

$$-1 < \theta < 1 \quad \Rightarrow \quad \sum_{k=0}^{\infty} \theta^k = \frac{1}{1-\theta} \quad \sum_{k=1}^{\infty} \theta^k = \frac{\theta}{1-\theta} \quad \frac{d}{dx} \log x = \frac{1}{x}$$

$$\frac{1}{ad-bc} \begin{pmatrix} d & -b \\ -c & a \end{pmatrix} \quad \frac{-b \pm \sqrt{b^2 - 4ac}}{2a} \quad \text{Binomial Thm } (x+y)^n = \sum_{k=0}^n \binom{n}{k} x^k y^{n-k}$$

$$\text{Cov}(X, Y) = E[(X - EX)(Y - EY)] = E(XY) - E(X)E(Y)$$

$$\gamma(h) = \text{Cov}(X_{t+h}, X_t) = E(X_{t+h}X_t) \text{ when } EX = 0 \quad \gamma(0) = \text{Var } X = \sigma_X^2$$

$$\hat{\gamma}(h) = \frac{1}{n} \sum_{t=1}^{n-h} (x_{t+h} - \bar{x})(x_t - \bar{x}) \quad \nabla_{12} X_t = X_t - X_{t-12} \quad \nabla^d = (1 - B)^d$$

$$\cos A \cos B + \sin A \sin B = \cos(A - B) = \cos(B - A)$$

$$\cos 2\lambda = 2 \cos^2 \lambda - 1 \quad 2 \cos \theta = e^{i\theta} + e^{-i\theta} \quad 2i \sin \theta = e^{i\theta} - e^{-i\theta} \quad (-1)^{|h|} = \cos(\pi h)$$

$$\text{The noise band is } \pm 1.96/\sqrt{n} \quad \text{Prediction bounds } \sigma_n^2(h) \approx \sigma^2 \sum_{j=0}^{h-1} \psi_j^2$$

$$\boxed{\text{Estimation of } \mu}$$

$$\sqrt{n}(\bar{X} - \mu) \sim N\left(0, \sum_{h=-\infty}^{\infty} \gamma(h)\right) \quad \Rightarrow \quad \bar{X} \sim N\left(\mu, \frac{1}{n} \sum_{h=-\infty}^{\infty} \gamma(h)\right)$$

$$\text{The 95\% confidence interval for } \mu \text{ is } \bar{x} \pm 1.96\sqrt{v/n} \text{ where } v = \sum_{h=-\infty}^{\infty} \gamma(h)$$

$$\text{AR}(1) \quad \Rightarrow \quad v = \frac{\sigma^2}{(1-\phi)^2} \quad \text{MA}(1) \quad \Rightarrow \quad v = \sigma^2(1 + 2\theta + \theta^2)$$

$$\boxed{\text{Best Linear Predictor}}$$

$$\begin{pmatrix} a_1 \\ a_2 \end{pmatrix} = \begin{pmatrix} \gamma(0) & \gamma(1) \\ \gamma(1) & \gamma(0) \end{pmatrix}^{-1} \begin{pmatrix} \gamma(1) \\ \gamma(2) \end{pmatrix} \quad a_0 = \mu(1 - a_1 - a_2)$$

$$P_n X_{n+h} = \mu + a_1(X_n - \mu) + \cdots + a_n(X_1 - \mu) = a_0 + a_1 X_n + \cdots + a_n X_1$$

$$E(X_{n+h} - P_n X_{n+h})^2 = \gamma(0) - \mathbf{a}_n^T \gamma_n(h)$$

$$E(X_{n+h} - P_n X_{n+h}) = 0$$

$$E[(X_{n+h} - P_n X_{n+h})X_j] = 0, \quad j = 1, \dots, n$$

$$E(X_{n+1} - \hat{X}_{n+1})^2 = \sigma^2 r_n = v_n$$

$$\boxed{\text{Innovations}}$$

$$v_0 = E[(X_1 - \hat{X}_1)^2] = \gamma(0) \quad \theta_{11} = \frac{\gamma(1)}{\gamma(0)}$$

$$v_1 = E[(X_2 - \hat{X}_2)^2] = \gamma(0) - \frac{\gamma(1)^2}{\gamma(0)} \quad \theta_{22} = \frac{\gamma(2)}{\gamma(0)} \quad \theta_{21} = \frac{\gamma(1)[\gamma(0) - \gamma(2)]}{\gamma(0)^2 - \gamma(1)^2} \quad \theta_{33} = \frac{\gamma(3)}{\gamma(0)}$$

$$\boxed{\text{MA}(\infty)} \quad X_t = \sum_{j=0}^{\infty} \psi_j Z_{t-j} \quad \boxed{\text{AR}(\infty)} \quad Z_t = \sum_{j=0}^{\infty} \pi_j X_{t-j}$$

$$\boxed{\text{AR}(1)} \quad X_t - \phi X_{t-1} = Z_t \quad \{Z_t\} \sim \text{WN}(0, \sigma^2) \quad \gamma_X(h) = \frac{\sigma^2 \phi^{|h|}}{1 - \phi^2} \quad \rho_X(h) = \phi^{|h|}$$

$$\hat{\phi} = \frac{\hat{\gamma}(1)}{\hat{\gamma}(0)} \quad \hat{\sigma}^2 = \hat{\gamma}(0) - \hat{\gamma}(1)\hat{\phi} \quad \bar{x} \pm 1.96 \frac{\sigma}{(1 - \phi)\sqrt{n}} \quad E(X_{n+h} - P_n X_{n+h})^2 = \frac{\sigma^2(1 - \phi^{2h})}{1 - \phi^2}$$

$$|\phi| < 1 \Rightarrow X_t = \sum_{j=0}^{\infty} \phi^j Z_{t-j} \quad |\phi| > 1 \Rightarrow X_t = - \sum_{j=1}^{\infty} \phi^{-j} Z_{t+j} \quad \sum_{h=-\infty}^{\infty} \gamma_X(h) = \frac{\sigma^2}{(1 - \phi)^2}$$

$$f(\lambda) = \frac{\sigma^2}{2\pi} (1 - 2\phi \cos \lambda + \phi^2)^{-1}$$

$$\boxed{\text{MA}(1)} \quad X_t = Z_t + \theta Z_{t-1} \quad \{Z_t\} \sim \text{WN}(0, \sigma^2) \quad \gamma_X(h) = \begin{cases} \sigma^2(1 + \theta^2) & h = 0 \\ \sigma^2\theta & h = \pm 1 \\ 0 & |h| > 1 \end{cases}$$

$$\bar{x}_n \pm 1.96\sqrt{v/n} \quad v = \sum_{h=-\infty}^{\infty} \gamma(h) = \sigma^2(1 + 2\theta + \theta^2) \quad f(\lambda) = \frac{\sigma^2}{2\pi} (1 + 2\theta \cos \lambda + \theta^2)$$

$$\boxed{\text{ARMA}(1,1)}$$

$$\psi_0 = 1 \quad \psi_j = (\phi + \theta)\phi^{j-1}, j = 1, 2, \dots$$

$$\pi_0 = 1 \quad \pi_j = -(\phi + \theta)(-\theta)^{j-1}, j = 1, 2, \dots$$

$$\gamma(0) = \sigma^2 \left[ 1 + \frac{(\theta + \phi)^2}{1 - \phi^2} \right] \quad \gamma(1) = \sigma^2 \left[ \theta + \phi + \frac{(\theta + \phi)^2 \phi}{1 - \phi^2} \right] \quad \gamma(h) = \phi^{h-1} \gamma(1), h \geq 2$$

$$\boxed{\text{ARMA}(p,q)}$$

$$\phi(z) = 1 - \phi_1 z - \dots - \phi_p z^p \quad \text{Roots of } \phi(z) \text{ greater than 1 implies causality.}$$

$$\theta(z) = 1 + \theta_1 z + \dots + \theta_q z^q \quad \text{Roots of } \theta(z) \text{ greater than 1 implies invertibility.}$$

$$\psi(z) = \frac{\theta(z)}{\phi(z)} = 1 + \psi_1 z + \psi_2 z^2 + \psi_3 z^3 + \dots$$

$$\psi_j = \theta_j + \sum_{k=1}^p \phi_k \psi_{j-k} \quad \pi_j = -\phi_j - \sum_{k=1}^q \theta_k \pi_{j-k} \quad f(\lambda) = \frac{\sigma^2}{2\pi} \frac{|\theta(e^{-i\lambda})|^2}{|\phi(e^{-i\lambda})|^2} \quad -\pi \leq \lambda \leq \pi$$

$$\boxed{\text{Yule-Walker}} \quad \begin{pmatrix} \hat{\phi}_1 \\ \hat{\phi}_2 \end{pmatrix} = \begin{pmatrix} \hat{\gamma}(0) & \hat{\gamma}(1) \\ \hat{\gamma}(1) & \hat{\gamma}(0) \end{pmatrix}^{-1} \begin{pmatrix} \hat{\gamma}(1) \\ \hat{\gamma}(2) \end{pmatrix} \quad \hat{\sigma}^2 = \hat{\gamma}(0) - \hat{\gamma}(1)\hat{\phi}_1 - \hat{\gamma}(2)\hat{\phi}_2$$

### 3.2.1 Calculation of the ACVF

$$X_t = \sum_{j=0}^{\infty} \psi_j Z_{t-j}, \{Z_t\} \sim \text{WN}(0, \sigma^2) \Rightarrow \gamma_X(h) = E(X_{t+h} X_t) = \sigma^2 \sum_{j=0}^{\infty} \psi_j \psi_{j+|h|}$$

$$\gamma_X(0) = \sigma^2 \sum_{j=0}^{\infty} \psi_j^2 \quad \text{MA}(q) \Rightarrow \gamma_X(h) = \begin{cases} \sigma^2 \sum_{j=0}^{q-|h|} \theta_j \theta_{j+|h|} & |h| \leq q \\ 0 & |h| > q \end{cases}$$

### 4.1 Spectral Densities

$$\text{WN} \Rightarrow f(\lambda) = \frac{\sigma^2}{2\pi}$$

$$\lambda = \text{radians} \quad \omega = \text{radians/sec} \quad \text{period} = 2\pi \text{ radians}/\omega = \text{seconds}$$

$$f(\lambda) = \text{spectral density function} \quad F(\lambda) = \text{spectral distribution function}$$

$$f(\lambda) = \frac{1}{2\pi} \sum_{h=-\infty}^{\infty} e^{-ih\lambda} \gamma(h) = \frac{1}{2\pi} \left[ \gamma(0) + \sum_{h=1}^{\infty} (e^{-ih\lambda} + e^{ih\lambda}) \gamma(h) \right], -\infty < \lambda < \infty$$

$$f(\lambda) \geq 0 \quad f(-\lambda) = f(\lambda) \quad f(\lambda + 2\pi) = f(\lambda) \quad \int_{-\pi}^{\pi} f(\lambda) d\lambda = \gamma(0)$$

$$\gamma(h) = \int_{-\pi}^{\pi} e^{ih\lambda} f(\lambda) d\lambda = \int_{-\pi}^{\pi} \cos(h\lambda) f(\lambda) d\lambda$$

$$F(\lambda) = \int_{-\pi}^{\lambda} f(y) dy \quad F(-\pi) = 0 \quad F(\pi) = \gamma(0) \quad dF(\lambda) = f(\lambda) d\lambda$$

$$X_t = A \cos(\omega t) + B \sin(\omega t); A, B \sim (0, \nu^2); \gamma_X(h) = \nu^2 \cos(\omega h); F_X(\lambda) = \begin{cases} 0 & \lambda < -\omega \\ \nu^2/2 & -\omega \leq \lambda < \omega \\ \nu^2 & \lambda \geq \omega \end{cases}$$

### 4.3 Time-Invariant Linear Filters

$$\text{Filter } Y_t = \sum_{j=-\infty}^{\infty} \psi_j X_{t-j} \Rightarrow f_Y(\lambda) = |\Psi(e^{-i\lambda})|^2 f_X(\lambda) = \Psi(e^{-i\lambda}) \Psi(e^{i\lambda}) f_X(\lambda)$$

$$\text{Transfer function } \Psi(e^{-i\lambda}) = \sum_{j=-\infty}^{\infty} \psi_j e^{-ij\lambda} \quad \text{Power transfer function } |\Psi(e^{-i\lambda})|^2$$

$$Y_t = \frac{1}{2q+1} \sum_{j=-q}^q X_{t-j} \Rightarrow f_Y(\lambda) = \frac{1}{2q+1} \left( \sum_{j=-q}^q e^{-ij\lambda} \right)^2 f_X(\lambda)$$

$$q = 1 \Rightarrow f_Y(\lambda) = \frac{1}{3} (e^{i\lambda} + 1 + e^{-i\lambda})^2 f_X(\lambda)$$

### 5.3 Diagnostic Checking

$$\hat{W}_t = \frac{X_t - \hat{X}_t}{\sqrt{r_{t-1}}} \quad t = 1, \dots, n \quad \text{The noise band is } \pm 1.96/\sqrt{n}$$

### 5.4 Forecasting

$$P_n X_{n+1} = \phi_1 X_n + \dots + \phi_p X_{n+1-p} + \sum_{j=1}^q \theta_{nj} (X_{n+1-j} - \hat{X}_{n+1-j}) \quad \theta_{nj} \approx \theta_j$$

$$X_{n+h} - \hat{X}_{n+h} = 0 \text{ for } h \geq 1 \quad v_n = \sigma^2 r_n = E[(X_{n+1} - P_n X_{n+1})^2] = \sigma_n^2(1) \approx \sigma^2 \sum_{j=0}^0 \psi_j^2 = \sigma^2$$

### 5.5 Order Selection

$$\text{AICC} = -2 \log L + 2(p+q+1)n/(n-p-q-2)$$

$$L = \frac{1}{\sqrt{(2\pi)^n v_0 \dots v_{n-1}}} \exp \left[ -\frac{1}{2} \sum_{j=1}^n \frac{(X_j - \hat{X}_j)^2}{v_{j-1}} \right] \quad v_n = \sigma^2 r_n$$

$$\ell = \log L = -\frac{n}{2} \log(2\pi) - \frac{n}{2} \log \sigma^2 - \frac{1}{2} \sum_{j=1}^n \log r_{j-1} - \frac{1}{2\sigma^2} \sum_{j=1}^n \frac{(X_j - \hat{X}_j)^2}{r_{j-1}}$$

$$\text{Solve } \frac{\partial \ell}{\partial \sigma^2} = 0 \text{ to obtain } \hat{\sigma}^2 = \frac{1}{n} \sum_{j=1}^n \frac{(X_j - \hat{X}_j)^2}{r_{j-1}}$$

$$\text{Maximize reduced log likelihood } \ell^* = -\frac{n}{2} \log \hat{\sigma}^2 - \frac{1}{2} \sum_{j=1}^n \log r_{j-1}$$

$$\text{Or multiply by } -2/n \text{ and minimize } \ell^* = \log \hat{\sigma}^2 + \frac{1}{n} \sum_{j=1}^n \log r_{j-1} \quad (5.2.12) \text{ on p. 160}$$

$$E(Y_{n+1} - \hat{\phi}_1 Y_n - \dots - \hat{\phi}_p Y_{n+1-p})^2 \approx \sigma^2 \left(1 + \frac{p}{n}\right)$$

### Classical Decomposition

$$X_t = m_t + s_t + Y_t \quad EY_t = 0 \quad s_{t+d} = s_t \quad \sum_{j=1}^d s_j = 0$$

$$\text{No distortion} \Rightarrow m_t = \sum_j a_j m_{t-j} \text{ for all polynomials } m_t = c_0 + c_1 t + \dots + c_k t^k$$

$$\text{No distortion} \Rightarrow \sum_j a_j = 1 \text{ and } \sum_j j^r a_j = 0 \text{ for } j = 1, \dots, k$$

$$\text{Eliminate seasonal components} \Rightarrow \sum_j a_j s_{t-j} = \text{const} \times \sum_{j=1}^d s_j = 0$$

Model	ACF	PACF
AR( $p$ )	Decays	Zero for $h > p$
MA( $q$ )	Zero for $h > q$	Decays
ARMA( $p, q$ )	Decays	Decays