

# 7. Model error and structural uncertainty

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Bayesian Methods in Health Economics, Lausanne

- Probabilistic sensitivity analysis to structural assumptions
  - Model average vs model comparison
  - Structural PSA using DIC
  - Example

## References


 *Bayesian Methods in Health Economics*, chapters 3.6, 4.7.2

 [Book website \(CRC\)](#)

[Book website](#)

[Code](#)

 *The BUGS Book*, chapter 8.6

 [Book website](#)

## Problem

- All the methods discussed so far assume that, while there may be uncertainty in the current knowledge, the model is “correct”
  - All variables that need to be considered have been considered
  - All distributional assumptions are reasonable (in fact, “correct”)
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- Possible (potential) solutions
  - Bayesian model averaging
  - Model comparison

- “Proper” Bayesian way to deal with the issue
  - Identify a set of *finite and exhausting models*  $\mathcal{M} = (\mathcal{M}_1, \dots, \mathcal{M}_H)$
  - Each of the  $H$  models is characterised by a set of parameters  $\boldsymbol{\theta}$  and a suitable prior  $p(\boldsymbol{\theta} \mid \mathcal{M}_h)$
  - In addition, define a prior  $p(\mathcal{M}_h)$  that model  $h$  is the “true” one
  - Update these as

$$p(\mathcal{M}_h \mid y) \propto p(\mathcal{M}_h) \int p(y \mid \boldsymbol{\theta}, \mathcal{M}_h) p(\boldsymbol{\theta} \mid \mathcal{M}_h) d\boldsymbol{\theta}$$

and use the posterior probabilities to compute a weighted average for any function of the parameters (eg utilities) over the space of models

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- **NB:** requires that *all* possible models are completely specified and given a prior probability

- Compare a (not necessarily exhaustive!) set of models in terms of their out-of-sample prediction
  - Quantifies how well the predictive distribution for a given model would fit a replicated dataset based on the observed data
  - **NB:** especially in health economic evaluations, the possible models considered as merely a (rough) approximation to the complex phenomenon under study – so no guarantee that any of those should be the “true” one!



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- **One** possibility is the Deviance Information Criterion (DIC)

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## Structural PSA via DIC

- Consider  $\mathcal{M} = (\mathcal{M}_1, \dots, \mathcal{M}_H)$ 
  - 1 Compute the value  $\text{DIC}_h$  for each of them
  - 2 Derive the respective weights (model posterior probabilities) by simply re-proportioning them, eg

$$w_h = \frac{\exp(-0.5\Delta\text{DIC}_h)}{\sum_{h=1}^H \exp(-0.5\Delta\text{DIC}_h)}, \quad \Delta\text{DIC}_h = |\min_h(\text{DIC}_h) - \text{DIC}_h|$$

- 3 Use the weights  $w_h$  to build an average model accounting for the (limited set) of plausible models

Review: Statins  
Comparison: 91 Placebo-controlled studies: CHD death plus non-fatal MI  
Outcome: 01 CHD death plus non-fatal MI

Study or subcategory	Treatment n/N	Control n/N
<b>01 Atorvastatin</b>		
DAL <sup>86</sup>	0/145	1/72
ASCOT-LLA <sup>102</sup>	100/5168	154/5137
Mohler <sup>21</sup>	7/240	3/114
CARDS <sup>103</sup>	43/1428	65/1410
Subtotal (95% CI)	6981	6733
Total events: 150 (treatment), 223 (control)		
Test for heterogeneity: $\chi^2 = 1.32$ , $df = 3$ ( $p = 0.973$ ), $I^2 = 0\%$		
Test for overall effect: $z = 4.09$ ( $p < 0.0001$ )		
<b>02 Fluvastatin</b>		
FLARE <sup>108</sup>	6/409	17/425
LISA <sup>93</sup>	2/187	5/178
LIPS <sup>110</sup>	42/844	60/833
Subtotal (95% CI)	1440	1436
Total events: 50 (treatment), 82 (control)		
Test for heterogeneity: $\chi^2 = 1.91$ , $df = 2$ ( $p = 0.39$ ), $I^2 = 0\%$		
Test for overall effect: $z = 2.76$ ( $p = 0.006$ )		
<b>03 Pravastatin</b>		
PMSG <sup>96</sup>	0/530	7/532
KAPS <sup>133</sup>	5/224	8/223
WOSCOPS <sup>82</sup>	174/3302	248/3293
CAIUS <sup>107</sup>	2/151	2/154
CARE <sup>111</sup>	212/2081	274/2078
LIPID <sup>112</sup>	557/4512	715/4502
PROSPER <sup>81</sup>	292/2891	356/2913
Subtotal (95% CI)	13,691	13,695
Total events: 1242 (treatment), 1610 (control)		
Test for heterogeneity: $\chi^2 = 4.95$ , $df = 6$ ( $p = 0.55$ ), $I^2 = 0\%$		
Test for overall effect: $z = 7.19$ ( $p < 0.00001$ )		
<b>04 Simvastatin</b>		
4S <sup>97</sup>	431/2221	622/2223
CIS <sup>98</sup>	2/129	7/125
HPS <sup>74</sup>	898/10,269	1212/10,267
Subtotal (95% CI)	12,619	12,615
Total events: 1331 (treatment), 1841 (control)		
Test for heterogeneity: $\chi^2 = 2.38$ , $df = 2$ ( $p = 0.30$ ), $I^2 = 15.9\%$		
Test for overall effect: $z = 8.29$ ( $p < 0.00001$ )		
<b>Total (95% CI)</b>	<b>34,731</b>	<b>34,479</b>
Total events: 2773 (treatment), 3756 (control)		
Test for heterogeneity: $\chi^2 = 15.28$ , $df = 16$ ( $p = 0.50$ ), $I^2 = 0\%$		
Test for overall effect: $z = 12.94$ ( $p < 0.00001$ )		

- Based on published data on RCTs comparing statins to placebo
- Complex formulation: in particular, model the response for the controls as

$$y_{sj}^{\text{ctr}} \sim \text{Binomial}(\theta_s, n_{sj}^{\text{ctr}})$$

$$\text{logit}(\theta_s) = \alpha_s$$

and then use 2 formulations for the prior on  $\alpha_s$

## 1 Normal

- $\alpha_s \sim \text{Normal}(\mu_\alpha, \tau_\alpha)$
- $\mu_\alpha \sim \text{Normal}(0, 0.00001)$
- $\sigma_\alpha = \tau_\alpha^{-2} \sim \text{Uniform}(0, 20)$

## 2 Half-Cauchy (robust alternative)

- $\alpha_s = \mu_\alpha + \xi \eta_s$
- $\mu_\alpha \sim \text{Normal}(0, 0.00001)$
- $\xi \sim \text{Normal}(0, \tau_\xi)$
- $\eta_s \sim \text{Normal}(0, \tau_\eta)$
- $\tau_\eta \sim \text{Gamma}(0.5, 0.5); \tau_\xi = 12^{-2}$

## Model $\mathcal{M}_1$

```
> ...  
> ## Priors for the hyperparameters  
> ## Exchangeable normal prior  
> for (s in 1:Nstatins) {  
+   alpha[s] ~ dnorm(mu.alpha, tau.alpha)  
+ }  
> sigma.alpha ~ dunif(0,20)  
> tau.alpha <- pow(sigma.alpha,-2)  
> mu.alpha ~ dnorm(0,0.0001)  
> sigma.alpha <- abs(xi) / sqrt(tau.eta)  
> prior.scale <- 12  
> mu.alpha ~ dnorm(0,.0001)  
> ...
```

## Model $\mathcal{M}_2$

```
> ...  
> ## Priors for the hyperparameters  
> ## Half-Cauchy prior  
> for (s in 1:Nstatins) {  
+   alpha[s] <- mu.alpha + xi*eta[s]  
+   eta[s] ~ dnorm(0,tau.eta)  
+ }  
> xi ~ dnorm(0,tau.xi)  
> tau.xi <- pow(prior.scale,-2)  
> tau.eta ~ dgamma(.5,.5)  
> sigma.alpha <- abs(xi) / sqrt(tau.eta)  
> prior.scale <- 12  
> mu.alpha ~ dnorm(0,.0001)  
> ...
```

`print(m1)` → R object containing the MCMC simulations for  $\mathcal{M}_1$

Inference for Bugs model at "model.txt", fit using jags , 2 chains, each with 100000 iterations (first 9500 discarded), n.thin = 181 n.sims = 1000 iterations saved

	mean	sd	2.5%	25%	50%	75%	97.5%	Rhat	n.eff
cost.hosp[1]	238.7	137.0	91.7	168.6	213.5	278.2	482.0	1	980
cost.hosp[2]	315.6	168.8	124.2	219.2	290.2	372.7	668.5	1	1000
cost.hosp[3]	523.1	451.0	144.0	297.0	427.9	619.9	1358.0	1	1000

DIC info (using the rule,  $pD = \text{var}(\text{deviance})/2$ )

$pD = 71.1$  and  $DIC = 2233.9$

DIC is an estimate of expected predictive error (lower deviance is better)

`print(m2)` → R object containing the MCMC simulations for  $\mathcal{M}_2$

Inference for Bugs model at "model.txt", fit using jags , 2 chains, each with 100000 iterations (first 9500 discarded), n.thin = 181 n.sims = 1000 iterations saved

	mean	sd	2.5%	25%	50%	75%	97.5%	Rhat	n.eff
cost.hosp[1]	228.9	108.0	82.9	161.6	209.3	272.9	502.5	1	1000
cost.hosp[2]	301.5	143.0	101.2	211.2	276.6	366.5	621.8	1	960
cost.hosp[3]	469.8	277.1	128.6	280.5	402.1	583.2	1184.6	1	1000

DIC info (using the rule,  $pD = \text{var}(\text{deviance})/2$ )

$pD = 63.9$  and  $DIC = 2226.0$

DIC is an estimate of expected predictive error (lower deviance is better)

```
> library(BCEA)
> # Objects containing the MCMC simulations from the posteriors for the two models
> m1.sims <- m1$sims.list;
> m2.sims <- m2$sims.list
>
> # Defines suitable variables of clinical benefits & costs
> e1 <- m1.sims$effect;
> e2 <- m2.sims$effect
> c1 <- m1.sims$cost.tot;
> c2 <- m2.sims$cost.tot
>
> # Runs BCEA's function to do PSA to structural assumptions
> avg <- struct.psa(list(m1,m2),list(e1,e2),list(c1,c2),ref=2)
```

```
> # Weights associated with each model (based on DIC)
> avg$w
```

```
[1] 0.01901127 0.98098873
```

```
> # Actual DIC computed for each model
> avg$DIC
```

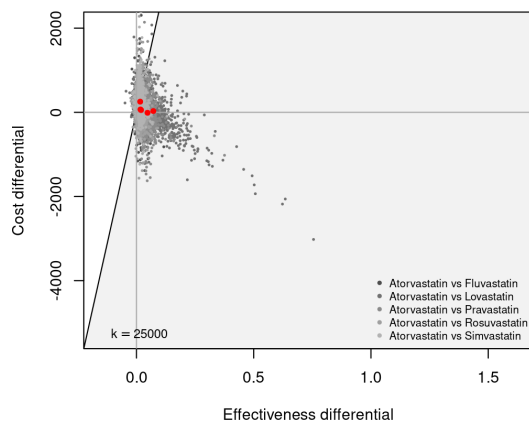
```
[1] 2233.875 2225.988
```

Model 1 (Normal prior)

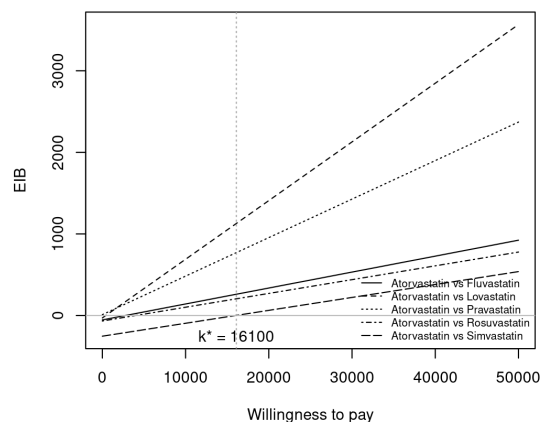
Model 2 (HC prior)

Model average

Cost-Effectiveness Plane

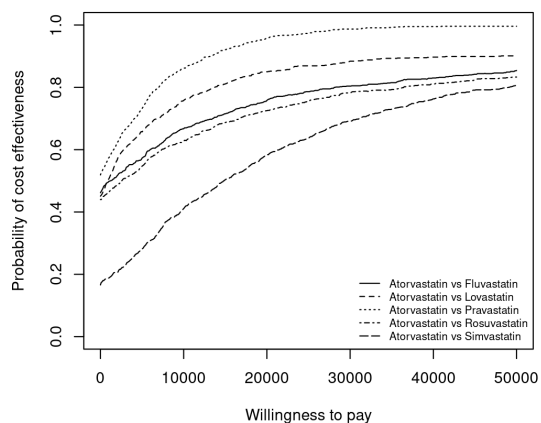


Expected Incremental Benefit

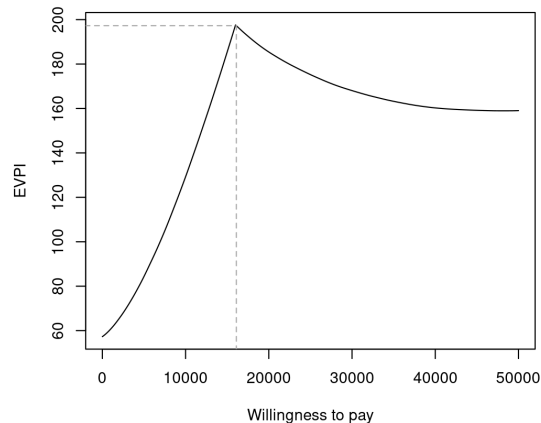


- $DIC = 2233.875$
- $w_1 = 0.019$

Cost Effectiveness Acceptability Curve



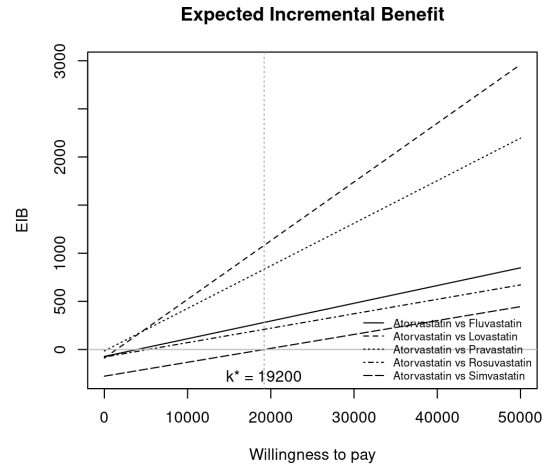
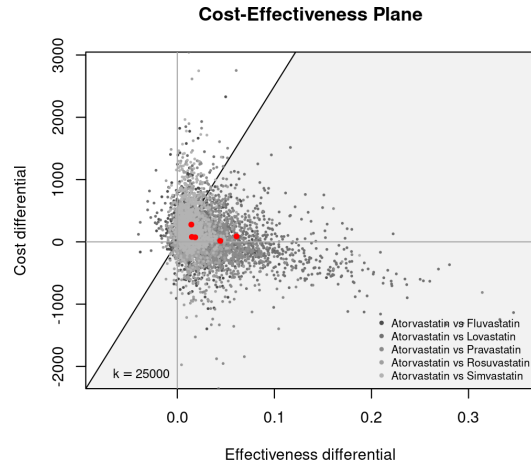
Expected Value of Information



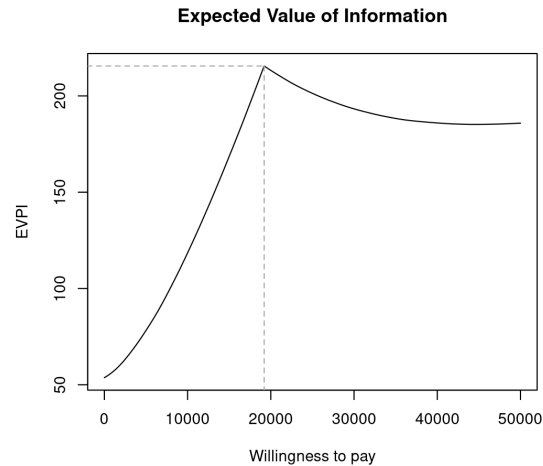
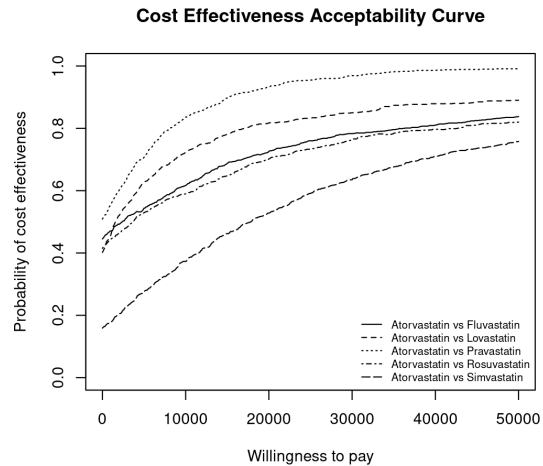
Model 1 (Normal prior)

Model 2 (HC prior)

Model average



- $DIC = 2225.988$
- $w_2 = 0.981$

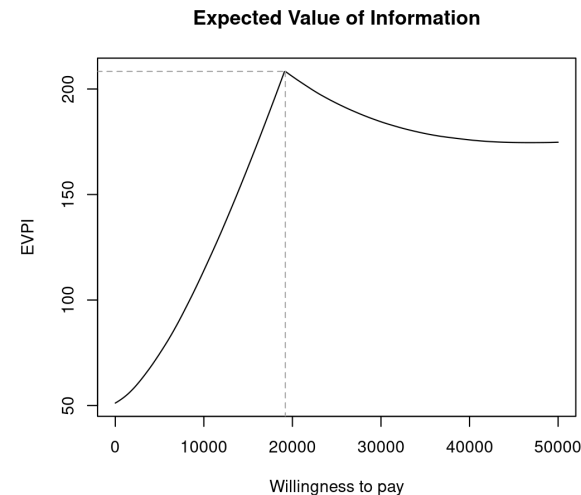
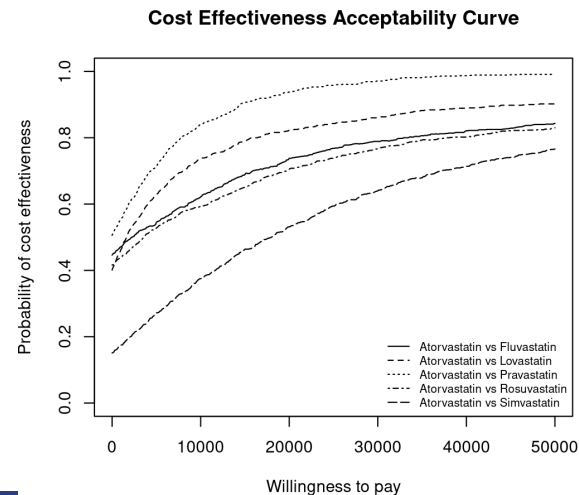
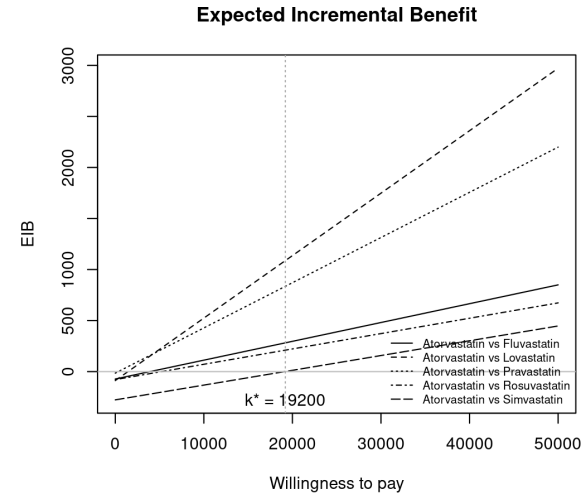
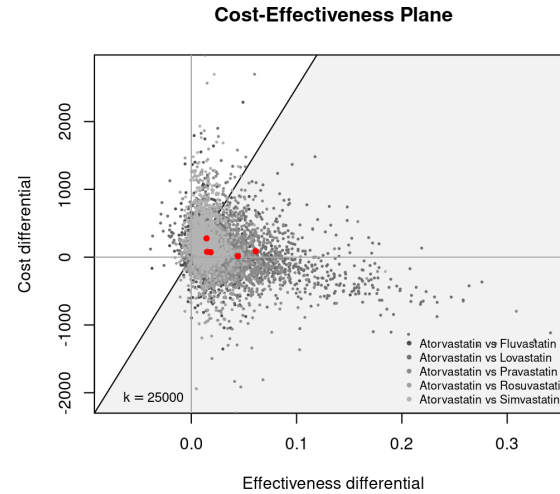




Model 1 (Normal prior)

Model 2 (HC prior)

Model average



## Model $\mathcal{M}_2$

Analysis for willingness to pay parameter  $k = 25000$

	Expected utility
Atorvastatin	22824
Fluvastatin	22437
Lovastatin	21389
Pravastatin	21734
Rosuvastatin	22528
Simvastatin	22741

	EIB	CEAC	ICER
Atorvastatin vs Fluvastatin	387.804	0.760	3943.45
Atorvastatin vs Lovastatin	1435.830	0.833	1455.71
Atorvastatin vs Pravastatin	1090.236	0.955	348.42
Atorvastatin vs Rosuvastatin	296.181	0.733	5254.11
Atorvastatin vs Simvastatin	83.889	0.591	19195.85

Optimal intervention (max expected utility)  
for  $k = 25000$ : Atorvastatin

EVPI 201.33

## Model average

Analysis for willingness to pay parameter  $k = 25000$

	Expected utility
Atorvastatin	22823
Fluvastatin	22435
Lovastatin	21381
Pravastatin	21731
Rosuvastatin	22526
Simvastatin	22738

	EIB	CEAC	ICER
Atorvastatin vs Fluvastatin	388.695	0.768	3919.60
Atorvastatin vs Lovastatin	1442.154	0.843	1432.44
Atorvastatin vs Pravastatin	1092.152	0.958	336.72
Atorvastatin vs Rosuvastatin	297.304	0.738	5227.07
Atorvastatin vs Simvastatin	85.002	0.594	19129.49

Optimal intervention (max expected utility)  
for  $k = 25000$ : Atorvastatin

EVPI 193.22