

RESEARCH

Predicting the spread and persistence of genetically modified dominant sterile male mosquitoes

Adrien Ickowicz, Scott D Foster, Geoffrey R Hosack and Keith R Hayes

Full list of author information is available at the end of the article

Abstract

Background: Reproductive containment provides an opportunity to implement a staged-release strategy for genetic control of malaria vectors, in particular allowing predictions about the spread and persistence of (self-limiting) sterile and male-biased strains to be compared to outcomes before moving to (self-sustaining) gene-drive strains.

Methods: We: (i) describe a diffusion-advection-reaction model of the spread and persistence of a single cohort of male mosquitoes; (ii) elicit informative prior distributions for model parameters, for wild type (WT) and genetically modified dominant sterile strains (DSM); (iii) estimate posterior distributions for wild type strains using data from published Mark Recapture Release (MRR) experiments, with inference performed through a delayed rejection adaptive Metropolis algorithm; and, (iv) weight prior distributions in order to make predictions about genetically modified strains using Bayes factors calculated for the wild type strains.

Results: If a single cohort of 5000 genetically modified dominant sterile male mosquitoes are released at the same location as the previous MRR experiments with WT, there is a 90% probability that the expected number of released mosquitoes will fall below 1 in 10 days, and that by 12 days there will be a 99% probability that no mosquitoes will be found more than 150 meters from the release location.

Conclusions: Spread and persistence models should form a key component of risk assessments of novel genetic control strategies for malaria vectors. Our predictions, used in an independent risk assessment, suggest genetically modified sterile male mosquitoes will remain within the locality of the release site, and persist for a very limited amount of time. Data gathered following the release of these mosquitoes will enable us to test the accuracy of these predictions and also provide a means to update parameter distributions for genetic strains in a coherent (Bayesian) framework. We anticipate this will provide additional insights about how to conduct probabilistic risk assessments for the stage-released of genetically modified mosquitoes.

Keywords: entomological survey data; sterile male; expert elicitation; Bayesian hierarchical model; monitoring

1 Background

Over the last thirty years conventional control strategies have reduced the global incidence rate of malaria. The efficacy of these strategies, however, appears to be waning. Global malaria incidence rates have barely changed since 2014, stalling

around 57 cases per 1000 population at risk despite increased expenditure on elimination and prevention between from 2014 to 2017 [?]. Consequently, the World Health Organisation predicts that morbidity targets for 2020, 2025 and 2030 will not to be met without significant improvements in vector control and chemoprevention.

Laboratory studies and modelling indicate that novel control strategies that use gene drives to force effector genes through vector populations to suppress them, or render them unable to transmit plasmodium parasites, could augment conventional strategies and significantly improve current rates of malaria control [? ? ? ?]. As with many novel technologies, however, the safety of genetic control methods is uncertain, and the potential for spread and ecosystem-wide impacts makes gene-drives contentious [? ?].

To date, all research on gene-drive modified mosquitoes (GDMMs) has been conducted in laboratories under strict physical, ecological, reproductive and/or molecular containment [?]. Current guidance recommends that if containment is to be relaxed it should be done in a “phased-release” strategy wherein data is gathered, risks are evaluated, and containment is gradually lifted in a step-wise fashion if and when the risks are deemed acceptable [? ?]. The possibility of long-distance dispersal [?] and biosecurity concerns with large semi-field enclosure, however, may erode confidence in strategies that try to gradually lift ecological or physical containment.

Genetic and reproductive containment strategies on the other hand may be more amenable to a gradual-release strategy, and several approaches are currently being pursued [? ? ?]. In this paper we focus on the spatio-temporal dynamics that govern the spread and persistence of a single cohort of genetically modified, Dominant Sterile Male (DSM) *Anopheles coluzzii* mosquitoes released in a small village in Burkina Faso. These mosquitoes are reproductively contained, and represent the first stage in a three-stage pathway to malaria vector control using a gene drive that results in a male-biased sex ratio [? ?].

In our analysis the spatio-temporal dynamics of DSM mosquitoes are driven by: (i) a diffusion process that accounts for the dispersal capacity of the insects; (ii) an advection process that accounts for the attraction of individuals to swarms, and (iii) a reaction process that accounts for the death of sterile male mosquitoes. These processes can be described mathematically by a partial differential equation (PDE) of the advection-diffusion-reaction type, which models how the distribution of mosquitoes in time and space is affected by a chemo-attractant present in the environment. Models of this type have also been widely used within the environmental and ecological literature. In particular, such models have been used before to model mosquitoes, for example the dispersal of *Aedes albopictus* in Reunion Island using a similar type of parabolic PDE [?].

Our objective is to predict the spread and persistence of the genetically modified mosquitoes ahead of a proposed field release. These predictions are an important component of an independent risk assessment conducted prior to the field release [?], and could also help to inform the design of a post-release monitoring strategy. Our approach implements the Bayesian paradigm; we elicit the prior distribution of model parameters from relevant experts, for both the wild type mosquito (WT) and its DSM counterpart. We obtain the posterior model parameter values for the WT parameters using the results of mark-release-recapture (MRR) experiments

conducted by [?]. We then use Bayesian model averaging to predict the spread and persistence of DSM mosquitoes.

2 Methods

2.1 PDE model

The literature review conducted by [?] suggests that the behaviour of host-seeking mosquitoes can be categorised as: (i) *plume finding* in which flight direction is either random (kinesis), determined by visual features in the absence of wind or deliberately upwind, downwind or crosswind (anemotaxis) when wind is present; and, (ii) *plume tracking*, where once the mosquito detects an odour plume, flight direction is deliberately upwind, or possibly determined by the odour gradient in windless conditions, in order to find the odour source (positive chemotaxis). More details on the properties of these search strategies can be found in [?].

In this analysis, we assume that village compounds (houses) provide a source of CO₂ that acts as an attractant for female *Anopheles gambiae sensu strictu* and *Anopheles coluzzii* mosquitoes [? ?]. We further assume that male mosquitoes swarm in places where the probability of encountering receptive females is highest [?], hence swarms occur in and around compounds (as evident in Figure 1 of [?]). Consequently, high concentrations of attractant result in flights that (on average) are closer to the direction of the nearest compound for both males and females. Conversely, low concentrations have less influence on the mosquitoes and result in greater variability in flight direction. We do not incorporate the effect of wind speed or direction into the analysis (we assume the local dispersal dynamics are not greatly affected by the wind at this spatial scale), and also assume that the attraction process is not influenced by the number of mosquitoes.

With these assumptions we developed an advection-diffusion-reaction model to describe the expected number of male WT and DSM mosquitoes,

$$\lambda_t = D \left[\lambda_{ss} - \alpha \cdot \nabla_s (\lambda U(s)) \right] - \mu \lambda, \quad (1)$$

with the following boundary conditions (Neumann reflecting conditions):

$$\frac{\partial \lambda}{\partial s} = 0 \quad (2)$$

where $U(s)$ is a function that describes the strength of mosquito attraction at location s , and α enables us to modulate the effect of the utility function on the overall attraction. Utility functions like this have been used in similar ecological contexts to describe an attractive or repulsive flux [?].

In this model we use swarm sites as centers of attraction and assume that the strength of attraction decays according to a squared-sum exponential decay kernel

$$U(s) = \sum_{s_\ell \in \mathcal{L}} \exp \left[- \frac{|s - s_\ell|^2}{\sigma^2} \right] \quad (3)$$

where \mathcal{L} are the known swarm locations (identified prior to the MRR experiment [?]) and σ is an unknown parameter that controls the range of attraction - that

is the distance beyond which the attraction of a swarm site s_ℓ for a mosquito is negligible.

The theoretical solution to the partial differential equation 1 is the expected number of male mosquitoes at time t and location $\mathbf{s} = (\text{easting}, \text{northing})$. The partial derivatives are denoted λ_t (with respect to time) and λ_{ss} (with respect to location). Because we use a numerical solver, we define a grid over the spatial domain. Each grid cell is therefore one areal unit. In this study, an areal unit is a square of 100 m² of area. All process model parameters are summarised in Table 1.

Name	Description	Units
α	Attractiveness of swarms	$\text{reward}^{-1} \cdot \text{m}^{-1}$
μ	Male mosquito mortality rate	day^{-1}
λ	The expected number of male mosquitoes	Number of mosquitoes per areal unit
D	Isotropic diffusion coefficient	$\text{m}^2 \cdot \text{day}^{-1}$
σ	Decay of attraction to swarm sites	m

Table 1 The parameters and units of the diffusion-advection-reaction model for sterile male mosquitoes.

2.2 Data

In this analysis we use the results of five MRR experiments [?] wherein approximately 5000 adult male WT mosquitoes, marked with a colored powder, were released two hours before swarming (around 4pm), at three different locations, on five different dates (Table 2). Mosquitoes were recaptured in three ways via: (i) swarm sampling in the evening (times were not specified, but we assume the sampling occurred between 7pm and 9pm); (ii) Pyrethroid Spray Catches in the morning inside houses; and, (iii) placement of humidified clay pots within rooms within houses, that were subsequently checked in the morning (between 6am-7am). All releases and re-captures sites were geo-located allowing for a spatial model to be used.

Recaptures were performed for seven days after release for all experiments except the first where recaptures were performed for five days. We used the results of the first four MRR experiments to derive posterior estimates of male WT model parameters, reserving the results of the fifth experiment to compare with model predictions.

MRR	Date	Release site	Longitude	Latitude	Number released
1	2013-10-09	A	-4.4724	11.2347	1146
1	-	B	-4.4755	11.2342	1158
1	-	C	-4.4718	11.2318	1103
2	2014-05-07	A	-4.4724	11.2347	1878
2	-	B	-4.4755	11.2342	1655
2	-	C	-4.4718	11.2318	1734
3	2014-09-04	A	-4.4724	11.2347	1665
3	-	B	-4.4755	11.2342	1673
3	-	C	-4.4718	11.2318	1684
4	2015-04-09	A	-4.4724	11.2347	2107
4	-	B	-4.4755	11.2342	2013
4	-	C	-4.4718	11.2318	1953
5	2015-10-09	C	-4.4724	11.2347	5992

Table 2 Summary of the MRR experiments conducted by [?] in Burkina Faso. We use the results of the first four experiments to parameterise our model and calculate Bayes factor for expert-derived priors, and use the fifth MRR experiment is to validate the model predictions.

2.3 Priors

2.3.1 Process model

A key challenge when conducting risk assessments of a novel technology is the lack of empirical information. Without any operational history from which success and failure rates might be estimated, quantitative risk predictions must (at least initially), be based on the testable predictions of domain experts. The Bayesian inference paradigm encourages the careful elicitation of prior information, and provides a coherent mechanism for updating this information as data becomes available [?]. The experimental observations and expert elicitation, however, must be independent but carefully aligned for this process to work.

For this analysis we carefully elicited opinions on the mortality rate and dispersal distance of WT and DSM male mosquitoes from four independent experts who were not involved in the development of the DSM mosquitoes [?]. Prior to the elicitation, experts were provided information on the genetic construct (originally incorporated into the G3 laboratory colony strain of *An. gambiae*) and were told that the genetically modified strain would be repeatedly backcrossed with wild type mosquitoes captured in the vicinity of a biosecure laboratory in order to gradually replace the G3 genetic background with that of the locally originated strain. The location of the laboratory was provided to each expert but the species of local WT mosquitoes was not specified. For the mortality rate, experts were proposed to answer questions about either the probability of mortality, the probability of survival, or the life expectancy. All responses were converted to an estimate of the daily probability of mortality which was then converted to the mortality rate parameter used in the PDE model.

The model's diffusion parameter (D) was not elicited directly but calculated using the experts' prior opinion on the average daily dispersal distance (d) of male mosquitoes. All experts agreed that a log-normal distribution provided the most appropriate description of their uncertainty for both WT and DSM strains. The model's prior diffusion parameter was subsequently calculated using the relation $D = d^2/\pi t$ (see Appendix for a detailed derivation). The geographical and temporal context of the elicitation was carefully prescribed, including a standardized time frame (dispersal over a one day period), so that all experts' opinions could be combined.

The model's two chemotaxis-related parameters are: (i) the relative strength of attraction towards the source of the attractant that a mosquito experiences at a location (α), with units $m^{-1}R^{-1}$ where m is the distance in meters between the location and the source and R is a measure of the "reward" at the source, often expressed as the concentration of CO_2 for female mosquitoes; and, (ii) the distance beyond which a mosquito is negligibly attracted to a swarm site, called the range (σ) with units m . The attractiveness of a swarm site at distance σ from a mosquito is roughly 38% of its attractiveness at zero distance.

Finding prior information on these parameters proved difficult for the particular setting of this model, in particular for α . We therefore chose a very broad prior distribution for α that spans several orders of magnitude, from 1% (1st percentile) to 140% (99th percentile) of the attraction at the source. For the range σ defined the same way we do, different articles on related species provide values ranging

from 3m [?], to 18m for single baiting or over 36m for double baiting [?], or even 40m [?]. We therefore used a log-normal distribution for this prior with positive probability across all these possibilities (Table 3).

2.3.2 Observation model

An important advantage of the Bayesian paradigm is that it allows us to develop more realistic hierarchical models that capture uncertainty in the biological and physical processes that drive a phenomenon, as well as the uncertainty associated with the measurement process with which these processes are observed. In this context, this allows us to reflect the uncertainty about how many mosquitoes any particular trap will catch given a known number of mosquitoes in the proximity of the trap.

Here we represent the observation model uncertainty using three, trap-specific, catchability parameters, denoted p_{cfr} for clay-pots, p_{psc} for insecticide spraying and p_{ss} for swarm sampling. These parameters define the probability of catching a mosquito given that it was in close proximity of the trap at the time of sampling [?]. Note that prior and posterior estimates of these parameters depend on the model resolution because the definition of “close proximity” is determined by the resolution of the raster (i.e. the grid cell size) over which the PDE model is numerically solved. Hence all predictions must be made at the same resolution used for solving the PDE and updating the prior distributions.

The priors for the catchability parameters were chosen according to the following principles:

- For p_{ss} , a beta distribution with parameters $\alpha_{ss} = 10$ and $\beta_{ss} = 20$ (meaning on average one third of the mosquitoes in the swarm are caught, following the standard operating procedures of the MRR experiment, which requires one third of the swarm to be captured).
- For p_{psc} , a beta distribution with parameters $\alpha_{psc} = 1.4$ and $\beta_{psc} = 1$. This prior was chosen because the PSC procedure involves spraying an entire room and capturing all the dead mosquitoes in the room. The PDE model, however, is specified at a resolution of the compounds which on average have 1.4 rooms. We assume that each room has on average the same number of male mosquitoes.
- For p_{cfr} , assuming this is the least efficient method but not having more information we deliberately selected a fairly low informative prior. So we chose a beta distribution with parameters $\alpha_{cfr} = 1$ and $\beta_{cfr} = 1$.

A complete list of process and observation model priors and their sources is summarised in Table 3.

2.4 Inference

2.4.1 WT male parameters

A Bayesian Hierarchical Model (BHM) is classically written:

$$\pi(\theta, \lambda | y) \propto \underbrace{\pi(y | \lambda)}_{data} \underbrace{\pi(\lambda | \theta)}_{process} \underbrace{\pi(\theta)}_{prior} \quad (4)$$

Parameter	Prior parameters	Prior distribution	Source
Mortality (μ)	$\theta_\mu = \{\alpha_\mu, \beta_\mu\}$	Beta	Experts
Diffusion (D)	$\theta_D = \{m_D, \sigma_D^2\}$	Log-normal	Experts
Swarm attraction (α)	$\theta_\alpha = \{m_\alpha, \sigma_\alpha^2\}$	Log-normal	Weakly informative
Swarm range (σ)	$\theta_\sigma = \{m_\sigma, \sigma_\sigma^2\}$	Log-normal	Literature
Catchability cfr (p_{cfr})	$\theta_{cfr} = \{\alpha_{cfr}, \beta_{cfr}\}$	Beta	Weakly informative
Catchability psc (p_{psc})	$\theta_{psc} = \{\alpha_{psc}, \beta_{psc}\}$	Beta	SOP
Catchability ss (p_{ss})	$\theta_{ss} = \{\alpha_{ss}, \beta_{ss}\}$	Beta	SOP

Table 3 Process (PDE model) and observation model parameters, prior distributions and sources. SOP: Standard Operating Procedures: the procedure used during the MRR experiments are described in a set of SOP that were made available to us as part of an independent risk assessment process.

where here θ is the vector of parameters of the PDE model and the trap catchability parameters, λ refers to the expected abundance of male mosquitoes in the vicinity of the traps, and y refers to the MRR data collected following the twelve releases ($r = 1, \dots, 12$), observed with three different collection (trap) techniques ($c = 1, \dots, 3$). Assuming conditional independence of the observations, the posterior distribution can be re-written as:

$$\pi(\theta_d, \lambda_r, p_c | y) \propto \prod_c \prod_r \left[\underbrace{\pi(y_{r,c} | \lambda_r, p_c)}_{data} \underbrace{\pi(\lambda_r | \theta_d)}_{process} \right] \underbrace{\pi(\theta_d, p_c)}_{prior} \quad (5)$$

where θ_d is the list of parameters defined in Table 1.

For a given release r , we describe the number of available released mosquitoes at time t as a Poisson random variable with expectation $\lambda_r(t)$, where the dependence of this expectation on location is suppressed:

$$n_r(t) | \lambda_r(t) \sim \text{Poisson}(\lambda_r(t)) \quad (6)$$

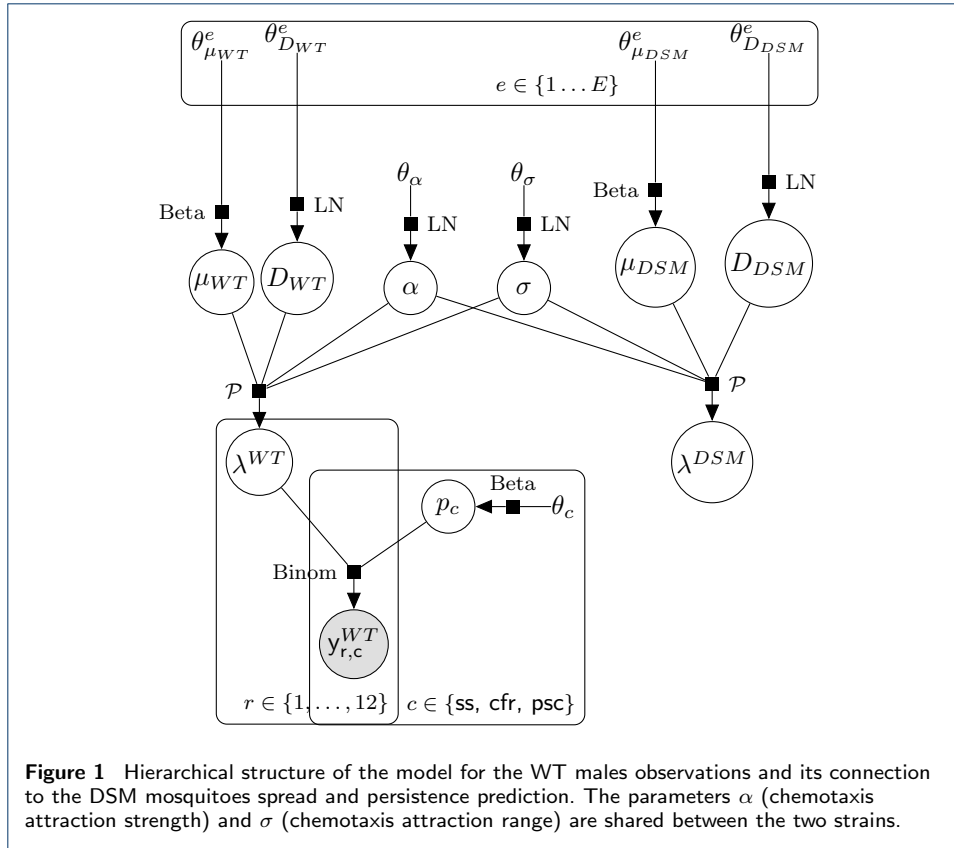
where $\lambda_r(t)$ is given by the PDE model described in Eq. 1. The data y are the observed count of male mosquitoes in a given trap. Since the traps catch only a portion of the mosquitoes that occur in their immediate vicinity, a natural trap observation model is the Binomial distribution:

$$y_{r,c} | n_r, p_c \sim \text{Binomial}(n_r, p_c) \quad (7)$$

where n_r is a realization of the Poisson random variable λ_r (the expected number of released mosquitoes in the trap's immediate vicinity), and p_c is the catchability parameter, or “trap efficiency”, which is trap dependent. In this study we do not account for the possibility of false positive (insects or mosquito species incorrectly identified as released male mosquitoes) or false negative (released male *An. gambiae* mosquitoes incorrectly identified as something else) probabilities as they are assumed to be equal to 0 [?].

Given the binomial likelihood, the posterior distribution is obtained by integrating over the possible Poisson outcomes:

$$\pi(\theta_d, \lambda_r, p_c | y) \propto \prod_c \prod_r \left[\sum_{n_r} \underbrace{\pi(y_{r,c} | n_r, p_c) \pi(n_r | \lambda_r)}_{data} \right] \underbrace{\pi(\lambda_r | \theta_d)}_{process} \underbrace{\pi(\theta_d)}_{prior} \quad (8)$$



Markov Chain Monte Carlo (MCMC) inference methods have been successfully applied to hierarchical Bayesian models, similar to the model described here, on many occasions [see for example ? ? ? ?]. We found, however, that the standard random-walk Metropolis MCMC routine was too slow to mix. Furthermore in this context, more advanced methods such as Hamiltonian Monte Carlo sampling, which require repeated likelihood computations along the proposal path, are inefficient because of the high computational costs entailed by the need to numerically solve the PDE model for each new proposal.

We found a reasonable compromise through the delayed rejection adaptive Metropolis algorithm DRAM, proposed by [?]. This MCMC algorithm uses a multivariate proposal distribution that automatically adapts to allow for posterior correlations between the parameters and identifies the directions of principal change along the ridges in the posterior surface. The acceptance rate of the DRAM algorithm is also improved by using a delayed rejection scheme where, instead of immediately advancing the chain following rejection of a parameter set, a second proposal is made that depends on both the current position of the chain and the rejected parameter set. We implemented DRAM by using the function modMCMC in the FME package [?] in R (R Core Team, 2015), using one delayed rejection step, and updating the proposal distribution every 200 iterations. We run a total of 3 MCMC chains of 15000 samples each, with 5000 used as burn-in. The convergence of the chains was assessed using Gelman's \hat{R} criterion [see ?, chapter 11].

2.5 Posterior prediction validation

We evaluated the accuracy of the posterior model predictions by comparing them against the observed recaptures in the MRR experiment that we deliberately withheld from the inference procedure. The fifth MRR experiment was conducted at the same location as the first four experiments, and under similar conditions with two exceptions: (i) all marked male mosquitoes were released at a single location; and, (ii) a much larger number of mosquitoes were released (Table 2). The experimental conditions and population dynamics might therefore be considered to be slightly different than those that prevailed during the first four experiments. If the model is nonetheless able to make sufficiently accurate predictions, then we may be more confident in its ability to be generalized to other similar release scenarios.

2.6 Bayesian model averaging

The field data for the WT male mosquito allows us to calculate the posterior distribution for the WT mortality and dispersal distance parameters. This in turn allows us to measure how well each expert's prior distribution matches the posterior distribution, and weight each expert's response accordingly. We do this by considering each expert's prior as an alternative model, and use the theory of model evidence [?] to calculate the Bayes factor [?].

We assume that experts who are good at predicting outcomes with WT mosquitoes (i.e. those whose prior distributions are close to the posterior distributions) will also be good at predicting outcomes with DSM mosquitoes, and weight the experts' DSM priors by the posterior probability of the models using Bayesian model averaging to obtain a weighted linear pool of expert opinion for the DSM mortality and dispersal distance.

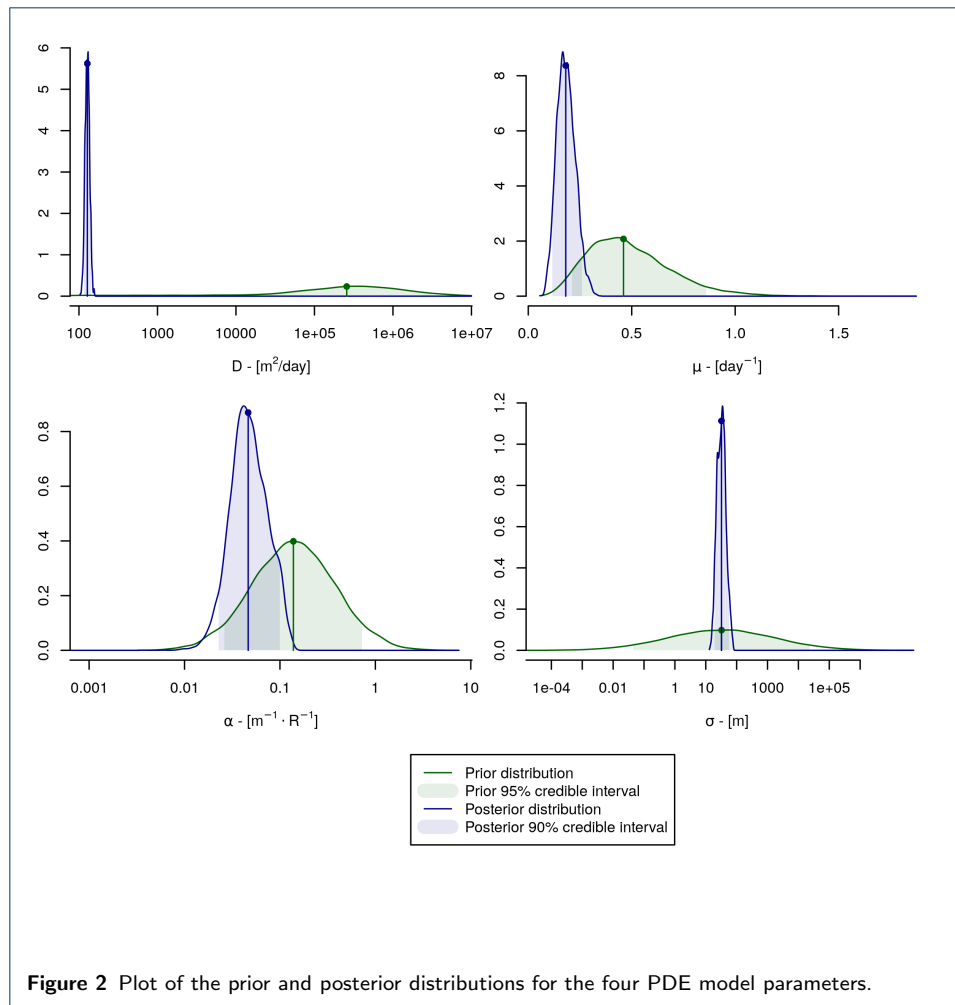
3 Results

3.1 WT parameters

Table 4 provides the summary statistics of the posterior distributions of the PDE model parameters (Figure 2 provides both the priors and posteriors for comparison). The use of the MRR data allowed us to provide refined estimates for the parameters of the PDE model. The diffusion coefficient in particular has seen its uncertainty decrease, yielding a mean value of 127 m² per day. The chemotaxis component has a posterior mean value of about 0.07, while the range parameter for the mosquito attraction is about 33.9 m.

Parameter	Mean	Q05	Q95
Mortality (μ)	0.16	0.11	0.24
Diffusion (D)	127.0	113.7	140.7
Swarm attraction (α)	0.07	0.03	0.10
Swarm range (σ)	33.9	19.8	56.1
Catchability psc (p_{psc})	0.18	0.11	0.25
Catchability cfr (p_{cfr})	0.03	0.02	0.05
Catchability ss (p_{ss})	0.29	0.24	0.34

Table 4 Summary statistics for the posterior distributions of the PDE model parameters inferred from the WT MRRs.



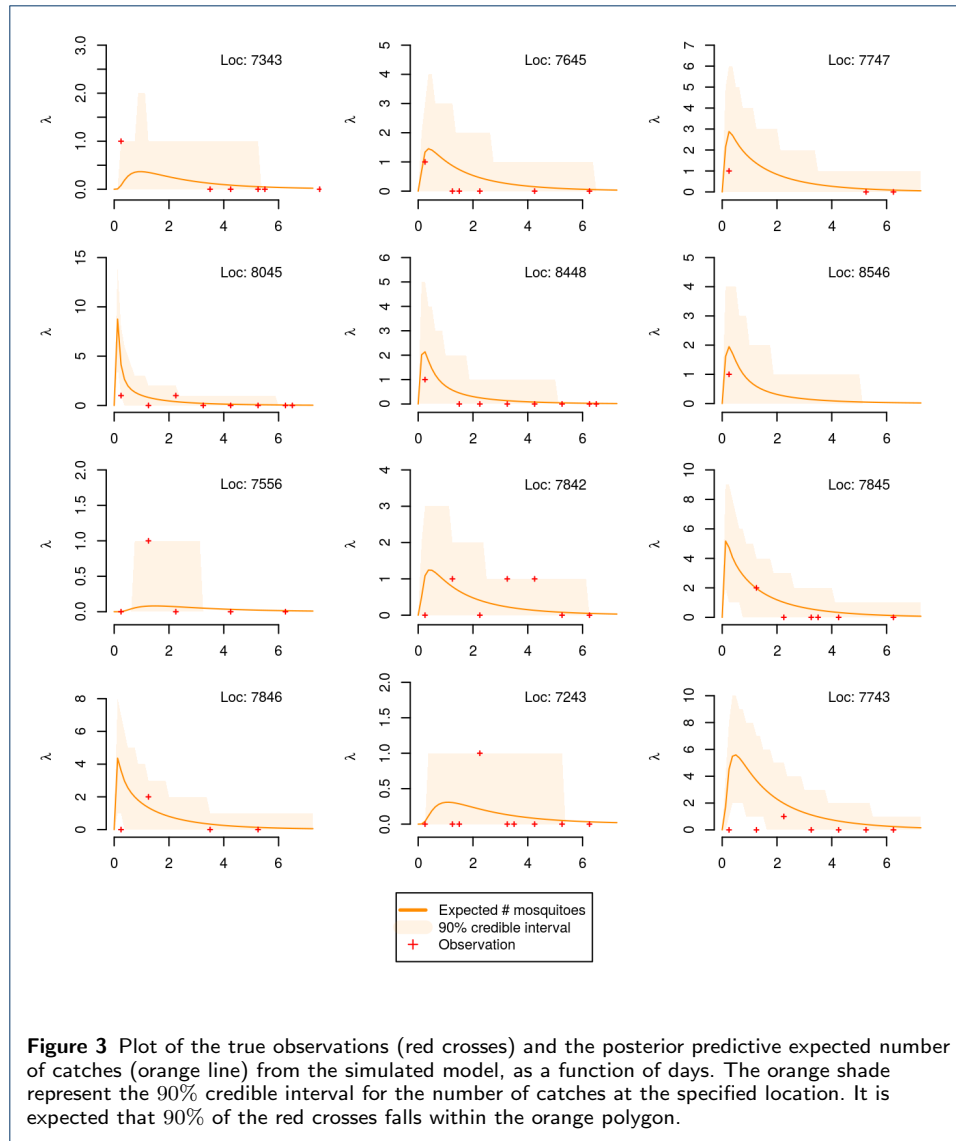
3.2 Spread and persistence of male WT mosquitoes

Figure 3 provides an overview of the expected evolution of the number of mosquitoes caught in the set of traps set for the fifth MRR. When comparing the model outcomes to the actual capture data, we note that of the 70 observations (number of mosquitoes caught in a given trap at a given time), only 2 are outside the 90% credible interval given by the model, providing a good level of confidence on the ability of the model to capture the general dynamic, and its ability to make usable predictions both in space and time about the spread and dispersal of WT strain mosquitoes.

3.3 DSM parameters

The weighted linear pool priors for DSM mortality and diffusion (weighted according to the Bayesian model averaging approach detailed above) are summarised in Table 5. The two parameters are quite different from the WT posterior estimates, with the DSM mortality prior six times higher than the WT posterior, and the dispersal multiple orders of magnitude bigger.

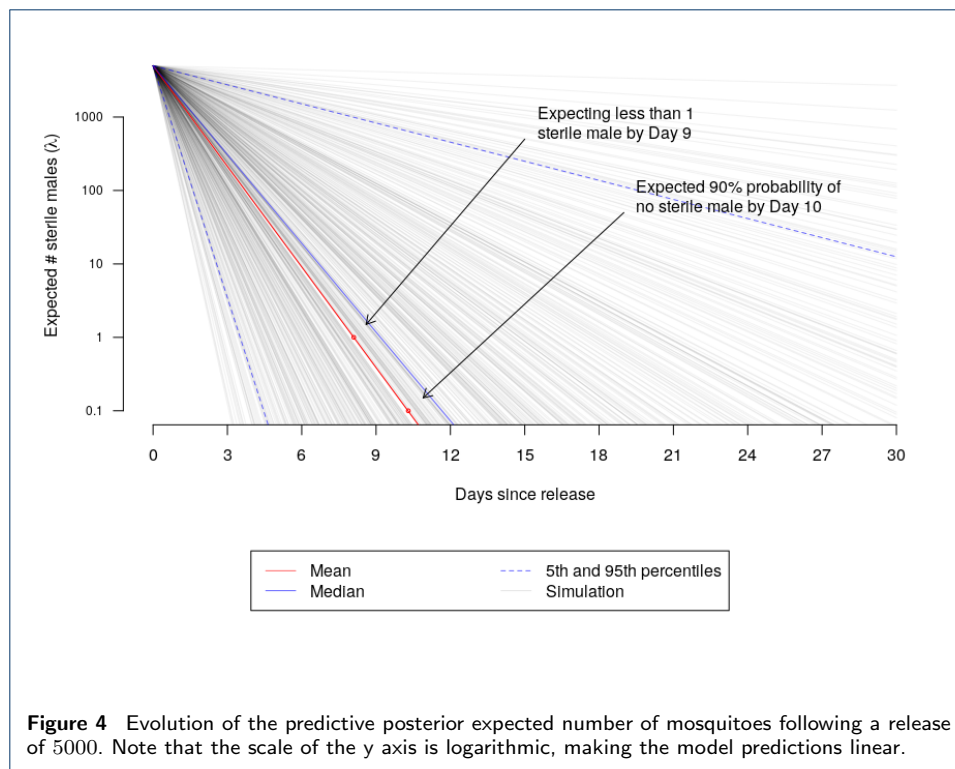
The difference between the posterior distributions for WT and DSM mortality reflects the effect of the Bayesian model averaging but also the significant reduction



in uncertainty that occurs when moving from prior to posterior distributions (see Figure 2) as often occurs in a Bayesian analysis. The linear pool of expert prior distribution for the dispersal distance of WT mosquitoes was also very much higher than its posterior distribution estimated using the data. It is possible therefore that the weighted linear pool prior for DSM dispersal will also prove to be an overestimate, noting of course that the posterior estimates of WT dispersal are conditional on the design of the MRR experiments which was established before this analysis.

Parameter	Mean	Q05	Q95
Mortality (μ)	1.05	0.19	2.42
Diffusion (D)	753×10^3	8.77	112×10^4

Table 5 Updated statistics for the PDE model parameters using the WT MRRs results and the DSM priors.



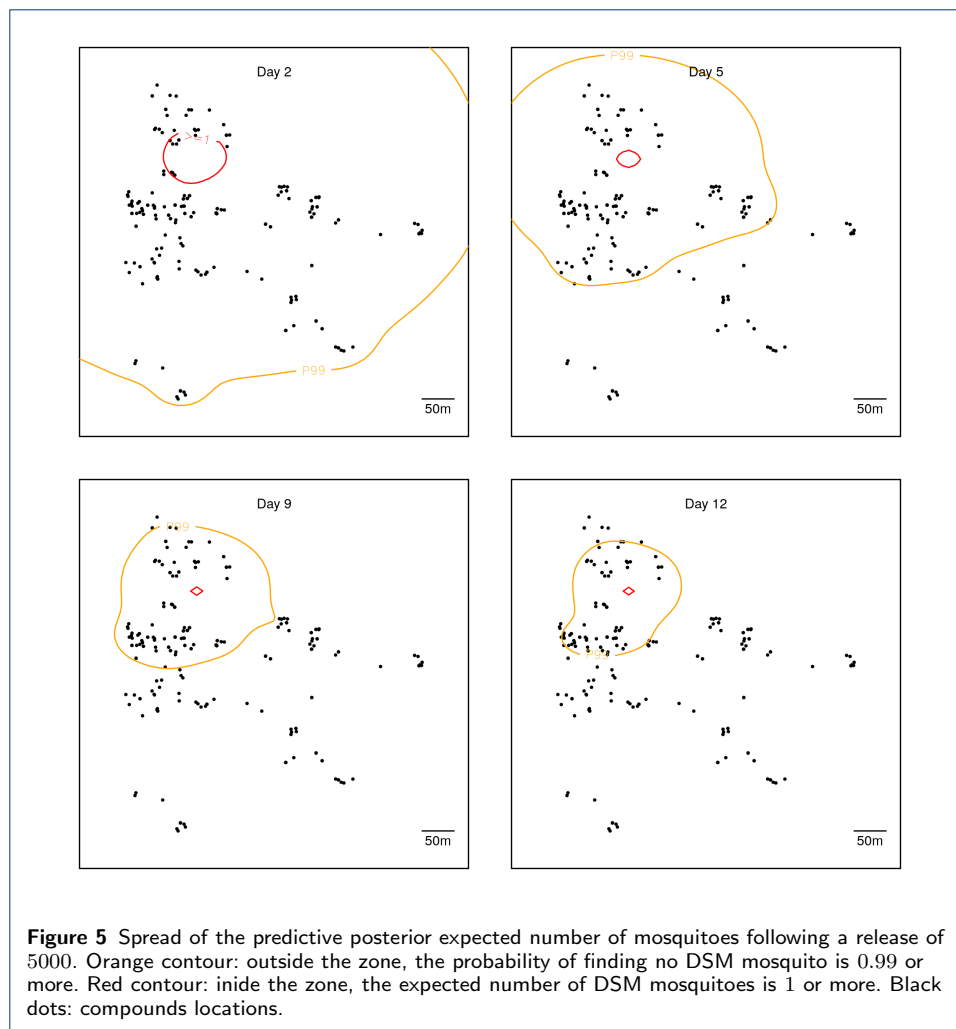
3.4 Spread and persistence of male DSM mosquitoes

The predicted persistence of DSM mosquitoes following a release of a single cohort of 5000 males is summarised in Figure 4. The large uncertainty captured in Table 5 is clearly reflected in these predictions: the mean expected survival is estimated to be about 9 to 10 days, with a 90% probability that the expected number of male DSM mosquitoes falls below 1 by Day 10. The 90% credible interval, however, is large, and there exists a very low probability (~ 0.05) that it could take as long about two months for this to occur.

The dispersal of the DSM mosquitoes is balanced by the mortality rate. Because of this, the DSM are not really expected to disperse far from their release site, with the expected number of mosquitoes per cell delineated by the red and orange contours in Figure 5. For instance, while the spread is expected to extend on average to up to 500m away from the release location by day 2 (with 1% chance that it extends further), the extent then contracts quickly to a limited area (less than 250m from release location) by day 5, to finally about 150m away from the release location at day 12.

4 Discussion

In this analysis we have used a combination of mathematical modeling, Bayesian inference, and expert elicitation to predict the spread and persistence of genetically modified DSM mosquitoes following the release of a single cohort of 5000 individuals. These predictions were made as part of an independent risk assessment that was completed in May 2018 prior to the proposed field release. The field release itself occurred in July 2019 (<https://targetmalaria.org/>



target-malaria-proceeded-with-a-small-scale-release-of-genetically-modified-sterile-male-mosq

We anticipate that the results from this field experiment will soon be forthcoming and provide an important opportunity to test the accuracy of the predictions presented here.

In this example we were able to validate the model predictions for wild type mosquitoes by holding back a proportion of the observation data. This also enabled us to use Bayesian model averaging to identify experts who were more accurate in their prior predictions. By allowing the opinions of these experts to carry greater weight we were also able to reduce uncertainty in the prior predictions for DSM mosquitoes. This approach assumes that experts who make more accurate predictions about wild type mosquitoes will also make better predictions about genetically modified mosquitoes. We will also be able to test this assumption once the results of the DSM field release are published.

Our results demonstrate how field observations greatly reduce the uncertainty between the prior information elicited with independent domain experts and the posterior distribution. Despite their relatively large uncertainty, however, our experience is that expert-derived prior distributions are essential when attempting to

run inference over the multi-dimensional parameter space that dynamic population models demand, and that careful elicitation will therefore continue to be an essential component of future risk assessment studies.

Our posterior estimates of WT dispersal are conditional on a series of MRR experiments that were conducted prior to the elicitation of the dispersal prior distributions. These priors played no part in the experimental design, and the difference between the prior and posterior distributions depends on this design. This difference may reflect conservative prior estimates of mosquitoes' mean daily dispersal when in the vicinity of attractants (swarm locations) but it may also be an artefact of the MRR design that focussed all recaptures in and around village compounds. As [?] notes further intensive sampling studies outside or around villages will be useful to ensure that posterior estimates of dispersal are not blind to long range dispersal outcomes that are not witnessed because of the recapture site design.

We believe that dynamic population models will form a central component of any risk-based governance system for gene drive modified mosquitoes. We therefore suspect that confidence in this governance structure will likely depend on the extent to which risk assessments are able to predict the spread and persistence of mosquitoes carrying gene drive constructs, within the bounds of adequate accuracy. Critically, a staged release strategy provides the opportunity to compare the predictions of these types of models against observed outcomes, learn how to approach the modeling and inference challenges and gradually define the bounds of accuracy that are regulators and stakeholders believe are adequate.

This staged-learning is borne out by this analysis: the reproductive containment strategy utilised in the DSM mosquitoes provides a good, relatively simple, starting point for this type of analysis. The reaction dynamics in our model are greatly simplified by virtue of the fact that the released mosquitoes are sterile. The dynamic models for the second stage in Target Malaria's pathway, namely self-limiting male bias strains, will be more complicated because they must accommodate the birth processes and genetic dynamics that are not relevant here, and also other potentially important inter-specific interactions [?]. The models for the third and final stage, that is self-sustaining driving male bias strains, will be more complicated again particularly because the inter-specific interactions are likely to be more important and because the spatio-temporal scope of the analysis will be significantly larger.

Our analysis demonstrates the strength of the Bayesian approach in this context. Given its qualities, we believe this paradigm is the most appropriate to handle the prediction and risk assessment challenges that novel, gene drive-based strategies for controlling malaria vectors will pose in the coming years.

Competing interests

The authors declare that they have no competing interests.

Author's contributions

AI carried out the original draft preparation, proposed the statistical model and ran the statistical analyses; SF wrote the code for the PDE model. KH and GH participated in expert elicitation. All authors participated on drafting and revision of the manuscript. All authors gave final approval for publication and agree to be held accountable for the work performed herein.

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Ethics

All elicitations were conducted with the informed consent of independent experts under ethics approval from the CSIRO Social Science Human Research Ethics Committee (reference LR06/2014). The mark release capture data was kindly provided by the Institut de Recherche en Sciences de la Santé, Bobo-Dioulasso, Burkina Faso, and was collected with approval from the local institutional ethics committee (Centre Muraz Institutional Ethics Committee), reference number 009–2012/CE-CM.

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Additional Files

4.1 Additional file 1 – Elicitation process, questions and answers

Elicitation of the mortality parameter The experts were interviewed in person and individually. They were asked the following question “What is the mortality rate of WT/DSM males mosquitoes once outside the insectary?”, where the mortality rate could be elicited in terms of probability of survival or death. Experts were able to also choose which quantiles of the PRIOR parameter distribution they wanted to elicit. Table 6 summarises the details of each expert's opinion. The experts answered either in terms of probability of survival (p_S) or death (p_D), which then we transformed into the mortality parameter μ of the PDE model using the relationship $\mu = -\log(1 - p_D)$, or in case of survival $\mu = -\log p_S$.

Mosquito	Parameter	Distribution	Quantiles	Values
Wild type	Mortality	beta	{0.1, 0.9}	{0.1, 0.2}
Wild type	Mortality	beta	{0.05, 0.95}	{0.055, 0.165}
Wild type	Survival	beta	{0.15, 0.85}	{0.5, 0.75}
Sterile male	Mortality	beta	{0.05, 0.95}	{0.066, 0.198}
Sterile male	Mortality	beta	{0.05, 0.95}	{0.2, 0.9}
Sterile male	Survival	beta	{0.15, 0.85}	{0.3, 0.6}

Table 6 Elicited values for the probability of death / survival.

Elicitation of the dispersion parameter The experts were asked the following question “What is the average daily dispersal distance of a WT/DSM male mosquito?”, where the dispersal distance was understood as the distance from the origin a mosquito could be found after a duration of 1 day. Experts were able to choose which quantiles of the prior parameter distribution they wanted to elicit. Table 7 summarises the details of each expert's opinion. The relationship between the elicited parameter d and the diffusion coefficient D is given below.

Mosquito	Distribution	Quantiles	Values
Wild type	lognormal	{0.05, 0.95}	{0.01, 2.9}
Wild type	lognormal	{0.25, 0.75}	{0.02, 0.2}
Sterile male	lognormal	{0.05, 0.95}	{0.008, 2.32}
Sterile male	lognormal	{0.25, 0.75}	{0.01, 0.1}

Table 7 Elicited values for the average daily dispersal.

4.2 Additional file 2 – Derivation of the relationship between the diffusion coefficient and the average dispersal distance

Let d be the average dispersal distance after t days. In order to determine what prior on D the prior elicited from d gives us, we need to establish the relationship between these two variables. Given the elicitation question, we have

$$d = \int_{\mathbb{R}_+} \sqrt{\langle x, x \rangle} \times p(x) dx$$

where $p(x)$ is the probability to find the mosquito at location x , and $\langle x, x \rangle$ is the Euclidean norm. This can be empirically estimated by

$$d = \frac{1}{n} \sum_{i=1}^n \sqrt{\langle x_i, x_i \rangle}$$

where x_i is the location of a mosquito i after t days, given it has been released from $(0, 0)$. If we note $f(x)$ the daily dispersal distance :

$$f(x) = \sqrt{\langle x, x \rangle}$$

we can then express the above equation in terms of expectation (in the random variable sense):

$$d = E[f(x)]$$

Because we assume that the mosquito behaviour is governed by the PDE model described in the model section of the article, we can write:

$$p(X_t = x) = \frac{1}{4\pi Dt} \exp \left\{ -\frac{\langle x, x \rangle}{4Dt} \right\}$$

where X_t is the random variable describing the location of that same released mosquito at time t , and x is a given location in \mathbb{R}^2 . So we have:

$$E[f(x)] = \int_{\mathbb{R}^2} \frac{\sqrt{\langle x, x \rangle}}{4\pi Dt} \exp \left\{ -\frac{\langle x, x \rangle}{4Dt} \right\} dx$$

We can perform a change of variable, $x_1 = r \cos(\theta)$ and $x_2 = r \sin(\theta)$ where $r \in [0, +\infty)$ and $\theta \in [0, 2\pi]$:

$$d = \int_{\mathbb{R}_+} \int_{[0, 2\pi]} \frac{r^2}{4\pi Dt} \exp\left\{-\frac{r^2}{4Dt}\right\} dr d\theta$$

which yields:

$$d = \int_{\mathbb{R}_+} \frac{r^2}{2Dt} \exp\left\{-\frac{r^2}{4Dt}\right\} dr$$

If we write $s^2 = 2Dt$, we recognize the expectation of a Rayleigh distribution, $E[r] = s\sqrt{\frac{\pi}{2}}$, and so we have:

$$d = \sqrt{Dt\pi}$$

or equivalently

$$D = \frac{d^2}{\pi t}$$

Note: it is often found in the literature that the diffusion coefficient can also be expressed in terms of mean squared dispersal distance d_2 . It can be approximated by $d_2 = \frac{1}{n} \sum^n \langle x_i, x_i \rangle$ - note that $d_2 \neq d^2$ - and using the same development we have $D = d_2/4t$.