

The above analysis is valid only if the model is correctly specified: The true density is truly described by a parametric family  $f(x/\theta)$ ,  $\theta \in \Theta$  for one specific  $\theta_0$  which we are trying to determine from our sampling

Otherwise the model is mis-specified:

There is a true  $\theta_0$  such that the  $f(x/\theta_0)$  is the true density.

However, it can be seen as a good approximation of the true model. Then, the pseudo-true parameter  $\theta_0 \in \Theta$  is selected so that the approximate  $f(x/\theta_0)$  to diverge as little as possible from the true one:  $f(x)$

Kullback-Leibler divergence

$$KLIC(f, f_\theta) = \int f(x) \cdot \log \frac{f(x)}{f_\theta(x)} dx$$

$$= \int f(x) \cdot \log f(x) dx - \underbrace{\int f(x) \log f_\theta(x) dx}_{l(\theta)}$$

$\rightarrow KLIC(f, f) = 0$   
 $\rightarrow KLIC(f, g) \geq 0$   
 $\rightarrow \operatorname{argmin} KLIC(f, g) = f$

pseudo-true parameter:  $\theta_0 \equiv \operatorname{argmin}_{\theta \in \Theta} KLIC(f, f_\theta) \equiv \operatorname{argmax}_{\theta \in \Theta} l(\theta)$

$\Rightarrow$  For a mis-specified model the MLE- $\hat{\theta}$  is the sample analog of the pseudo-true parameter

So the whole analysis remains unaltered except for the difference that for mis-specified model:

$$\boxed{\theta_0 \neq \theta_0} \quad \text{thus} \quad \boxed{\sqrt{n}(\hat{\theta} - \theta_0) \xrightarrow{d} N(0, \frac{1}{H_0} \cdot F_0 \cdot \frac{1}{H_0})}$$

Practical comments

- to estimate the (pseudo) true  $\theta_0$ : calculate the MLE- $\hat{\theta} \equiv \operatorname{argmax}_{\theta} \bar{l}(\theta)$
- to estimate the variance of  $\hat{\theta}$ , assume that  $n$  is big enough for the normal distribution to be valid. Thus  $\operatorname{var}[\hat{\theta}] \equiv H_0^{-1} \cdot F_0 \cdot H_0^{-1}$  or  $\operatorname{var}[\hat{\theta}] \equiv F_0^{-1}$  for the pseudo-true case.
- the previous exact expression must be replaced by appropriate estimates of  $H_0, F_0$ .
- After calculating the estimate- $\hat{V}$ , our experimental estimation of  $\theta_0$  is MLE- $\hat{\theta}$  with standard error  $\sqrt{\hat{V}}$ .

$\rightarrow F_0^{-1} = H_0^{-1} \cdot F_0 \cdot H_0^{-1} \Rightarrow \oplus$  smaller BUT  $\operatorname{var}[\hat{V}]$  smaller  $\oplus$   
 $\rightarrow H_0^{-1} \cdot F_0 \cdot H_0^{-1} \Rightarrow \ominus$   $\operatorname{var}[\hat{V}]$  bigger BUT  $\ominus$  condition  $\oplus$