[3.11] Let N(+) be the number of arrivals in [01t]. (a) Ni(t) be the number of arrivals in [oit] in line 1 N2(t) be the number of arrivals in [0,t] in line 2

Note that N(+) = N,(+)+ N2(+).

$$P(N_{1}(t)=n, N_{2}(t)=m) = P(N_{1}(t)=n, N_{2}(t)=m, N(t)=n+m)$$

$$= P(N(t)=n+m) P(N_{1}(t)=n, N_{2}(t)=m | N(t)=n+m)$$

= N(t) is a Poisson process with rate = 2 » $P(N(t) = v + w) = \frac{(v + w)!}{(v + w)!} e^{-yt}$

Define $X_i = \begin{cases} 1 & \text{if the ith arrival is routed to line 1} \\ 0 & \text{if the ith arrival is routed to line 2} \\ R.V. \end{cases}$

$$\Rightarrow x_i = \begin{cases} 1 & \text{with prob. } P \\ 0 & \text{with prob. } I-P \end{cases}$$

$$P(N_1(t)=n, N_2(t)=m \mid N(t)=n+m) = P(\sum_{i=1}^{n+m} x_i = n)$$

or each packet is routed independently

$$\sum_{i=1}^{n+m} x_i$$
 has a binomial distribution

i.e.
$$\mathbb{P}\left(\sum_{i=1}^{n+m} x_i = m\right) = \binom{n+m}{n} p^n (1-p)^m = \frac{(n+m)!}{n! m!} p^n (1-p)^m$$

Therefore
$$P(N_{1}(t)=n, N_{2}(t)=m) = \frac{(\lambda t)^{n+m}}{(n+m)!} e^{-\lambda t} \cdot \frac{(n+m)!}{n!} p^{n} (1-p)^{m}$$

$$= \frac{(\beta \lambda t)^{n}}{n!} e^{-\beta \lambda t} \cdot \frac{(n+m)!}{n!} e^{-(1-p)\lambda t}^{m} e^{-(1-p)\lambda t}$$

$$P(N_{1}(t)=n) = \sum_{m=0}^{\infty} P(N_{1}(t)=n, N_{2}(t)=m) = \frac{(\beta \lambda t)^{n}}{n!} e^{-\beta \lambda t} \longrightarrow Poisson(rate=p\lambda)$$

$$P(N_{1}(t)=m) = \sum_{m=0}^{\infty} P(N_{1}(t)=n, N_{2}(t)=m) = \frac{(n+m)!}{n!} e^{-(1-p)\lambda t} Poisson(rate=p\lambda)$$

 $P(N_2(+) = m) = \sum_{n=0}^{\infty} P(N_1(+) = n, N_2(+) = m) = ((1-p)\lambda +)^m e^{-(1-p)\lambda +} - Poisson$ (rate = (1-p)\lambda)

NI(t) and N2(t) are independent