost due to the low-rank nature of $\mathbf{B}\mathbf{B}^\top$. Nevertheless, we show that we can recover the community structure in \mathbf{S} by applying spectral clustering directly on \mathbf{C}_x , or its estimate $\hat{\mathbf{C}}_x$.

¹Minor variations to the method presented here co-exist in the literature.

Algorithm 1 Community detection from low-rank excitation.

- 1: **Input:** Graph signals $\{\mathbf{y}^t\}_{t=1}^T$; desired number of clusters K .
- 2: Use $\{\mathbf{y}^t\}_{t=1}^T$ to compute the sample covariance $\hat{\mathbf{C}}_y$ as in (9).
- 3: Find the K eigenvectors to $\hat{\mathbf{C}}_y$ associated with the largest K eigenvalues. Denote the set of eigenvectors as $\hat{\mathbf{P}}_K \in \mathbb{R}^{N \times K}$.
- 4: Perform K -means clustering (e.g., [26]), which optimizes:

$$\min_{\mathcal{C}_1, \dots, \mathcal{C}_K} \sum_{i=1}^K \sum_{j \in \mathcal{C}_i} \left\| \hat{\mathbf{p}}_j - \frac{1}{|\mathcal{C}_i|} \sum_{q \in \mathcal{C}_i} \hat{\mathbf{p}}_q \right\|_2^2 \text{ s.t. } \mathcal{C}_i \subseteq V, \quad (8)$$

where $\hat{\mathbf{p}}_i := [\hat{\mathbf{P}}_K]_{i,:} \in \mathbb{R}^K$. Let the solution be $\hat{\mathcal{C}}_1, \dots, \hat{\mathcal{C}}_K$.

- 5: **Output:** Partition of V into K communities, $\hat{\mathcal{C}}_1, \dots, \hat{\mathcal{C}}_K$.
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3.1. Community detection algorithm

From the observed graph signals $\{\mathbf{y}^t\}_{t=1}^T$ [cf. (5)] we construct the empirical sample covariance

$$\hat{\mathbf{C}}_y = (1/T) \sum_{t=1}^T \mathbf{y}^t (\mathbf{y}^t)^\top, \quad (9)$$

and apply the spectral clustering procedure presented in Section 2 but only based on the eigenvectors of $\hat{\mathbf{C}}_y$ associated to the K largest eigenvalues. This algorithm is summarized in Algorithm 1.

We first discuss the intuition behind the algorithm we propose. First notice that the filter $\mathcal{H}(\mathbf{S})$ is a matrix polynomial on the symmetric matrix \mathbf{S} [cf. (1)] and, as such, preserves its eigenvectors. Moreover, since $R < N$, the matrix $\mathcal{H}(\mathbf{S})\mathbf{B} \in \mathbb{R}^{N \times R}$ is a ‘sketch’ of $\mathcal{H}(\mathbf{S})$ with \mathbf{B} being the *sketching matrix*. On the other hand, if $R \geq K$ and the graph filter is (K, η) -separable with a small η , the top- K eigenvectors of $\mathcal{H}(\mathbf{S})$ can be preserved in the sketched version since $\mathcal{H}(\mathbf{S})$ is approximately rank K . Indeed, when $\mathcal{R}(\mathbf{B}) = \mathcal{R}(\mathbf{V}_K)$, it is easy to observe that $\mathcal{R}(\mathcal{H}(\mathbf{S})\mathbf{B}) = \mathcal{R}(\mathbf{V}_K)$. In this case, applying Algorithm 1 based on the covariance $\mathbf{C}_x = \mathcal{H}(\mathbf{S})\mathbf{B}\mathbf{B}^\top \mathcal{H}(\mathbf{S})^\top$ yields the same output as applying spectral clustering on \mathbf{S} .

In most cases, however, Algorithm 1 *will not* return the same communities as those obtained by spectral clustering on the GSO \mathbf{S} . There are two sources for this discrepancy. First, the sketching matrix \mathbf{B} does not share the range of \mathbf{V}_K in general, thus there will be a discrepancy with respect to the eigenvectors in $\mathcal{H}(\mathbf{S})$. Second, having access to a finite set of graph signals, we do not observe \mathbf{C}_x perfectly but rather an estimate $\hat{\mathbf{C}}_x$ of it. In the next section we bound the error introduced by these two sources of inaccuracy.

3.2. Theoretical Guarantees

Denote by $\{\hat{\mathcal{C}}_1, \dots, \hat{\mathcal{C}}_K\}$ the community structure obtained from Algorithm 1. Our goal is to find out how close these communities are to the ones obtained when \mathbf{S} is available. To do so, we measure the optimality of the clustering result by comparing the following objective — let $\mathcal{C}_1, \dots, \mathcal{C}_K$ be any partition of V ,

$$F(\mathcal{C}_1, \dots, \mathcal{C}_K) := \sum_{i=1}^K \sum_{j \in \mathcal{C}_i} \left\| \mathbf{v}_j - \frac{1}{|\mathcal{C}_i|} \sum_{q \in \mathcal{C}_i} \mathbf{v}_q \right\|_2^2, \quad (10)$$

where $\mathbf{v}_i := [\mathbf{V}_K]_{i,:}$, $i = 1, \dots, N$. Notice that (10) is the same objective function used in step 4 of Algorithm 1, but applied to the eigenvectors of the GSO \mathbf{S} instead. In general, $F(\mathcal{C}_1, \dots, \mathcal{C}_K)$ can only be minimized when \mathbf{S} is known perfectly.

Let us introduce the following definitions. The singular value decomposition of $\mathcal{H}(\mathbf{S})\mathbf{B}$ is given by $\mathcal{H}(\mathbf{S})\mathbf{B} := \mathbf{P}\mathbf{\Sigma}\mathbf{Q}^\top$ with

$\sigma_1 \geq \sigma_2 \geq \dots$, the difference between the covariance and the sample covariance is $\mathbf{\Delta} := \hat{\mathbf{C}}_y - \mathbf{C}_x$. We have:

Theorem 1 *If the following conditions hold:*

1. $\mathcal{H}(\mathbf{S})$ is a low-pass and (K, η) -separable filter [cf. (3)],
2. Step 4 of Algorithm 1 finds a solution $\hat{\mathcal{C}}_1, \dots, \hat{\mathcal{C}}_K$ that exactly minimizes the problem (8),
3. $\text{rank}(\mathbf{V}_K \text{diag}(\tilde{\mathbf{h}}_K) \mathbf{V}_K^\top \mathbf{B} \mathbf{Q}_K) = K$,
4. $\sigma_K > 0$,
5. $\exists \delta > 0$ such that $\|\mathbf{\Delta}\|_2 + \delta \leq \beta_K(\mathbf{C}_x) - \beta_{K+1}(\mathbf{C}_x)$,

then, denoting by $F^* := \min_{\mathcal{C}_1, \dots, \mathcal{C}_K \subseteq V} F(\mathcal{C}_1, \dots, \mathcal{C}_K)$ as the optimal cost associated with spectral clustering with perfect knowledge of \mathbf{S} , we have that

$$\sqrt{F(\hat{\mathcal{C}}_1, \dots, \hat{\mathcal{C}}_K)} - \sqrt{F^*} \leq \sqrt{8K} \left(\sqrt{\frac{\gamma^2}{1 + \gamma^2}} + \frac{\|\mathbf{\Delta}\|_2}{\delta} \right), \quad (11)$$

where γ is bounded by

$$\gamma \leq \eta \cdot \|\mathbf{V}_{N-K}^\top \mathbf{B} \mathbf{Q}_K\|_2 \|(\mathbf{V}_K^\top \mathbf{B} \mathbf{Q}_K)^{-1}\|_2. \quad (12)$$

The proof of Theorem 1 – partially inspired by [22] – is omitted here due to space limitations but can be found in an online appendix.²

As desired, (11) bounds the difference between the optimal cost F^* associated with running spectral clustering in the ideal case of perfect information and the cost achieved by the communities $\{\hat{\mathcal{C}}_1, \dots, \hat{\mathcal{C}}_K\}$ identified from our algorithm. This optimality gap consists of the two summands on the right-hand side of (11). The first summand, which depends on γ , captures the distortion introduced by the sketching operation. From (12) it follows that this optimality gap depends on the parameter η of the low pass graph filter, which might be improved for higher filter order L , as discussed in Section 2. Furthermore, expression (12) reveals that γ depends on the angle between $\mathcal{R}(\mathbf{B} \mathbf{Q}_K)$ and $\mathcal{R}(\mathbf{V}_K)$. In particular, if $\mathcal{R}(\mathbf{B} \mathbf{Q}_K) = \mathcal{R}(\mathbf{V}_K)$, then $\|\mathbf{V}_{N-K}^\top \mathbf{B} \mathbf{Q}_K\|_2 = 0$ and $\gamma = 0$.

The second summand in the right-hand side of (11) bounds the error resulting from the fact that we have access to a finite number of signals. This bound follows from applying the classical Weyl’s inequality combined with Davis-Kahan theorem [27]. Theorem 1 shows that when the magnitude of the error $\mathbf{\Delta}$ is small compared to the relevant *spectral gap* of \mathbf{C}_x [cf. Condition 5], then the additive error introduced will simply be proportional to $\|\mathbf{\Delta}\|_2$. Moreover, using standard results from concentration inequalities and under mild assumptions on the statistics of \mathbf{z}^t , the error $\|\mathbf{\Delta}\|_2$ decays to σ at the rate of $\mathcal{O}(1/\sqrt{T})$ with high probability. Overall, Theorem 1 shows that when the graph filter has a favorable frequency response, and the number of collected samples is sufficient, we expect Algorithm 1 to produce communities comparable to those obtained by applying spectral clustering on \mathbf{S} .

4. NUMERICAL EXAMPLES

This section presents numerical results to illustrate the performance of our low-rank community detection approach. We first focus on a controlled setting where the GSO and the associated graph signals are generated synthetically. More precisely, we consider an undirected graph G generated according to a stochastic block model (SBM) with K communities, i.e. $G \sim \text{SBM}(N, K, a, b)$ such that a

²http://www.public.asu.edu/~hwai2/pdf/gsp_app.pdf

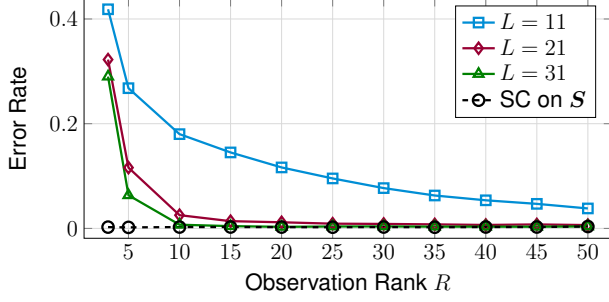


Fig. 2. (Effect of filter order L and observation rank R). Matrix S is adjacency matrix of SBM graphs with $N = 150$ nodes and $K = 3$ communities, and $\mathcal{H}(S) = (I - \alpha S)^{L-1}$, under a noiseless observation setting with $\hat{C}_x = C_x$.

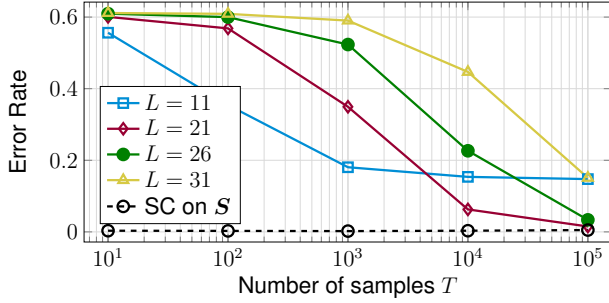


Fig. 3. (Effect of sample size T). Similar settings as in Fig. 2 with $R = 15$. We consider noisy observations and only a finite number of samples are observed T .

and b are the probabilities that an edge is formed within a community and between the communities, respectively, and we have that $a > b$. The graph filter used is $\mathcal{H}(S) = (I - \alpha S)^{L-1}$ for $\alpha = 1/(2d_{\max})$ where $d_{\max} = \max_i d_i$ and d_i is the degree of node i . The filter order is given by L . It is immediate to check that the graph filter is low-pass and (K, η) -separable. The corresponding η is given as:

$$\eta = \left(\frac{1 - \alpha \lambda_{K+1}}{1 - \alpha \lambda_K} \right)^{L-1}. \quad (13)$$

Since $\alpha < 1/\lambda_N$, we have $1 - \alpha \lambda_{K+1} < 1 - \alpha \lambda_K$ if $\lambda_{K+1} > \lambda_K$. Therefore, we anticipate η to decrease to zero as L increases. Furthermore, the low-rank excitation signal is generated as $z^t = B \tilde{z}^t$ where $\tilde{z}^t \in \mathbb{R}^R$ satisfies $\tilde{z}^t \sim \mathcal{N}(0, I)$, B is a row-sparse binary matrix with only R non-zero rows, chosen at random, and each row has only $\lceil (R/N)d_i \rceil$ ones, also chosen at random. This corresponds to the setting when the graph is excited only on R nodes and each is driven by a few independent sources. The results reported here correspond to the average of 1000 Monte-Carlo simulations. For step 4 of Algorithm 1, we use the `kmeans` function in MATLAB and normalized the vectors $\{\hat{p}_j\}$ as unit norm as suggested in [25].

The first example considers a noiseless setting where the algorithm has access to $\hat{C}_y = C_x$. We compare the average error rate (compared to the ground truth communities of the SBM) when applying Algorithm 1 for different filter orders L and ranks R of the sketching matrix B . The SBM graph has $N = 150$ nodes, $K = 3$ communities and the parameters are set as $a = 8 \log N/N$ and $b = \log N/N$; see Fig. 2. We observe that the error rate decays as we increase the filter order L . This is consistent with our result in Theorem 1. Indeed, for larger values of L , we have that η in (12) decreases, thus leading to better performance. Furthermore, the per-

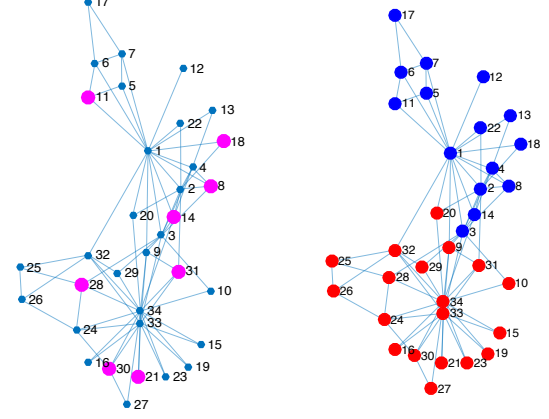


Fig. 4. (Community Detection on Real Network). The Zachary Karate Club network with $N = 34$ nodes. (Left) Highlighted nodes are the locations of the non-zero rows of B . (Right) Snapshot of the detected communities by proposed method with rank $R = 8$ observations and filter order $L = 8$.

formance improves with the observation rank R and approaches that of applying spectral clustering on the original S . This is reasonable since a higher observation rank R gives more degree of freedom for the proposed method to detect the K communities. With a filter order of $L = 21$ and rank $R \approx 15$, the proposed method achieves a performance comparable to that of having perfect knowledge of S .

The second example considers the finite sample setting and compares the error rate against the number of samples observed T . We focus on the same simulation settings as in the previous example and we fix $R = 15$, yet only a sample covariance \hat{C}_y is observed and the observation noise has a standard deviation of $\sigma = 10^{-1}$ [cf. (5) and (9)]. From Fig. 3, we observe that the error rate generally decreases as the number of samples increases. This is as predicted by Theorem 1, and the performance is ultimately limited by the filter order L and observation noise variance σ^2 . Interestingly, we observe that to achieve the same error rate, more samples are also required when the filter order L is high. This is because for higher filter orders, the magnitude of the eigenvalue $\beta_K(C_x)$ decreases, which forces δ in Theorem 1 to decrease [cf. Condition 5].

The last example deals with detecting communities in a real network, namely Zachary's Karate Club [28]. The graph consists of $N = 34$ nodes and is depicted in Fig. 4 (Left). For the proposed method, we set $K = 2$ and the graph filter is the same as in the previous examples with $L = 8$. The low-rank excitation matrix B has rank $R = 8$ with a row-sparsity pattern given by the active nodes in Fig. 4 (Left). We consider the noisy observation setting with $T = 10^3$ and $\sigma = 10^{-1}$. For this case, we have an average error rate $\approx 8.8\%$ when compared to the communities recovered using spectral clustering on the original S . An example of the recovered communities is depicted in Fig. 4 (Right).

Conclusions. We studied a method for recovering community structures from the outputs of a low-pass graph filter subject to low-rank excitation. The proposed method relies on applying spectral clustering on the output (sample) covariance matrix. We characterized the error of such procedure compared to the case where the true graph shift operator is available. Our analysis shows that the performance hinges on the spectral gap property of the graph filter involved. Numerical experiments were performed on synthetic and real networks samples to verify our results.

5. REFERENCES

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6. PROOF OF THEOREM 1

To simplify the notations while proving the theorem, let us define the following indicator matrices for the communities structures found. Firstly, the matrix $\hat{\mathbf{X}} \in \mathbb{R}^{N \times K}$ is associated with the communities $\{\hat{\mathcal{C}}_1, \dots, \hat{\mathcal{C}}_K\}$ found with Algorithm 1 such that

$$\hat{X}_{ij} := \begin{cases} 1/|\hat{\mathcal{C}}_j|, & \text{if } i \in \hat{\mathcal{C}}_j \\ 0, & \text{otherwise.} \end{cases} \quad (14)$$

Moreover, let $v^* = F(\mathcal{C}_1^*, \dots, \mathcal{C}_K^*)$ be the optimal set of community found by minimizing $F(\mathcal{C}_1, \dots, \mathcal{C}_K)$ [cf. (10)], similarly we can define $\mathbf{X}^* \in \mathbb{R}^{N \times R}$ such that:

$$X_{ij}^* := \begin{cases} 1/|\mathcal{C}_j^*|, & \text{if } i \in \mathcal{C}_j^* \\ 0, & \text{otherwise.} \end{cases} \quad (15)$$

If we define by \mathcal{X} the set of all possible indicator matrices over all the combinations of communities, it can be verified that:

$$\begin{aligned} \|\hat{\mathbf{P}}_K - \hat{\mathbf{X}} \hat{\mathbf{X}}^\top \hat{\mathbf{P}}_K\|_F^2 &= \min_{\mathbf{X} \in \mathcal{X}} \|\hat{\mathbf{P}}_K - \mathbf{X}(\mathbf{X})^\top \hat{\mathbf{P}}_K\|_F^2 \\ &= \min_{\mathcal{C}_1, \dots, \mathcal{C}_K} \sum_{i=1}^K \sum_{j \in \mathcal{C}_i} \left\| \hat{\mathbf{p}}_j - \frac{1}{|\mathcal{C}_i|} \sum_{q \in \mathcal{C}_i} \hat{\mathbf{p}}_q \right\|_2^2 \text{ s.t. } \mathcal{C}_i \subseteq V, \end{aligned} \quad (16)$$

and similarly

$$\begin{aligned} \|\mathbf{V}_K - \mathbf{X}^*(\mathbf{X}^*)^\top \mathbf{V}_K\|_F^2 &= \min_{\mathbf{X} \in \mathcal{X}} \|\mathbf{V}_K - \mathbf{X}(\mathbf{X})^\top \mathbf{V}_K\|_F^2 \\ &= \min_{\mathcal{C}_1, \dots, \mathcal{C}_K} F(\mathcal{C}_1, \dots, \mathcal{C}_K) = F^*, \end{aligned} \quad (17)$$

furthermore, we have

$$\|\mathbf{V}_K - \hat{\mathbf{X}} \hat{\mathbf{X}}^\top \mathbf{V}_K\|_F^2 = F(\hat{\mathcal{C}}_1, \dots, \hat{\mathcal{C}}_K). \quad (18)$$

Let us begin the proof of Theorem 1. Define an error matrix as $\mathbf{E} = \mathbf{V}_K \mathbf{V}_K^\top - \hat{\mathbf{P}}_K \hat{\mathbf{P}}_K^\top$. We observe the following chain:

$$\begin{aligned} \|\mathbf{V}_K - \hat{\mathbf{X}} \hat{\mathbf{X}}^\top \mathbf{V}_K\|_F &= \|(\mathbf{I} - \hat{\mathbf{X}} \hat{\mathbf{X}}^\top) \mathbf{V}_K \mathbf{V}_K^\top\|_F = \|(\mathbf{I} - \hat{\mathbf{X}} \hat{\mathbf{X}}^\top)(\hat{\mathbf{P}}_K \hat{\mathbf{P}}_K^\top + \mathbf{E})\|_F \\ &\leq \|(\mathbf{I} - \hat{\mathbf{X}} \hat{\mathbf{X}}^\top) \hat{\mathbf{P}}_K \hat{\mathbf{P}}_K^\top\|_F + \|(\mathbf{I} - \hat{\mathbf{X}} \hat{\mathbf{X}}^\top) \mathbf{E}\|_F \\ &\leq \|(\mathbf{I} - \hat{\mathbf{X}} \hat{\mathbf{X}}^\top) \hat{\mathbf{P}}_K \hat{\mathbf{P}}_K^\top\|_F + \|\mathbf{E}\|_F, \end{aligned}$$

where the second equality is due to the orthogonality of \mathbf{V}_K^\top and the last inequality is due to the fact that $(\mathbf{I} - \hat{\mathbf{X}} \hat{\mathbf{X}}^\top)$ is a projection matrix. Using (16), we proceed as

$$\begin{aligned} &\|(\mathbf{I} - \hat{\mathbf{X}} \hat{\mathbf{X}}^\top) \hat{\mathbf{P}}_K \hat{\mathbf{P}}_K^\top\|_F + \|\mathbf{E}\|_F \\ &\leq \|(\mathbf{I} - \mathbf{X}^*(\mathbf{X}^*)^\top) \hat{\mathbf{P}}_K \hat{\mathbf{P}}_K^\top\|_F + \|\mathbf{E}\|_F \\ &= \|(\mathbf{I} - \mathbf{X}^*(\mathbf{X}^*)^\top)(\mathbf{V}_K \mathbf{V}_K^\top - \mathbf{E})\|_F + \|\mathbf{E}\|_F \\ &\leq \|(\mathbf{I} - \mathbf{X}^*(\mathbf{X}^*)^\top) \mathbf{V}_K \mathbf{V}_K^\top\|_F + 2\|\mathbf{E}\|_F \\ &= \sqrt{F^*} + 2\|\mathbf{E}\|_F. \end{aligned} \quad (19)$$

To bound $\|\mathbf{E}\|_F = \|\mathbf{V}_K \mathbf{V}_K^\top - \hat{\mathbf{P}}_K \hat{\mathbf{P}}_K^\top\|_F$, we invoke the following lemma:

Lemma 1 [22, Lemma 7] For any $\mathbf{A}, \mathbf{B} \in \mathbb{R}^{m \times n}$ with $m \geq n$ and $\mathbf{A}^\top \mathbf{A} = \mathbf{B}^\top \mathbf{B} = \mathbf{I}$, it holds that:

$$\|\mathbf{A} \mathbf{A}^\top - \mathbf{B} \mathbf{B}^\top\|_F^2 \leq 2n \|\mathbf{A} \mathbf{A}^\top - \mathbf{B} \mathbf{B}^\top\|_2^2. \quad (20)$$

As a result, we obtain the bound that:

$$\|\mathbf{E}\|_F \leq \sqrt{2K} \|\mathbf{V}_K \mathbf{V}_K^\top - \hat{\mathbf{P}}_K \hat{\mathbf{P}}_K^\top\|_2, \quad (21)$$

which can be further bounded using the following propositions —

Proposition 1 Under Condition 1, 3 and 4 in Theorem 1, we have

$$\|\mathbf{P}_K \mathbf{P}_K^\top - \mathbf{V}_K \mathbf{V}_K^\top\|_2^2 = \frac{\gamma^2}{1 + \gamma^2}, \quad (22)$$

where the columns of \mathbf{P}_K is the top K eigenvectors of \mathbf{C}_x and γ is bounded as stated in (12).

Proposition 2 Under Condition 5 in Theorem 1, it holds that

$$\|\mathbf{P}_K \mathbf{P}_K^\top - \hat{\mathbf{P}}_K \hat{\mathbf{P}}_K^\top\|_2 \leq \frac{|\beta_1(\Delta)|}{\delta}. \quad (23)$$

Obviously, combining (22), (23) and using the triangular inequality yields the desired result in expression (11). The proof is thus concluded.

6.1. Proof of Proposition 1

Denote the rank- K approximation to $\mathcal{H}(\mathbf{S})$ as $[\mathcal{H}(\mathbf{S})]_K := \mathbf{V}_K \text{diag}(\tilde{\mathbf{h}}_K) \mathbf{V}_K^\top$ (this expression is valid since $\mathcal{H}(\mathbf{S})$ is low pass). Define the shorthand notation $\tilde{\mathbf{B}} := \mathbf{B} \mathbf{Q}_K$, we observe that

$$\mathcal{R}([\mathcal{H}(\mathbf{S})]_K) = \mathcal{R}([\mathcal{H}(\mathbf{S})]_K \tilde{\mathbf{B}}), \quad (24)$$

where the equality is due to condition 3 of Theorem 1 such that the linear transformation on the right does not modify the range space of $[\mathcal{H}(\mathbf{S})]_K$. If we denote the columns of $\tilde{\mathbf{V}}_K$ as the top K left singular vectors of $[\mathcal{H}(\mathbf{S})]_K \tilde{\mathbf{B}}$, the above shows that the two products are equal $\mathbf{V}_K \mathbf{V}_K^\top = \tilde{\mathbf{V}}_K \tilde{\mathbf{V}}_K^\top$. Notice that \mathbf{V}_K corresponds to the K eigenvectors of \mathbf{S} with the K smallest eigenvalues [cf. Section 2] and it is required by the spectral clustering method applied on \mathbf{S} .

Similarly, we define $[\mathbf{C}_x]_K := \mathbf{P}_K \text{diag}(\sigma_K)^2 \mathbf{P}_K^\top$ as the rank K approximation to \mathbf{C}_x and observe that

$$\mathcal{R}([\mathbf{C}_x]_K) = \mathcal{R}([\mathcal{H}(\mathbf{S}) \mathbf{B}]_K) = \mathcal{R}(\mathcal{H}(\mathbf{S}) \tilde{\mathbf{B}}) \quad (25)$$

where the last equality is due to the fact that $\mathcal{H}(\mathbf{S}) \tilde{\mathbf{B}} = \mathcal{H}(\mathbf{S}) \mathbf{B} \mathbf{Q}_K = \mathbf{P}_K \text{diag}(\sigma_K) \mathbf{P}_K^\top$ as the columns of \mathbf{Q}_K are the top K right singular vectors. Likewise, if the columns of $\tilde{\mathbf{P}}_K$ are the top K left singular vectors of $\mathcal{H}(\mathbf{S}) \tilde{\mathbf{B}}$, then $\mathbf{P}_K \mathbf{P}_K^\top = \tilde{\mathbf{P}}_K \tilde{\mathbf{P}}_K^\top$.

Furthermore, we observe that

$$\mathcal{R}([\mathcal{H}(\mathbf{S})]_K \tilde{\mathbf{B}}) \perp \mathcal{R}((\mathcal{H}(\mathbf{S}) - [\mathcal{H}(\mathbf{S})]_K) \tilde{\mathbf{B}}). \quad (26)$$

Invoking [29, Lemma 8] through setting $\mathbf{D} = \mathcal{H}(\mathbf{S}) \tilde{\mathbf{B}}$, $\mathbf{C} = [\mathcal{H}(\mathbf{S})]_K \tilde{\mathbf{B}}$ and $\mathbf{E} = [\mathcal{H}(\mathbf{S})]_{N-K} \tilde{\mathbf{B}}$ therein, shows the following:

$$\begin{aligned} &\|\tilde{\mathbf{V}}_K \tilde{\mathbf{V}}_K^\top - \tilde{\mathbf{P}}_K \tilde{\mathbf{P}}_K^\top\|_2^2 \\ &= 1 - \beta_K \left([\mathcal{H}(\mathbf{S})]_K \tilde{\mathbf{B}} ((\mathcal{H}(\mathbf{S}) \tilde{\mathbf{B}})^\top \mathcal{H}(\mathbf{S}) \tilde{\mathbf{B}})^\dagger ([\mathcal{H}(\mathbf{S})]_K \tilde{\mathbf{B}})^\top \right). \end{aligned}$$

Denote the matrix in the middle of the expression above as $\Pi := (\mathcal{H}(\mathbf{S}) \tilde{\mathbf{B}})^\top \mathcal{H}(\mathbf{S}) \tilde{\mathbf{B}}$. Now, under condition 4 of Theorem 1 that

$\sigma_K > 0$, the $K \times K$ matrix $\mathbf{\Pi}$ is non-singular. We can observe the following chain:

$$\begin{aligned} & \beta_K \left([\mathcal{H}(\mathbf{S})]_K \tilde{\mathbf{B}} \mathbf{\Pi}^{-1} ([\mathcal{H}(\mathbf{S})]_K \tilde{\mathbf{B}})^\top \right) \\ &= \beta_K \left(\text{diag}(\tilde{\mathbf{h}}_K) \mathbf{V}_K^\top \tilde{\mathbf{B}} \mathbf{\Pi}^{-1} (\text{diag}(\tilde{\mathbf{h}}_K) \mathbf{V}_K^\top \tilde{\mathbf{B}})^\top \right) \\ &= \frac{1}{\beta_1 \left((\text{diag}(\tilde{\mathbf{h}}_K) \mathbf{V}_K^\top \tilde{\mathbf{B}})^{-\top} \mathbf{\Pi} (\text{diag}(\tilde{\mathbf{h}}_K) \mathbf{V}_K^\top \tilde{\mathbf{B}})^{-1} \right)}, \end{aligned} \quad (27)$$

where the first equality is due to $\beta_K(\mathbf{V} \mathbf{A} \mathbf{V}^\top) = \beta_K(\mathbf{A})$ for any $\mathbf{V} \in \mathbb{R}^{N \times K}$ with orthogonal columns. Moreover, observe that $\mathbf{\Pi}$ has the following decomposition:

$$\begin{aligned} \mathbf{\Pi} &= (\mathcal{H}(\mathbf{S}) \mathbf{B} \mathbf{Q}_K)^\top \mathcal{H}(\mathbf{S}) \mathbf{B} \mathbf{Q}_K \\ &= (\text{diag}(\tilde{\mathbf{h}}_K) \mathbf{V}_K^\top \mathbf{B} \mathbf{Q}_K)^\top (\text{diag}(\tilde{\mathbf{h}}_K) \mathbf{V}_K^\top \mathbf{B} \mathbf{Q}_K) \\ &\quad + \mathbf{Q}_K^\top \mathbf{B}^\top \mathbf{V}_{N-K} \text{diag}(\tilde{\mathbf{h}}_{N-K})^2 \mathbf{V}_{N-K}^\top \mathbf{B} \mathbf{Q}_K. \end{aligned} \quad (28)$$

This yields:

$$\begin{aligned} & \beta_K \left([\mathcal{H}(\mathbf{S})]_K \tilde{\mathbf{B}} \mathbf{\Pi}^{-1} ([\mathcal{H}(\mathbf{S})]_K \tilde{\mathbf{B}})^\top \right) \\ &= \left(1 + \beta_1 \left((\text{diag}(\tilde{\mathbf{h}}_K) \mathbf{V}_K^\top \tilde{\mathbf{B}})^{-\top} \tilde{\mathbf{B}}^\top \mathbf{V}_{N-K} \right. \right. \\ &\quad \left. \left. \text{diag}(\tilde{\mathbf{h}}_{N-K})^2 \mathbf{V}_{N-K}^\top \tilde{\mathbf{B}} (\text{diag}(\tilde{\mathbf{h}}_K) \mathbf{V}_K^\top \tilde{\mathbf{B}})^{-1} \right) \right)^{-1} \\ &= \frac{1}{1 + \|\text{diag}(\tilde{\mathbf{h}}_{N-K}) \mathbf{V}_{N-K}^\top \tilde{\mathbf{B}} (\text{diag}(\tilde{\mathbf{h}}_K) \mathbf{V}_K^\top \tilde{\mathbf{B}})^{-1}\|_2^2} \\ &= \left(1 + \gamma^2 \right)^{-1}, \end{aligned} \quad (29)$$

where we have defined γ such that:

$$\begin{aligned} \gamma &:= \|\text{diag}(\tilde{\mathbf{h}}_{N-K}) \mathbf{V}_{N-K}^\top \tilde{\mathbf{B}} (\text{diag}(\tilde{\mathbf{h}}_K) \mathbf{V}_K^\top \tilde{\mathbf{B}})^{-1}\|_2 \\ &\leq \left(\frac{\tilde{h}_{K+1}}{\tilde{h}_K} \right) \cdot \|\mathbf{V}_{N-K}^\top \mathbf{B} \mathbf{Q}_K\|_2 \|(\mathbf{V}_K^\top \mathbf{B} \mathbf{Q}_K)^{-1}\|_2, \end{aligned} \quad (30)$$

as desired. This concludes the proof of our claim.

6.2. Proof of Proposition 2

Observe that the left hand side of (23) can be written as:

$$\|\mathbf{P}_K \mathbf{P}_K^\top - \hat{\mathbf{P}}_K \hat{\mathbf{P}}_K^\top\|_2 = \|\hat{\mathbf{P}}_{N-K}^\top \mathbf{P}_K\|_2, \quad (31)$$

where the last equality is due to [30, Theorem 2.6.1].

Now, observe that Condition 5 of Theorem 1 implies that the largest eigenvalue in $\hat{\Sigma}_{N-K}$ will never exceed $\beta_K(\mathbf{C}_x) - \delta$ since

$$\begin{aligned} \beta_{\max}(\hat{\Sigma}_{N-K}) &= \beta_{K+1}(\hat{\mathbf{C}}_y) \leq \beta_{K+1}(\mathbf{C}_x) + \beta_1(\mathbf{\Delta}) \\ &\leq \beta_{K+1}(\mathbf{C}_x) + \|\mathbf{\Delta}\|_2, \end{aligned} \quad (32)$$

where the second last inequality is due to Weyl's inequality. The perturbed matrix $\hat{\mathbf{C}}_y$ thus satisfies the requirement of the Davis-Kahan's $\sin(\Theta)$ theorem [27], which gives:

$$\|\hat{\mathbf{P}}_{N-K}^\top \mathbf{P}_K\|_2 \leq \frac{\|\hat{\mathbf{P}}_{N-K}^\top \mathbf{\Delta} \mathbf{P}_K\|_2}{\delta}. \quad (33)$$

The last inequality in (23) is obtained by observing that both \mathbf{P}_K and $\hat{\mathbf{P}}_{N-K}$ are orthogonal matrices.