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Description hSDM is an R package for estimating parameters of hierarchical Bayesian species distribution models. Such models allow interpreting the observations (occurrence and abur of a species) as a result of several hierarchical processes including ecological processes (habitat suitability, spatial dependence and anthropogenic disturbance) and observation pro (species detectability). Hierarchical species distribution models are essential for accurately acterizing the environmental response of species, predicting their probability of occurrence assessing uncertainty in the model results.	ocesses / char-
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hierarchical Bayesian species distribution models

Description

hSDM is an R package for estimating parameters of hierarchical Bayesian species distribution models. Such models allow interpreting the observations (occurrence and abundance of a species) as a result of several hierarchical processes including ecological processes (habitat suitability, spatial dependence and anthropogenic disturbance) and observation processes (species detectability). Hierarchical species distribution models are essential for accurately characterizing the environmental response of species, predicting their probability of occurrence, and assessing uncertainty in the model results.

Details

Package: hSDM Type: Package Version: 1.2

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hSDM.binomial

The hSDM. binomial function performs a Binomial logistic regression model in a Bayesian framework.

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Description

The hSDM.binomial function calls a Gibbs sampler written in C code which uses a Metropolis algorithm to estimate the conditional posterior distribution of model's parameters.

Usage

```
hSDM.binomial(presences, trials, suitability, data, burnin = 5000, mcmc = 10000, thin = 10, beta.start, mubeta = 0, Vbeta = 1e+06, seed = 1234, verbose = 1)
```

Arguments

presences A vector indicating the number of successes (or presences) for each observation. trials A vector indicating the number of trials for each observation. t_n should be superior or equal to y_n , the number of successes for observation n. If $t_n = 0$, then $y_n = 0$. suitability A one-sided formula of the form '~x1+...+xp' with p terms specifying the explicative variables for the suitability process of the model. A data frame containing the model's explicative variables. data burnin The number of burnin iterations for the sampler. The number of Gibbs iterations for the sampler. Total number of Gibbs iterations mcmc is equal to burnin+mcmc. burnin+mcmc must be divisible by 10 and superior or equal to 100 so that the progress bar can be displayed. thin The thinning interval used in the simulation. The number of mcmc iterations must be divisible by this value. Starting values for beta parameters of the suitability process. If beta.start beta.start takes a scalar value, then that value will serve for all of the betas. mubeta Means of the priors for the β parameters of the suitability process. mubeta must be either a scalar or a p-length vector. If mubeta takes a scalar value, then that value will serve as the prior mean for all of the betas. The default value is set to 0 for an uninformative prior. Vbeta Variances of the Normal priors for the β parameters of the suitability process. Vbeta must be either a scalar or a p-length vector. If Vbeta takes a scalar value, then that value will serve as the prior variance for all of the betas. The default variance is large and set to 1.0E6 for an uninformative flat prior.

Value

seed

verbose

mcmc An mcmc object that contains the posterior sample. This object can be summa-

The seed for the random number generator. Default to 1234.

(in %) reached by the Gibbs sampler.

rized by functions provided by the coda package. The posterior sample of the

A switch (0,1) which determines whether or not the progress of the sampler is

printed to the screen. Default is 1: a progress bar is printed, indicating the step

deviance D, with $D = -2\log(\prod_i P(y_i, n_i|\beta))$, is also provided.

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prob.p.pred Predictive posterior mean of the probability associated to the suitability process for each spatial cell.

Author(s)

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References

Latimer, A. M.; Wu, S. S.; Gelfand, A. E. and Silander, J. A. (2006) Building statistical models to analyze species distributions. *Ecological Applications*, 16, 33-50.

Gelfand, A. E.; Schmidt, A. M.; Wu, S.; Silander, J. A.; Latimer, A. and Rebelo, A. G. (2005) Modelling species diversity through species level hierarchical modelling. *Applied Statistics*, 54, 1-20.

See Also

```
plot.mcmc, summary.mcmc
```

```
## Not run:
# hSDM.binomial()
# Example with simulated data
#=======
#== Preambule
library(hSDM)
#========
#== Data simulation
# Set seed for repeatability
set.seed(1234)
# Constants
trials <- rpois(nobs,5) # Number of trial associated to each observation</pre>
# Covariates for "suitability" process
X1 <- rnorm(n=nobs,0,1)</pre>
X2 <- rnorm(n=nobs,0,1)</pre>
X <- cbind(rep(1,nobs),X1,X2)</pre>
# Target parameters
beta.target <- matrix(c(0.2,0.1,0.1),ncol=1) # fixed effects
```

```
#== Simulating latent variables
# Suitability
logit.theta <- X %*% beta.target</pre>
theta <- inv.logit(logit.theta)</pre>
#== Simulating response variable
Y <- rbinom(n=nobs,size=trials,prob=theta)</pre>
#== Data-set
Data <- data.frame(Y,trials,X1,X2)</pre>
str(Data)
par(mfrow=c(2,1))
plot(Data$X1,logit.theta)
plot(Data$X2,logit.theta)
#== Statistical modelling
#== glm resolution to compare
mod.glm <- glm(cbind(Y,trials-Y)~X1+X2,family="binomial",data=Data)</pre>
summary(mod.glm)
beta.hat.glm <- coef(mod.glm)</pre>
#== hSDM
mod.hSDM.binomial <- hSDM.binomial(presences=Data$Y,</pre>
                                   trials=Data$trials,
                                   suitability=~X1+X2,
                                   data=Data, burnin=1000,
                                   mcmc=1000, thin=1,
                                   beta.start=beta.hat.glm,
                                   mubeta=0, Vbeta=1.0E6,
                                   seed=1234, verbose=1)
#=======
#== Outputs
summary(mod.hSDM.binomial$mcmc)
pdf(file="Posteriors_hSDM.binomial.pdf")
plot(mod.hSDM.binomial$mcmc)
dev.off()
summary(mod.hSDM.binomial$prob.p.pred)
## End(Not run)
```

hSDM.binomial.iCAR

The hSDM.binomial.iCAR function performs a Binomial logistic regression model in a hierarchical Bayesian framework. The suitability process includes a spatial correlation process. The spatial correlation is modelled using an intrinsic CAR model.

Description

The hSDM.binomial.iCAR function calls a Gibbs sampler written in C code which uses a Metropolis algorithm to estimate the conditional posterior distribution of model's parameters.

Usage

```
hSDM.binomial.iCAR(presences, trials, suitability, cells, n.neighbors, neighbors, data, burnin = 5000, mcmc = 10000, thin = 10, beta.start, Vrho.start, mubeta = 0, Vbeta = 1e+06, priorVrho = "1/Gamma", shape = 0.5, rate = 0.0005, Vrho.max = 1000, seed = 1234, verbose = 1, save.rho = 0)
```

Arguments

rguments	
presences	A vector indicating the number of successes (or presences) for each observation.
trials	A vector indicating the number of trials for each observation. t_n should be superior or equal to y_n , the number of successes for observation n . If $t_n = 0$, then $y_n = 0$.
suitability	A one-sided formula of the form '~x1++xp' with p terms specifying the explicative variables for the suitability process of the model.
cells	A vector indicating the spatial cell identifier (from 1 to total number of cell) for each observation. Several observations can occur in one spatial cell.
n.neighbors	A vector of integers indicating the number of neighbors (adjacent cells) of each spatial cell.
neighbors	A vector of integers indicating the neighbors (adjacent cells) of each spatial cell. Must be of the form c(neighbors of cell 1, neighbors of cell 2,, neighbors of the last cell). Length of the neighbors vector should be equal to sum(data\$num).
data	A data frame containing the model's explicative variables.
burnin	The number of burnin iterations for the sampler.
mcmc	The number of Gibbs iterations for the sampler. Total number of Gibbs iterations is equal to burnin+mcmc. burnin+mcmc must be divisible by 10 and superior or equal to 100 so that the progress bar can be displayed.
thin	The thinning interval used in the simulation. The number of mcmc iterations must be divisible by this value.
beta.start	Starting values for beta parameters.
Vrho.start	Positive scalar indicating the starting value for the variance of the spatial random effects.
mubeta	Means of the priors for the β parameters of the suitability process. mubeta must be either a scalar or a p-length vector. If mubeta takes a scalar value, then that value will serve as the prior mean for all of the betas. The default value is set to 0 for an uninformative prior.
Vbeta	Variances of the Normal priors for the β parameters of the suitability process. Vbeta must be either a scalar or a p-length vector. If Vbeta takes a scalar value, then that value will serve as the prior variance for all of the betas. The default variance is large and set to 1.0E6 for an uninformative flat prior.

priorVrho

Type of prior for the variance of the spatial random effects. Can be set to a fixed positive scalar, or to an inverse-gamma distribution ("1/Gamma") with parameters shape and rate, or to a uniform distribution ("Uniform") on the interval [0,Vrho.max]. Default to "1/Gamma".

Shape

The shape parameter for the Gamma prior on the precision of the spatial random

effects. Default value is shape=0.05 for uninformative prior.

The rate (1/scale) parameter for the Gamma prior on the precision of the spatial

random effects. Default value is rate=0.0005 for uninformative prior.

Vrho.max Upper bound for the uniform prior of the spatial random effect variance. Default

to 1000.

seed The seed for the random number generator. Default to 1234.

verbose A switch (0,1) which determines whether or not the progress of the sampler is

printed to the screen. Default is 1: a progress bar is printed, indicating the step

(in %) reached by the Gibbs sampler.

save.rho A switch (0,1) which determines whether or not the sampled values for rhos

are saved. Default is 0: the posterior mean is computed and returned in the rho.pred vector. Be careful, setting save.rho to 1 might require a large amount

of memory.

Value

rate

mcmc An mcmc object that contains the posterior sample. This object can be summa-

rized by functions provided by the coda package. The posterior sample of the

deviance D, with $D = -2 \log(\prod_i P(y_i, n_i | \beta, \rho_i))$, is also provided.

rho.pred If save.rho is set to 0 (default), rho.pred is the predictive posterior mean of

the spatial random effect associated to each spatial cell. If save.rho is set to 1, rho.pred is an mcmc object with sampled values for each spatial random effect

associated to each spatial cell.

prob.p.pred Predictive posterior mean of the probability associated to the suitability process

for each spatial cell.

Author(s)

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References

Latimer, A. M.; Wu, S. S.; Gelfand, A. E. and Silander, J. A. (2006) Building statistical models to analyze species distributions. *Ecological Applications*, 16, 33-50.

Gelfand, A. E.; Schmidt, A. M.; Wu, S.; Silander, J. A.; Latimer, A. and Rebelo, A. G. (2005) Modelling species diversity through species level hierarchical modelling. *Applied Statistics*, 54, 1-20.

See Also

plot.mcmc, summary.mcmc

```
## Not run:
# hSDM.binomial.iCAR()
# Example with simulated data
#=======
#== Preambule
library(mvtnorm)
library(lme4) # To compare with nonspatial random effects model
library(hSDM)
#========
#== Data simulation
# Set seed for repeatability
set.seed(1234)
# Constants
ncell <- 150 # Number of cells
nobs <- 10*ncell # Number of observation for the *binomial* random variable
trials <- rpois(nobs,5) # Number of trial associated to each observation
cell <- rep(c(1:ncell),each=nobs/ncell)</pre>
# Covariates for "suitability" process
X1 <- rnorm(n=nobs,0,1)</pre>
X2 <- rnorm(n=nobs,0,1)</pre>
X <- cbind(rep(1,nobs),X1,X2)</pre>
# Target parameters
beta.target <- matrix(c(0.2,0.1,0.1), ncol=1) # fixed effects
Vrho.target <- 10 # Spatial Variance
# Generate symmetric adjacency matrix, A
A <- matrix(0,ncell,ncell)
A[upper.tri(A,diag=F)] <- rbinom(ncell*(ncell-1)/2,1,.05)
A \leftarrow A+t(A)
n.neighbors <- apply(A,1,sum)</pre>
f.adjacent <- function (x) {</pre>
 which(x==1)
adj <- unlist(apply(A,1,f.adjacent))</pre>
# Spatial effects
d <- 1 # Spatial dependence parameter = 1 for intrinsic CAR
Q <- diag(n.neighbors)-d*A + diag(.0001,ncell) # Add small constant to make Q non-singular
covrho <- Vrho.target*solve(Q) # Covariance of rhos</pre>
rho <- c(rmvnorm(1,sigma=covrho)) # Spatial Random Effects</pre>
rho <- rho-mean(rho) # Centering rhos on zero</pre>
```

```
#== Simulating latent variables
# Suitability
logit.theta <- vector()</pre>
for (n in 1:nobs) {
  logit.theta[n] <- X[n,]%*%beta.target+rho[cell[n]]</pre>
theta <- inv.logit(logit.theta)</pre>
#== Simulating response variable
Y <- rbinom(nobs,trials,theta)
#== Data-set
Data <- data.frame(Y,trials,cell,X1,X2)</pre>
str(Data)
#== Statistical modelling
#== glm resolution to compare
mod.glm <- glm(cbind(Y,trials-Y)~X1+X2,data=Data,family="binomial")</pre>
summary(mod.glm)
#== glmm (mixed models) resolution to compare
#== cell random effects without spatial correlation
mod.glmm <- lmer(cbind(Y,trials-Y)~X1+X2+(1|cell),data=Data,family="binomial") # Non-spatial random effect fit</pre>
summary(mod.glmm)
beta.hat.glmm <- fixef(mod.glmm)</pre>
#== hSDM
mod.hSDM.binomial.iCAR <- hSDM.binomial.iCAR(presences=Data$Y,</pre>
                                              trials=Data$trials,
                                              suitability=~X1+X2,
                                              cells=Data$cell,
                                              n.neighbors=n.neighbors,
                                              neighbors=adj,
                                              data=Data, burnin=1000,
                                              mcmc=1000, thin=1,
                                              beta.start=beta.hat.glmm,
                                              Vrho.start=10,
                                              priorVrho="1/Gamma",
                                              #priorVrho="Uniform",
                                              #priorVrho=10,
                                              mubeta=0, Vbeta=1.0E6,
                                              shape=0.5, rate=0.0005,
                                              Vrho.max=1000,
                                              seed=1234, verbose=1, save.rho=0)
#=======
#== Outputs
summary(mod.hSDM.binomial.iCAR$mcmc)
pdf(file="Posteriors_hSDM.binomial.iCAR.pdf")
```

```
plot(mod.hSDM.binomial.iCAR$mcmc)
dev.off()
summary(mod.hSDM.binomial.iCAR$prob.p.pred)
summary(mod.hSDM.binomial.iCAR$rho.pred)
## End(Not run)
```

hSDM.hierarchical.binomial

The hSDM.hierarchical.binomial function can be used to model species distribution including different processes in a hierarchical Bayesian framework: (i) a Bernoulli suitability process (refering to environmental suitability) which takes into account the spatial dependence of the observations, (ii) an alteration process (refering to anthropogenic disturbances), and (iii) a Binomial observability process (refering to various ecological and methodological issues explaining the species presence). The model reduces to a ZIB (Zero-Inflated Binomial) model with spatial dependence if the alteration is set to zero.

Description

The hSDM.hierarchical.binomial function calls a Gibbs sampler written in C code which uses a Metropolis algorithm to estimate the conditional posterior distribution of hierarchical model's parameters.

Usage

```
hSDM.hierarchical.binomial(presences, trials, suitability, cells, n.neighbors, neighbors, alteration, observability, data, burnin = 5000, mcmc = 10000, thin = 10, beta.start, gamma.start, Vrho.start, mubeta = 0, Vbeta = 1e+06, mugamma = 0, Vgamma = 1e+06, priorVrho = "1/Gamma", shape = 0.5, rate = 0.0005, Vrho.max=1000, seed = 1234, verbose = 1, save.rho = 0)
```

Arguments

presences	A vector indicating the number of successes (or presences) for each observation.
trials	A vector indicating the number of trials for each observation. t_n should be superior or equal to y_n , the number of successes for observation n . If $t_n = 0$, then $y_n = 0$.
suitability	A one-sided formula of the form ' \sim x1++xp' with p terms specifying the explicative variables for the suitability process of the model.
cells	A vector indicating the spatial cell identifier (from 1 to total number of cell) for each observation. Several observations can occur in one spatial cell.

n.neighbors A vector of integers that indicates the number of neighbors (adjacent cells) of

each spatial cell.

neighbors A vector of integers indicating the neighbors (adjacent cells) of each spatial

cell. Must be of the form c(neighbors of cell 1, neighbors of cell 2, \dots , neighbors of the last cell). Length of the neighbors vector should be equal

to sum(data\$num).

alteration A vector indicating the proportion of area in the spatial cell which is transformed

(by anthropogenic activities for example) for each observation. Must be between

0 and 1.

observability A one-sided formula of the form '~x1+...+xq' with q terms specifying the ex-

plicative variables for the observability process of the model.

data A data frame containing the model's variables.

burnin The number of burnin iterations for the sampler.

mcmc The number of Gibbs iterations for the sampler. Total number of Gibbs iterations

is equal to burnin+mcmc. burnin+mcmc must be divisible by 10 and superior or

equal to 100 so that the progress bar can be displayed.

thin The thinning interval used in the simulation. The number of mcmc iterations

must be divisible by this value.

beta.start Starting values for beta parameters.
gamma.start Starting values for gamma parameters.

Vrho. start Positive scalar indicating the starting value for the variance of the spatial random

effects

mubeta Means of the priors for the β parameters of the suitability process. mubeta must

be either a scalar or a p-length vector. If mubeta takes a scalar value, then that value will serve as the prior mean for all of the betas. The default value is set to

0 for an uninformative prior.

Vbeta Variances of the Normal priors for the β parameters of the suitability process.

Vbeta must be either a scalar or a p-length vector. If Vbeta takes a scalar value, then that value will serve as the prior variance for all of the betas. The default

variance is large and set to 1.0E6 for an uninformative flat prior.

mugamma Means of the Normal priors for the γ parameters of the observability process.

mugamma must be either a scalar or a p-length vector. If mugamma takes a scalar value, then that value will serve as the prior mean for all of the gammas. The

default value is set to 0 for an uninformative prior.

Vgamma Variances of the Normal priors for the γ parameters of the observability process.

Vgamma must be either a scalar or a p-length vector. If Vgamma takes a scalar value, then that value will serve as the prior variance for all of the gammas. The

default variance is large and set to 1.0E6 for an uninformative flat prior.

priorVrho Type of prior for the variance of the spatial random effects. Can be set to a fixed

positive scalar, or to an inverse-gamma distribution ("1/Gamma") with parameters shape and rate, or to a uniform distribution ("Uniform") on the interval

[0,Vrho.max]. Default to "1/Gamma".

shape The shape parameter for the Gamma prior on the precision of the spatial random

effects. Default value is shape=0.05 for uninformative prior.

The rate (1/scale) parameter for the Gamma prior on the precision of the spatial

random effects. Default value is rate=0.0005 for uninformative prior.

Vrho.max Upper bound for the uniform prior of the spatial random effect variance. Default

to 1000.

seed The seed for the random number generator. Default to 1234.

verbose A switch (0,1) which determines whether or not the progress of the sampler is

printed to the screen. Default is 1: a progress bar is printed, indicating the step

(in %) reached by the Gibbs sampler.

save. rho A switch (0,1) which determines whether or not the sampled values for rhos

are saved. Default is 0: the posterior mean is computed and returned in the rho.pred vector. Be careful, setting save.rho to 1 might require a large amount

of memory.

Value

mcmc An mcmc object that contains the posterior sample. This object can be summa-

rized by functions provided by the coda package. The posterior sample of the deviance D, with $D = -2 \log(\prod_i P(y_i, n_i | u_i, \beta, \gamma, \rho_i))$, is also provided.

rho.pred If save.rho is set to 0 (default), rho.pred is the predictive posterior mean of

the spatial random effect associated to each spatial cell. If save.rho is set to 1, rho.pred is an meme object with sampled values for each spatial random effect

associated to each spatial cell.

prob.p.pred Predictive posterior mean of the probability associated to the suitability process

for each spatial cell.

prob.q.pred Predictive posterior mean of the probability associated to the observability pro-

cess for each spatial cell.

Author(s)

Ghislain Vieilledent <ghislain.vieilledent@cirad.fr>

References

Latimer, A. M.; Wu, S. S.; Gelfand, A. E. and Silander, J. A. (2006) Building statistical models to analyze species distributions. *Ecological Applications*, 16, 33-50.

Gelfand, A. E.; Schmidt, A. M.; Wu, S.; Silander, J. A.; Latimer, A. and Rebelo, A. G. (2005) Modelling species diversity through species level hierarchical modelling. *Applied Statistics*, 54, 1-20.

Gilks, W. R., Best, N. G. and Tan, K. K. C. (1995) Adaptive rejection Metropolis sampling. *Applied Statistics*, 44, 455-472.

See Also

plot.mcmc, summary.mcmc

```
## Not run:
# hSDM.hierarchical.binomial()
# Example with simulated data
#=======
#== Preambule
library(mvtnorm)
library(hSDM)
#=========
#== Data simulation
# Set seed for repeatability
set.seed(1234)
# Constants
ncell <- 150 # Number of cells
nobs <- 100*ncell # Number of observation for the *binomial* random variable
trials <- rpois(nobs,5) # Number of trial associated to each observation</pre>
cell <- rep(c(1:ncell),each=nobs/ncell)</pre>
# Covariates for "suitability" process
X1 <- rnorm(n=nobs,0,1)</pre>
X2 <- rnorm(n=nobs,0,1)</pre>
X <- cbind(rep(1,nobs),X1,X2)</pre>
# Alteration
U <- runif(n=nobs,min=0,max=1)</pre>
# Covariates for "observability" process
W1 <- rnorm(n=nobs,0,1)
W2 <- rnorm(n=nobs,0,1)
W <- cbind(rep(1,nobs),W1,W2)</pre>
# Target parameters
beta.target <- matrix(c(0.2,0.1,0.1), ncol=1) # fixed effects
gamma.target <- matrix(c(0.3,0.1,0.1), ncol=1) # fixed effects
Vrho.target <- 10 # Spatial Variance</pre>
# Generate symmetric adjacency matrix, A
A <- matrix(0,ncell,ncell)
A[upper.tri(A,diag=F)] <- rbinom(ncell*(ncell-1)/2,1,.05)
A \leftarrow A+t(A)
n.neighbors <- apply(A,1,sum)</pre>
f.adjacent <- function (x) {</pre>
 which(x==1)
}
```

```
adj <- unlist(apply(A,1,f.adjacent))</pre>
# Spatial effects
d <- 1 # Spatial dependence parameter = 1 for intrinsic CAR
Q <- diag(n.neighbors)-d*A + diag(.0001,ncell) # Add small constant to make Q non-singular
covrho <- Vrho.target*solve(Q) # Covariance of rhos</pre>
rho <- c(rmvnorm(1,sigma=covrho)) # Spatial Random Effects</pre>
rho <- rho-mean(rho) # Centering rhos on zero</pre>
#== Simulating latent variables
# Suitability
logit.theta.1 <- vector()</pre>
for (n in 1:nobs) {
  logit.theta.1[n] <- X[n,]%*%beta.target+rho[cell[n]]</pre>
theta.1 <- inv.logit(logit.theta.1)</pre>
y.1 <- rbinom(nobs,1,theta.1)</pre>
# Alteration
u <- rbinom(nobs,1,U)</pre>
# Observability
logit.theta.2 <- W%*%gamma.target</pre>
theta.2 <- inv.logit(logit.theta.2)</pre>
y.2 <- rbinom(nobs,trials,theta.2)</pre>
#== Simulating response variable
Y \leftarrow y.2*(1-u)*y.1
#== Data-set
Data <- data.frame(Y,trials,cell,X1,X2,W1,W2,U)</pre>
str(Data)
#== Statistical modelling with hSDM
model <- hSDM.hierarchical.binomial(presences=Data$Y,</pre>
                                      trials=Data$trials,
                                      suitability=~X1+X2,
                                      cells=Data$cell,
                                      n.neighbors=n.neighbors,
                                      neighbors=adj,
                                      alteration=Data$U,
                                      observability=~W1+W2,
                                      data=Data, burnin=500,
                                      mcmc=500, thin=1,
                                      beta.start=0,
                                      gamma.start=0,
                                      Vrho.start=1,
                                      priorVrho="1/Gamma",
                                      #priorVrho="Uniform",
                                      #priorVrho=10,
```

```
mubeta=0, Vbeta=1.0E6,
mugamma=0, Vgamma=1.0E6,
shape=0.5, rate=0.0005,
Vrho.max=1000,
seed=1234, verbose=1, save.rho=0)
```

```
#=======
#== Outputs
summary(model$mcmc)
pdf(file="Posteriors_hSDM.hierarchical.binomial.pdf")
plot(model$mcmc)
dev.off()
summary(model$prob.p.pred)
summary(model$rho.pred)
## End(Not run)
```

hSDM.hierarchical.poisson

The hSDM.hierarchical.poisson function can be used to model species distribution including different processes in a hierarchical Bayesian framework: (i) a Bernoulli suitability process (refering to environmental suitability) which takes into account the spatial dependence of the observations, (ii) an alteration process (refering to anthropogenic disturbances), and (iii) a Poisson observability process (refering to various ecological and methodological issues explaining the species abundance). The model reduces to a ZIP (Zero-Inflated Poisson) model with spatial dependence if the alteration is set to zero.

Description

The hSDM.hierarchical.poisson function calls a Gibbs sampler written in C code which uses a Metropolis algorithm to estimate the conditional posterior distribution of model's parameters.

Usage

```
hSDM.hierarchical.poisson(counts, visits, suitability, cells, n.neighbors, neighbors, alteration, observability, data, burnin = 5000, mcmc = 10000, thin = 10, beta.start, gamma.start, Vrho.start, mubeta = 0, Vbeta = 1e+06, mugamma = 0, Vgamma = 1e+06, priorVrho = "1/Gamma", shape = 0.5, rate = 0.0005, Vrho.max=1000, seed = 1234, verbose = 1, save.rho = 0)
```

Arguments

counts

A vector indicating the count (or abundance) for each observation.

visits A vector indicating if the observation point n has been visited (visits[n]=1) or

not (visits[n]=0). If visits[n]=0, then counts[n]=0.

suitability A one-sided formula of the form '~x1+...+xp' with p terms specifying the ex-

plicative variables for the suitability process of the model.

cells A vector indicating the spatial cell identifier (from 1 to total number of cell) for

each observation. Several observations can occur in one spatial cell.

n.neighbors A vector of integers that indicates the number of neighbors (adjacent cells) of

each spatial cell.

neighbors A vector of integers indicating the neighbors (adjacent cells) of each spatial

cell. Must be of the form c(neighbors of cell 1, neighbors of cell 2, \dots , neighbors of the last cell). Length of the neighbors vector should be equal

to sum(data\$num).

alteration A vector indicating the proportion of area in the spatial cell which is transformed

(by anthropogenic activities for example) for each observation. Must be between

0 and 1.

observability A one-sided formula of the form '~x1+...+xq' with q terms specifying the ex-

plicative variables for the observability process of the model.

data A data frame containing the model's explicative variables.

burnin The number of burnin iterations for the sampler.

mcmc The number of Gibbs iterations for the sampler. Total number of Gibbs iterations

is equal to burnin+mcmc. burnin+mcmc must be divisible by 10 and superior or

equal to 100 so that the progress bar can be displayed.

thin The thinning interval used in the simulation. The number of mcmc iterations

must be divisible by this value.

beta.start Starting values for beta parameters.

gamma.start Starting values for gamma parameters.

Vrho. start Positive scalar indicating the starting value for the variance of the spatial random

effects.

mubeta Means of the priors for the β parameters of the suitability process. mubeta must

be either a scalar or a p-length vector. If mubeta takes a scalar value, then that value will serve as the prior mean for all of the betas. The default value is set to

0 for an uninformative prior.

Vbeta Variances of the Normal priors for the β parameters of the suitability process.

Vbeta must be either a scalar or a p-length vector. If Vbeta takes a scalar value, then that value will serve as the prior variance for all of the betas. The default

variance is large and set to 1.0E6 for an uninformative flat prior.

mugamma Means of the Normal priors for the γ parameters of the observability process.

mugamma must be either a scalar or a p-length vector. If mugamma takes a scalar value, then that value will serve as the prior mean for all of the gammas. The

default value is set to 0 for an uninformative prior.

Vgamma Variances of the Normal priors for the γ parameters of the observability process.

Vgamma must be either a scalar or a p-length vector. If Vgamma takes a scalar value, then that value will serve as the prior variance for all of the gammas. The

default variance is large and set to 1.0E6 for an uninformative flat prior.

priorVrho	Type of prior for the variance of the spatial random effects. Can be set to a fixed positive scalar, or to an inverse-gamma distribution ("1/Gamma") with parameters shape and rate, or to a uniform distribution ("Uniform") on the interval [0,Vrho.max]. Default to "1/Gamma".
shape	The shape parameter for the Gamma prior on the precision of the spatial random effects. Default value is shape=0.05 for uninformative prior.
rate	The rate (1/scale) parameter for the Gamma prior on the precision of the spatial random effects. Default value is rate=0.0005 for uninformative prior.
Vrho.max	Upper bound for the uniform prior of the spatial random effect variance. Default to 1000.
seed	The seed for the random number generator. Default to 1234.
verbose	A switch (0,1) which determines whether or not the progress of the sampler is printed to the screen. Default is 1: a progress bar is printed, indicating the step (in %) reached by the Gibbs sampler.
save.rho	A switch (0,1) which determines whether or not the sampled values for rhos are saved. Default is 0: the posterior mean is computed and returned in the rho.pred vector. Be careful, setting save.rho to 1 might require a large amount of memory.

Value

mcmc	An meme object that contains the posterior sample. This object can be summarized by functions provided by the coda package. The posterior sample of the deviance D , with $D = -2\log(\prod_i P(y_i, n_i u_i, \beta, \gamma, \rho_i))$, is also provided.
rho.pred	If save.rho is set to 0 (default), rho.pred is the predictive posterior mean of the spatial random effect associated to each spatial cell. If save.rho is set to 1, rho.pred is an mcmc object with sampled values for each spatial random effect associated to each spatial cell.
prob.p.pred	Predictive posterior mean of the probability associated to the suitability process for each spatial cell.
prob.q.pred	Predictive posterior mean of the probability associated to the observability process for each spatial cell.

Author(s)

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References

Latimer, A. M.; Wu, S. S.; Gelfand, A. E. and Silander, J. A. (2006) Building statistical models to analyze species distributions. *Ecological Applications*, 16, 33-50.

Gelfand, A. E.; Schmidt, A. M.; Wu, S.; Silander, J. A.; Latimer, A. and Rebelo, A. G. (2005) Modelling species diversity through species level hierarchical modelling. *Applied Statistics*, 54, 1-20.

Gilks, W. R., Best, N. G. and Tan, K. K. C. (1995) Adaptive rejection Metropolis sampling. *Applied Statistics*, 44, 455-472.

See Also

```
plot.mcmc, summary.mcmc
```

```
## Not run:
# hSDM.hierarchical.poisson()
# Example with simulated data
#========
#== Preambule
library(mvtnorm)
library(hSDM)
#=========
#== Data simulation
# Set seed for repeatability
set.seed(1234)
# Constants
ncell <- 150 # Number of cells
nobs <- 100*ncell # Number of observation for the *binomial* random variable</pre>
cell <- rep(c(1:ncell),each=nobs/ncell)</pre>
# Covariates for "suitability" process
X1 <- rnorm(n=nobs,0,1)</pre>
X2 <- rnorm(n=nobs,0,1)</pre>
X <- cbind(rep(1,nobs),X1,X2)</pre>
# Alteration
U <- runif(n=nobs,min=0,max=1)</pre>
# Covariates for "observability" process
W1 <- rnorm(n=nobs,0,1)
W2 <- rnorm(n=nobs,0,1)
W <- cbind(rep(1,nobs),W1,W2)</pre>
# Target parameters
beta.target <- matrix(c(0.2,0.1,0.1),ncol=1) # fixed effects
gamma.target <- matrix(c(0.3,0.1,0.1),ncol=1) # fixed effects
Vrho.target <- 10 # Spatial Variance
# Generate symmetric adjacency matrix, A
A <- matrix(0,ncell,ncell)
A[upper.tri(A,diag=F)] <- rbinom(ncell*(ncell-1)/2,1,.05)
A \leftarrow A+t(A)
n.neighbors <- apply(A,1,sum)</pre>
```

```
f.adjacent <- function (x) {</pre>
  which(x==1)
adj <- unlist(apply(A,1,f.adjacent))</pre>
\# Spatial effects, non-intrinsic CAR with rho \sim 1
d <- 1 # Spatial dependence parameter = 1 for intrinsic CAR
Q \leftarrow diag(n.neighbors) - d*A + diag(.0001,ncell) # Add small constant to make Q non-singular
covrho <- Vrho.target*solve(Q) # Covariance of rhos</pre>
rho <- c(rmvnorm(1, sigma=covrho)) # Spatial Random Effects</pre>
rho <- rho-mean(rho) # Centering rhos on zero</pre>
#== Simulating latent variables
# Suitability
logit.theta.1 <- vector()</pre>
for (n in 1:nobs) {
  logit.theta.1[n] <- X[n,]%*%beta.target+rho[cell[n]]</pre>
}
theta.1 <- inv.logit(logit.theta.1)</pre>
y.1 <- rbinom(nobs,1,theta.1)</pre>
# Alteration
u <- rbinom(nobs,1,U)</pre>
# Observability
log.theta.2 <- W%*%gamma.target</pre>
theta.2 <- exp(log.theta.2)</pre>
y.2 <- rpois(nobs,theta.2)</pre>
#== Simulating response variable
Y <- y.2*(1-u)*y.1
V <- rep(1,nobs)</pre>
#== Data-set
Data <- data.frame(Y,V,cell,X1,X2,W1,W2,U)</pre>
str(Data)
#==============
#== Statistical modelling with hSDM
model <- hSDM.hierarchical.poisson(counts=Data$Y,</pre>
                                      visits=Data$V,
                                      suitability=~X1+X2,
                                      cells=Data$cell,
                                      n.neighbors=n.neighbors,
                                      neighbors=adj,
                                      alteration=Data$U,
                                      observability=~W1+W2,
                                      data=Data, burnin=500,
                                      mcmc=500, thin=1,
                                      beta.start=0,
                                      gamma.start=0,
```

```
Vrho.start=1,
                                   priorVrho="1/Gamma",
                                   #priorVrho="Uniform",
                                   #priorVrho=10,
                                   mubeta=0, Vbeta=1.0E6,
                                   mugamma=0, Vgamma=1.0E6,
                                    shape=0.5, rate=0.0005,
                                   Vrho.max=1000,
                                   seed=1234, verbose=1, save.rho=0)
#=======
#== Outputs
summary(model$mcmc)
pdf(file="Posteriors_hSDM.hierarchical.binomial.pdf")
plot(model$mcmc)
dev.off()
summary(model$prob.p.pred)
summary(model$rho.pred)
```

hSDM.Nmixture

End(Not run)

The hSDM. Nmixture function can be used to model species distribution including different processes in a hierarchical Bayesian framework: a Poisson suitability process (refering to environmental suitability) which takes into account the spatial dependence of the observations, and a Binomial observability process (refering to various ecological and methodological issues explaining the species detection).

Description

The hSDM.Nmixture function calls a Gibbs sampler written in C code which uses a Metropolis algorithm to estimate the conditional posterior distribution of hierarchical model's parameters.

Usage

```
# Starting values
beta.start,
gamma.start,
Vrho.start,
# Priors
mubeta = 0, Vbeta = 1.0E6,
mugamma = 0, Vgamma = 1.0E6,
priorVrho = "1/Gamma",
shape = 0.5, rate = 0.0005,
Vrho.max = 1000,
# Various
seed = 1234, verbose = 1,
save.rho = 0, save.p = 0, save.N = 0)
```

Arguments

A vector indicating the count (or abundance) for each observation. counts

observability A one-sided formula of the form $\sim w_1 + ... + w_q$ with q terms specifying the explicative variables for the observability process.

spatial.entity A vector indicating the spatial entity identifier (from one to the total number of entities) for each observation. Several observations can occur in one spatial entity. A spatial entity can be a raster cell for example.

data.observability

A data frame containing the model's variables for the observability process.

suitability A one-sided formula of the form $\sim x_1 + ... + x_p$ with p terms specifying the explicative variables for the suitability process.

data.suitability

A data frame containing the model's variables for the suitability process.

A vector of integers that indicates the number of neighbors (adjacent entities) of each spatial entity. length(n.neighbors) indicates the total number of spatial

entities.

neighbors A vector of integers indicating the neighbors (adjacent entities) of each spatial entity. Must be of the form c(neighbors of entity 1, neighbors of entity 2, ..., neighbors of the last entity). Length of the neighbors vector should be equal to

sum(n.neighbors).

suitability.pred

An optional data frame in which to look for variables with which to predict. If NULL, the observations are used.

spatial.entity.pred

An optional vector indicating the spatial entity identifier (from one to the total number of entities) for predictions. If NULL, the vector spatial.entity for observations is used.

burnin The number of burnin iterations for the sampler.

The number of Gibbs iterations for the sampler. Total number of Gibbs iterations mcmc is equal to burnin+mcmc. burnin+mcmc must be divisible by 10 and superior or

equal to 100 so that the progress bar can be displayed.

n.neighbors

thin The thinning interval used in the simulation. The number of mcmc iterations must be divisible by this value. Starting values for β parameters of the suitability process. This can either be a beta.start scalar or a p-length vector. gamma.start Starting values for β parameters of the observability process. This can either be a scalar or a q-length vector. Vrho.start Positive scalar indicating the starting value for the variance of the spatial random effects. mubeta Means of the priors for the β parameters of the suitability process. mubeta must be either a scalar or a p-length vector. If mubeta takes a scalar value, then that value will serve as the prior mean for all of the betas. The default value is set to 0 for an uninformative prior. Vbeta Variances of the Normal priors for the β parameters of the suitability process. Vbeta must be either a scalar or a p-length vector. If Vbeta takes a scalar value, then that value will serve as the prior variance for all of the betas. The default variance is large and set to 1.0E6 for an uninformative flat prior. mugamma Means of the Normal priors for the γ parameters of the observability process. mugamma must be either a scalar or a p-length vector. If mugamma takes a scalar value, then that value will serve as the prior mean for all of the gammas. The default value is set to 0 for an uninformative prior. Variances of the Normal priors for the γ parameters of the observability process. Vgamma Vgamma must be either a scalar or a p-length vector. If Vgamma takes a scalar value, then that value will serve as the prior variance for all of the gammas. The default variance is large and set to 1.0E6 for an uninformative flat prior. priorVrho Type of prior for the variance of the spatial random effects. Can be set to a fixed positive scalar, or to an inverse-gamma distribution ("1/Gamma") with parameters shape and rate, or to a uniform distribution ("Uniform") on the interval [0,Vrho.max]. Default set to "1/Gamma". shape The shape parameter for the Gamma prior on the precision of the spatial random effects. Default value is shape=0.05 for uninformative prior. The rate (1/scale) parameter for the Gamma prior on the precision of the spatial rate random effects. Default value is rate=0.0005 for uninformative prior. Vrho.max Upper bound for the uniform prior of the spatial random effect variance. Default set to 1000. seed The seed for the random number generator. Default set to 1234. verbose A switch (0,1) which determines whether or not the progress of the sampler is printed to the screen. Default is 1: a progress bar is printed, indicating the step (in %) reached by the Gibbs sampler. A switch (0,1) which determines whether or not the sampled values for rhos save.rho are saved. Default is 0: the posterior mean is computed and returned in the rho.pred vector. Be careful, setting save.rho to 1 might require a large amount of memory.

A switch (0,1) which determines whether or not the sampled values for predic-

tions are saved. Default is 0: the posterior mean is computed and returned in

save.p

the prob.p.pred vector. Be careful, setting save.p to 1 might require a large amount of memory.

save.N

A switch (0,1) which determines whether or not the sampled values for the latent count variable N for each observed cells are saved. Default is 0: the mean (rounded to the closest integer) is computed and returned in the N.pred vector. Be careful, setting save. N to 1 might require a large amount of memory.

Details

The model integrates two processes, an ecological process associated to the presence or absence of the species due to habitat suitability and an observation process that takes into account the fact that the probability of detection of the species is inferior to one. The ecological process includes an intrinsic conditional autoregressive model (iCAR) model for spatial autocorrelation between observations, assuming that the probability of presence of the species at one site depends on the probability of presence of the species on neighboring sites.

Ecological process:

$$N_i \sim \mathcal{P}oisson(\lambda_i)$$

$$log(\lambda_i) = X_i \beta + \rho_i$$

 ρ_i : spatial random effect

Spatial autocorrelation:

An intrinsic conditional autoregressive model (iCAR) is assumed:

$$\rho_i \sim \mathcal{N}ormal(\mu_i, V_{\rho}/n_i)$$

 μ_i : mean of $\rho_{i'}$ in the neighborhood of i.

 V_{ρ} : variance of the spatial random effects.

 n_i : number of neighbors for spatial entity i.

Observation process:

$$y_{it} \sim \mathcal{B}inomial(N_i, \delta_{it})$$

$$logit(\delta_{it}) = W_{it}\gamma$$

Value

mcmc An mcmc object that contains the posterior sample. This object can be summa-

rized by functions provided by the coda package. The posterior sample of the

deviance D, with $D = -2 \log(\prod_{i,t} P(y_{i,t}, N_i|...))$, is also provided.

 $\mbox{rho.pred}$ If save. rho is set to 0 (default), rho. pred is the predictive posterior mean of

the spatial random effect associated to each spatial entity. If save.rho is set to 1, rho.pred is an mcmc object with sampled values for each spatial random

effect associated to each spatial entity.

prob.p.pred If save.p is set to 0 (default), prob.p.pred is the predictive posterior mean

of the probability associated to the suitability process for each prediction. If save.p is set to 1, prob.p.pred is an mcmc object with sampled values of the

probability associated to the suitability process for each prediction.

N.pred If save.N is set to 0 (default), N.pred is the posterior mean (rounded to the closest integer) of the latent count variable N for each observed cell. If save.N is set to 1, N.pred is an mcmc object with sampled values of the latent count variable N for each observed cell.
 prob.p.latent Predictive posterior mean of the probability associated to the suitability process for each observation.

Predictive posterior mean of the probability associated to the observability process for each observation.

Author(s)

prob.q.latent

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References

Gelfand, A. E.; Schmidt, A. M.; Wu, S.; Silander, J. A.; Latimer, A. and Rebelo, A. G. (2005) Modelling species diversity through species level hierarchical modelling. *Applied Statistics*, 54, 1-20

Gilks, W. R., Best, N. G. and Tan, K. K. C. (1995) Adaptive rejection Metropolis sampling. *Applied Statistics*, 44, 455-472.

Latimer, A. M.; Wu, S. S.; Gelfand, A. E. and Silander, J. A. (2006) Building statistical models to analyze species distributions. *Ecological Applications*, 16, 33-50.

MacKenzie, D. I.; Nichols, J. D.; Lachman, G. B.; Droege, S.; Andrew Royle, J. and Langtimm, C. A. (2002) Estimating site occupancy rates when detection probabilities are less than one. *Ecology*, 83, 2248-2255.

See Also

```
plot.mcmc, summary.mcmc
```

```
#== Data simulation
# Set seed for repeatability
seed <- 1234
# Target parameters
beta.target <- matrix(c(0.2,0.5,0.5),ncol=1)
gamma.target <- matrix(c(0.2),ncol=1) #= This means that theta.2\sim0.5: 50% of chance to detect the species
Vrho.target <- 1 # Spatial Variance</pre>
# Landscape
Landscape <- raster(ncol=20,nrow=20,crs='+proj=utm +zone=1')</pre>
ncell <- ncell(Landscape)</pre>
# Neighbors
neighbors.mat <- adjacent(Landscape, cells=c(1:ncell), directions=8, pairs=TRUE, sorted=TRUE)</pre>
n.neighbors <- as.data.frame(table(as.factor(neighbors.mat[,1])))[,2]</pre>
adj <- neighbors.mat[,2]</pre>
# Generate symmetric adjacency matrix, A
A <- matrix(0,ncell,ncell)
index.start <- 1</pre>
for (i in 1:ncell) {
    index.end <- index.start+n.neighbors[i]-1</pre>
    A[i,adj[c(index.start:index.end)]] <- 1
    index.start <- index.end+1</pre>
}
# Spatial effects
d <- 1 # Spatial dependence parameter = 1 for intrinsic CAR
Q <- diag(n.neighbors)-d*A + diag(.0001,ncell) # Add small constant to make Q non-singular
covrho <- Vrho.target*solve(Q) # Covariance of rhos</pre>
set.seed(seed)
rho <- c(rmvnorm(1,sigma=covrho)) # Spatial Random Effects</pre>
rho <- rho-mean(rho) # Centering rhos on zero</pre>
# Visited cells
n.visited <- 150 # Compare with 400, 350 and 100 for example
set.seed(seed)
visited.cells <- sort(sample(1:ncell,n.visited,replace=FALSE)) # Draw visited cells at random
notvisited.cells <- c(1:ncell)[-visited.cells]</pre>
# Number of observations
nobs <- 300
# Cell vector
set.seed(seed)
cells <- c(visited.cells,sample(visited.cells,nobs-n.visited,replace=TRUE))</pre>
coords <- xyFromCell(Landscape,cells) # Get coordinates</pre>
# Covariates for "suitability" process
set.seed(seed)
X1 <- rnorm(n=ncell,\emptyset,1) # Here, we decided to define one habitat by spatial entity (raster cell)
```

```
set.seed(2*seed)
X2 <- rnorm(n=ncell,0,1)</pre>
X <- cbind(rep(1,ncell),X1,X2)</pre>
# Covariates for "observability" process
W <- cbind(rep(1,nobs))</pre>
#== Simulating latent variables
# Suitability
log.lambda <- vector()</pre>
log.lambda <- X%*%beta.target+rho</pre>
lambda <- exp(log.lambda)</pre>
set.seed(seed)
N <- rpois(ncell,lambda)
# Observability
logit.theta <- W%*%gamma.target</pre>
theta <- inv.logit(logit.theta)</pre>
#== Simulating response variable
set.seed(seed)
Y <- rbinom(nobs,N[cells],theta)</pre>
#== Data-sets
Data.obs <- data.frame(Y,cells,W1,W2)</pre>
Data.suit <- data.frame(cells=c(1:ncell),X1,X2)</pre>
Data.obs <- SpatialPointsDataFrame(coords=coords,data=Data.obs)</pre>
plot(Data.obs)
Data.pred <- Data.suit
#== N-mixture model
mod.hSDM.Nmixture <- hSDM.Nmixture(# Observations</pre>
                                    counts=Data.obs$Y, observability=~1,
                              spatial.entity=Data.obs$cells, data.observability=Data.obs,
                                    # Habitat
                                    suitability=~X1+X2, data.suitability=Data.suit,
                                    # Spatial structure
                                    n.neighbors=n.neighbors, neighbors=adj,
                                    # Predictions
                             suitability.pred=Data.pred, spatial.entity.pred=Data.pred$cells,
                                    # Chains
                                    burnin=10000, mcmc=10000, thin=10,
                                    # Starting values
                                    beta.start=0,
                                    gamma.start=0,
                                    Vrho.start=1,
                                    # Priors
                                    mubeta=0, Vbeta=1.0E6,
                                    mugamma=0, Vgamma=1.0E6,
                                    priorVrho="1/Gamma",
```

```
# shape=0.5, rate=0.0005,
                                     shape=0.1, rate=0.1,
                                     Vrho.max=1000,
                                     # Various
                                     seed=seed, verbose=1,
                                     save.rho=0, save.p=0, save.N=1)
#=======
#== Outputs
#= Parameter estimates
summary(mod.hSDM.Nmixture$mcmc)
#= MCMC and posteriors
pdf(file="Posteriors_hSDM.Nmixture.pdf")
plot(mod.hSDM.Nmixture$mcmc)
dev.off()
pdf(file="Posteriors.rho_hSDM.Nmixture.pdf")
plot(mod.hSDM.Nmixture$rho.pred)
dev.off()
pdf(file="Posteriors.N_hSDM.Nmixture.pdf")
plot(mod.hSDM.Nmixture$N.pred)
dev.off()
#= Check that Ns are correctly estimated
M <- as.matrix(mod.hSDM.Nmixture$N.pred)</pre>
N.est <- apply(M,2,mean)</pre>
Y.by.cell <- tapply(Data.obs$Y,Data.obs$cells,mean)</pre>
pdf(file="Check_N.pdf",width=10,height=5)
par(mfrow=c(1,2))
plot(Y.by.cell, N.est) ## More individuals are expected in reality (N > Y) due to detection process
abline(a=0,b=1,col="red")
plot(N[visited.cells], N.est) ## N are well estimated
abline(a=0,b=1,col="red")
cor(N[visited.cells], N.est) ## Close to 1
dev.off()
#= Summary plots
# rho
r.rho <- r.rho.pred <- r.visited <- Landscape</pre>
r.rho[] <- rho
rho.pred <- mod.hSDM.Nmixture$rho.pred # apply(mod.hSDM.Nmixture$rho.pred,2,mean)</pre>
r.rho.pred[] <- rho.pred</pre>
r.visited[] <- 0</pre>
r.visited[visited.cells] <- 1</pre>
# prob.p
r.prob.p <- Landscape</pre>
r.prob.p[] <- mod.hSDM.Nmixture$prob.p.pred</pre>
pdf(file="Summary_hSDM.Nmixture.pdf")
```

hSDM.poisson

hSDM.poisson

The hSDM.poisson function performs a Poisson regression in a Bayesian framework.

Description

The hSDM.poisson function calls a Gibbs sampler written in C code which uses a Metropolis algorithm to estimate the conditional posterior distribution of model's parameters.

Usage

```
hSDM.poisson(counts, visits, suitability, data, burnin = 5000, mcmc = 10000, thin = 10, beta.start, mubeta = 0, Vbeta = 1e+06, seed = 1234, verbose = 1)
```

Arguments

counts	A vector indicating the count (or abundance) for each observation.
visits	A vector indicating if the observation point n has been visited (visits[n]=1) or not (visits[n]=0). If visits[n]=0, then counts[n]=0.
suitability	A one-sided formula of the form '~x1++xp' with p terms specifying the explicative covariates for the suitability process of the model.
data	A data frame containing the model's explicative variables.
burnin	The number of burnin iterations for the sampler.
mcmc	The number of Gibbs iterations for the sampler. Total number of Gibbs iterations is equal to burnin+mcmc. burnin+mcmc must be divisible by 10 and superior or equal to 100 so that the progress bar can be displayed.

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thin The thinning interval used in the simulation. The number of meme iterations

must be divisible by this value.

beta.start Starting values for beta parameters of the suitability process. If beta.start

takes a scalar value, then that value will serve for all of the betas.

mubeta Means of the priors for the β parameters of the suitability process. mubeta must

be either a scalar or a p-length vector. If mubeta takes a scalar value, then that value will serve as the prior mean for all of the betas. The default value is set to

0 for an uninformative prior.

Vbeta Variances of the Normal priors for the β parameters of the suitability process.

Vbeta must be either a scalar or a p-length vector. If Vbeta takes a scalar value, then that value will serve as the prior variance for all of the betas. The default

variance is large and set to 1.0E6 for an uninformative flat prior.

seed The seed for the random number generator. Default to 1234.

verbose A switch (0,1) which determines whether or not the progress of the sampler is

printed to the screen. Default is 1: a progress bar is printed, indicating the step

(in %) reached by the Gibbs sampler.

Value

mcmc An mcmc object that contains the posterior sample. This object can be summa-

rized by functions provided by the coda package. The posterior sample of the

deviance D, with $D = -2\log(\prod_i P(y_i, n_i|\beta))$, is also provided.

prob.p.pred Predictive posterior mean of the probability associated to the suitability process

for each spatial cell.

Author(s)

Ghislain Vieilledent <ghislain.vieilledent@cirad.fr>

References

Latimer, A. M.; Wu, S. S.; Gelfand, A. E. and Silander, J. A. (2006) Building statistical models to analyze species distributions. *Ecological Applications*, 16, 33-50.

Gelfand, A. E.; Schmidt, A. M.; Wu, S.; Silander, J. A.; Latimer, A. and Rebelo, A. G. (2005) Modelling species diversity through species level hierarchical modelling. *Applied Statistics*, 54, 1-20.

See Also

```
plot.mcmc, summary.mcmc
```

hSDM.poisson

```
# Example with simulated data
#== Preambule
library(hSDM)
#=========
#== Data simulation
# Set seed for repeatability
set.seed(1234)
# Constants
nobs <- 1000
# Covariates for "suitability" process
X1 <- rnorm(n=nobs,0,1)</pre>
X2 <- rnorm(n=nobs,0,1)</pre>
X <- cbind(rep(1,nobs),X1,X2)</pre>
# Target parameters
beta.target <- matrix(c(0.2,0.1,0.1),ncol=1) # fixed effects
#== Simulating latent variables
# Suitability
log.theta <- X %*% beta.target</pre>
theta <- exp(log.theta)</pre>
#== Simulating response variable
Y <- rpois(n=nobs,lambda=theta)
V <- rep(1,nobs)</pre>
#== Data-set
Data <- data.frame(Y,V,X1,X2)</pre>
str(Data)
par(mfrow=c(2,1))
plot(Data$X1,log.theta)
plot(Data$X2,log.theta)
#==========
#== Statistical modelling
#== glm resolution to compare
mod.glm <- glm(Y~X1+X2,family="poisson",data=Data)</pre>
summary(mod.glm)
beta.hat.glm <- coef(mod.glm)</pre>
#== hSDM
mod.hSDM.poisson <- hSDM.poisson(counts=Data$Y,</pre>
                                 visits=Data$V,
                                 suitability=~X1+X2,
                                 data=Data, burnin=1000,
```

hSDM.poisson.iCAR

```
mcmc=1000, thin=1,
beta.start=beta.hat.glm,
mubeta=0, Vbeta=1.0E6,
seed=1234, verbose=1)
```

```
#=======
#== Outputs
summary(mod.hSDM.poisson$mcmc)
pdf(file="Posteriors_hSDM.poisson.pdf")
plot(mod.hSDM.poisson$mcmc)
dev.off()
summary(mod.hSDM.poisson$prob.p.pred)
## End(Not run)
```

hSDM.poisson.iCAR

The hSDM.poisson.iCAR function performs a Poisson regression in a hierarchical Bayesian framework. The suitability process includes a spatial correlation process. The spatial correlation is modelled using an intrinsic CAR model.

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Description

The hSDM. poisson. iCAR function calls a Gibbs sampler written in C code which uses a Metropolis algorithm to estimate the conditional posterior distribution of model's parameters.

Usage

```
hSDM.poisson.iCAR(counts, visits, suitability, cells, n.neighbors, neighbors, data, burnin = 5000, mcmc = 10000, thin = 10, beta.start, Vrho.start, mubeta = 0, Vbeta = 1e+06, priorVrho = "1/Gamma", shape = 0.5, rate = 0.0005, Vrho.max = 1000, seed = 1234, verbose = 1, save.rho = 0)
```

Arguments

counts	A vector indicating the count (or abundance) for each observation.
visits	A vector indicating if the observation point n has been visited (visits[n]=1) or not (visits[n]=0). If visits[n]=0, then counts[n]=0.
suitability	A one-sided formula of the form '~x1++xp' with p terms specifying the explicative variables for the suitability process of the model.
cells	A vector indicating the spatial cell identifier (from 1 to total number of cell) for each observation. Several observations can occur in one spatial cell.
n.neighbors	A vector of integers that indicates the number of neighbors (adjacent cells) of each spatial cell.

neighbors A vector of integers indicating the neighbors (adjacent cells) of each spatial

> cell. Must be of the form c(neighbors of cell 1, neighbors of cell 2, ..., neighbors of the last cell). Length of the neighbors vector should be equal

to sum(data\$num).

data A data frame containing the model's explicative variables.

burnin The number of burnin iterations for the sampler.

The number of Gibbs iterations for the sampler. Total number of Gibbs iterations mcmc

is equal to burnin+mcmc. burnin+mcmc must be divisible by 10 and superior or

equal to 100 so that the progress bar can be displayed.

thin The thinning interval used in the simulation. The number of mcmc iterations

must be divisible by this value.

beta.start Starting values for beta parameters.

Positive scalar indicating the starting value for the variance of the spatial random Vrho.start

effects.

mubeta Means of the priors for the β parameters of the suitability process. mubeta must

> be either a scalar or a p-length vector. If mubeta takes a scalar value, then that value will serve as the prior mean for all of the betas. The default value is set to

0 for an uninformative prior.

Vbeta Variances of the Normal priors for the β parameters of the suitability process.

> Vbeta must be either a scalar or a p-length vector. If Vbeta takes a scalar value, then that value will serve as the prior variance for all of the betas. The default

variance is large and set to 1.0E6 for an uninformative flat prior.

priorVrho Type of prior for the variance of the spatial random effects. Can be set to a fixed

positive scalar, or to an inverse-gamma distribution ("1/Gamma") with parameters shape and rate, or to a uniform distribution ("Uniform") on the interval

[0,Vrho.max]. Default to "1/Gamma".

The shape parameter for the Gamma prior on the precision of the spatial random shape

effects. Default value is shape=0.05 for uninformative prior.

rate The rate (1/scale) parameter for the Gamma prior on the precision of the spatial

random effects. Default value is rate=0.0005 for uninformative prior.

Upper bound for the uniform prior of the spatial random effect variance. Default Vrho.max

to 1000.

seed The seed for the random number generator. Default to 1234.

A switch (0,1) which determines whether or not the progress of the sampler is verbose

printed to the screen. Default is 1: a progress bar is printed, indicating the step

(in %) reached by the Gibbs sampler.

save.rho A switch (0,1) which determines whether or not the sampled values for rhos

> are saved. Default is 0: the posterior mean is computed and returned in the rho.pred vector. Be careful, setting save.rho to 1 might require a large amount

of memory.

Value

An mcmc object that contains the posterior sample. This object can be summamcmc

rized by functions provided by the coda package. The posterior sample of the

deviance D, with $D = -2\log(\prod_i P(y_i, n_i|\beta, \rho_i))$, is also provided.

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rho.pred If save.rho is set to 0 (default), rho.pred is the predictive posterior mean of

the spatial random effect associated to each spatial cell. If save.rho is set to 1, rho.pred is an mcmc object with sampled values for each spatial random effect $\,$

associated to each spatial cell.

prob.p.pred Predictive posterior mean of the probability associated to the suitability process

for each spatial cell.

Author(s)

Ghislain Vieilledent <ghislain.vieilledent@cirad.fr>

References

Latimer, A. M.; Wu, S. S.; Gelfand, A. E. and Silander, J. A. (2006) Building statistical models to analyze species distributions. *Ecological Applications*, 16, 33-50.

Gelfand, A. E.; Schmidt, A. M.; Wu, S.; Silander, J. A.; Latimer, A. and Rebelo, A. G. (2005) Modelling species diversity through species level hierarchical modelling. *Applied Statistics*, 54, 1-20.

See Also

```
plot.mcmc, summary.mcmc
```

```
## Not run:
# hSDM.poisson.iCAR()
# Example with simulated data
#== Preambule
library(mvtnorm)
library(lme4)
               # To compare with nonspatial random effects model
library(hSDM)
#========
#== Data simulation
# Set seed for repeatability
set.seed(1234)
# Constants
ncell <- 150 # Number of cells
nobs <- 10*ncell # Number of observation for the *poisson* random variable
cell <- rep(c(1:ncell),each=nobs/ncell)</pre>
# Covariates for "suitability" process
X1 <- rnorm(n=nobs,0,1)</pre>
```

hSDM.poisson.iCAR

```
X2 <- rnorm(n=nobs,0,1)</pre>
X <- cbind(rep(1,nobs),X1,X2)</pre>
# Target parameters
beta.target <- matrix(c(0.2,0.1,0.1), ncol=1) # fixed effects
Vrho.target <- 10 # Spatial Variance</pre>
# Generate symmetric adjacency matrix, A
A <- matrix(0,ncell,ncell)
A[upper.tri(A,diag=F)] <- rbinom(ncell*(ncell-1)/2,1,.05)
A \leftarrow A+t(A)
n.neighbors <- apply(A,1,sum)</pre>
f.adjacent <- function (x) {</pre>
  which(x==1)
adj <- unlist(apply(A,1,f.adjacent))</pre>
# Spatial effects, phi
d <- 1 # Spatial dependence parameter = 1 for intrinsic CAR
Q <- diag(n.neighbors)-d*A + diag(.0001,ncell) # Add small constant to make Q non-singular
covrho <- Vrho.target*solve(Q) # Covariance of rhos</pre>
rho <- c(rmvnorm(1,sigma=covrho)) # Spatial Random Effects</pre>
rho <- rho-mean(rho) # Centering rhos on zero</pre>
#== Simulating latent variables
# Suitability
log.theta <- vector()</pre>
for (n in 1:nobs) {
 log.theta[n] <- X[n,]%*%beta.target+rho[cell[n]]</pre>
theta <- exp(log.theta)</pre>
#== Simulating response variable
Y <- rpois(nobs,theta)</pre>
V <- rep(1,nobs)</pre>
#== Data-set
Data <- data.frame(Y,V,cell,X1,X2)</pre>
str(Data)
#==========
#== Statistical modelling
#== glm resolution to compare
mod.glm <- glm(Y~X1+X2,data=Data,family="poisson")</pre>
summary(mod.glm)
\#== glmm (mixed models) resolution to compare
#== cell random effects without spatial correlation
summary(mod.glmm)
beta.hat.glmm <- fixef(mod.glmm)</pre>
```

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```
#== hSDM
mod.hSDM.poisson.iCAR <- hSDM.poisson.iCAR(counts=Data$Y,</pre>
                                            visits=Data$V,
                                            suitability=~X1+X2,
                                            cells=Data$cell,
                                            n.neighbors=n.neighbors,
                                            neighbors=adj,
                                            data=Data, burnin=1000,
                                            mcmc=1000, thin=1,
                                            beta.start=beta.hat.glmm,
                                            Vrho.start=10,
                                            priorVrho="1/Gamma",
                                            #priorVrho="Uniform",
                                            #priorVrho=10,
                                            mubeta=0, Vbeta=1.0E6,
                                            shape=0.5, rate=0.0005,
                                            Vrho.max=1000,
                                            seed=1234, verbose=1, save.rho=0)
#=======
#== Outputs
summary(mod.hSDM.poisson.iCAR$mcmc)
pdf(file="Posteriors_hSDM.poisson.iCAR.pdf")
plot(mod.hSDM.poisson.iCAR$mcmc)
dev.off()
summary(mod.hSDM.poisson.iCAR$prob.p.pred)
summary(mod.hSDM.poisson.iCAR$rho.pred)
## End(Not run)
```

hSDM.site.occupancy

The hSDM.site.occupancy function can be used to model species distribution including different processes in a hierarchical Bayesian framework: a Bernoulli suitability process (refering to environmental suitability) which takes into account the spatial dependence of the observations, and a Binomial observability process (refering to various ecological and methodological issues explaining the species detection).

Description

The hSDM. site.occupancy function calls a Gibbs sampler written in C code which uses a Metropolis algorithm to estimate the conditional posterior distribution of hierarchical model's parameters.

Usage

```
hSDM.site.occupancy(presences, trials, suitability,
observability, spatial.entity, data, n.neighbors, neighbors,
suitability.pred=NULL, spatial.entity.pred=NULL, burnin = 5000, mcmc =
10000, thin = 10, beta.start, gamma.start, Vrho.start, mubeta = 0, Vbeta
= 1e+06, mugamma = 0, Vgamma = 1e+06, priorVrho = "1/Gamma", shape =
0.5, rate = 0.0005, Vrho.max=1000, seed = 1234, verbose = 1, save.rho =
0, save.p = 0)
```

Arguments

presences A vector indicating the number of successes (or presences) for each observation. trials A vector indicating the number of trials for each observation. t_i should be superior to zero and superior or equal to y_i , the number of successes for observation i.

A one-sided formula of the form $\sim x_1 + ... + x_p$ with p terms specifying the explicative variables for the suitability process.

observability A one-sided formula of the form $\sim w_1 + ... + w_q$ with q terms specifying the explicative variables for the observability process.

spatial.entity A vector indicating the spatial entity identifier (from one to the total number of entities) for each observation. Several observations can occur in one spatial entity. A spatial entity can be a raster cell for example.

data A data frame containing the model's variables.

n.neighbors A vector of integers that indicates the number of neighbors (adjacent entities) of each spatial entity. length(n.neighbors) indicates the total number of spatial entities.

> A vector of integers indicating the neighbors (adjacent entities) of each spatial entity. Must be of the form c(neighbors of entity 1, neighbors of entity 2, ..., neighbors of the last entity). Length of the neighbors vector should be equal to

suitability.pred

An optional data frame in which to look for variables with which to predict. If NULL, the observations are used.

spatial.entity.pred

An optional vector indicating the spatial entity identifier (from one to the total number of entities) for predictions. If NULL, the vector spatial.entity for observations is used.

burnin The number of burnin iterations for the sampler.

The number of Gibbs iterations for the sampler. Total number of Gibbs iterations is equal to burnin+mcmc. burnin+mcmc must be divisible by 10 and superior or equal to 100 so that the progress bar can be displayed.

thin The thinning interval used in the simulation. The number of mcmc iterations must be divisible by this value.

> Starting values for β parameters of the suitability process. This can either be a scalar or a p-length vector.

suitability

neighbors

sum(n.neighbors).

mcmc

beta.start

Vrho.start

mubeta

gamma.start Starting values for β parameters of the observability process. This can either be a scalar or a q-length vector.

Positive scalar indicating the starting value for the variance of the spatial random effects.

Means of the priors for the β parameters of the suitability process. mubeta must be either a scalar or a p-length vector. If mubeta takes a scalar value, then that

value will serve as the prior mean for all of the betas. The default value is set to

0 for an uninformative prior.

Vbeta Variances of the Normal priors for the β parameters of the suitability process.

Vbeta must be either a scalar or a p-length vector. If Vbeta takes a scalar value, then that value will serve as the prior variance for all of the betas. The default

variance is large and set to 1.0E6 for an uninformative flat prior.

mugamma Means of the Normal priors for the γ parameters of the observability process.

mugamma must be either a scalar or a p-length vector. If mugamma takes a scalar value, then that value will serve as the prior mean for all of the gammas. The

default value is set to 0 for an uninformative prior.

Vgamma Variances of the Normal priors for the γ parameters of the observability process.

Vgamma must be either a scalar or a p-length vector. If Vgamma takes a scalar value, then that value will serve as the prior variance for all of the gammas. The

default variance is large and set to 1.0E6 for an uninformative flat prior.

priorVrho Type of prior for the variance of the spatial random effects. Can be set to a fixed positive scalar, or to an inverse-gamma distribution ("1/Gamma") with param-

eters shape and rate, or to a uniform distribution ("Uniform") on the interval

[0,Vrho.max]. Default set to "1/Gamma".

shape The shape parameter for the Gamma prior on the precision of the spatial random

effects. Default value is shape=0.05 for uninformative prior.

The rate (1/scale) parameter for the Gamma prior on the precision of the spatial

random effects. Default value is rate=0.0005 for uninformative prior.

Vrho.max Upper bound for the uniform prior of the spatial random effect variance. Default

set to 1000.

seed The seed for the random number generator. Default set to 1234.

verbose A switch (0,1) which determines whether or not the progress of the sampler is

printed to the screen. Default is 1: a progress bar is printed, indicating the step

(in %) reached by the Gibbs sampler.

save.rho A switch (0,1) which determines whether or not the sampled values for rhos

are saved. Default is 0: the posterior mean is computed and returned in the rho.pred vector. Be careful, setting save.rho to 1 might require a large amount

of memory.

save.p A switch (0,1) which determines whether or not the sampled values for predic-

tions are saved. Default is 0: the posterior mean is computed and returned in the prob.p.pred vector. Be careful, setting save.p to 1 might require a large

amount of memory.

Details

The model integrates two processes, an ecological process associated to the presence or absence of the species due to habitat suitability and an observation process that takes into account the fact that the probability of detection of the species is inferior to one. The ecological process includes an intrinsic conditional autoregressive model (iCAR) model for spatial autocorrelation between observations, assuming that the probability of presence of the species at one site depends on the probability of presence of the species on neighboring sites.

Ecological process:

$$z_i \sim \mathcal{B}ernoulli(\theta_i)$$

 $logit(\theta_i) = X_i\beta + \rho_{j(i)}$

 ρ_i : spatial random effect

j(i): index of the spatial entity for observation i.

Spatial autocorrelation:

An intrinsic conditional autoregressive model (iCAR) is assumed:

$$\rho_j \sim \mathcal{N}ormal(\mu_j, V_\rho/n_j)$$

 μ_j : mean of $\rho_{j'}$ in the neighborhood of j.

 V_{ρ} : variance of the spatial random effects.

 n_i : number of neighbors for spatial entity j.

Observation process:

$$y_i \sim \mathcal{B}inomial(z_i * \delta_i, t_i)$$

$$logit(\delta_i) = W_i \gamma$$

Value

mcmc	An meme object that contains the posterior sample. This object can be summarized by functions provided by the coda package. The posterior sample of the deviance D , with $D = -2 \log(\prod_i P(y_i, z_i))$, is also provided.
rho.pred	If save.rho is set to 0 (default), rho.pred is the predictive posterior mean of the spatial random effect associated to each spatial entity. If save.rho is set to 1, rho.pred is an mcmc object with sampled values for each spatial random effect associated to each spatial entity.
prob.p.pred	If save.p is set to 0 (default), prob.p.pred is the predictive posterior mean of the probability associated to the suitability process for each prediction. If save.p is set to 1, prob.p.pred is an mcmc object with sampled values of the probability associated to the suitability process for each prediction.
prob.p.latent	Predictive posterior mean of the probability associated to the suitability process for each observation.
prob.q.latent	Predictive posterior mean of the probability associated to the observability pro-

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cess for each observation.

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References

Gelfand, A. E.; Schmidt, A. M.; Wu, S.; Silander, J. A.; Latimer, A. and Rebelo, A. G. (2005) Modelling species diversity through species level hierarchical modelling. *Applied Statistics*, 54, 1-20.

Gilks, W. R., Best, N. G. and Tan, K. K. C. (1995) Adaptive rejection Metropolis sampling. *Applied Statistics*, 44, 455-472.

Latimer, A. M.; Wu, S. S.; Gelfand, A. E. and Silander, J. A. (2006) Building statistical models to analyze species distributions. *Ecological Applications*, 16, 33-50.

MacKenzie, D. I.; Nichols, J. D.; Lachman, G. B.; Droege, S.; Andrew Royle, J. and Langtimm, C. A. (2002) Estimating site occupancy rates when detection probabilities are less than one. *Ecology*, 83, 2248-2255.

See Also

plot.mcmc, summary.mcmc

```
## Not run:
# hSDM.binomial.iCAR()
# Example with simulated data
#=======
#== Preambule
library(lme4)
library(hSDM)
library(raster)
library(sp)
library(mvtnorm)
#=========
#== Data simulation
# Set seed for repeatability
seed <- 1234
# Target parameters
beta.target <- matrix(c(0.2,0.5,0.5),ncol=1)
gamma.target <- matrix(c(0.2),ncol=1) #= This means that theta.2~0.5: 50% of chance to detect the species
Vrho.target <- 1 # Spatial Variance
# Landscape
Landscape <- raster(ncol=20,nrow=20,crs='+proj=utm +zone=1')</pre>
ncell <- ncell(Landscape)</pre>
# Neighbors
```

```
neighbors.mat <- adjacent(Landscape, cells=c(1:ncell), directions=8, pairs=TRUE, sorted=TRUE)</pre>
n.neighbors <- as.data.frame(table(as.factor(neighbors.mat[,1])))[,2]</pre>
adj <- neighbors.mat[,2]</pre>
# Generate symmetric adjacency matrix, A
A <- matrix(0,ncell,ncell)
index.start <- 1</pre>
for (i in 1:ncell) {
    index.end <- index.start+n.neighbors[i]-1</pre>
    A[i,adj[c(index.start:index.end)]] <- 1
    index.start <- index.end+1</pre>
}
# Spatial effects
d <- 1 # Spatial dependence parameter = 1 for intrinsic CAR
Q \leftarrow diag(n.neighbors)-d*A + diag(.0001,ncell) # Add small constant to make Q non-singular
covrho <- Vrho.target*solve(Q) # Covariance of rhos</pre>
set.seed(seed)
rho <- c(rmvnorm(1,sigma=covrho)) # Spatial Random Effects</pre>
rho <- rho-mean(rho) # Centering rhos on zero</pre>
# Visited cells
n.visited <- 150 # Compare with 400, 350 and 100 for example
set.seed(seed)
visited.cells <- sort(sample(1:ncell,n.visited,replace=FALSE)) # Draw visited cells at random
notvisited.cells <- c(1:ncell)[-visited.cells]</pre>
# Number of observations
nobs <- 300
# Cell vector
set.seed(seed)
cells <- c(visited.cells,sample(visited.cells,nobs-n.visited,replace=TRUE))</pre>
coords <- xyFromCell(Landscape,cells) # Get coordinates</pre>
# Covariates for "suitability" process
set.seed(seed)
X1.cell \leftarrow rnorm(n=ncell,0,1) # Here, we decided to define one habitat by spatial entity (raster cell)
set.seed(2*seed)
X2.cell <- rnorm(n=ncell,0,1)</pre>
X1 <- X1.cell[cells]</pre>
X2 <- X2.cell[cells]</pre>
X <- cbind(rep(1,nobs),X1,X2)</pre>
# Covariates for "observability" process
W <- cbind(rep(1,nobs))</pre>
#== Simulating latent variables
# Suitability
logit.theta.1 <- vector()</pre>
for (n in 1:nobs) {
  logit.theta.1[n] <- X[n,]%*%beta.target+rho[cells[n]]</pre>
```

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```
theta.1 <- inv.logit(logit.theta.1)</pre>
set.seed(seed)
y.1 <- rbinom(nobs,1,theta.1)</pre>
# Observability
set.seed(seed)
trials <- rpois(nobs,5) # Number of trial associated to each observation
trials[trials==0] <- 1</pre>
logit.theta.2 <- W%*%gamma.target</pre>
theta.2 <- inv.logit(logit.theta.2)</pre>
set.seed(seed)
y.2 <- rbinom(nobs,trials,theta.2)</pre>
#== Simulating response variable
Y \leftarrow y.2*y.1
#== Data-set
Data <- data.frame(Y,trials,cells,X1,X2,W1,W2)</pre>
Data <- SpatialPointsDataFrame(coords=coords,data=Data)</pre>
plot(Data)
#== Data-set for predictions (suitability on each spatial cell)
Data.pred <- data.frame(X1=X1.cell,X2=X2.cell,cells=c(1:ncell))</pre>
#== Site-occupancy model
mod.hSDM.site.occupancy <- hSDM.site.occupancy(presences=Data$Y,</pre>
                                                 trials=Data$trials,
                                                 suitability=~X1+X2,
                                                 observability=~1,
                                                 spatial.entity=Data$cells,
                                                 data=Data,
                                                 n.neighbors=n.neighbors,
                                                 neighbors=adj,
                                                 ## suitability.pred=NULL,
                                                 ## spatial.entity.pred=NULL,
                                                 suitability.pred=Data.pred,
                                                 spatial.entity.pred=Data.pred$cells,
                                                 burnin=5000, mcmc=5000, thin=5,
                                                 beta.start=0,
                                                 gamma.start=0,
                                                 Vrho.start=10,
                                                 priorVrho="1/Gamma",
                                                 #priorVrho="Uniform",
                                                 #priorVrho=10,
                                                 mubeta=0, Vbeta=1.0E6,
                                                 shape=0.5, rate=0.0005,
                                                 Vrho.max=1000,
                                                 seed=1234, verbose=1,
                                                 save.rho=1, save.p=0)
```

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```
#== Outputs
#= Parameter estimates
summary(mod.hSDM.site.occupancy$mcmc)
#= MCMC and posteriors
pdf(file="Posteriors_hSDM.site.occupancy.pdf")
plot(mod.hSDM.site.occupancy$mcmc)
dev.off()
pdf(file="Posteriors.rho_hSDM.site.occupancy.pdf")
plot(mod.hSDM.site.occupancy$rho.pred)
dev.off()
#= Summary plots
# rho
r.rho <- r.rho.pred <- r.visited <- Landscape</pre>
r.rho[] <- rho
rho.pred <- apply(mod.hSDM.site.occupancy$rho.pred,2,mean)</pre>
r.rho.pred[] <- rho.pred</pre>
r.visited[] <- 0
r.visited[visited.cells] <- 1</pre>
# prob.p
r.prob.p <- Landscape</pre>
r.prob.p[] <- mod.hSDM.site.occupancy$prob.p.pred</pre>
pdf(file="Summary_hSDM.site.occupancy.pdf")
par(mfrow=c(3,2))
plot(r.rho, main="rho target")
plot(r.visited,main="Visited cells and presences")
plot(Data[Y>0,],add=TRUE,pch=16,cex=0.5)
plot(r.rho.pred, main="rho estimated")
plot(rho[visited.cells], rho.pred[visited.cells],
     xlab="rho target",
     ylab="rho estimated")
points(rho[notvisited.cells],rho.pred[notvisited.cells],pch=16,col="blue")
legend(x=-4,y=3.5,legend="Unvisited cells",col="blue",pch=16,bty="n")
abline(a=0,b=1,col="red")
plot(r.prob.p,main="Proba of presence")
plot(Data[Y>0,],add=TRUE,pch=16,cex=0.5)
dev.off()
## End(Not run)
```

logit 43

Description

Compute generalized logit and generalized inverse logit functions.

Usage

```
logit(x, min = 0, max = 1)
inv.logit(x, min = 0, max = 1)
```

Arguments

x value(s) to be transformedmin Lower end of logit intervalmax Upper end of logit interval

Details

The generalized logit function takes values on [min, max] and transforms them to span [-Inf,Inf] it is defined as:

$$y = log(\frac{p}{(1-p)})$$

where

$$p = \frac{(x - min)}{(max - min)}$$

The generized inverse logit function provides the inverse transformation:

$$x = p'(max - min) + min$$

where

$$p' = \frac{exp(y)}{(1 + exp(y))}$$

Value

Transformed value(s).

Author(s)

Gregory R. Warnes <greg@warnes.net>

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```
## Not run:
    x <- seq(0,10, by=0.25)
    xt <- logit(x, min=0, max=10)
    cbind(x,xt)

y <- inv.logit(xt, min=0, max=10)
    cbind(x,xt,y)

## End(Not run)</pre>
```

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