# Medical Cost Personal Dataset Analysis

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#### 1 Introduction

The goal of this project is to explore the Medical Insurance Dataset to predict accurately the medical costs based on some explanatory variables. We are interested to find out which variables affect the costs and We will apply linear regression model to select the best fit for our analysis.

# 2 Knowing the Data

The *Medical Cost Personal Dataset* has been obtained from Kaggle provided from "Machine Learning with R" by Brett Lantz.

This dataset contains 1338 observations and 7 variables including:

- age: age of individuals
- sex: gender; female or male
- bmi: Body mass index, a measure of body fat based on height and weight
- children: Number of children of individuals
- smoker: smoker or non-smoker
- region: the residential area of individuals in the US, northeast, southeast, southwest, northwest
- charges: Individual medical costs billed by health insurance

The response variable (dependent) that we are going to predict is "charges".

# 3 Loading libraries

```
library(readr)
library("dplyr")

## Warning: package 'dplyr' was built under R version 4.1.3

##

## Attaching package: 'dplyr'

## The following objects are masked from 'package:stats':

##

## filter, lag

## The following objects are masked from 'package:base':

##

## intersect, setdiff, setequal, union

library(ggplot2)

## Warning: package 'ggplot2' was built under R version 4.1.3
```

```
library(ggcorrplot)
## Warning: package 'ggcorrplot' was built under R version 4.1.3
library(leaps)
## Warning: package 'leaps' was built under R version 4.1.3
library(tinytex)
## Warning: package 'tinytex' was built under R version 4.1.3
library(broom)
## Warning: package 'broom' was built under R version 4.1.3
library(modelr)
## Warning: package 'modelr' was built under R version 4.1.3
## ## Attaching package: 'modelr'
## The following object is masked from 'package:broom':
## ## bootstrap
```

# 4 Importing Dataset

```
insurance <- read_csv("D:/_UniPD/Semster 2/Statistical Learning B/PROJECT/Dataset/insurance.csv")

## Rows: 1338 Columns: 7

## -- Column specification -------

## Delimiter: ","

## chr (3): sex, smoker, region

## dbl (4): age, bmi, children, charges

##

## i Use 'spec()' to retrieve the full column specification for this data.

## i Specify the column types or set 'show_col_types = FALSE' to quiet this message.

attach(insurance)</pre>
```

## 5 Knowing the Dataset

```
str(insurance)
  spec_tbl_df [1,338 x 7] (S3: spec_tbl_df/tbl_df/tbl/data.frame)
              : num [1:1338] 19 18 28 33 32 31 46 37 37 60 ...
              : chr [1:1338] "female" "male" "male" "male" ...
##
    $ sex
    $ bmi
              : num [1:1338] 27.9 33.8 33 22.7 28.9 ...
    $ children: num [1:1338] 0 1 3 0 0 0 1 3 2 0 ...
    $ smoker : chr [1:1338] "yes" "no" "no" "no" ...
##
    $ region : chr [1:1338] "southwest" "southeast" "southeast" "northwest" ...
    $ charges : num [1:1338] 16885 1726 4449 21984 3867 ...
##
    - attr(*, "spec")=
##
##
     .. cols(
##
          age = col_double(),
##
          sex = col_character(),
##
          bmi = col_double(),
     . .
          children = col_double(),
##
##
          smoker = col_character(),
          region = col_character(),
##
##
          charges = col_double()
     . .
##
     ..)
    - attr(*, "problems")=<externalptr>
```

According tostr() we find out that we are dealing with 3 categorical variables including sex, smoker, region and 4 numerical variables. We convert them to factors and take a look at summary(insurance) to explore data further.

```
# converting categorical variables to factors
insurance$sex <- as.factor(insurance$sex)
insurance$smoker <- as.factor(insurance$smoker)
insurance$region <- as.factor(insurance$region)</pre>
summary(insurance)
```

```
##
                                         bmi
                                                        children
                                                                      smoker
         age
                          sex
           :18.00
                     female:662
                                           :15.96
                                                            :0.000
    Min.
                                   Min.
                                                     Min.
                                                                      no:1064
##
    1st Qu.:27.00
                     male :676
                                   1st Qu.:26.30
                                                     1st Qu.:0.000
                                                                      yes: 274
    Median :39.00
                                                     Median :1.000
##
                                   Median :30.40
##
    Mean
            :39.21
                                           :30.66
                                                            :1.095
                                   Mean
                                                     Mean
                                                     3rd Qu.:2.000
##
    3rd Qu.:51.00
                                   3rd Qu.:34.69
##
    Max.
            :64.00
                                           :53.13
                                                            :5.000
                                   Max.
                                                     {\tt Max.}
##
          region
                         charges
##
    northeast:324
                             : 1122
##
    northwest:325
                     1st Qu.: 4740
##
    southeast:364
                     Median: 9382
##
    southwest:325
                     Mean
                             :13270
##
                     3rd Qu.:16640
##
                             :63770
                     Max.
```

Briefly, from the summary() it is evident that sex and region variables are evenly distributed. The age of individuals is between 18 and 64 years old. Furthermore, the number of non-smokers is approximately 4

times more than the number of smokers. The number of children is from 0 to 5. Also, the average charge is 13279 USD.

# 6 Data Cleaning

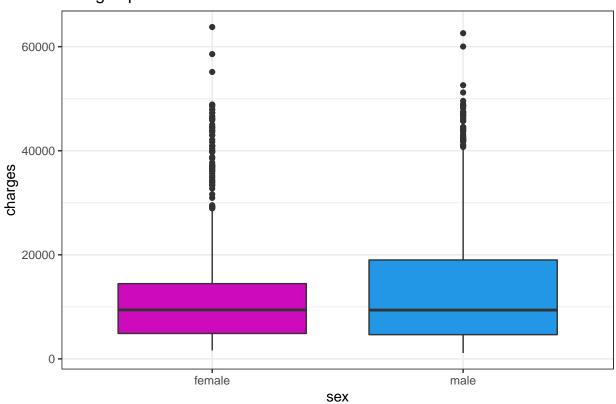
We do not have any NA's and there was only one duplicated row which is removed.

# 7 Exploratory Data Analysis (EDA)

#### 7.1 Boxplots

```
ggplot(data = insurance,aes(sex,charges)) + geom_boxplot(fill = c(6,4)) +
    theme_bw() + ggtitle("charges per sex")
```

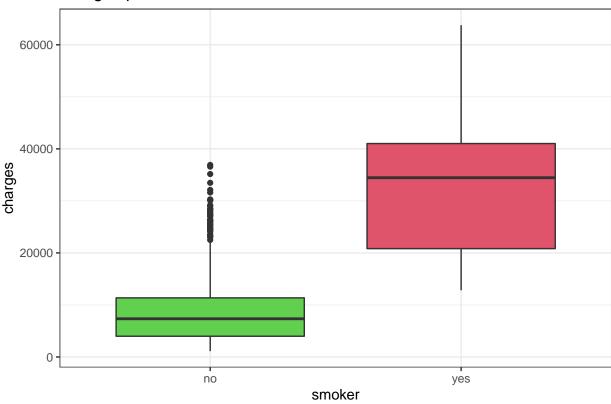
# charges per sex



*Insight*: the plot shows that there is not a notable difference in charges for male and females. So, we can conclude that the charges is not affected by gender.

```
ggplot(data = insurance,aes(smoker,charges)) + geom_boxplot(fill = c(3,2)) +
theme_bw() + ggtitle("charges per smoker")
```

# charges per smoker



*Insight*: smokers have higher medical charges compared with non-smokers; therefore, it seems to be a significant predictor for our analysis.

```
ggplot(data = insurance,aes(as.factor(children),charges)) + geom_boxplot(fill = c(2:7)) +
    theme_bw() + ggtitle("charges per children") +
    xlab("Num of Children")
```

# charges per children 40000 20000 0

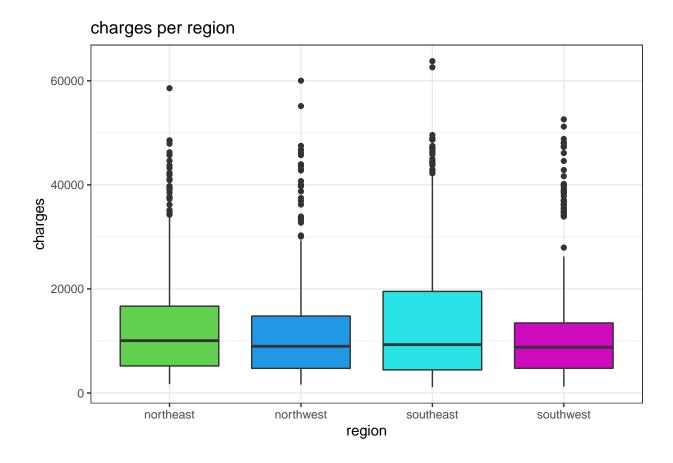
Insight: We normally expect to witness more medical charges for people with more children but, as we can see from this plot, people with 5 children, on average, have lower medical costs in comparison with other people with 0 to 4 children. It is surprising to know that groups with no child have higher insurance charges than other groups.

Num of Children

5

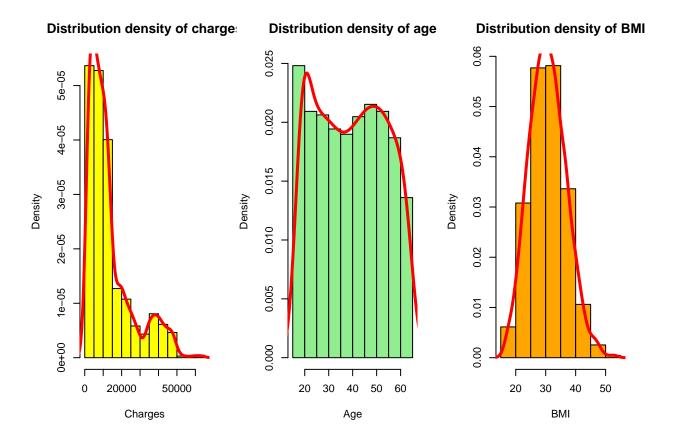
4

```
ggplot(data = insurance,aes(region,charges)) + geom_boxplot(fill = c(3:6)) +
theme_bw() + ggtitle("charges per region")
```



*Insight*: There is no impact on charges based on different regions. Thus, region seems to be an insignificant variable to our analysis.

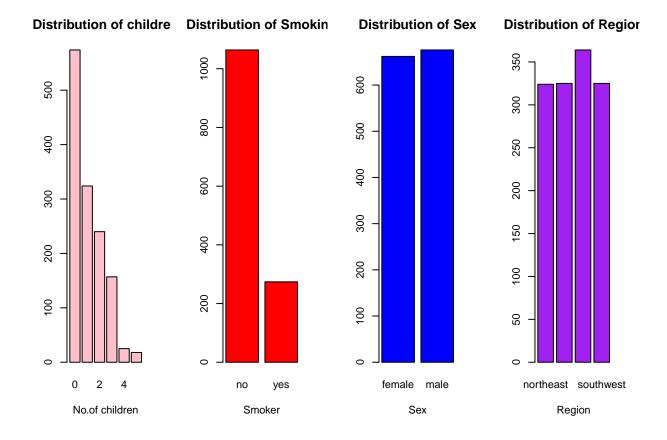
#### 7.2 Histograms and Density



#### par(mfrow = c(1,1))

In sight:

- Distribution of charges is right-skewed indicating the mean larger than median. The highest density of charges is between 1000 to 13000
- Ages range from 18 to 64 years old. observations are more frequent in ages between 18 to 20
- bmi has a normal distribution so we do not have many individuals who are extremely obese or underweight



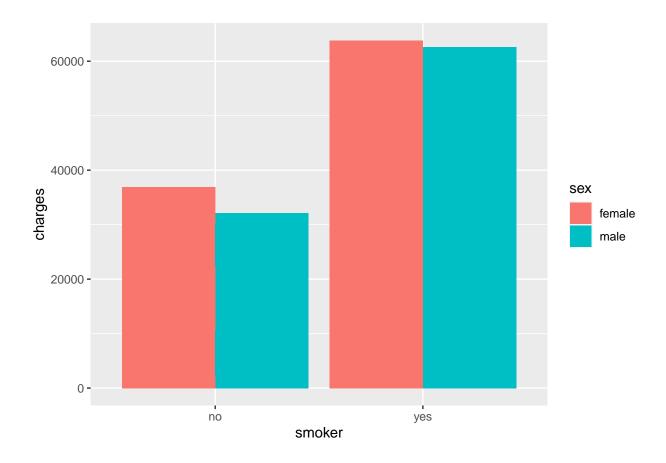
```
par(mfrow = c(1,1))
```

#### In sight:

- the number of individuals with no child is much more than the ones with 5 children.
- the number of non-smokers is almost 4 times the number of smokers
- population of females and males are almost equal
- 4 different regions have almost the same number of observations with a little more in SouthEast region

Moreover, it is useful to take a look at the charges for smoker and non-smoker based on sex:

```
ggplot(insurance, aes(x = smoker, y = charges, fill = sex)) +
geom_col(position = "dodge")
```

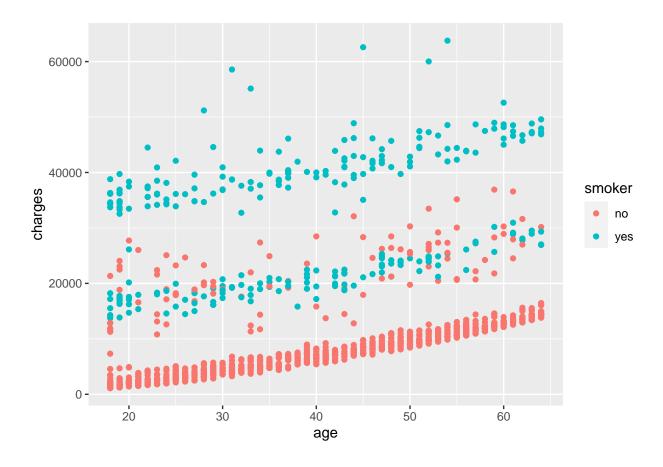


Insight: generally, gender does not affect the charges no matter they are smoker or non-smoker. However, females display a slightly higher charges.

# 7.3 scatter plots

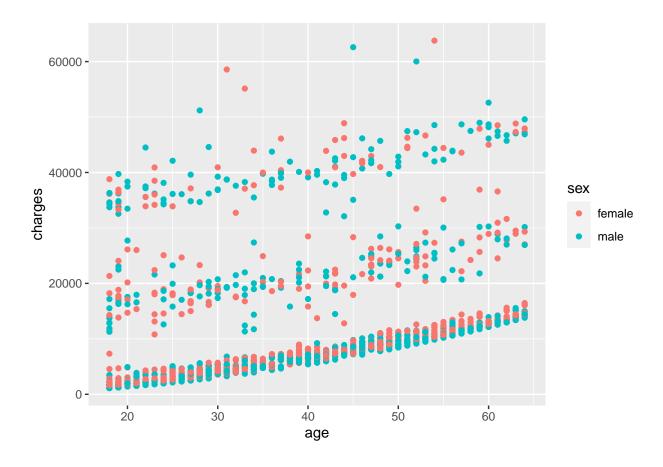
we are exploring the interaction between bmi, age and smoker and their effect on medical costs

```
ggplot(insurance, aes(age, charges, color = smoker)) +
  geom_point()
```



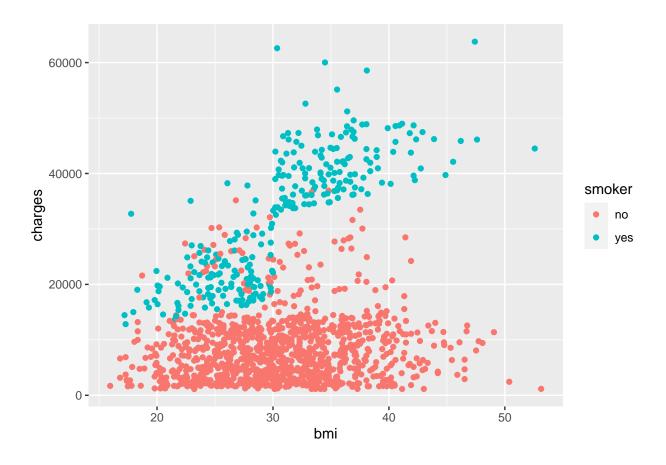
Insight: we can infer that older people spend more on medical expenses; more specifically, smokers' medical costs is massively higher than non-smokers as age increases.

```
ggplot(insurance, aes(age, charges, color = sex)) +
  geom_point()
```



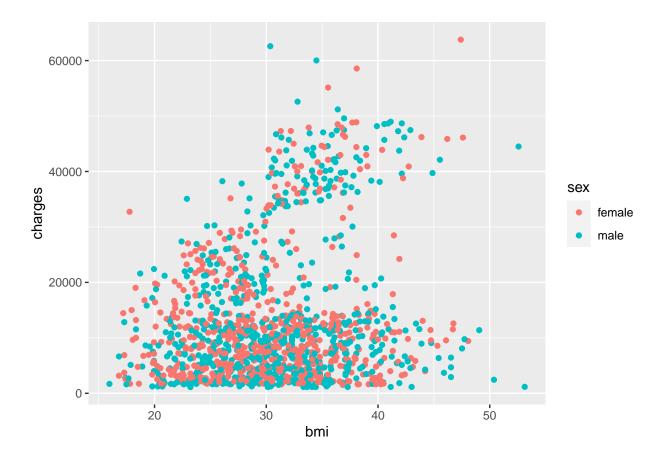
Insight: from this plot we can deduce that charges does not depend on sex because there is no specific pattern.

```
ggplot(insurance, aes(bmi, charges, color = smoker)) +
  geom_point()
```



Insight: medical cost for smokers steeply augments for people with higher bmi specially for those who are obese (bmi>30). In contrast, non-smokers' bmi does not significantly affect medical charges.

```
ggplot(insurance, aes(bmi, charges, color = sex)) +
geom_point()
```



Insight: there is no evident increase on charges considering bmi in association with sex.

#### 7.4 Correlation between variables

In order to check the correlation between variables we need to convert variables to numeric first.

```
ins <- insurance
ins$sex <- as.factor(ins$sex)
ins$smoker <- as.factor(ins$smoker)
ins$region <- as.factor(ins$region)
ins$children <- as.factor(ins$children)

ins$sex <- as.numeric(ins$sex)
ins$smoker <- as.numeric(ins$smoker)
ins$region <- as.numeric(ins$region)
ins$children <- as.numeric(ins$children)</pre>
```

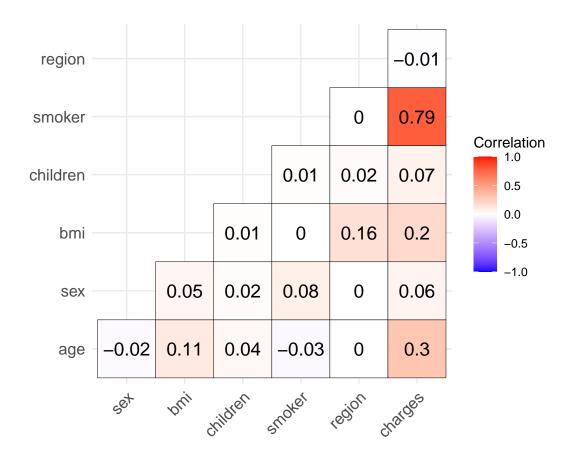
```
## age sex bmi children smoker

## age 1.00000000 -0.019813556 0.109343610 0.041536210 -0.025587118

## sex -0.019813556 1.000000000 0.046397060 0.017847542 0.076595850

## bmi 0.109343610 0.046397060 1.000000000 0.012754658 0.003746217
```

```
## children 0.041536210 0.017847542 0.012754658 1.000000000 0.007331222
## smoker
            -0.025587118
                         0.076595850 0.003746217 0.007331222 1.000000000
            0.001626251
                         0.004936187 0.157574500 0.016258480 -0.002358321
## region
## charges
            0.298308213
                         0.058044496 0.198400831 0.067389351 0.787234367
##
                 region
                              charges
            0.001626251
                         0.298308213
## age
## sex
            0.004936187
                         0.058044496
                         0.198400831
## bmi
            0.157574500
## children
            0.016258480
                         0.067389351
            -0.002358321
                         0.787234367
## smoker
## region
            1.00000000 -0.006546563
           -0.006546563 1.000000000
## charges
```



Based on the correlation plot above, the strongest correlation is between smoker and charges with the value of 0.79; age and bmi variables also tend to have a a correlation with charges but the relationship is not so strong. There is no collinearity between independent variables.

#### 7.5 introdicuing the new variable "obese"

Indeed, based on the scatter plot charges~bmi, it is noteworthy to introduce a new column named "obese"(bmi>30) to use in our analysis.

```
# Introducing a new column "obese"
insurance$obese <- ifelse(insurance$bmi >= 30 , "yes" , "no")
```



##

### 8 Regression Models

To begin our analysis, we apply a Full model where all covariates are included in the linear model. Then, we try to select the most significant variables through comparing the models.

First of all we split our data into Train and Test: 75% data is considered as the train-set and 25% as the test-set

```
# Spllitting data into Train and Test
set.seed(134)
partition <- floor(0.75*nrow(insurance))
train.numbers <- sample(seq_len(nrow(insurance)), partition, replace = FALSE)
train <- insurance[train.numbers, ]
test <- insurance[-train.numbers, ]</pre>
```

```
# MODEL 1 (Full Model)
model.1 <- lm(charges ~. , data = train)
summary(model.1)</pre>
```

```
## Call:
## lm(formula = charges ~ ., data = train)
## Residuals:
##
        Min
                  1Q
                       Median
                                    3Q
                                            Max
## -11719.8 -3615.0
                        -38.1
                                1619.5
                                        28194.5
##
## Coefficients:
##
                   Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                   -7728.68
                               1472.67
                                        -5.248 1.88e-07 ***
                     261.80
                                 13.75 19.042 < 2e-16 ***
## age
## sexmale
                      31.98
                                         0.083 0.93376
                                384.66
                     137.37
## bmi
                                 53.52
                                         2.567 0.01041 *
                                         2.696 0.00713 **
## children
                     425.47
                                157.81
## smokeryes
                   23626.39
                                474.60 49.781
                                                < 2e-16 ***
                                        -0.218 0.82735
## regionnorthwest
                    -119.92
                                549.69
## regionsoutheast
                    -647.77
                                556.63
                                        -1.164
                                                0.24481
## regionsouthwest
                    -966.34
                                553.09
                                        -1.747 0.08092
## obeseyes
                                        4.988 7.21e-07 ***
                    3164.91
                                634.53
## ---
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
```

```
##
## Residual standard error: 6053 on 992 degrees of freedom
## Multiple R-squared: 0.7536, Adjusted R-squared: 0.7513
## F-statistic:
                     337 on 9 and 992 DF, p-value: < 2.2e-16
BIC(model.1)
## [1] 20360.81
par(mfrow = c(2,2))
plot(model.1)
                                                     Standardized residuals
                                                                          Normal Q-Q
                 Residuals vs Fitted
      -10000 30000
Residuals
             0
                  10000
                                  30000
                                                                      -2
                                                                                                3
                      Fitted values
                                                                       Theoretical Quantiles
/IStandardized residuals
                                                     Standardized residuals
                    Scale-Location
                                                                    Residuals vs Leverage
      0.0 1.0 2.0
                                                              0.000
             0
                  10000
                                                                         0.010
                                  30000
                                                                                   0.020
                                                                                              0.030
                      Fitted values
                                                                             Leverage
par(mfrow = c(1,1))
# Metrics for model 1
glance(model.1) %>%
  dplyr::select(adj.r.squared, sigma, AIC, BIC, p.value)
## # A tibble: 1 x 5
      adj.r.squared sigma
                                 AIC
                                         {\tt BIC}
                                                p.value
               <dbl> <dbl> <dbl> <dbl> <dbl>
## 1
               0.751 6053. 20307. 20361. 1.74e-294
```

```
# RMSE and MAE
data.frame(
    R2 = rsquare(model.1, data = insurance),
    RMSE = rmse(model.1, data = insurance),
    MAE = mae(model.1, data = insurance)
)
```

```
## R2 RMSE MAE
## 1 0.7553929 5987.339 4283.494
```

#### Interpretation:

The result of high value of F-statistic: 337 and a very low P-value: < 2.2e-16 shows that the null hypothesis is rejected therefore we can conclude that there is a potential relationship between covariates and the response variable. significant covariates, based on their small P-values, are smoker, age, and obese respectively. Furthermore, two independent variables of sex and region are not significant based on their P-value which is higher than %5. Adjusted R-squared is about 0.75 which shows a pretty decent fit for our model. It means that 75% of variability of data can be explained by this model.

According to smoker's coefficient, we expect the charges of smokers to be, on average, 23626.39 US Dollars more than non-smokers.

Diagnostic plots depict that our model is not predicting data very well because of existence of outliers. the residual vs. fitted values exhibits a U-shaped pattern which provides an evidence of non-linearity in the data; therefor, we try to transform the response variable to tackle this problem and compare the behavior of the model.

#### 8.1 Transformation of the response variable

#### 8.1.1 log transformation

```
# MODEL 2 (Full Model_Log Transformation)
model.2 <- lm(log(charges) ~. , data = train)
summary(model.2)</pre>
```

```
##
## Call:
## lm(formula = log(charges) ~ ., data = train)
##
## Residuals:
##
                  1Q
                       Median
                                     3Q
                                             Max
## -0.95621 -0.20583 -0.05165 0.07547
                                        2.11486
##
## Coefficients:
                    Estimate Std. Error t value Pr(>|t|)
##
                    7.123537
                               0.106688
                                        66.770 < 2e-16 ***
## (Intercept)
                    0.034250
                                          34.387
## age
                               0.000996
                                                 < 2e-16 ***
## sexmale
                   -0.056611
                               0.027867
                                          -2.031 0.042472 *
## bmi
                    0.009419
                                           2.429 0.015308 *
                               0.003877
## children
                    0.100676
                               0.011432
                                           8.806
                                                 < 2e-16 ***
                               0.034383 44.798 < 2e-16 ***
## smokeryes
                    1.540279
## regionnorthwest -0.059152
                               0.039822
                                         -1.485 0.137759
```

```
## regionsoutheast -0.143690
                                 0.040325 -3.563 0.000384 ***
## regionsouthwest -0.148147
                                 0.040069 -3.697 0.000230 ***
                                 0.045969
                                            1.603 0.109152
## obeseyes
                     0.073709
##
                    0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
## Signif. codes:
##
## Residual standard error: 0.4385 on 992 degrees of freedom
## Multiple R-squared: 0.7703, Adjusted R-squared: 0.7682
## F-statistic: 369.7 on 9 and 992 DF, p-value: < 2.2e-16
BIC(model.2)
## [1] 1257.321
par(mfrow = c(2,2))
plot(model.2)
                                                Standardized residuals
                Residuals vs Fitted
                                                                   Normal Q-Q
Residuals
     1.0
     -1.0
           8.0
                    9.0
                            10.0
                                    11.0
                                                                                   2
                                                                                       3
                                                                         0
                                                                Theoretical Quantiles
                    Fitted values
                                                              Residuals vs Leverage
                  Scale-Location
```

```
Scale-Location

Scale-Location

Scale-Location

Standardized residuals vs Leverage

Residuals vs Leverage

0.000 0.010 0.020 0.030

Fitted values

Leverage
```

```
par(mfrow = c(1,1))

# Metrics for model 2
glance(model.2) %>%
  dplyr::select(adj.r.squared, sigma, AIC, BIC, p.value)
```

## # A tibble: 1 x 5

```
##
     adj.r.squared sigma
                           AIC
                                        p.value
                                  BIC
##
             <dbl> <dbl> <dbl> <dbl> <
                                           <dbl>
## 1
             0.768 0.438 1203. 1257. 1.24e-309
data.frame(
  R2 = rsquare(model.2, data = insurance),
 RMSE = rmse(model.2, data = insurance),
 MAE = mae(model.2, data = insurance)
)
##
            R2
                    RMSE
                                MAE
## 1 0.7680326 0.4423151 0.2825195
```

By log transformation of the response variable the behavior of the model improved. We witnessed increased value of Adjusted R-squared to 0.77 and F-Statistic to 369.7 and BIC from 20360.81 to 1257.321; so this model appears to be a much better fit but there is still some problem in diagnostic plots.

#### 8.1.2 Check for another transformation

```
# MODEL 3 (Full Model_ Transformation (charges)^(0.5))
model.3 <- lm((charges)^(1/2) ~. , data = train)
summary(model.3)
##
##</pre>
```

```
## Call:
## lm(formula = (charges)^(1/2) ~ ., data = train)
## Residuals:
##
      Min
                1Q
                   Median
                                3Q
                                       Max
  -39.530 -12.134
                   -4.334
##
                             3.812 102.987
##
## Coefficients:
##
                   Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                   12.14872
                               5.41304
                                         2.244 0.025030 *
## age
                    1.40113
                               0.05054 27.726 < 2e-16 ***
## sexmale
                   -1.07087
                               1.41388 -0.757 0.448990
## bmi
                                        2.468 0.013760 *
                    0.48549
                               0.19672
## children
                    3.17899
                               0.58005
                                        5.481 5.38e-08 ***
## smokeryes
                               1.74448 51.598 < 2e-16 ***
                   90.01167
## regionnorthwest -1.66916
                               2.02048
                                        -0.826 0.408934
## regionsoutheast -4.63597
                               2.04599
                                        -2.266 0.023674 *
## regionsouthwest -5.50311
                               2.03298
                                        -2.707 0.006908 **
## obeseyes
                                        3.744 0.000191 ***
                    8.73266
                               2.33233
## ---
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
## Residual standard error: 22.25 on 992 degrees of freedom
## Multiple R-squared: 0.7834, Adjusted R-squared: 0.7814
## F-statistic: 398.7 on 9 and 992 DF, p-value: < 2.2e-16
```

```
BIC(model.3)
## [1] 9126.347
par(mfrow = c(2,2))
plot(model.3)
                                                      Standardized residuals
                  Residuals vs Fitted
                                                                           Normal Q-Q
                                                           2
Residuals
      20
      -50
                                                           7
            50
                     100
                               150
                                        200
                                                                       -2
                                                                                  0
                                                                                            2
                                                                                                 3
                                                                  -3
                      Fitted values
                                                                        Theoretical Quantiles
||Standardized residuals
                                                      Standardized residuals
                    Scale-Location
                                                                     Residuals vs Leverage
      2.0
      0.0 1.0
                                                                                                   0
             50
                     100
                               150
                                        200
                                                               0.000
                                                                          0.010
                                                                                    0.020
                                                                                               0.030
                      Fitted values
                                                                              Leverage
par(mfrow = c(1,1))
# Metrics for model 3
glance(model.3) %>%
  dplyr::select(adj.r.squared, sigma, AIC, BIC, p.value)
## # A tibble: 1 x 5
      adj.r.squared sigma
                                       BIC p.value
                               AIC
               <dbl> <dbl> <dbl> <dbl> <
##
## 1
               0.781 22.2 9072. 9126.
data.frame(
  R2 = rsquare(model.3, data = insurance),
```

RMSE = rmse(model.3, data = insurance),
MAE = mae(model.3, data = insurance)

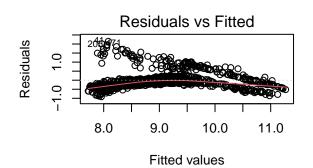
)

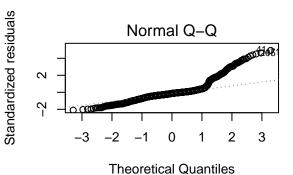
```
## R2 RMSE MAE
## 1 0.7816839 22.30563 14.96255
```

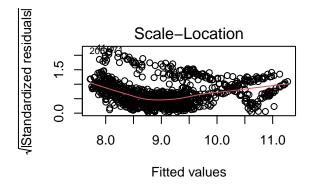
From the plot, it is obvious that transforming the response variable to (charges)^0.5 did not provide a noticeable improvement compared to log transformation, it even resulted in a much more BIC equal to 9126.347. Consequently, we keep the log transformation of the dependent variable.

Besides, we try a polynomial regression model to find out if it is going to be a better fit.

```
# MODEL 4 (Full Model Polynomial)
model.4 \leftarrow lm(log(charges) \sim . + poly(age , 2) , data = train)
summary(model.4)
##
## Call:
## lm(formula = log(charges) ~ . + poly(age, 2), data = train)
##
## Residuals:
##
        Min
                  1Q
                       Median
                                    3Q
                                            Max
## -0.91297 -0.20112 -0.05713 0.06626
                                        2.14759
## Coefficients: (1 not defined because of singularities)
##
                    Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                    7.140152
                               0.106718 66.907 < 2e-16 ***
                    0.034274
                               0.000994 34.480 < 2e-16 ***
## age
## sexmale
                   -0.056449
                               0.027809
                                         -2.030 0.042635 *
## bmi
                    0.008939
                               0.003875
                                          2.307 0.021276 *
## children
                    0.092383
                               0.011981
                                          7.711 3.04e-14 ***
## smokeryes
                               0.034314 44.916 < 2e-16 ***
                    1.541237
## regionnorthwest -0.057854
                               0.039744
                                         -1.456 0.145804
## regionsoutheast -0.140772
                               0.040262 -3.496 0.000492 ***
## regionsouthwest -0.145699
                               0.040000
                                         -3.642 0.000284 ***
## obeseyes
                    0.082214
                                          1.786 0.074368 .
                               0.046026
## poly(age, 2)1
                          NA
                                     NA
                                             NA
                                                      NA
## poly(age, 2)2
                   -1.045655
                               0.461383 -2.266 0.023645 *
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
##
## Residual standard error: 0.4376 on 991 degrees of freedom
## Multiple R-squared: 0.7715, Adjusted R-squared: 0.7692
## F-statistic: 334.6 on 10 and 991 DF, p-value: < 2.2e-16
BIC(model.4)
## [1] 1259.051
par(mfrow = c(2,2))
plot(model.4)
```



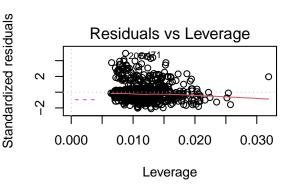




## misleading

## misleading

## misleading



```
par(mfrow = c(1,1))
# Metrics for model 4
glance(model.4) %>%
  dplyr::select(adj.r.squared, sigma, AIC, BIC, p.value)
## # A tibble: 1 x 5
     adj.r.squared sigma
                           AIC
                                  BIC
                                        p.value
##
             <dbl> <dbl> <dbl> <dbl> <
                                          <dbl>
## 1
             0.769 0.438 1200. 1259. 1.91e-309
data.frame(
 R2 = rsquare(model.4, data = insurance),
 RMSE = rmse(model.4, data = insurance),
 MAE = mae(model.4, data = insurance)
```

## Warning in predict.lm(model, data): prediction from a rank-deficient fit may be

## Warning in predict.lm(model, data): prediction from a rank-deficient fit may be

## Warning in predict.lm(model, data): prediction from a rank-deficient fit may be

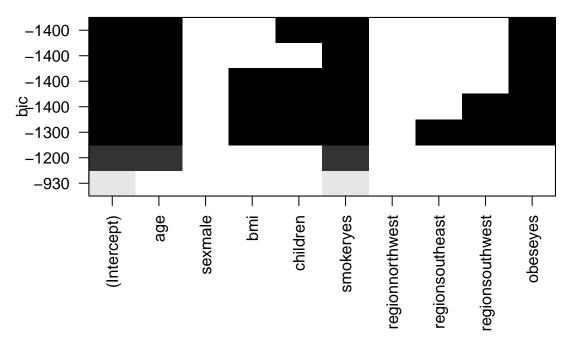
```
## R2 RMSE MAE
## 1 0.769781 0.440647 0.2821346
```

• To summarize: Applying polynomial regression model did not contribute to increasing Adjusted R-squared, F-statistic, or decreasing BIC to a great extent. As a result, we still select MODEL 2 (Full Model\_log Transformation) to address the non-linearity problem instead of choosing a polynomial model.

The next step is to select the most significant variables.

#### 8.2 Variable Selection Stepwise Backward Elimination

```
regfit.bwd <- regsubsets((charges) ~. , data = train, nvmax=7, method="backward")
bwd.summary <- summary(regfit.bwd)
plot(regfit.bwd, scale="bic")</pre>
```



```
#Looking at the Outmat Matrix
bwd.summary$outmat
```

```
## 3 (1) "*" "
                                  11 🕌 11
                                  "*"
    (1)"*"
                                  "*"
    (1) "*" "
## 6 (1) "*" "
                                  "*"
     (1)"*"""
                      "*" "*"
                                  11 * 11
## 7
                                                           11 * 11
##
           regionsouthwest obeseyes
    (1)""
## 1
     (1)""
                          11 11
## 2
## 3
     (1)""
                          11 * 11
## 4 (1)""
                          "*"
## 5 (1)""
                          "*"
## 6 (1) "*"
                          "*"
    (1)"*"
                          "*"
## 7
```

bwd.summary\$bic

```
## [1] -926.7341 -1242.5269 -1358.5605 -1359.2961 -1357.9305 -1353.4822 -1348.0765
```

```
which.min(bwd.summary$bic)
```

#### ## [1] 4

With this method we ended up with the model in row 4 which selected the following variables as the most significant ones: age, children, smoker, obese

```
# MODEL 5 Reduced Model Based on selected variables by Backward Elimination
model.5 <- lm(log(charges) ~ age + children + smoker + obese , data = train)
summary(model.5)</pre>
```

```
##
## Call:
## lm(formula = log(charges) ~ age + children + smoker + obese,
##
      data = train)
##
## Residuals:
                      Median
##
                 1Q
                                   3Q
## -0.98223 -0.20022 -0.04999 0.08519 2.04687
## Coefficients:
##
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 7.253519 0.045004 161.175 < 2e-16 ***
              0.034427
                         0.001002 34.345 < 2e-16 ***
## age
                                    8.747 < 2e-16 ***
## children
              0.100802
                         0.011524
## smokeryes
              1.530233
                         0.034522 44.327 < 2e-16 ***
## obeseyes
              0.140335
                         0.028178
                                    4.980 7.48e-07 ***
## ---
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
## Residual standard error: 0.443 on 997 degrees of freedom
## Multiple R-squared: 0.7644, Adjusted R-squared: 0.7635
## F-statistic: 808.7 on 4 and 997 DF, p-value: < 2.2e-16
```

```
BIC(model.5)
## [1] 1248.247
par(mfrow = c(2,2))
plot(model.5)
                                                       Standardized residuals
                  Residuals vs Fitted
                                                                            Normal Q-Q
Residuals
      -1.0
            8.0
                      9.0
                               10.0
                                         11.0
                                                                                    0
                                                                                              2
                                                                                                   3
                                                                         -2
                       Fitted values
                                                                         Theoretical Quantiles
||Standardized residuals
                                                       Standardized residuals
                    Scale-Location
                                                                      Residuals vs Leverage
      2.0
      0.0 1.0
            8.0
                      9.0
                               10.0
                                         11.0
                                                                0.000
                                                                          0.005
                                                                                     0.010
                                                                                               0.015
                       Fitted values
                                                                               Leverage
par(mfrow = c(1,1))
# Metrics for model 5
glance(model.5) %>%
  dplyr::select(adj.r.squared, sigma, AIC, BIC, p.value)
## # A tibble: 1 x 5
      adj.r.squared sigma
                                AIC
                                        {\tt BIC}
                                               p.value
               <dbl> <dbl> <dbl> <dbl> <
##
## 1
               0.763 0.443 1219. 1248. 4.02e-311
data.frame(
  R2 = rsquare(model.5, data = insurance),
```

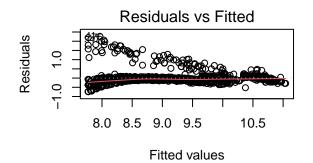
RMSE = rmse(model.5, data = insurance),
MAE = mae(model.5, data = insurance)

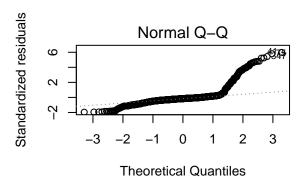
)

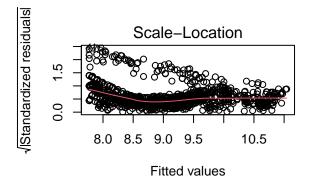
```
## R2 RMSE MAE
## 1 0.7623534 0.4476937 0.2845681
```

# 8.3 Including intraction effect of smoker\*obese

```
# MODEL 6 Reduced Model Based on Backward Elimination including
# interaction of smoker*age and smoker*obese
model.6 <- lm(log(charges) ~ age + children + smoker*age + smoker*obese , data = train)
summary(model.6)
##
## Call:
## lm(formula = log(charges) ~ age + children + smoker * age + smoker *
      obese, data = train)
##
## Residuals:
##
       Min
                1Q
                    Median
                                3Q
                                        Max
## -0.74072 -0.13731 -0.06067 0.01030 2.26494
## Coefficients:
                     Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                    7.0351928 0.0426020 165.138 <2e-16 ***
## age
                    0.1097584 0.0098862 11.102 <2e-16 ***
## children
## smokeryes
                    2.4658538 0.0892119 27.640
                                                 <2e-16 ***
## obeseyes
                   -0.0132487 0.0271611 -0.488
                                                0.626
## age:smokeryes
                    ## smokeryes:obeseyes 0.7062960 0.0594444 11.882
                                                <2e-16 ***
## ---
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
##
## Residual standard error: 0.3795 on 995 degrees of freedom
## Multiple R-squared: 0.8274, Adjusted R-squared: 0.8264
## F-statistic: 795 on 6 and 995 DF, p-value: < 2.2e-16
BIC(model.6)
## [1] 950.2432
par(mfrow = c(2,2))
plot(model.6)
```





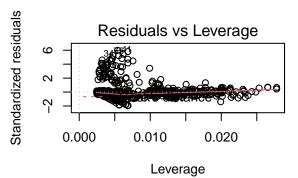


R2

## 1 0.8227052 0.3866913 0.2056036

RMSE

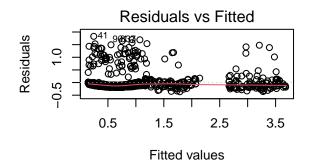
MAE

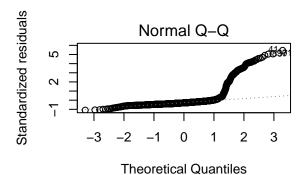


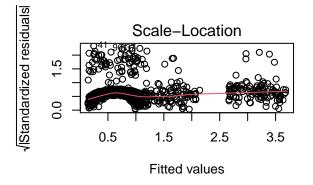
```
par(mfrow = c(1,1))
# Metrics for model 6
glance(model.6) %>%
  dplyr::select(adj.r.squared, sigma, AIC, BIC, p.value)
## # A tibble: 1 x 5
     adj.r.squared sigma
                           AIC
                                 BIC p.value
             <dbl> <dbl> <dbl> <dbl> <
## 1
             0.826 0.380 911.
                                950.
data.frame(
  R2 = rsquare(model.6, data = insurance),
  RMSE = rmse(model.6, data = insurance),
  MAE = mae(model.6, data = insurance)
)
```

#### 8.4 Concave transformation of the response variable with smoker\*obese

```
# MODEL 7 Concave transformation of the response variable with smoker*obese
model.7 <- lm(charges/mean(charges) ~ age + children + smoker*obese , data = train)</pre>
summary(model.7)
##
## Call:
## lm(formula = charges/mean(charges) ~ age + children + smoker *
      obese, data = train)
##
##
## Residuals:
               1Q Median
                                 3Q
## -0.36309 -0.14079 -0.09875 -0.04068 1.82897
## Coefficients:
                     Estimate Std. Error t value Pr(>|t|)
##
## (Intercept)
                   ## age
                    0.0432727 0.0087591
## children
                                         4.940 9.14e-07 ***
## smokeryes
                    1.0020317  0.0374485  26.758  < 2e-16 ***
## obeseyes
                     0.0152741 0.0240367
                                          0.635
                                                  0.525
## smokeryes:obeseyes 1.5028468 0.0524683 28.643 < 2e-16 ***
## ---
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
## Residual standard error: 0.3364 on 996 degrees of freedom
## Multiple R-squared: 0.8636, Adjusted R-squared: 0.8629
## F-statistic: 1261 on 5 and 996 DF, p-value: < 2.2e-16
BIC(model.7)
## [1] 702.6071
par(mfrow = c(2,2))
plot(model.7)
```







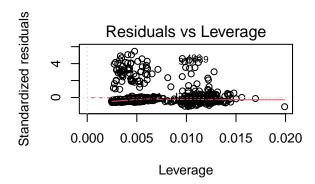
##

R2

## 1 0.8609659 0.3399417 0.1863047

RMSE

MAE



```
# Metrics for model 7
glance(model.7) %>%
  dplyr::select(adj.r.squared, sigma, AIC, BIC, p.value)
## # A tibble: 1 x 5
##
     adj.r.squared sigma
                           AIC
                                 BIC p.value
##
             <dbl> <dbl> <dbl> <dbl> <
                                       <dbl>
             0.863 0.336 668.
                               703.
## 1
data.frame(
  R2 = rsquare(model.7, data = insurance),
  RMSE = rmse(model.7, data = insurance),
  MAE = mae(model.7, data = insurance)
)
```

#### 9 Remarks

So, the model we selected as the optimal model is MODEL 7 with Adjusted R-squared: 0.8264 which has been significantly improved compared to other 5 models.it implies that 0.8264 variation of charges can be explained by our select independent variables in MODEL 6.

- The Residual Vs. Fitted Values plot shows a straight line indicating that the relationship is linear
- In the normal Q-Q plot residuals are not normally distributed
- The Scale-Location plot shows the assumption of equal variances is satisfied (homoscedasticity)
- The Residuals vs Leverage plot does not display any influential points within the cook's distance

We are going to use this model for our prediction of charges on test set as follows;

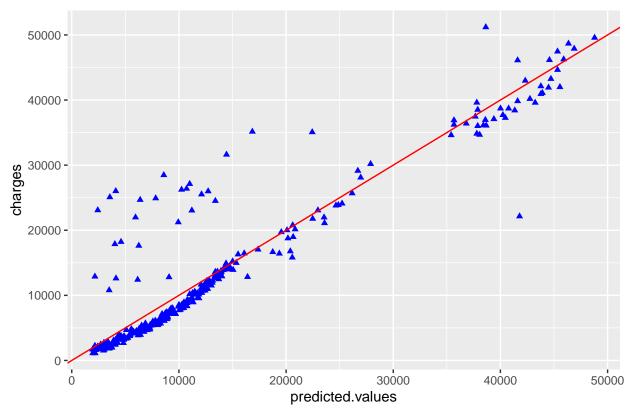
#### 10 Predicted Values vs Actual Values

```
test$predicted <- predict(model.7, newdata = test)

predicted.values <- (test$predicted)* mean(charges)

test %>%
    ggplot(aes(x = predicted.values , y = charges)) +
    geom_point(shape = "triangle", color = "blue") +
    geom_abline(color = "red") +
    ggtitle("Predicted Values vs. Actual Values")
```

#### Predicted Values vs. Actual Values



This plot shows that our model (MODEL 7) is a good fit for our test set and provides a relatively robust prediction even though it may struggle due to presence of outliers.

The best model for predicting the medical costs is with interaction terms. The best BIC and adjusted R-squared achieved are 702.60 and 0.86 respectively.

In our analysis, we found out that the most significant predictors of charges were the variables related to the age, children, and interaction of obese and smoker.

our model (MODEL 7) is a good fit for our test set and provides a relatively robust prediction even though it may struggle due to presence of outliers.

As we have seen, some observations had relatively high insurance charges, although they were young and non-smokers; they might have had an accident, surgery or something that affected their charges in such a manner.

As a consequence, insurance companies need to collect more accurate data regarding such observations and even include more variables such as their medical history, the hospitals they visited.

# 11 Bibliography

[1] MIRI CHOI. (2018). Medical Cost Personal Datasets. Retrieved [25/06/22] from https://www.kaggle.com/datasets/mirichoi0218/insurance