$$-1 < \theta < 1$$
 \Rightarrow $\sum_{k=0}^{\infty} \theta^k = \frac{1}{1-\theta}$ $\sum_{k=1}^{\infty} \theta^k = \frac{\theta}{1-\theta}$ $\frac{d}{dx} \log x = \frac{1}{x}$

$$\frac{1}{ad-bc}\begin{pmatrix} d & -b \\ -c & a \end{pmatrix} \qquad \frac{-b\pm\sqrt{b^2-4ac}}{2a} \qquad \text{Binomial Thm } (x+y)^n = \sum_{k=0}^n \binom{n}{k} x^k y^{n-k}$$

$$Cov(X, Y) = E[(X - EX)(Y - EY)] = E(XY) - E(X)E(Y)$$

$$\gamma(h) = \text{Cov}(X_{t+h}, X_t) = E(X_{t+h}X_t) \text{ when } EX = 0 \qquad \gamma(0) = \text{Var } X = \sigma_X^2$$

$$\hat{\gamma}(h) = \frac{1}{n} \sum_{t=1}^{n-h} (x_{t+h} - \bar{x})(x_t - \bar{x}) \qquad \nabla_{12} X_t = X_t - X_{t-12} \qquad \nabla^d = (1 - B)^d$$

$$\cos A \cos B + \sin A \sin B = \cos(A - B) = \cos(B - A)$$

$$\cos 2\lambda = 2\cos^2 \lambda - 1 \qquad 2\cos \theta = e^{i\theta} + e^{-i\theta} \qquad 2i\sin \theta = e^{i\theta} - e^{-i\theta} \qquad (-1)^{|h|} = \cos(\pi h)$$

The noise band is $\pm 1.96/\sqrt{n}$ Prediction bounds $\sigma_n^2(h) \approx \sigma^2 \sum_{j=0}^{h-1} \psi_j^2$

Estimation of μ

$$\sqrt{n} \left(\bar{X} - \mu \right) \sim N \left(0, \sum_{h = -\infty}^{\infty} \gamma(h) \right) \quad \Rightarrow \quad \bar{X} \sim N \left(\mu, \frac{1}{n} \sum_{h = -\infty}^{\infty} \gamma(h) \right)$$

The 95% confidence interval for μ is $\bar{x} \pm 1.96\sqrt{v/n}$ where $v = \sum_{h=-\infty}^{\infty} \gamma(h)$

$$AR(1) \Rightarrow v = \frac{\sigma^2}{(1-\phi)^2} \quad MA(1) \Rightarrow v = \sigma^2(1+2\theta+\theta^2)$$

Best Linear Predictor

$$\begin{pmatrix} a_1 \\ a_2 \end{pmatrix} = \begin{pmatrix} \gamma(0) & \gamma(1) \\ \gamma(1) & \gamma(0) \end{pmatrix}^{-1} \begin{pmatrix} \gamma(1) \\ \gamma(2) \end{pmatrix} \qquad a_0 = \mu(1 - a_1 - a_2)$$

$$P_n X_{n+h} = \mu + a_1 (X_n - \mu) + \dots + a_n (X_1 - \mu) = a_0 + a_1 X_n + \dots + a_n X_1$$

$$E(X_{n+h} - P_n X_{n+h})^2 = \gamma(0) - \mathbf{a}_n^T \gamma_n(h)$$

$$E(X_{n+h} - P_n X_{n+h}) = 0$$

$$E[(X_{n+h} - P_n X_{n+h})X_j] = 0, j = 1, \dots, n$$

$$E(X_{n+1} - \hat{X}_{n+1})^2 = \sigma^2 r_n = v_n$$

Innovations

$$v_0 = E[(X_1 - \hat{X}_1)^2] = \gamma(0)$$
 $\theta_{11} = \frac{\gamma(1)}{\gamma(0)}$

$$v_1 = E[(X_2 - \hat{X}_2)^2] = \gamma(0) - \frac{\gamma(1)^2}{\gamma(0)} \qquad \theta_{22} = \frac{\gamma(2)}{\gamma(0)} \qquad \theta_{21} = \frac{\gamma(1)[\gamma(0) - \gamma(2)]}{\gamma(0)^2 - \gamma(1)^2} \qquad \theta_{33} = \frac{\gamma(3)}{\gamma(0)}$$

$$\boxed{\mathrm{MA}(\infty)}$$
 $X_t = \sum_{j=0}^{\infty} \psi_j Z_{t-j}$ $\boxed{\mathrm{AR}(\infty)}$ $Z_t = \sum_{j=0}^{\infty} \pi_j X_{t-j}$

AR(1)
$$X_t - \phi X_{t-1} = Z_t$$
 $\{Z_t\} \sim WN(0, \sigma^2)$ $\gamma_X(h) = \frac{\sigma^2 \phi^{|h|}}{1 - \phi^2}$ $\rho_X(h) = \phi^{|h|}$

$$\hat{\phi} = \frac{\hat{\gamma}(1)}{\hat{\gamma}(0)} \qquad \hat{\sigma}^2 = \hat{\gamma}(0) - \hat{\gamma}(1)\hat{\phi} \qquad \bar{x} \pm 1.96 \frac{\sigma}{(1-\phi)\sqrt{n}} \qquad E(X_{n+h} - P_n X_{n+h})^2 = \frac{\sigma^2(1-\phi^{2h})}{1-\phi^2}$$

$$|\phi| < 1 \Rightarrow X_t = \sum_{j=0}^{\infty} \phi^j Z_{t-j} \quad |\phi| > 1 \Rightarrow X_t = -\sum_{j=1}^{\infty} \phi^{-j} Z_{t+j} \quad \sum_{h=-\infty}^{\infty} \gamma_X(h) = \frac{\sigma^2}{(1-\phi)^2}$$

$$f(\lambda) = \frac{\sigma^2}{2\pi} \left(1 - 2\phi \cos \lambda + \phi^2 \right)^{-1}$$

$$\bar{x}_n \pm 1.96\sqrt{v/n}$$
 $v = \sum_{h=-\infty}^{\infty} \gamma(h) = \sigma^2(1 + 2\theta + \theta^2)$ $f(\lambda) = \frac{\sigma^2}{2\pi} \left(1 + 2\theta \cos \lambda + \theta^2\right)$

ARMA(1,1)

$$\psi_0 = 1$$
 $\psi_j = (\phi + \theta)\phi^{j-1}, j = 1, 2, \dots$

$$\pi_0 = 1$$
 $\pi_j = -(\phi + \theta)(-\theta)^{j-1}, j = 1, 2, \dots$

$$\gamma(0) = \sigma^2 \left[1 + \frac{(\theta + \phi)^2}{1 - \phi^2} \right] \qquad \gamma(1) = \sigma^2 \left[\theta + \phi + \frac{(\theta + \phi)^2 \phi}{1 - \phi^2} \right] \qquad \gamma(h) = \phi^{h-1} \gamma(1), \ h \ge 2$$

ARMA(p,q)

$$\phi(z) = 1 - \phi_1 z - \dots - \phi_p z^p$$
 Roots of $\phi(z)$ greater than 1 implies causality.

$$\theta(z) = 1 + \theta_1 z + \dots + \theta_q z^q$$
 Roots of $\theta(z)$ greater than 1 implies invertibility.

$$\psi(z) = \frac{\theta(z)}{\phi(z)} = 1 + \psi_1 z + \psi_2 z^2 + \psi_3 z^3 + \cdots$$

$$\psi_{j} = \theta_{j} + \sum_{k=1}^{p} \phi_{k} \psi_{j-k} \qquad \pi_{j} = -\phi_{j} - \sum_{k=1}^{q} \theta_{k} \pi_{j-k} \qquad f(\lambda) = \frac{\sigma^{2}}{2\pi} \frac{\left|\theta\left(e^{-i\lambda}\right)\right|^{2}}{\left|\phi\left(e^{-i\lambda}\right)\right|^{2}} \qquad -\pi \leq \lambda \leq \pi$$

$$\underbrace{\text{Yule-Walker}} \begin{pmatrix} \hat{\phi}_1 \\ \hat{\phi}_2 \end{pmatrix} = \begin{pmatrix} \hat{\gamma}(0) & \hat{\gamma}(1) \\ \hat{\gamma}(1) & \hat{\gamma}(0) \end{pmatrix}^{-1} \begin{pmatrix} \hat{\gamma}(1) \\ \hat{\gamma}(2) \end{pmatrix} \qquad \hat{\sigma}^2 = \hat{\gamma}(0) - \hat{\gamma}(1)\hat{\phi}_1 - \hat{\gamma}(2)\hat{\phi}_2$$

3.2.1 Calculation of the ACVF

$$X_t = \sum_{j=0}^{\infty} \psi_j Z_{t-j}, \{Z_t\} \sim \text{WN}(0, \sigma^2) \quad \Rightarrow \quad \gamma_X(h) = E(X_{t+h} X_t) = \sigma^2 \sum_{j=0}^{\infty} \psi_j \psi_{j+|h|}$$

$$\gamma_X(0) = \sigma^2 \sum_{j=0}^{\infty} \psi_j^2 \quad \text{MA}(q) \Rightarrow \quad \gamma_X(h) = \begin{cases} \sigma^2 \sum_{j=0}^{q-|h|} \theta_j \theta_{j+|h|} & |h| \le q \\ 0 & |h| > q \end{cases}$$

4.1 Spectral Densities WN
$$\Rightarrow f(\lambda) = \frac{\sigma^2}{2\pi}$$

 $\lambda = \text{radians}$ $\omega = \text{radians/sec}$ period = $2\pi \text{ radians/}\omega = \text{seconds}$

 $f(\lambda) = \text{spectral density function}$ $F(\lambda) = \text{spectral distribution function}$

$$f(\lambda) = \frac{1}{2\pi} \sum_{h=-\infty}^{\infty} e^{-ih\lambda} \gamma(h) = \frac{1}{2\pi} \left[\gamma(0) + \sum_{h=1}^{\infty} (e^{-ih\lambda} + e^{ih\lambda}) \gamma(h) \right], -\infty < \lambda < \infty$$

$$f(\lambda) \ge 0$$
 $f(-\lambda) = f(\lambda)$ $f(\lambda + 2\pi) = f(\lambda)$ $\int_{-\pi}^{\pi} f(\lambda) d\lambda = \gamma(0)$

$$\gamma(h) = \int_{-\pi}^{\pi} e^{ih\lambda} f(\lambda) \, d\lambda = \int_{-\pi}^{\pi} \cos(h\lambda) f(\lambda) \, d\lambda$$

$$F(\lambda) = \int_{-\pi}^{\lambda} f(y) \, dy$$
 $F(-\pi) = 0$ $F(\pi) = \gamma(0)$ $dF(\lambda) = f(\lambda) \, d\lambda$

$$X_t = A\cos(\omega t) + B\sin(\omega t); \ A, B \sim (0, \nu^2); \ \gamma_X(h) = \nu^2 \cos(\omega h); \ F_X(\lambda) = \begin{cases} 0 & \lambda < -\omega \\ \nu^2/2 & -\omega \le \lambda < \omega \\ \nu^2 & \lambda \ge \omega \end{cases}$$

4.3 Time-Invariant Linear Filters

Filter
$$Y_t = \sum_{j=-\infty}^{\infty} \psi_j X_{t-j} \quad \Rightarrow \quad f_Y(\lambda) = \left| \Psi \left(e^{-i\lambda} \right) \right|^2 f_X(\lambda) = \Psi \left(e^{-i\lambda} \right) \Psi \left(e^{i\lambda} \right) f_X(\lambda)$$

Transfer function $\Psi\left(e^{-i\lambda}\right) = \sum_{j=-\infty}^{\infty} \psi_{j} e^{-ij\lambda}$ Power transfer function $\left|\Psi\left(e^{-i\lambda}\right)\right|^{2}$

$$Y_t = \frac{1}{2q+1} \sum_{j=-q}^{q} X_{t-j} \quad \Rightarrow \quad f_Y(\lambda) = \frac{1}{2q+1} \left(\sum_{j=-q}^{q} e^{-ij\lambda} \right)^2 f_X(\lambda)$$

$$q = 1$$
 \Rightarrow $f_Y(\lambda) = \frac{1}{3} \left(e^{i\lambda} + 1 + e^{-i\lambda} \right)^2 f_X(\lambda)$

5.3 Diagnostic Checking

$$\hat{W}_t = \frac{X_t - \hat{X}_t}{\sqrt{r_{t-1}}}$$
 $t = 1, \dots, n$ The noise band is $\pm 1.96/\sqrt{n}$

5.4 Forecasting
$$P_n X_{n+1} = \phi_1 X_n + \dots + \phi_p X_{n+1-p} + \sum_{j=1}^q \theta_{nj} (X_{n+1-j} - \hat{X}_{n+1-j}) \quad \theta_{nj} \approx \theta_j$$

$$X_{n+h} - \hat{X}_{n+h} = 0 \text{ for } h \ge 1$$
 $v_n = \sigma^2 r_n = E[(X_{n+1} - P_n X_{n+1})^2] = \sigma_n^2(1) \approx \sigma^2 \sum_{i=0}^{n} \psi_i^2 = \sigma^2$

5.5 Order Selection AICC =
$$-2 \log L + 2(p+q+1)n/(n-p-q-2)$$

$$L = \frac{1}{\sqrt{(2\pi)^n v_0 \cdots v_{n-1}}} \exp\left[-\frac{1}{2} \sum_{j=1}^n \frac{(X_j - \hat{X}_j)^2}{v_{j-1}}\right] \qquad v_n = \sigma^2 r_n$$

$$\ell = \log L = -\frac{n}{2}\log(2\pi) - \frac{n}{2}\log\sigma^2 - \frac{1}{2}\sum_{j=1}^n \log r_{j-1} - \frac{1}{2\sigma^2}\sum_{j=1}^n \frac{(X_j - \hat{X}_j)^2}{r_{j-1}}$$

Solve
$$\frac{\partial \ell}{\partial \sigma^2} = 0$$
 to obtain $\hat{\sigma}^2 = \frac{1}{n} \sum_{j=1}^n \frac{(X_j - \hat{X}_j)^2}{r_{j-1}}$

Maximize reduced log likelihood
$$\ell^* = -\frac{n}{2} \log \hat{\sigma}^2 - \frac{1}{2} \sum_{i=1}^n \log r_{i-1}$$

Or multiply by
$$-2/n$$
 and minimize $\ell^* = \log \hat{\sigma}^2 + \frac{1}{n} \sum_{j=1}^n \log r_{j-1}$ (5.2.12) on p. 160

$$E(Y_{n+1} - \hat{\phi}_1 Y_n - \dots - \hat{\phi}_p Y_{n+1-p})^2 \approx \sigma^2 \left(1 + \frac{p}{n}\right)$$

Classical Decomposition
$$X_t = m_t + s_t + Y_t$$
 $EY_t = 0$ $s_{t+d} = s_t$ $\sum_{j=1}^{d} s_j = 0$

No distortion
$$\Rightarrow m_t = \sum_j a_j m_{t-j}$$
 for all polynomials $m_t = c_0 + c_1 t + \cdots + c_k t^k$

No distortion
$$\Rightarrow \sum_{i} a_{i} = 1$$
 and $\sum_{j} j^{r} a_{j} = 0$ for $j = 1, \dots, k$

Eliminate seasonal components
$$\Rightarrow \sum_{j} a_j s_{t-j} = \text{const} \times \sum_{j=1}^d s_j = 0$$

Model	ACF	PACF
AR(p)	Decays	Zero for $h > p$
MA(q)	Zero for $h > q$	Decays
ARMA(p,q)	Decays	Decays