## Lecture 29

## Dynamic Linear Model: Part 2

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#### DLM: State equation (Recap)

▶ DLM in its basic form, employs an order one, *p*-dimensional vector autoregression as the state equation,

$$\mathbf{X}_t = \mathbf{\Phi} \mathbf{X}_{t-1} + \mathbf{W}_t,$$

where  $\boldsymbol{W}_{t} \stackrel{\textit{IID}}{\sim} \mathcal{N}_{p}(\boldsymbol{0}, \boldsymbol{Q})$ .

- In the DLM, we assume the process starts with a normal vector  $\mathbf{X}_0$ , such that  $\mathbf{X}_0 \sim \mathcal{N}_p(\mu_0, \Sigma_0)$ .
- ► Here *p* is called state dimension.

#### DLM: Observation equation (Recap)

- We do not observe the state vector  $\mathbf{X}_t$  directly, but only a linear transformed version of it with noise added, say  $\mathbf{Y}_t = \mathbf{A}_t \mathbf{X}_t + \mathbf{V}_t$ , where  $\mathbf{A}_t$  is a  $q \times p$  measurement or observation matrix; this equation is called the observation equation.
- The observed data vector,  $\mathbf{Y}_t$ , is q-dimensional, which can be larger than or smaller than p, the state dimension. The additive observation noise is  $\mathbf{V}_t \stackrel{\textit{IID}}{\sim} \mathcal{N}_q(\mathbf{0}, \mathbf{R})$ .
- ▶ For simplicity, we initially assume  $X_0$ ,  $\{W_t\}$ , and  $\{V_t\}$  are uncorrelated.

### Regression with Autocorrelated Errors (Recap)

- We discuss the modifications to the regression model when the errors are correlated.
- ► That is, consider the regression model

$$Y_t = \sum_{j=1}^r \beta_j u_{tj} + X_t$$

- ▶ Here  $X_t$  is a process with some covariance function  $\gamma_X(s, t)$ .
- In ordinary least squares, the assumption is that  $X_t$  is white Gaussian noise, in which case  $\gamma_X(s,t)=0$  for  $s\neq t$  and  $\gamma_X(t,t)=\sigma^2$ , independent of t.
- If this is not the case, then weighted least squares should be used.

#### Multivariate ARMAX models

- Suppose we have a multivariate time series  $Y_t$ , t = 1, 2, ...
- ▶ That is, consider the regression model

$$Y_{it} = \sum_{j=1}^r \beta_{ij} u_{tj} + X_{it}, \quad i = 1, \ldots, q.$$

In vector-matrix notation

$$\mathbf{Y}_t = \mathbf{B}\mathbf{u}_t + \mathbf{X}_t.$$

- ► Here  $X_t$  is a multivariate ARMA process with some covariance function matrix  $\Gamma_X(s,t)$ .
- We have not covered multivariate ARMA but have discussed multivariate time series and univariate ARMA separately; this is a combination of the concepts.

#### DLM with covariates

► The ARMAX model involves covariates that may enter into the states or into the observations.

▶ In this case, we suppose we have an  $r \times 1$  vector of inputs  $\mathbf{u}_t$ , and write the model as

$$egin{aligned} oldsymbol{X}_t &= \Phi oldsymbol{X}_{t-1} + \gamma oldsymbol{u}_t + oldsymbol{W}_t \ oldsymbol{Y}_t &= oldsymbol{A}_t oldsymbol{X}_t + \Gamma oldsymbol{u}_t + oldsymbol{V}_t \end{aligned}$$

▶ Here  $\gamma$  is  $p \times r$  and  $\Gamma$  is  $q \times r$ ; either of these matrices may be the zero matrix.

### Example of DLM (Global temperature data)

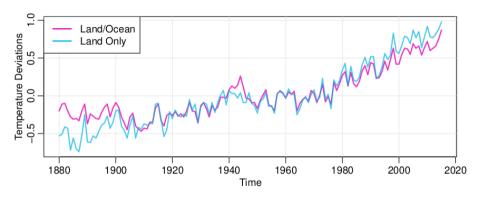


Fig. 6.3. Annual global temperature deviation series, measured in degrees centigrade, 1880–2015. The series differ by whether or not ocean data is included.

#### Example of DLM (Global temperature data, contd.)

► They show two different estimators for the global temperature series from 1880 to 2015.

- First series are the global mean land-ocean temperature index data.
- ► The second series are the surface air temperature index data using only meteorological station data.
- Conceptually, both series should be measuring the same underlying climatic signal, and we may consider the problem of extracting this underlying signal.

### Example of DLM (Global temperature data, contd.)

We suppose both series are observing the same signal with different noises; that is,  $Y_{t1} = X_t + V_{t1}$  and  $Y_{t2} = X_t + V_{t2}$  or more compactly as

$$(Y_{t1}, Y_{t2})' = (1, 1)'X_t + (V_{t1}, V_{t2})',$$

where  $\mathbf{R} = \text{Var}[(V_{t1}, V_{t2})'].$ 

It is reasonable to suppose that the unknown common signal  $X_t$  can be modeled as a random walk with drift of the form

$$X_t = \delta + X_{t-1} + W_t,$$

with  $Q = Var(W_t)$ .

▶ In this example, p = 1, q = 2,  $\Phi = 1$ , and  $\gamma = \delta$  with  $u_t = 1$ .



### An AR(1) Process with Observational Noise

Consider a univariate state-space model where the observations are noisy,

$$Y_t = X_t + V_t$$

► The signal (state) is an AR(1) process,

$$X_t = \phi X_{t-1} + W_t$$

- ► Here  $V_t \stackrel{\text{IID}}{\sim} \mathcal{N}(0, \sigma_V^2)$ ,  $W_t \stackrel{\text{IID}}{\sim} \mathcal{N}(0, \sigma_W^2)$ , and  $X_0 \sim \mathcal{N}(0, (1 \phi^2)^{-1} \sigma_W^2)$
- ▶ Besides,  $X_0$ ,  $\{W_t\}$ , and  $\{V_t\}$  are independent.

#### An AR(1) Process with Observational Noise (contd.)

ightharpoonup The autocovariance function of  $X_t$  is

$$\gamma_X(h) = (1 - \phi^2)^{-1} \sigma_W^2 \phi^h, \quad h = 0, 1, 2, \dots$$

ightharpoonup The marginal variance of  $Y_t$  is

$$\gamma_{Y}(0) = \text{Var}(X_t + V_t) = \text{Var}(X_t) + \text{Var}(V_t) = (1 - \phi^2)^{-1} \sigma_W^2 + \sigma_V^2$$

ightharpoonup The autocovariance function of  $Y_t$  is

$$\gamma_{Y}(h) = \text{Cov}(X_{t+h} + V_{t+h}, X_t + V_t) = \text{Cov}(X_{t+h}, X_t) = (1 - \phi^2)^{-1} \sigma_W^2 \phi^h, \ h = 1, 2, ...$$

ightharpoonup The ACF of  $Y_t$  is

$$\rho_Y(h) = \gamma_Y(h)/\gamma_Y(0) = (1 + \sigma_V^2/\sigma_W^2(1 - \phi^2))^{-1}\phi^h, \ h = 1, 2, \dots$$

▶ The ACF of  $Y_t$  is identical to the ACF of an ARMA(1,1) process.

### Things to cover

- We will cover the concepts of
  - prediction

filtering

smoothing

state space models and include their derivations.

# Thank you!