

get a model of the form  $gas = b_0 + b_1 distance$ .

If you were to use E[y] for your predictions here, you would calculate the mean gas volume for all trips and say "regardless of the distance we're going to travel, I anticipate that we will use E[gas] gallons of fuel." But, because we have the regression model, we can say "given that we will travel d miles, we expect to use  $E[gas|distance=d]=b_0+b_1d$  gallons of fuel."

If most of your trips are about the same distance, just using E[gas] might actually give you decent results. But given that we know there is a relationship between distance traveled and gas used, we anticipate that the conditional expectation will give us much better predictions than the blanket prediction of the unconditioned expectation.

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answered Nov 11 '13 at 5:27



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There are two basic versions of a regression specification, regarding the assumptions about the nature of the variables involved:

In the first, the regresses are assumed deterministic, and so the actual matrix contains



elements that are not realizations of random variables. This comes from an "experimental design" tradition (and hence sometimes the regressor matrix is called the "design matrix"). In such a set up the regressors are not random variables and so it is not meaningful to "condition" on them. So we have to formulate the model in terms of unconditional expected values, like the ones OP wrote.

The second version mainly relates to a non-experimental setting, like economics for example, we treat the regressors as random variables, while the specific sample/regressor matrix as being comprised of realizations of random variables. Here, we condition on X to be humble: we do not claim that we know the unconditional behavior of Y or of the error, but we make "narrower" assumptions, conditional assumptions. Hence  $E(u\mid X)=0$  etc. Intuitively we restrict ourselves to "what can I say about the relation between Y and X, given the specific data that I have in hand?" Apart for humbleness, it also cleverly permits us to derive all sorts of useful results (which we could not if we assumed stochastic regressors and considered at the same time assumptions on unconditional expectations).

For example, to stick to basics, in proving unbiasedness of the OLS estimator we write

$$\hat{\beta} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'(\mathbf{X}'\beta + \mathbf{u}) = \beta + (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{u}$$

Assume that we treat the regressors as stochastic. If you want to check unbiasedness, you will examine

$$E(\hat{eta}) = \beta + E((\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{u})$$

Where does that leave you? Nowhere - it is difficult to justify an a priori assumption like  $E((\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{u}) = 0$ . But note that, by the properties of expected values

$$\begin{split} E\left((\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{u}\right) &= E\left[E\left((\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{u}\right) \mid X\right] \\ &= E\left[(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'E\left(\mathbf{u} \mid X\right)\right] \end{split}$$

Now, to have unbiasedness, we only need to a priori assume that  $E(\mathbf{u}\mid X)=0$ , which can at least be logically argued upon.

Finally, note that there is a drawback in assuming deterministic, non-stochastic regressors, which is sometimes overlooked: If our model is

$$y_i = a + bx_i + u_i, \quad i = 1, \dots, n$$

then under non-stochastic regressors we have

$$E(y_i) = a + bx_i$$

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But this means that each  $y_i$  has in principle a different expected value: so the  $y_i$ 's here do not come from an identically distributed population. If they don't, then our sample  $\{Y, X\}$ , that contains as a

But this means that each  $y_i$  has in principle a different expected value: so the  $y_i$ 's here do not come from an identically distributed population. If they don't, then our sample  $\{Y,X\}$ , that contains as a random variable only the Y is not "random" (i.e. it is not i.i.d), due to the assumption that the X's are deterministic.

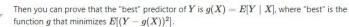
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Let's say you observe a RV X and you want to predict a second RV Y where your predictor of Y is g(X).





So  $E[Y\mid X]$  will always be at least as good as, if not better than E[Y].

To prove that E[Y|X] is the "best" predictor possible we just need to show that

$$E[(Y - g(X))^2] \ge E[(Y - E[Y \mid X])^2]$$

Let's start off by looking at  $E[(Y-g(X))^2\mid X]$  which is not quite the same as the LHS above but not to worry, at the end we can leverage the property that E[E[X|Y]]=E[X]. So,

$$\begin{split} E[(Y-g(X))^2 \mid X] &= E[(Y-E[Y \mid X] + E[Y \mid X] - g(X))^2 \mid X] \\ &= E[(Y-E[Y \mid X]))^2 \mid X] \\ &\quad + E[(E[Y \mid X] - g(X))^2 \mid X] \\ &\quad + 2E[(Y-E[Y \mid X])(E[Y \mid X] - g(X)) \mid X] \end{split}$$

Where the above is basically doing

$$(x+y)^2 = (x-c+c+y)^2 = (x-c)^2 + (x+c)^2 + 2(x-c)(x+c)$$

and using the linearity of expectations.

Focusing on the third term of the expansion only, we can show that it equals zero.

$$2E[(Y - E[Y \mid X])(E[Y \mid X] - g(X)) \mid X] = 0$$

This is because given X, the term  $E[Y\mid X]=g(X)$  is like a constant and can be pulled out of the expectation using the property that E[aX]=aE[X]. Letting the constant we pulled out equal c we have

$$\begin{split} c \cdot E[Y - E[Y \mid X] \mid X] &= c \cdot \left( E[Y \mid X] - E[E[Y \mid X] \mid X] \right) \\ &= c \cdot \left( E[Y \mid X] - E[Y \mid X] \right) = 0 \end{split}$$

From the property  $E[g(X) \mid X] = g(X)$  we know that  $E[E[Y \mid X] \mid X] = E[Y \mid X]$  since  $E[Y \mid X]$  is a function of X.

Therefore going back to the expansion, we have

$$E[(Y - g(X))^2 \mid X] = E[(Y - E[Y \mid X]))^2 \mid X] + E[(E[Y \mid X] - g(X))^2 \mid X]$$

Noting that the second term on the RHS must always be positive, we can say

$$E[(Y - g(X))^2 \mid X] \ge E[(Y - E[Y \mid X]))^2 \mid X]$$

and then taking the expected value of both sides, leveraging the property that  $E[X] = E[E[X \mid Y]]$ , we get

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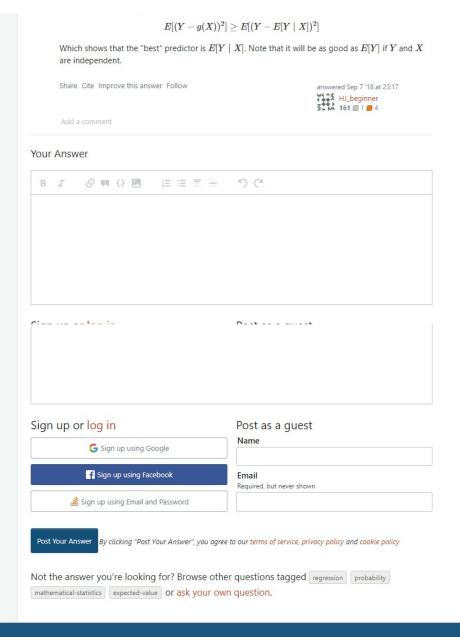
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