Introduction to multiparameter models - part 1

Data analytics

What do we mean by multiple parameters? And what to do with them?

- Multiple unknown or unobservable quantities
- Usually we are interested only in few
- Advantage of Bayesian statistics marginalization

We can average over 'nuisance parameters'

- Split of vector of parameters $\theta = (\theta_1, \theta_2)$
- For example, in normal distribution

$$y \mid \mu, \sigma^2 \propto N(\mu, \sigma^2)$$

we are often interested only in mean

Marginalization

The formal statement

Let us consider joint posterior density

$$p(\theta_1, \theta_2 | y) \propto p(y | \theta_1, \theta_2) p(\theta_1, \theta_2)$$

• We can average over θ_2

$$p(\theta_1 | y) = \int p(\theta_1, \theta_2 | y) d\theta_2$$
$$= \int p(\theta_1 | \theta_2, y) p(\theta_2 | y) d\theta_2$$

What is marginal distribution?

- Posterior distribution of interest, $p(\theta_1 | y)$ is a mixture of the conditional posterior distributions given the nuisance parameter, θ_2 , where $p(\theta_2 | y)$ is a weighting function for the different possible values of θ_2 .
- Again compromise between data and prior knowledge
- θ_2 can be discrete (multiple sub models)
- Usually we do not compute marginal distributions directly

Normal data with a non-informative prior distribution

- we consider a vector y of n independent observations from a univariate normal distribution, $N(\mu, \sigma^2)$
- We assume, that prior distributions are independent and uniform on $(\mu, \log \sigma)$
- This can be generally realized by

$$p(\mu, \sigma^2) \propto \frac{1}{\sigma^2}$$

Joint posterior distribution is proportional to likelihood multiplied by $1/\sigma^2$

$$p(\mu, \sigma^{2} | y) \propto \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^{2}} \sum_{i=1}^{n} (y_{i} - \mu)^{2}\right)$$

$$= \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^{2}} \left[\sum_{i=1}^{n} (y_{i} - \bar{y})^{2} + n(\bar{y} - \mu)^{2}\right]\right)$$

$$= \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^{2}} \left[(n-1)s^{2} + n(\bar{y} - \mu)^{2}\right]\right)$$

$$\bar{y} = \frac{1}{n} \sum_{i=1}^{n} y_i, \ s^2 = \frac{1}{n-1} \sum_{i=1}^{n} (y_i - \bar{y})^2$$

We can factorize the joint posterior distribution

• The conditional posterior distribution of μ is

$$\mu \mid \sigma^2, y \sim N(\bar{y}, \sigma^2, n)$$

• The marginal posterior of σ^2

$$p(\sigma^2 \mid y) \propto \int \sigma^{-n-2} \exp\left(-\frac{1}{2\sigma^2} \left[(n-1)s^2 + n(\bar{y} - \mu)^2 \right] \right) d\mu$$

Marginal posterior for σ^2 is a normal integral

$$p(\sigma^2 | y) \propto \sigma^{-n-2} \exp\left(-1\frac{1}{2\sigma^2}(n-1)s^2\right) \sqrt{2\pi\sigma^2/n}$$
$$\propto \sigma^{-(n+1)/2} \exp\left(-\frac{(n-1)s^2}{2\sigma^2}\right)$$

Which can be recognized as a form of scaled inverse chi-square distribution

$$\sigma^2 | y \sim \text{Inv} - \chi^2(n - 1, s^2)$$

Factorization allows for easy sampling

Joint posterior has a form

$$p(\mu, \sigma^2 | y) = p(\mu | \sigma^2, y)p(\sigma^2 | y)$$

- It is easy to draw samples from the joint posterior distribution:
 - first draw σ^2 from $p(\sigma^2 | y)$,
 - then draw μ from $p(\mu \mid \sigma^2, y)$
- Through repeated simulation we can marginalize over σ^2

This case is actually analytically solvable A rarity!

$$p(\mu \mid y) = \int_0^\infty p(\mu, \sigma^2 \mid y) \mathrm{d}\sigma^2$$
 Integration by substitution $z = \frac{A}{2\sigma^2}$, where $A = (n-1)s^2 + n(\mu - \bar{y})^2$
$$p(\mu \mid y) \propto A^{-n/2} \int_0^\infty z^{(n-2)/2} \exp(-z) \mathrm{d}z$$

$$\propto \left[(n-1)s^2 + n(\mu - \bar{y})^2 \right]^{-n/2}$$

$$\propto \left[1 + \frac{n(\mu - \bar{y})^2}{(n-1)s^2} \right]^{-n/2}$$

$$= t_{n-1}(\bar{y}, s^2/n)$$

Posterior predictive distribution for a future observation

$$p(\tilde{y}|y) = \iint p(\tilde{y}|\mu, \sigma^2, y) p(\mu, \sigma^2|y) d\mu d\sigma^2$$

The first of the two factors in the integral is just the normal distribution for the future observation given the values of (μ, σ^2) , and does not depend on y at all. To draw from the posterior predictive distribution, first draw μ, σ^2 from their joint posterior distribution and then simulate

$$\tilde{y} \sim N(\mu, \sigma^2)$$

Newcomb's experiment