

## Chapter 2

The Monte-Carlo Method and applications to option pricing

### 2.1 The Monte-Carlo method

The basic principle of the Monte-Carlo method is to implement on a computer the strong law of large numbers (SLLN): if  $(X_k)_{k \geq 1}$  is a sequence, defined on a prob. space  $(\Omega, \mathcal{A}, P)$ , of independent random variables with the same distribution as  $X$ , then:

$$P(\text{d}w)\text{-a.s. } X_M = \frac{X_1(w) + \dots + X_M(w)}{M} \xrightarrow{M \rightarrow \infty} m_X = E(X)$$

the sequence  $(X_n)$  is also called an i.i.d. sequence of random variables with distribution  $\nu = P_X$ .

Conditions to implement this MC method on a computer:

- \* generate pseudo-random numbers  $(U_k)_{k \geq 1}$  of law  $U(0, 1)$
  - \* generate  $X$  as a function of  $U_1, U_2, \dots$ :
- $$X = \varphi_\tau(U_1, \dots, U_\tau) \quad (\tau \text{ is a finite stopping rule})$$

#### 2.1.1. Rates of convergence

The (weak) rate of convergence in the SLLN is governed by the Central Limit Theorem (CLT) which says that if  $X \in L^2(\mathbb{R})$  then:

$$\sqrt{M} (X_M - m_X) \xrightarrow[M \rightarrow \infty]{\text{law}} W(0; \sigma_X^2)$$

where  $\sigma_X^2 = \text{Var}(X)$

the mean quadratic error is exactly:

$$\|\bar{X}_M - m_x\|_2 = \sqrt{E((\bar{X}_M - m_x)^2)} = \frac{\sigma_x}{\sqrt{M}}$$

(indeed:  $E((\bar{X}_M - m_x)^2) = E\left(\frac{1}{M} \sum_{i=1}^M (X_i - m_x)^2\right)$

$$= \frac{1}{M} \sum_{i=1}^M E((X_i - m_x)^2)$$

$$= \frac{M\sigma_x^2}{M^2} = \frac{\sigma_x^2}{M}$$

This error is also known as the RMSE (Root Mean Square Error)

### 2.1.2. Confidence level and confidence interval

The CLT also reads:  $\sqrt{M} \left( \frac{\bar{X}_M - m_x}{\sigma_x} \right) \xrightarrow[M \rightarrow \infty]{\text{law}} N(0,1)$

so for all  $a, b \in \mathbb{R}$  ( $a < b$ ):

$$\lim_{M \rightarrow \infty} P\left(\sqrt{M} \left( \frac{\bar{X}_M - m_x}{\sigma_x} \right) \in (a; b)\right) = \Phi(b) - \Phi(a)$$

where  $\Phi$  is the cumulative distribution function of  $N(0,1)$

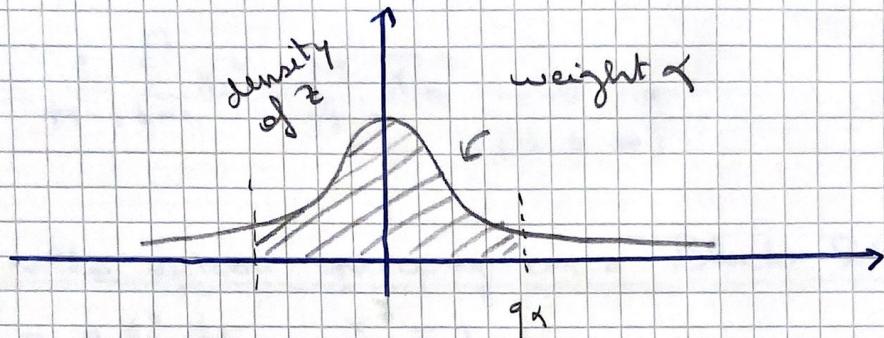
$$\Phi(x) = \int_{-\infty}^x \frac{e^{-t^2/2}}{\sqrt{2\pi}} dt$$

We usually assume  $\sqrt{M} \left( \frac{\bar{X}_M - m_x}{\sigma_x} \right)$  is of law  $N(0,1)$ .

Let  $\alpha \in (0,1)$  denote a confidence level (close to 1)  
and let  $q_\alpha$  be the two-sided quantile defined as the  
unique sol. to:  $\Phi(q_\alpha) = 1 - \alpha$

$$P(|Z| \leq q_\alpha) = \alpha \quad \text{i.e.} \quad 2\Phi(q_\alpha) - 1 = \alpha$$

$(Z \sim N(0,1))$



Small computation: we want  $\mathbb{P}(|z| \leq q_\alpha) = 1 - \alpha$

This is:  $\mathbb{P}(z > q_\alpha) + \mathbb{P}(z < -q_\alpha) = 1 - \alpha$

↑  
These two are equal by symmetry

$$2\mathbb{P}(z > q_\alpha) = 1 - \alpha$$

$$2(1 - \Phi(q_\alpha)) = 1 - \alpha$$

$$2\Phi(q_\alpha) - 1 = \alpha$$

$$\Phi(q_\alpha) = \frac{1+\alpha}{2}$$

We take  $a = q_\alpha$  and  $b = -q_\alpha$  and we can define a theoretical confidence interval at level  $\alpha$ :

$$J_M^\alpha = \left[ \bar{X}_M - q_\alpha \frac{\sigma_x}{\sqrt{M}}, \bar{X}_M + q_\alpha \frac{\sigma_x}{\sqrt{M}} \right]$$

which satisfies:  $\mathbb{P}(m_x \in J_M^\alpha)$

$$\begin{aligned} &\mathbb{P}\left(\left|\frac{\bar{X}_M - m_x}{\sigma_x}\right| \leq q_\alpha\right) \\ &\xrightarrow{(M \rightarrow \infty)} \mathbb{P}(|z| \leq q_\alpha) = \alpha \end{aligned}$$

However:  $J_M^\alpha$  is Theoretical because we do not know  $\sigma_x$ .

We usually replace  $\sigma_x^2$  by its Monte-Carlo estimate:

$$\begin{aligned} \hat{\sigma}_M^2 &= \frac{1}{M-1} \sum_{k=1}^M (X_k - \bar{X}_M)^2 = \frac{1}{M-1} \sum_{k=1}^M X_k^2 + \frac{M}{M-1} \bar{X}_M^2 \\ &\quad - 2 \bar{X}_M \sum_{k=1}^M X_k \end{aligned}$$

$$\dots = \frac{1}{M-1} \sum_{k=1}^M X_k^2 - \frac{M}{M-1} \bar{X}_M^2 \xrightarrow{(M \rightarrow \infty)} \sigma_X^2$$

### 2.1.3 Vanilla option pricing in a Black-Scholes model

$$\begin{cases} dX_t^0 = r X_t^0 dt, & X_0^0 = 1 \\ dX_t^1 = X_t^1 (rdt + \sigma_1 dW_t^1), & X_0^1 = x_0^1 \\ dX_t^2 = X_t^2 (rdt + \sigma_2 dW_t^2), & X_0^2 = x_0^2 \end{cases}$$

$W = (W^1, W^2)$  is a bi-dim. Brownian motion such that  $\langle W^1, W^2 \rangle_t = \gamma t$  ( $\gamma \in [-1, 1]$ ).

In other words, we can decompose

$$W_t^2 = PW_t^1 + \sqrt{1-P^2} \tilde{W}_t^2 \text{ where } (W^1, \tilde{W}^2) \text{ is a standard 2-dim. Brownian motion.}$$

filtration:  $(F_t)_{t \in [0, T]} = \sigma(W_s, 0 \leq s \leq t), \cup_{\mathbb{P}}$   
augmented with  $\mathbb{P}$ -negligible sets

$\mathbb{P}$  is the risk-neutral probability, meaning that

$$e^{-rt} X_t = e^{-rt} \begin{pmatrix} X_t^1 \\ X_t^2 \end{pmatrix} \text{ is a } (\mathbb{P}, (F_t))\text{-martingale}$$

$$\text{For all } t: X_t^0 = e^{-rt}, X_t^i = x_0^i \exp \left[ \left( r - \frac{\sigma_i^2}{2} \right) t + \sigma_i W_t^i \right] \quad (i=1, 2)$$

(This is an application of Itô's Lemma)

A European vanilla option with maturity  $T > 0$  is an option related to a European payoff:

$$h_T = h(X_T)$$

which only depends on  $X$  at time  $T$ . In such a complete market, the option premium at time  $t \in (0, T)$  is given by  $V_t = e^{-r(T-t)} \mathbb{E}(h(X_T) | F_t)$   
 $(V_0 = e^{-rT} \mathbb{E}(h(X_T)))$

As  $W$  has independent stationary increments, we

have:  $\left( \frac{X_T^i}{X_t^i} \right)_{i=1,2} \stackrel{\text{law}}{=} \left( \frac{X_{T-t}^i}{X_0^i} \right)_{i=1,2}$ .

We define  $v(x_0, T) = e^{-rt} E(h(X_T))$  and we compute

$$\begin{aligned} V_t &= e^{-r(T-t)} E(h(X_T) | \mathcal{F}_t) \\ &= e^{-r(T-t)} E\left(h\left(\left[X_T^i, \frac{X_T^i}{X_t^i}\right]_{i=1,2}\right) | \mathcal{F}_t\right) \\ &= e^{-r(T-t)} \left[ E h\left(\left[x^i \times \frac{X_T^i}{x_0^i}\right]_{i=1,2}\right) \right]_{x^i = X_t^i, i=1,2} \\ &= v(x_t, T-t) \end{aligned}$$

Example:  $\times$  Vanilla call with strike price  $K$ :

$$h(x^1, x^2) = (x^1 - K)_+$$

There is a closed form: Black-Scholes formula.  
(for  $V_T$ )

- \* Best-of-call with strike price  $K$ :  $h_T = (\max(X_T^1, X_T^2) - K)_+$   
↳ a quasi-closed form exists but we will price it  
using a Monte-Carlo simulation

$$\begin{aligned} \text{We re-write: } e^{-rt} h_T &\stackrel{\text{law}}{=} \Psi(z^1, z^2) \\ &:= \left[ \max \left( x_0^1 \exp \left( -\frac{\sigma_1^2}{2} T + \sigma_1 \sqrt{T} z^1 \right), \right. \right. \\ &\quad \left. \left. x_0^2 \exp \left( -\frac{\sigma_2^2}{2} T + \sigma_2 \sqrt{T} (z^2 + \sqrt{1 - z^2} z^1) \right) \right. \right. \\ &\quad \left. \left. - K e^{-rt} \right) \right]_+ \end{aligned}$$

where  $z = (z^1, z^2) \sim \mathcal{N}(0, I_2)$

We take a  $M$ -sample  $(z_m)_{1 \leq m \leq M}$  and get

$$\begin{aligned} (\text{Best-of-call}) &= E \Psi(z^1, z^2) \\ &\approx \bar{\Psi}_M := \frac{1}{M} \sum_{m=1}^M \Psi(z_m) \end{aligned}$$

Estimate for the variance:

$$\bar{V}_M(\varphi) = \frac{1}{M-1} \sum_{m=1}^M \varphi(Z_m)^2 - \frac{M}{M-1} \bar{\varphi}_M^2 \approx \hat{v}_M(\varphi(Z))$$

(for  $M$  large enough).

Level of confidence:  $\alpha \in (0, 1)$ .

Confidence interval:

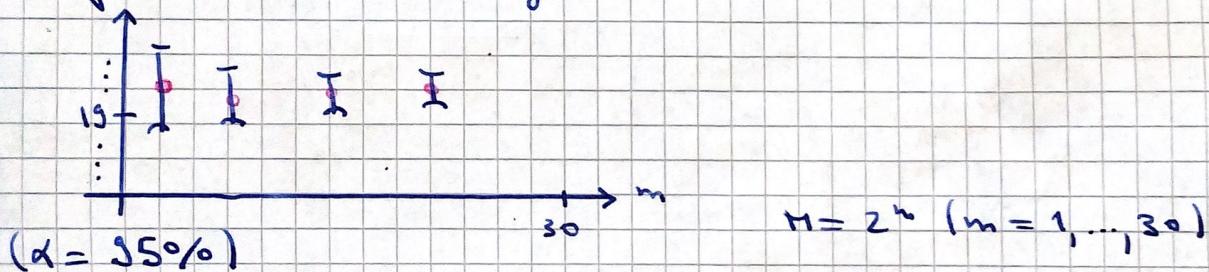
$$I_M^\alpha = \left[ \bar{\varphi}_M - q_\alpha \sqrt{\frac{\bar{V}_M(\varphi)}{M}}, \bar{\varphi}_M + q_\alpha \sqrt{\frac{\bar{V}_M(\varphi)}{M}} \right]$$

where  $q_\alpha$  is defined by:  $2\Phi(q_\alpha) - 1 = \alpha$ .

Numerical application:

$$n = 0.1, r_i = 0.2 = 20\%, \bar{\varphi} = 0.5, \bar{x}_0 = 100, T = 1, K = 100$$

Idea of what one should get



Methodology

Monte-Carlo simulation to compute  $m_x = \mathbb{E}(X)$

\* specification of a confidence interval  $\alpha \in (0, 1)$

\* simulation of a  $M$  sample, computation of empirical mean  $\bar{\varphi}_M$  and variance  $\bar{V}_M$

\* computation of the resulting confidence interval

In at confidence level  $\alpha$