

$$\textcircled{1} \quad P(|X_n - c| \geq \epsilon) \leq \frac{E(X_n - c)^2}{\epsilon^2} \quad \text{by Chebyshev's inequality}$$

$$= \frac{E(X_n - EX_n + EX_n - c)^2}{\epsilon^2}$$

$$= \frac{E(X_n - EX_n)^2 + (EX_n - c)^2}{\epsilon^2}$$

Since  $EX_n \rightarrow c$  and  $VX_n \rightarrow 0$  as  $n \rightarrow \infty$

$$\frac{E(X_n - EX_n)^2 + (EX_n - c)^2}{\epsilon^2} = \frac{V(X_n) + (EX_n - c)^2}{\epsilon^2}$$

$$\Rightarrow P(|X_n - c| \geq \epsilon) \rightarrow 0 \quad \text{as } n \rightarrow \infty \quad \forall \epsilon > 0.$$

$$\Rightarrow X_n \xrightarrow{P} c.$$

$$\textcircled{2} \quad S_n = \sum_{i=1}^n X_i \quad ; \quad \text{Take } a_n = \sum \mu_i \quad \text{and } b_n = n$$

$$\frac{S_n - a_n}{b_n} = \frac{\sum (X_i) - \sum \mu_i}{n}$$

$$P\left(\left|\frac{S_n - a_n}{b_n}\right| \geq \epsilon\right) = P\left(\left|\frac{\sum X_i - \sum \mu_i}{n}\right| \geq \epsilon\right)$$

$$\leq \frac{E(\sum X_i - \sum \mu_i)^2}{n^2 \epsilon^2}$$

$$= \frac{E(\sum X_i - E(\sum X_i))^2}{n^2 \epsilon^2}$$

$$= \frac{V(\sum X_i)}{n^2 \epsilon^2} \rightarrow 0 \quad \text{as } n \rightarrow \infty \quad \forall \epsilon > 0.$$

from the given condition

$$\Rightarrow \frac{s_n - a_n}{b_n} \xrightarrow{p} 0$$

$\Rightarrow$  WLLN holds for  $\{x_n\}$ .

$$\begin{aligned} \text{Furthermore, } \frac{s_n - a_n}{b_n} &= \frac{\sum x_i}{n} - \frac{\sum \mu_i}{n} \\ &= \bar{X}_n - \bar{\mu}_n \xrightarrow{p} 0 \\ \text{i.e. } \bar{X}_n &\xrightarrow{p} \bar{\mu}_n. \end{aligned}$$

③  $x_1, \dots, x_n$  i.i.d  $U(0,1)$

$$Y_n = \min(x_1, \dots, x_n) ; \quad Z_n = \max(x_1, \dots, x_n)$$

$$\begin{aligned} \text{d.f. } F_{Y_n}(y) &= P(\min(x_1, \dots, x_n) \leq y) \\ &= 1 - P(\min(x_1, \dots, x_n) > y) \\ &= 1 - (1 - F_X(y))^n \end{aligned}$$

$$\text{p.d.f. } f_{Y_n}(y) = n(1 - F_X(y))^{n-1} f_X(y) ; \quad 0 < y < 1$$

$$\text{i.e. } f_{Y_n}(y) = \begin{cases} n(1-y)^{n-1} & ; \quad 0 < y < 1 \\ 0 & \text{w.} \end{cases}$$

$$P(|Y_n| > \epsilon) \leq \frac{E Y_n^2}{\epsilon^2}$$

$$\begin{aligned} \text{Now } E Y_n^2 &= n \int_0^1 y^2 (1-y)^{n-1} dy = n \int_0^1 (1-x)^2 x^{n-1} dx \\ &= n \int_0^1 (1+x^2-2x) x^{n-1} dx = n \left( \frac{1}{n} + \frac{1}{n+2} - \frac{2}{n+1} \right) \\ &\quad \rightarrow 0 \text{ as } n \rightarrow \infty \end{aligned}$$

$$\Rightarrow \frac{E Y_n^2}{\epsilon^2} \rightarrow 0 \text{ as } n \rightarrow \infty \quad \forall \epsilon > 0. \text{ however small.}$$

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$$\Rightarrow P(|Y_n| > \epsilon) \rightarrow 0 \text{ as } n \rightarrow \infty$$

$$\Rightarrow Y_n \xrightarrow{P} 0 \Rightarrow \sqrt{Y_n} \xrightarrow{P} 0$$

p.d.f. of  $Z_n$  :  $f_{Z_n}(z) = \begin{cases} n z^{n-1}, & 0 < z < 1 \\ 0, & \text{o.w.} \end{cases}$

$$P(|Z_n - 1| > \epsilon) \leq \frac{E(Z_n - 1)^2}{\epsilon^2} = \frac{E Z_n^2 + 1 - 2E(Z_n)}{\epsilon^2}$$

$$E Z_n = n \int_0^1 z^n dz = \frac{n}{n+1}$$

$$E Z_n^2 = n \int_0^1 z^{n+1} dz = \frac{n}{n+2}$$

$$\Rightarrow \frac{E(Z_n - 1)^2}{\epsilon^2} = \frac{1}{\epsilon^2} \left( \frac{n}{n+2} + 1 - 2 \frac{n}{n+1} \right) \rightarrow 0 \text{ as } n \rightarrow \infty$$

$\forall \epsilon > 0$  fixed.

$$\Rightarrow P(|Z_n - 1| > \epsilon) \rightarrow 0 \text{ as } n \rightarrow \infty$$

$$\Rightarrow Z_n - 1 \xrightarrow{P} 0 \text{ i.e. } Z_n \xrightarrow{P} 1$$

$$\Rightarrow Z_n^2 \xrightarrow{P} 1 \quad \left( \begin{array}{l} \text{If } X_n \xrightarrow{P} X \\ \text{for } g \text{ cont} \\ g(X_n) \xrightarrow{P} g(X) \end{array} \right)$$

Since  $Y_n \xrightarrow{P} 0$  &  $Z_n^2 \xrightarrow{P} 1$

$$Y_n \cdot Z_n^2 \xrightarrow{P} 0 \quad \left( \begin{array}{l} \text{If } X_n \xrightarrow{P} x \text{ & } Y_n \xrightarrow{P} y \\ \text{then } X_n Y_n \xrightarrow{P} xy \end{array} \right)$$

④  $X_1, \dots, X_n$  i.i.d.  $N(0,1)$

④

$$\begin{aligned} \text{WLLN} &\Rightarrow \frac{1}{n} \sum X_i \xrightarrow{P} EX_1 = 0 \\ (\text{Khintchine's WLLN}) &\quad \text{i.e. } \bar{X}_n \xrightarrow{P} 0 \end{aligned}$$

$$S_n^2 = \frac{1}{n} \sum X_i^2 - \bar{X}_n^2$$

Note that  $X_1^2, X_2^2, \dots, X_n^2$  are i.i.d. with  $E(X_1^2) = 1$

$$\begin{aligned} \text{WLLN} &\Rightarrow \frac{1}{n} \sum X_i^2 \xrightarrow{P} EX_1^2 = 1 \\ (\text{Khintchine's WLLN}) & \end{aligned}$$

$$\text{Since } \bar{X}_n \xrightarrow{P} 0 \Rightarrow \bar{X}_n^2 \xrightarrow{P} 0.$$

$$\Rightarrow S_n^2 = \frac{1}{n} \sum X_i^2 - \bar{X}_n^2 \xrightarrow{P} 1 - 0 = 1$$

$$\Rightarrow S_n^{-1} \xrightarrow{P} 1$$

$$\text{Since } \bar{X}_n \xrightarrow{P} 0 \quad \& \quad S_n^{-1} \xrightarrow{P} 1$$

$$\bar{X}_n S_n^{-1} \xrightarrow{P} 0 (= o_p(1))$$

⑤

$$Y_n \sim \text{Bin}(n, p)$$

$$\begin{aligned} P\left(\left|\frac{Y_n}{n} - p\right| > \epsilon\right) &\leq \frac{E\left(\frac{Y_n}{n} - p\right)^2}{\epsilon^2} = \frac{E(Y_n - np)^2}{n^2 \epsilon^2} \\ &= \frac{npq}{n^2 \epsilon^2} = \frac{pq}{n \epsilon^2} \rightarrow 0 \quad \text{as } n \rightarrow \infty \end{aligned}$$

$$\Rightarrow \frac{Y_n}{n} \xrightarrow{P} p$$

$$\Rightarrow \left(1 - \frac{Y_n}{n}\right) \xrightarrow{P} \cancel{p} \cdot (1-p)$$

$$(6) \quad E(X_n) = 0; \quad V(X_n) = E X_n^2 = \frac{\sqrt{n}}{2} + \frac{\sqrt{n}}{2} = \sqrt{n}$$

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$$E \bar{X}_n = 0; \quad V(\bar{X}_n) = \frac{1}{n^2} \sum_{i=1}^n \sqrt{i} \leq \frac{n \sqrt{n}}{n^2} \rightarrow 0 \text{ as } n \rightarrow \infty$$

$$\Rightarrow P(|\bar{X}_n - 0| > \epsilon) \leq \frac{E \bar{X}_n^2}{\epsilon^2} = \frac{V \bar{X}_n}{\epsilon^2} \rightarrow 0 \text{ as } n \rightarrow \infty \forall \epsilon > 0$$

$$\Rightarrow \bar{X}_n \xrightarrow{P} 0.$$

$$(7) (a) \quad Y_n = \frac{2}{n(n+1)} \sum_{i=1}^n i X_i$$

$$E Y_n = \frac{2}{n(n+1)} \sum_{i=1}^n i \mu = \mu$$

$$V Y_n = \frac{4}{n^2(n+1)^2} \sum_{i=1}^n i^2 \sigma^2 = \frac{4\sigma^2}{n^2(n+1)^2} \cdot \frac{n(n+1)(2n+1)}{6} \rightarrow 0 \text{ as } n \rightarrow \infty$$

$$\Rightarrow P(|Y_n - \mu| > \epsilon) \leq \frac{E(Y_n - \mu)^2}{\epsilon^2} = \frac{V(Y_n)}{\epsilon^2} \rightarrow 0 \text{ as } n \rightarrow \infty$$

$$\Rightarrow Y_n \xrightarrow{P} \mu.$$

(b) By (a)

$$(8) \quad X_1, \dots, X_n \text{ i.i.d. } U(0,1)$$

$$\text{Let } T_i = -\log_3 X_i \quad i = 1, \dots, n$$

$$T_i \sim \text{Exp}(0,1) \quad f_{T_i}(t) = \begin{cases} e^{-t} & t > 0 \\ 0, & \text{o.w.} \end{cases}$$

$$E T_i = 1$$

$T_1, \dots, T_n$  are i.i.d.  $\text{Exp}(0,1)$  with  $E T_i = 1$

$$Z_n = \left( \prod_{i=1}^n X_i \right)^{1/n}$$

$$-\log Z_n = \frac{1}{n} \sum_{i=1}^n (-\log X_i) = \frac{1}{n} \sum_{i=1}^n T_i$$

$$\text{WLLN} \Rightarrow \frac{1}{n} \sum_{i=1}^n T_i \xrightarrow{P} E(T_i) = 1$$

$$\text{i.e.} \quad -\log Z_n \xrightarrow{P} 1$$

$$\Rightarrow Z_n \xrightarrow{P} e^{-1}.$$

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$$(9) \quad E X_i = \mu_i, \quad V(X_i) = \sigma_i^2$$

$$\text{if } \sum_{i=1}^n \sigma_i^2 \rightarrow \infty, \text{ take } a_n = \sum_{i=1}^n \mu_i$$

$$\text{ \& } b_n = \sum_{i=1}^n \sigma_i^2$$

$$\text{then } P\left(\left|\frac{S_n - a_n}{b_n}\right| > \epsilon\right) = P\left(\left|\frac{1}{\sum \sigma_i^2} \sum (X_i - \mu_i)\right| > \epsilon\right)$$

$$\leq \frac{E\left(\sum (X_i - \mu_i)\right)^2}{\left(\sum \sigma_i^2\right)^2 \epsilon^2}$$

$$= \frac{\sum \sigma_i^2}{\epsilon^2 \left(\sum \sigma_i^2\right)^2} = \frac{1}{\epsilon^2 \sum \sigma_i^2} \rightarrow 0 \text{ as } n \rightarrow \infty$$

$$\Rightarrow \frac{S_n - a_n}{b_n} \xrightarrow{P} 0$$

i.e. WLLN holds for  $\{X_n\}$ .

$$(10) \quad X_1, \dots, X_n \text{ i.i.d. } U(0,1) \text{ with } E X_i = \frac{1}{2}$$

$$\text{W.L.L.N } \Rightarrow \frac{1}{n} \sum X_i \xrightarrow{P} E X_1 = \frac{1}{2}$$

$$\text{i.e. } \bar{X}_n \xrightarrow{P} \frac{1}{2}$$

(6)

$$\begin{aligned}
 \textcircled{11} \quad F_{X_n}(x) &= P(X_n \leq x) \\
 &= P\left(\frac{X_n - 1/n}{\sqrt{1 - \frac{1}{n}}} \leq \frac{x - \frac{1}{n}}{\sqrt{1 - \frac{1}{n}}}\right) \\
 &= \Phi\left(\frac{x - \frac{1}{n}}{\sqrt{1 - \frac{1}{n}}}\right) \rightarrow \Phi(x) \text{ as } n \rightarrow \infty \\
 &\Rightarrow X_n \xrightarrow{d} X \sim N(0, 1)
 \end{aligned}$$

Alt m.g.f of  $X_n$

$$\begin{aligned}
 M_{X_n}(t) &= \exp\left(t/n + \frac{t^2}{2}\left(1 - \frac{1}{n}\right)\right) \\
 &\rightarrow e^{t^2/2} \leftarrow \text{m.g.f of } N(0, 1). \\
 &\Rightarrow X_n \xrightarrow{d} X \sim N(0, 1).
 \end{aligned}$$

$$\begin{aligned}
 \textcircled{12} \quad Y_i &= (X_i - \mu)^2 \\
 E(Y_i) &= E(X_i - \mu)^2 = \sigma^2 \\
 V(Y_i) &= E\left((X_i - \mu)^2 - \sigma^2\right)^2 \\
 &= E(X_i - \mu)^4 + \sigma^4 - 2\sigma^2 E(X_i - \mu)^2 \\
 &= (\sigma^4 + 1) + \sigma^4 - 2\sigma^4 = 1
 \end{aligned}$$

i.e.  $E(Y_i) = \sigma^2$ ;  $V(Y_i) = 1 \quad \forall i$  &  $Y_1, \dots, Y_n$  i.i.d

$$\begin{aligned}
 S_n &= \sum Y_i & ES_n &= n\sigma^2 \\
 & & VS_n &= n
 \end{aligned}$$

$$\text{CLT} \Rightarrow \frac{S_n - ES_n}{\sqrt{VS_n}} \xrightarrow{d} N(0, 1)$$

$$\text{i.e.} \quad \frac{(X_1 - \mu)^2 + \dots + (X_n - \mu)^2 - n\sigma^2}{\sqrt{n}} \xrightarrow{d} X \sim N(0, 1)$$

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$$\begin{aligned}
& \lim_{n \rightarrow \infty} P\left(\sigma^2 - \frac{1}{\sqrt{n}} \leq \frac{(X_1 - \mu)^2 + \dots + (X_n - \mu)^2}{n} \leq \sigma^2 + \frac{1}{\sqrt{n}}\right) \\
&= \lim_{n \rightarrow \infty} P\left(-\frac{1}{\sqrt{n}} \leq \frac{(X_1 - \mu)^2 + \dots + (X_n - \mu)^2 - n\sigma^2}{n} \leq \frac{1}{\sqrt{n}}\right) \\
&= \lim_{n \rightarrow \infty} P\left(-1 \leq \frac{(X_1 - \mu)^2 + \dots + (X_n - \mu)^2 - n\sigma^2}{\sqrt{n}} \leq 1\right) \\
&= \Phi(1) - \Phi(-1) = 2\Phi(1) - 1 = \dots
\end{aligned}$$

(13)  $E S_n = np$ ;  $V S_n = npq$

$$\begin{aligned}
P\left(\left|\frac{S_n}{n} - p\right| \geq t\right) &\leq \frac{E(S_n - np)^2}{t^2 n^2} = \frac{np(1-p)}{n^2 t^2} \leq \frac{1}{4n t^2} \leq 0.01 \quad (\text{given}) \\
\Rightarrow n &\geq \frac{1}{0.04 t^2} \quad \text{for } t = 0.01 \\
n &\geq \dots
\end{aligned}$$

(14)

CLT  $\Rightarrow \frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma} \xrightarrow{L} Z \sim N(0, 1)$

Also  $S_n^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X})^2 \xrightarrow{P} \sigma^2$

$$\begin{aligned}
S_n^2 &= \frac{1}{n} \sum X_i^2 - \bar{X}^2 \\
&\xrightarrow{P} \sigma^2 + \mu^2 \quad \xrightarrow{P} \sigma^2 \\
&\Rightarrow S_n \xrightarrow{P} \sigma \\
&\text{using Slutsky's lemma} \\
&\left( \begin{array}{l} X_n \xrightarrow{L} x ; Y_n \xrightarrow{P} c \\ X_n/Y_n \xrightarrow{L} x/c \end{array} \right) \\
&\frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma} \xrightarrow{L} X \sim N(0, 1) \\
&\text{i.e. } \frac{\sqrt{n}(\bar{X}_n - \mu)}{S_n} \xrightarrow{L} X \sim N(0, 1)
\end{aligned}$$



(8)

(15)  $X_1, \dots, X_{72}$  r.i.s. from  $f(x) = \begin{cases} \frac{1}{x^2} & x > 1 \\ 0 & \text{o/w} \end{cases}$

Define  $Y_i = \begin{cases} 1 & \text{if } X_i < 3 \\ 0 & \text{o/w} \end{cases}$

$$P(Y_i = 1) = P(X_i < 3) = \int_1^3 \frac{1}{x^2} dx = \frac{2}{3} = \theta \text{ say}$$

$Y_1, \dots, Y_{72}$  are i.i.d  $B(1, \theta)$

$$Y = \sum_{i=1}^{72} Y_i \sim B(72, \theta = 2/3)$$

$$CLT \Rightarrow \frac{Y - 72 \times \frac{2}{3}}{\sqrt{72 \times \frac{2}{3} \times \frac{1}{3}}} \xrightarrow{L} Z \sim N(0, 1)$$

$$\text{i.e. } \frac{Y - 48}{4} \xrightarrow{L} Z \sim N(0, 1)$$

$$P(Y > 50) = 1 - P(Y \leq 50) = 1 - P(Y \leq 50.5) \quad \leftarrow \text{continuity correction}$$

$$= 1 - P\left(\frac{Y - 48}{4} \leq \frac{50.5 - 48}{4}\right)$$

$$\approx 1 - \Phi\left(\frac{2.5}{4}\right) = \dots$$

(16)  $X_1, \dots, X_{100}$  i.i.d  $P(3)$

$$E(X_i) = 3 ; V(X_i) = 3$$

$$Y = \sum_{i=1}^{100} X_i \sim P(300) \Rightarrow \begin{matrix} E(Y) = 300 \\ V(Y) = 300 \end{matrix}$$

$$CLT \Rightarrow \frac{Y - 300}{10\sqrt{3}} \left( = \frac{S_n - ES_n}{\sqrt{VarS_n}} \right) \xrightarrow{L} N(0, 1)$$

$$P(100 \leq Y \leq 200) = P(99.5 \leq Y \leq 200.5) \quad \leftarrow \text{cont correction}$$

$$= P\left(\frac{99.5 - 300}{10\sqrt{3}} \leq \frac{Y - 300}{10\sqrt{3}} \leq \frac{200.5 - 300}{10\sqrt{3}}\right)$$

$$\approx \Phi\left(\frac{200.5 - 300}{10\sqrt{3}}\right) - \Phi\left(\frac{99.5 - 300}{10\sqrt{3}}\right)$$

$$= \dots$$

①⑦  $X \sim \text{bin}(100, 0.6)$

⑨

$$\text{CLT} \Rightarrow \frac{X - 100 \times 0.6}{\sqrt{100 \times 0.6 \times 0.4}} = \frac{X - 60}{\sqrt{24}} \xrightarrow{L} Z \sim N(0, 1)$$

$$\Rightarrow P(10 \leq X \leq 16) = P(9.5 \leq X \leq 16.5)$$

$$= P\left(\frac{9.5 - 60}{\sqrt{24}} \leq \frac{X - 60}{\sqrt{24}} \leq \frac{16.5 - 60}{\sqrt{24}}\right)$$

$$\approx \Phi\left(\frac{16.5 - 60}{\sqrt{24}}\right) - \Phi\left(\frac{9.5 - 60}{\sqrt{24}}\right)$$

$$= \dots$$

①⑧  $X_n$  has p.d.f

$$f_n(x) = \begin{cases} \frac{1}{n} e^{-x} x^{n-1} & x > 0 \\ 0 & x \leq 0 \end{cases}$$

m.g.f of  $X_n$

$$\begin{aligned} M_{X_n}(t) &= \frac{1}{n} \int_0^\infty e^{tx} e^{-x} x^{n-1} dx \\ &= \frac{1}{n} \int_0^\infty e^{-x(1-t)} x^{n-1} dx \\ &= \frac{1}{(1-t)^n} = (1-t)^{-n} \end{aligned}$$

m.g.f of  $Y_n = \frac{X_n}{n}$

$$M_{Y_n}(t) = E\left(e^{t \frac{X_n}{n}}\right) = \left(1 - \frac{t}{n}\right)^{-n}$$

$$\rightarrow e^t \text{ as } n \rightarrow \infty$$

$$Y_n \xrightarrow{L} X(\text{degen at } 1)$$

↑ m.g.f n.v. degenerate dengan at x=

(19)

$$f_{X_n}(x) = \begin{cases} \frac{1}{\Gamma_p \alpha^p} e^{-x/\alpha} x^{p-1} & x > 0 \\ 0 & \text{o/w} \end{cases}$$

$$\alpha = 2, p = 4$$

$$E(X_n) = \alpha p = 8 \quad ; \quad V(X_n) = \alpha^2 p = 16 = \sigma^2$$

$$E(\bar{X}) = 8 \quad ; \quad V(\bar{X}) = \frac{16}{n} = \frac{1}{4}$$

By CLT

$$\frac{\sqrt{n}(\bar{X} - \alpha p)}{\sqrt{\alpha^2 p}} \xrightarrow{L} Z \sim N(0, 1)$$

$$\text{i.e. } 2(\bar{X} - 8) \xrightarrow{L} Z \sim N(0, 1)$$

$$P(7 < \bar{X} < 9) = P(2(7-8) < 2(\bar{X} - 8) < 2(9-8))$$

$$\approx P(-2 < Z < 2)$$

$$= \Phi(2) - \Phi(-2) = 2\Phi(2) - 1$$

(20)

$$X_i \sim U(0, 2) \quad E X_i = \frac{1}{2} \int_0^2 x \, dx = 1$$

$$E X_i^2 = \frac{1}{2} \int_0^2 x^2 \, dx = \frac{4}{3} \quad ; \quad V(X_i) = \frac{1}{3}$$

$$X_1, \dots, X_n \quad \text{i.i.d. with } E X_1 = 1 \text{ \& } V X_1 = \frac{1}{3}$$

$$\text{By CLT } \sqrt{n}(\bar{X}_n - 1) \xrightarrow{L} N(0, \frac{1}{3}).$$

$$\text{i.e. } \sqrt{n}(\bar{Y}_n - 1) \xrightarrow{L} N(0, \frac{1}{3}).$$

(10)