

# Quant II

## Lab 3: DAG review, bootstrap, clustering

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February 9, 2023

# Today's plan

- DAG Review

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- Bootstrap

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- DAG Review
- Bootstrap
- Clustering

# Housekeeping

- Homework 1 due February 14 before class
- Email to me, cc Cyrus
- Send code and output (.Rmd or pdf + file)
- Math can be scanned and attached (or from tablet), no need to compile in LaTeX

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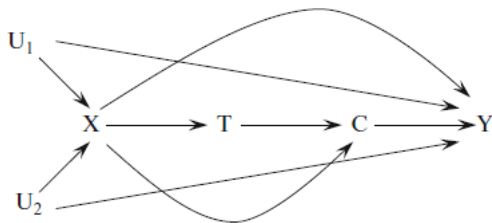
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- Acyclic: cannot come back to starting point

Important concepts:

- $A \rightarrow B$ : B is *children* of A, A is *parent* of B
- $A \rightarrow C \leftarrow B$ : C is a *collider*
- Causal path (relative to T): arrows go from T towards Y
- Non-causal path (relative to T): the others
- Backdoor path: non-causal path between T and Y
- Colliders block non-causal paths
- Conditioning on colliders open new non-causal paths

# DAG Review



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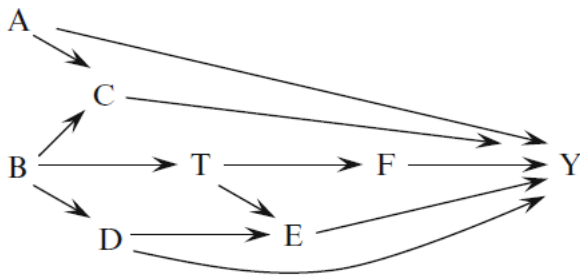
- If  $w$  is a collider, link all pairs of parents of  $w$  by drawing an undirected edge between them
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- Erase  $w$  from the graph and all the edges connected with  $w$

# DAG Review



# Collider bias

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- Random re-sampling from the sample approximates random sampling from the population
- The sampling distribution of statistics from re-samples approximates the true sampling distribution



# Non-parametric bootstrap

- ➊ From our sample of size  $N$ , draw a random sample of size  $N$  with replacement
- ➋ On this sample, compute the estimate
- ➌ Repeat many times (e.g. 1000)
- ➍ Obtain a distribution of estimates from the resamples (a bootstrapped sampling distribution)
- ➎ For inference, compute moments of the bootstrapped distribution:  
e.g. sd, quantiles

# Non-parametric bootstrap

## Advantages:

- No distributional assumptions
- Inference on functions of estimators

```
set.seed(123)
# Population
pop <- rnorm(10000)

# Sample
S <- sample(pop, 500)

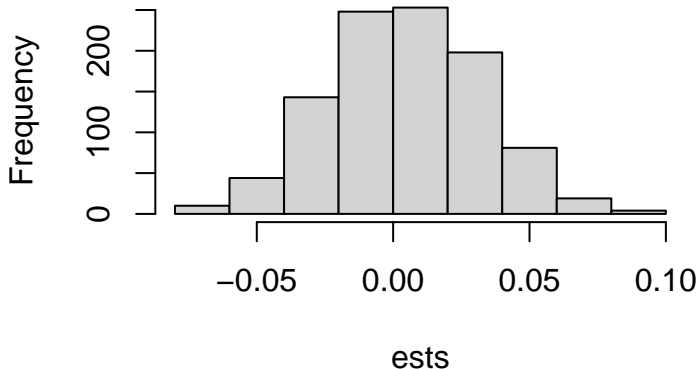
# Estimate Mean(x) - Median(x)
delta <- mean(S)-median(S)

# Bootstrap inference
nboot <- 1000
ests <- rep(NA, nboot)
for (i in 1:nboot){
  ids <- sample(1:length(S), length(S), replace=T)
  s <- S[ids]
  ests[i] <- mean(s)-median(s)
}
```

# Non-parametric bootstrap

```
# Bootstrapped distribution  
hist(ests, main="Bootstrapped sampling distribution")
```

## Bootstrapped sampling distribution



# Non-parametric bootstrap

```
# SE  
sd(ests)
```

```
## [1] 0.02840985
```

```
# CI  
delta + quantile(ests, c(0.025, 0.975))
```

```
##          2.5%          97.5%  
## -0.03465047  0.07370951
```

# Non-parametric bootstrap

We can also use the boot function

```
set.seed(123)
library(boot)
# Function for the bootstrap estimate
bootfn <- function(data, id){
  s <- data[id]
  out <- mean(s)-median(s)
  out
}

bs <- boot(S, statistic = bootfn, R=1000)
```

# Non-parametric bootstrap

```
# Output:
```

```
bs
```

```
##  
## ORDINARY NONPARAMETRIC BOOTSTRAP  
##  
##  
## Call:  
## boot(data = S, statistic = bootfn, R = 1000)  
##  
##  
## Bootstrap Statistics :  
##      original      bias      std. error  
## t1* 0.01479699 -0.01066407  0.03066007
```

```
# Bias:
```

```
mean(bs[["t"]]) - delta
```

```
## [1] -0.01066407
```

```
# Std. error
```

```
sd(bs[["t"]])
```

```
## [1] 0.03066007
```

# Bootstrap in regression

- Code manually
- `lm` + `sandwich`

```
library(sandwich);library(lmtest)
set.seed(000)
data("mtcars")
fit <- lm(mpg ~ qsec, mtcars)
coeftest(fit, vcov=vcovBS(fit, cluster=NULL, R=1000))
```

```
##
## t test of coefficients:
##
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -5.11404    8.47717  -0.6033 0.550862
## qsec        1.41212    0.48884   2.8887 0.007116 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

# Bootstrap in regression

## • lm + car

```
library(car)
set.seed(000)
Boot(fit, f=coef, R=1000)
```

```
##
## ORDINARY NONPARAMETRIC BOOTSTRAP
##
##
## Call:
## boot::boot(data = dd, statistic = boot.f, R = R, .fn = f, parallel = p_type,
##   ncpus = ncores, cl = cl2)
##
##
## Bootstrap Statistics :
##      original      bias    std. error
## t1* -5.114038 -1.07603420   9.1139840
## t2*  1.412125  0.05996816   0.5247496
# You can replace coef with more complex functions
```



# Bootstrap in regression

In Stata:

```
** Regression with bootstrap SEs
set seed 000
reg mpg qsec, vce(bootstrap, reps(1000))
```

Output:

Linear regression	Number of obs	=	32
	Replications	=	1,000
	Wald chi2(1)	=	9.01
	Prob > chi2	=	0.0027
	R-squared	=	0.1753
	Adj R-squared	=	0.1478
	Root MSE	=	5.5637

mpg	Observed Coef.	Bootstrap Std. Err.	z	P> z	Normal-based [95% Conf. Interval]	
qsec	1.412125	.4705516	3.00	0.003	.4898608	2.334389
_cons	-5.11404	8.153858	-0.63	0.531	-21.09531	10.86723

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- In both cases, relative to simple random process, units are contributing less information
- Extreme case: perfect correlation
- If we treat units as independent, we overstate the information in the data, and underestimate the variance of the estimator



# Guide to empirical practice

Source: [MacKinnon, Nielsen, and Webb \(JE, 2023\)](#)

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- Treatment assigned at a higher level: cluster at that level
  - E.g. individual-level data, geographic distances at place level
- Tests for cluster level (Ibragimov and Mueller): estimate model for each cluster and compare variation in estimates with the one under finer clusters. Only works if treatment varies within cluster

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  - Influential cluster (analogous to single observations)
- Recommended to plot distribution of estimates from exclusion of each cluster sequentially

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  - Example: Delaware contains 50+% of US incorporations
- The opposite (a few small clusters) is generally not problematic
- Only a few clusters have treated observations

# Summary guide

- ① List all possible clustering dimensions and levels for the data, make informed decision regarding clustering structure
- ② For each possible level, report the n. of clusters and the distribution of cluster sizes
- ③ For key regression specifications, report info on leverage, partial leverage and influence, especially the effective number of clusters for coefficients of interest  $\implies$  `summc` in Stata
  - Is leverage collinear with cluster size?
- ④ In addition to the CV1 estimator (the most common), use also the jack-knife based CV3 ( $\implies$  `summc`) and WCR ( $\implies$  `boottest`)
- ⑤ If treatment at cluster level and few treated/control clusters, check with other methods (e.g. randomization inference)

If you have many clusters of similar size and good treatment variation across clusters you are fine

# Summclust

```
summclust mw, yvar(hours2) xvar(black female) fevar(educ  
age year state) cluster(state) rho(0.5) jack
```

SUMMCLUST - MacKinnon, Nielsen, and Webb

Cluster summary statistics for mw when clustered by statefip.  
There are 492827 observations within 51 statefip clusters.

Regression Output

s.e.	Coeff	Sd. Err.	t-stat	P value	CI-lower	CI-upper
CV1	-0.153891	0.062314	-2.4696	0.0170	-0.279051	-0.028730
CV3	-0.153891	0.067127	-2.2925	0.0261	-0.288720	-0.019061
CV3J	-0.153891	0.067127	-2.2925	0.0261	-0.288720	-0.019062

Cluster Variability

Statistic	Ng	Leverage	Partial L.	beta no g
min	258.0	1.015250	0.000837	-0.177645
q1	2495.0	1.136757	0.003932	-0.156898
median	7082.0	1.413468	0.009442	-0.153682
mean	9663.3	1.549020	0.019608	-0.153896
q3	13481.0	1.759903	0.020038	-0.150470
max	35995.0	3.058464	0.160196	-0.122695
coefvar	0.9	0.331260	1.481140	0.062300

Effective Number of Clusters

-----  
G\*(0) is 16.18657

**Table 3**  
Summary statistics for cluster heterogeneity.

Clustering	$G$	$G_{\beta}^*(0)$	$\tilde{N}_g$	min.	1st quart.	Median	3rd quart.	max.
Hours data: $N = 492,827$								
State-year	765	79.4	644	6	176	480	860	3,052
State	51	16.2	9,663	258	2,495	7,082	13,481	35,995
Year	15	6.6	32,855	28,262	28,839	30,733	40,224	40,394
Region	9	7.5	54,759	27,849	37,396	50,489	65,389	96,337
Employment and student data: $N = 1,531,360$								
State-year	765	66.0	2,002	42	524	1,413	2,426	10,794
State	51	13.1	30,027	927	7,363	22,845	37,020	144,914
Year	15	6.5	102,091	92,701	95,589	102,319	108,858	110,528
Region	9	7.0	170,151	74,172	104,703	181,767	208,099	291,955

**Notes:** The values of  $G_{\beta}^*(0)$  are calculated using 28 regressors after the state dummies have been partialled out. The  $\beta$  subscript emphasizes the fact that they correspond to the coefficient  $\beta$  in (40). Because there are state fixed effects, values of  $G_{\beta}^*(1)$  are not reported; see [MacKinnon et al. \(2022b\)](#).



# Cluster-robust standard errors in regression

```
library(sandwich)
# Work with MNW example dataset
d <- read.csv("min_wage_teen_hours2.csv")

# lm + sandwich
fit <- lm(hours2 ~ mw + black + female + factor(educ) + factor(age) + factor(year) + factor(statefip), d)

coeftest(fit, vcov=vcovCL(fit, cluster = ~ statefip))["mw",]

##      Estimate Std. Error    t value    Pr(>|t|)
## -0.15389068  0.06231373 -2.46961125  0.01352633
```

# Cluster-robust standard errors in regression

```
# lfe
library(lfe)
fit <- felm(hours2 ~ mw + black + female | educ + age + year + statefip | 0 | statefip, d)
summary(fit)

##
## Call:
##   felm(formula = hours2 ~ mw + black + female | educ + age + year +      statefip | 0 | statefip, data = d)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -35.152  -6.934  -1.124   5.994  47.645
##
## Coefficients:
##      Estimate Cluster s.e. t value Pr(>|t|)
## mw      -0.15389      0.06231  -2.470   0.017 *
## black    0.98832      0.11714   8.437 3.54e-11 ***
## female  -2.61304      0.08368 -31.227 < 2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 9.887 on 492748 degrees of freedom
## Multiple R-squared(full model): 0.1927   Adjusted R-squared: 0.1926
## Multiple R-squared(proj model): 0.01753   Adjusted R-squared: 0.01738
## F-statistic(full model, *iid*): 1508 on 78 and 492748 DF, p-value: < 2.2e-16
## F-statistic(proj model): 357.1 on 3 and 50 DF, p-value: < 2.2e-16
## *** Standard errors may be too high due to more than 2 groups and exactDOF=FALSE
```

# Cluster-robust standard errors in regression

```
# fixest
library(fixest)
fit <- feols(hours2 ~ mw + black + female | educ + age + year + statefip, cluster=~statefip, data=d)
etable(fit, keep = c("mw"))
```

```
##                               fit
## Dependent Var.:             hours2
##
## mw                        -0.1539* (0.0623)
## Fixed-Effects:  -----
## educ                                Yes
## age                                Yes
## year                                Yes
## statefip                            Yes
## -----
## S.E.: Clustered      by: statefip
## Observations          492,827
## R2                     0.19270
## Within R2              0.01753
```

# Cluster-robust standard errors in regression

In Stata:

```
vce(cluster clustervar)
```

# Cluster bootstrap

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- Re-estimate the model for all parameters and find  $\hat{\beta}_b$
- Compute  $t$ -statistic  $t_b$  from the new estimates
- Repeat many times
- Compute bootstrap p-values by counting the share of simulated  $t_b$  to the left/right of the observed one  $t$

# Cluster bootstrap

- In Stata: `boottest`
- Computes bootstrap p-value and confidence interval, but not the standard error

```
. boottest mw, cluster(styear) noci reps(999) seed(123)
```

Overriding estimator's cluster/robust settings with `cluster(styear)`

Wild bootstrap-t, null imposed, 999 replications, Wald test, clustering by `styear`, bootstrap clustering by  
> `styear`, Rademacher weights:

```
mw
```

```
      t(764) =    -3.3823  
Prob>|t| =      0.0020
```

```
.
```

# Cluster bootstrap

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  - Function `boottest` works with objects of class `lm`, `felm`, `fixest`

# Cluster bootstrap

In R:

- Option `cluster` in `sandwich::vcovBS()`
- The package `fwildclusterboot` is a translation of Stata's `boottest` (same options)
  - Function `boottest` works with objects of class `lm`, `felm`, `fixest`
- Another option is the package `multiwaycov` and the function `cluster.boot` which can be used for post-estimation SE calculation (e.g. in `coeftest` or `stargazer`)