

Homework 4 - High Dimensional Probability

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Exercise 1

Let A be a subset of the sphere $\sqrt{n}S^{n-1}$ such that

$$\sigma(A) > 2 \exp(-cs^2) \quad \text{for some } s > 0$$

- (i) Prove that $\sigma(A_s) > 1/2$
- (ii) Deduce from this that for any $t \geq s$

$$\sigma(A_{2t}) \geq 1 - 2 \exp(-ct^2).$$

Solution:

- (i) $\sigma(A_s) > \frac{1}{2}$: suppose by contradiction that $\sigma(A_s) \leq \frac{1}{2}$. This would imply that $\sigma(B) \geq \frac{1}{2}$, where $B = (A_s)^c$. We could therefore apply the Blow Up Lemma¹ on B :

$$\sigma(B_s) \geq 1 - 2 \exp(-cs^2) > 1 - \sigma(A) \implies \sigma(B_s) + \sigma(A) > 1.$$

This is a contradiction since $B = (A_s)^c \iff B_s = A^c \implies \sigma(B_s) + \sigma(A) = 1$. Therefore $\sigma(A_s) > \frac{1}{2}$, and the first point is proven.

- (ii) Consider $t \geq s$. Since we proved that $\sigma(A_s) > \frac{1}{2}$, we can apply the Blow Up Lemma on A_s and obtain an inequality of this form:

$$\sigma((A_s)_\varepsilon) \geq 1 - 2 \exp(-c\varepsilon^2) \tag{1}$$

Where we want to put $(A_s)_\varepsilon = A_{2t}$ (this makes sense since $2t > s$). Now we only need to find the value of ε and plug it into inequality (1).

$$\varepsilon = 2t - s = t + (t - s) = t + \tilde{c}.$$

¹Let A be a subset of the sphere $\sqrt{n}S^{n-1}$ and let σ denote the normalized area on that sphere. If $\sigma(A) \geq 1/2$, then $\forall t \geq 0$, $\sigma(A_t) \geq 1 - 2 \exp(-ct^2)$.

Where $\tilde{c} \geq 0$ since $t \geq s$. Therefore we have:

$$\sigma(A_{2t}) \geq 1 - 2 \exp\left(-c(t + \tilde{c})^2\right) = 1 - 2 \exp\left(-ct^2 - c\tilde{c}^2 - 2tc\tilde{c}\right) \geq 1 - 2 \exp\left(-ct^2\right)$$

Since $c\tilde{c}^2, 2tc\tilde{c} \geq 0$.

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Exercise 2

Consider a random vector X taking values in some metric space (T, d) . Assume that there exists $K > 0$ such that

$$\|f(X) - \mathbb{E}f(X)\|_{\psi_2} \leq K\|f\|_{\text{Lip}} \quad (2)$$

for every Lipschitz function $f : T \rightarrow \mathbb{R}$. For a subset $A \subset T$, define

$\sigma(A) := \mathbb{P}(X \in A)$. (Then σ is a probability measure on T).

Show that if $\sigma(A) \geq \frac{1}{2}$, then, for every $t \geq 0$,

$$\sigma(A_t) \geq 1 - 2 \exp\left(-\frac{ct^2}{K^2}\right)$$

where $c > 0$ is an absolute constant.

Solution:

We apply the assumption to $f(x) := \text{dist}(x, A) = \inf\{d(x, y) : y \in A\}$ and we call M the median of $f(X)$. To find M we observe:

$$\sigma(A) \geq \frac{1}{2} \implies \mathbb{P}\{X \in A\} \geq \frac{1}{2} \iff \mathbb{P}\{f(X) = 0\} \geq \frac{1}{2} \implies M = 0.$$

Recalling the result of Exercise 5.1.13 in the book and combining it with (2) and the fact that $M = 0$, we obtain

$$\|f(X)\|_{\psi_2} \leq C\|f(X) - \mathbb{E}f(X)\|_{\psi_2} \leq CK\|f\|_{\text{Lip}}, \quad \exists C, K > 0. \quad (3)$$

Now we can conclude observing that

$$\begin{aligned} \sigma(A_t) &= \mathbb{P}\{X \in A_t\} = \mathbb{P}\{f(X) \leq t\} \\ &= \mathbb{P}\{|f(X)| \leq t\} \geq 1 - 2 \exp\left(-\frac{ct^2}{\|f(X)\|_{\psi_2}^2}\right) \\ &\geq 1 - 2 \exp\left(-\frac{ct^2}{K^2}\right), \end{aligned}$$

where the first inequality follows from the properties of a sub-gaussian random variable.

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Exercise 3

Let D be a $k \times m$ matrix and B be an $m \times n$ matrix. Prove that

$$\|DB\|_F \leq \|D\| \|B\|_F.$$

Solution:

$$\begin{aligned} \|DB\|_F &= \sqrt{\sum_{i=1}^{\min(k,n)} \sigma_i^2(DB)} \stackrel{(2)}{\leq} \sqrt{\sum_{i=1}^{\min(k,n)} \sigma_1^2(D) \sigma_i^2(B)} \\ &= \sigma_1(D) \cdot \sqrt{\sum_{i=1}^{\min(k,n)} \sigma_i^2(B)} \\ &= \|D\| \cdot \sqrt{\sum_{i=1}^{\min(k,n)} \sigma_i^2(B)} \\ &= \|D\| \cdot \|B\|_F \end{aligned}$$

where σ_i are the singular values, (2) holds since $\sigma_i^2(DB) \leq \sigma_1^2(D) \sigma_i^2(B)$ and the last equalities were obtained in Homework 3. ■

Exercise 4

Consider *i.i.d.* random variables $\delta_{ij} \sim \text{Ber}(p)$ where $i, j = 1, \dots, \mu$. Assuming that $pn \geq \log n$, show that

$$\mathbb{E} \max_i \sum_{j=1}^n (\delta_{ij} - p)^2 \leq Cpn.$$

Solution:

$$\begin{aligned} \exp \left(\mathbb{E} \max_i \sum_j (\delta_{ij} - p)^2 \right) &\stackrel{\text{Jensen}}{\leq} \mathbb{E} \exp \left(\max_i \sum_j (\delta_{ij} - p)^2 \right) \stackrel{*}{=} \mathbb{E} \max_i \exp \left(\sum_j (\delta_{ij} - p)^2 \right) \\ &\leq \mathbb{E} \sum_{i=1}^n \exp \left(\sum_j (\delta_{ij} - p)^2 \right) = \sum_{i=1}^n \mathbb{E} \exp \left(\sum_j (\delta_{ij} - p)^2 \right) \\ &= n \mathbb{E} \left[\prod_{j=1}^n \exp \left((\delta_{ij} - p)^2 \right) \right] = n \prod_{j=1}^n \mathbb{E} \left[\exp \left((\delta_{ij} - p)^2 \right) \right] \\ &= n \left(\mathbb{E} \left[\exp \left((\delta_{i1} - p)^2 \right) \right] \right)^n \\ &= n \left(p \exp \left((1-p)^2 \right) + (1-p) \exp \left(p^2 \right) \right)^n, \end{aligned}$$

where in \star we leveraged the monotonicity of the exponential and in the last equalities we the i.i.d. property of our variables.

We now show that $\exists k$ such that $p \exp((1-p)^2) + (1-p) \exp(p^2) \leq \exp(kp)$:

- If $p \geq 1-p$: $\exp((1-p)^2) \leq \exp(p^2) \leq \exp(p)$
 $\implies p \exp((1-p)^2) + (1-p) \exp(p^2) \leq \exp(p) \implies k = 1$;
- If $1-p > p$: $\exp((1-p)^2) + (1-p) \exp(p^2) \leq \exp((1-p)^2) \leq \exp(1-p)$,
which is smaller than $\exp(kp)$, for $k > \frac{1}{p}$.

Hence we have:

$$\exists k \quad s.t. \quad \exp \left(\mathbb{E} \max_i \sum_j (\delta_{ij} - p)^2 \right) \leq n e^{kp n}.$$

Finally, taking the $\log(\cdot)$ of both sides:

$$\mathbb{E} \max_i \sum_j (\delta_{ij} - p)^2 \leq \log(n) + kp n \leq pn + kp n = (1+k)pn = Cpn.$$

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