

WLLN

If X_1, \dots, X_n are iid RVs w/ mean μ & finite variance, then

$$P\left(\left|\frac{1}{n} \sum_{i=1}^n X_i - \mu\right| \geq \epsilon\right) \rightarrow 0$$

$\underbrace{\frac{1}{n} \sum_{i=1}^n X_i}_{\text{empirical mean}} \quad \underbrace{\mu}_{\text{actual mean}}$
 Def M_n

Proof: By Chebyshev, $P\left(\left|\frac{1}{n} \sum_{i=1}^n X_i\right| \leq \frac{\sigma^2}{n\epsilon^2}\right) \rightarrow 0$ as $n \rightarrow \infty$

The WLLN tells us that the empirical frequency estimate is a pretty good estimate of the actual mean

$$\Rightarrow \lim_{n \rightarrow \infty} P(|M_n - \mu| \geq \epsilon) = 0$$

Converging in probability:

$$\forall \epsilon, \delta, \exists n(\epsilon, \delta) \text{ s.t. } P(|M_n - \mu| \geq \epsilon) < \delta \quad \forall n > n(\epsilon, \delta)$$

Remarks:

- ϵ captures "accuracy" level
- δ captures "confidence" level
- $n(\epsilon, \delta)$ captures "threshold" for a given accuracy & confidence level

We say $M_n \xrightarrow{P} \mu$ M_n converges to μ in probability

Formally, let Y_1, \dots, Y_n be a sequence of RVs.

Then, we say the sequence converges to a number a in prob if

$$\forall \epsilon > 0 \quad \lim_{n \rightarrow \infty} P(|Y_n - a| > \epsilon) = 0.$$

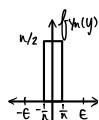
Ex 1: X_1, \dots, X_n iid Unif $[1, 1]$ RVs.



$1/Y_n = \frac{X_n}{n}$ converge in prob?

$$Y_n \leq y \Rightarrow X_n \leq ny \Rightarrow F_{Y_n}(y) = F_X(ny)$$

$$f_{Y_n}(y) = n f_X(ny)$$

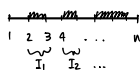


limit
 $P(|Y_n - 0| > \epsilon) = 0$ if $\frac{1}{n} < \epsilon$ or $n > \frac{1}{\epsilon}$

Ex 2: Suppose time in discrete units and $Y_n = 1$,

there's some arrival at time n & $Y_n = 0$ else.

$$\text{Def } I_k = \{2^k, 2^{k+1}, 2^{k+1}-1\}$$



Suppose exactly 1 arrival in each interval & it is equally likely in that interval.

$$\Rightarrow P(Y_n = 1) = \frac{1}{2^k} \text{ if } n \in I_k$$

$$\lim_{n \rightarrow \infty} P(Y_n = 1) = \lim_{k \rightarrow \infty} \frac{1}{2^k} = 0$$

$\hookrightarrow Y_n$ converges to 0 in prob.

Remark: Given any finite n , there are certain to be an infinite num of arrivals after n
 \Rightarrow this is the weakness of the weak law.

SLLN

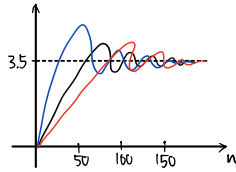
Let X_n be a seq. of iid RVs w/ μ mean,

then $\frac{1}{n} \sum_{i=1}^n X_i \rightarrow \mu$ as $n \rightarrow \infty$ w.p. 1.

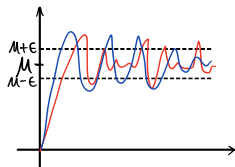
i.e. sample mean $M_n = \frac{1}{n} \sum_{i=1}^n X_i$ converges to the exp value μ w.p. 1.

$$\boxed{P\left(\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n X_i = \mu\right) = 1} \quad M_n \xrightarrow{a.s.} \mu$$

Rolling 6-sided die \rightarrow sample mean converges to 3.5



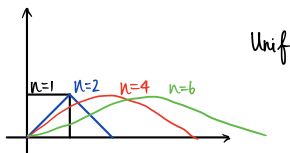
SLLN: Every realization converges to μ .



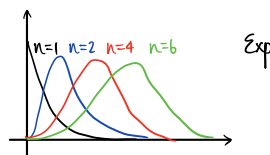
WLLN: Many diverge out of the bound.

CLT

Q/ What happens to $S_n = \sum_{i=1}^n X_i = X_1 + \dots + X_n$, X_i iid



Unif



Exp

Observe: Looks more like a normal as we take more copies of any distr

Var & mean of $S_n \rightarrow \infty$ as $n \rightarrow \infty$ (distr gets really flat & wide)

\hookrightarrow need to normalize

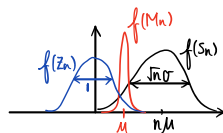
$$\mathbb{E}[S_n] = n\mu$$

$$\text{Var}(S_n) = n\sigma^2$$

$$M_n = \frac{S_n}{n}$$

$$Z_n = \frac{S_n - n\mu}{\sqrt{n}\sigma} \Rightarrow \mathbb{E}[Z_n] = 0$$

$$\text{Var}(Z_n) = 1$$



$$\boxed{\text{CLT: } \lim_{n \rightarrow \infty} P(Z_n \leq x) = \Phi(x) \quad \forall x}$$

where $\Phi(x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-z^2/2} dz$ is the CDF of $N(0,1)$

Note: If there are a # of indep small factors, the aggregate of these factors will be approximately normally distributed.

\Rightarrow models noise very well

Aside: If $Y \sim N(0,1) \Rightarrow \mathcal{M}_Y(s) = \mathbb{E}[e^{sY}] = e^{s^2/2}$

$$\log \mathcal{M}_Y(s) = \frac{s^2}{2}$$

Proof: Show that $M_Z(s)$ converges to $N(0,1)$.

X_1, \dots, X_n are iid RVs w/ mean 0 & var 1.

$$Z_n = \frac{\sum_{i=1}^n X_i}{\sqrt{n}}$$

$$M_{Z_n}(s) = \mathbb{E}[e^{sZ_n}] = \mathbb{E}[e^{s \sum_{i=1}^n X_i / \sqrt{n}}]$$

$$= \mathbb{E}[e^{\frac{s}{\sqrt{n}} X_1} \cdot e^{\frac{s}{\sqrt{n}} X_2} \dots e^{\frac{s}{\sqrt{n}} X_n}]$$

$$= \mathbb{E}[e^{\frac{s}{\sqrt{n}} X_1}]^n$$

$$M_{Z_n}(s) = [M_X(\frac{s}{\sqrt{n}})]^n$$

cumulant
MGF

$$\lim_{n \rightarrow \infty} \log M_{Z_n}(s) = \lim_{n \rightarrow \infty} n \log M_X(\frac{s}{\sqrt{n}})$$

$$= \lim_{n \rightarrow \infty} \frac{\log M_X(\frac{s}{\sqrt{n}})}{1/\sqrt{n}}$$

Let $y = \frac{1}{\sqrt{n}}$

$$= \lim_{y \rightarrow 0} \frac{\log M_X(sy)}{y^2}$$

Use L'Hos

$$= \lim_{y \rightarrow 0} \frac{\frac{1}{M_X(sy)}}{2y M_X'(sy)}$$

L'Hos

$$= \lim_{y \rightarrow 0} \frac{\frac{1}{M_X(sy)}}{2y M_X''(sy)}$$

$$= \frac{s^2}{2}$$

Recall: $M_X(0) = 1$, $M_X'(0) = 0$, $M_X''(0) = 1$

Ex: Polling

We ask n randomly chosen voters if they support

candidate $T \Rightarrow X_i = \begin{cases} 1 & \text{if yes} \\ 0 & \text{if no} \end{cases}$ X_i 's iid

Q/ We want 95% confidence that $|M_n - p| < \epsilon$ where p = true popularity of T .

M_n = empirical popularity of $T = \frac{1}{n} \sum_{i=1}^n X_i$

$$\text{Chebyshev: } P(|M_n - p| \geq a) \leq \frac{\text{Var}(M_n)}{a^2} = \frac{p(1-p)}{n a^2}$$

$$X_i \sim B(p)$$

$$\leq \frac{1}{4na^2}$$

$$\mathbb{E}[X_i] = p$$

$$a = 0.1: P(|M_n - p| \geq 0.1) \leq \frac{1}{4n(0.01)} = \frac{25}{n}$$

$$\text{Var}(X_i) = p(1-p)$$

$$\leq 0.05 \text{ (95\% confidence)}$$

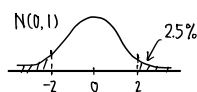
$$\Rightarrow n \geq 500$$

$$\text{CLT: } \frac{M_n - \mathbb{E}[M_n]}{\sqrt{\text{Var}(M_n)}} \xrightarrow{d} N(0,1)$$

$$\text{Want } P(|M_n - p| \geq 0.1) \leq 0.05$$

$$P\left(\underbrace{\frac{M_n - p}{1/\sqrt{n}}}_{N(0,1)} \geq \underbrace{\frac{0.1}{1/\sqrt{n}}}_{0.2\sqrt{n}}\right) \leq 0.05$$

\hookrightarrow now use Z table



$$0.2\sqrt{n} \leq 2$$

$$\Rightarrow n \leq 100$$